Kolmogorov Approximation

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1 1 Introduction

- 2 Many different approaches to approximation of probability distributions are studied in the literature [9,
- 3 12, 13]. The papers vary in the types random variables involved, how they are represented, and in
- 4 the criteria used for evaluation of the quality of the approximations. This paper is on approximating
- 5 discrete distributions represented as explicit probability mass functions with ones that are simpler to
- store and to manipulate. This is needed, for example, when a discrete distribution is given as a large
- data-set, obtained, e.g., by sampling, and we want to represent it approximately with a small table.
- 8 The main contribution of this paper is an efficient algorithm for computing the best possible approxi-
- 9 mation of a given random variable with a random variable whose complexity is not above a prescribed
- threshold, where the measures of the quality of the approximation and the complexity of the random
- variable are as specified in the following two paragraphs.
- 12 We measure the quality of an approximation by the distance between the original variable and the
- 13 approximate one. Specifically, we use the Kolmogorov distance which is one of the most used
- in statistical practice and literature. Given two random variables X and X' whose cumulative
- distribution functions (cdfs) are F_X and $F_{X'}$, respectively, the Kolmogorov distance between X and
- 16 X' is $d_K(X,X') = \sup_t |F_X(t) F_{X'}(t)|$ (see, e.g., [8]). We say taht X' is a good approximation
- of X if $d_K(X, X')$ is small.
- 18 The complexity of a random variable is measured by the size of its support, the number of values that
- 19 it can take, $|\operatorname{support}(X)| = |\{x \colon Pr(X = x) \neq 0\}|$. When distributions are maintained as explicit
- 20 tables, as done in many implementations of statistical software, the size of the support of a variable is
- 21 proportional to the amount of memory needed to store it and to the complexity of the computations
- 22 around it.
- 23 In summary, the exact notion of optimality of the approximation targeted in this paper is:
- **Definition 1.** A random variable X' is an optimal m-approximation of a random variable X if
- | support(X')| $\leq m$ and there is no random variable X'' such that $|\operatorname{support}(X'')| \leq m$ and
- 26 $d_k(X, X'') < d_k(X, X')$.
- The main contribution of the paper is an efficient algorithm that takes X and m as parameters and
- constructs an optimal m-approximation of X.
- 29 The rest of the paper is organized as follows. In Section 2 we describe how our work relates to other
- algorithms and problems studied in the literature. In Section 3 we detail the proposed algorithm,

analyze its properties, and prove Theorem ??. In Section 4 we demonstrate how the proposed approach performs on the problem of estimating the probability of hitting deadlines is plans and compare it to alternatives approximation approaches from the literature. We also demonstrate the performance of our approximation algorithm on some randomly generated random variables. The paper is concluded with a discussion in Section 5.

36 2 Related Work

The problem studied in this paper is related to the theory of Sparse Approximation (aka Sparse Representation) that deals with sparse solutions for systems of linear equations, as follows.

Given a matrix $D \in \mathbb{R}^{n \times p}$ and a vector $x \in \mathbb{R}^n$, the most studied sparse representation problem is finding the sparsest possible representation $\alpha \in \mathbb{R}^p$ satisfying $x = D\alpha$:

$$\min_{\alpha \in \mathbb{R}^p} \|\alpha\|_0$$
 subject to $x = D\alpha$.

where $\|\alpha\|_0 = |\{i: \alpha_i \neq 0, i = 1, \dots, p\}|$ is the ℓ_0 pseudo-norm, counting the number of non-zero coordinates of α . This problem is known to be NP-Hard with a reduction to NP-complete subset selection problems.

In these terms, using also the ℓ_{∞} norm that represents the maximal coordinate and the ℓ_1 norm that represents the sum of the coordinates, our problem can be phrased as:

$$\min_{\alpha \in [0,\infty)^p} \|x - D\alpha\|_{\infty} \text{ subject to } \|\alpha\|_0 = m \text{ and } \|\alpha\|_1 = 1.$$

where D is the all-ones triangular matrix (the entry at row i and column j is one if $i \leq j$ and zero otherwise), x is related to X such that the ith coordinate of x is $F_X(x_i)$ where $\operatorname{support}(X) = \{x_1 < x_2 < \cdots < x_n\}$ and α is related to X' such that the ith coordinate of α is $f_{X'}(x_i)$. The functions F_X and $f_{X'}$ represent, respectively, the cumulative distribution function of X and the mass distribution function of X'. This, of course, means that the coordinates of x are assumed to be positive and monotonically increasing and that the last coordinate of x is assumed to be one. We demonstrate an application for this specific sparse representation problem and show that it can be solve in $O(n^2m)$ time and $O(m^2)$ memory.

50 3 An Algorithm for Optimal Approximation

Let, in the scope of this section, X be a given random variable with a finite support of size n, and let $0 < m \le n$ be a given complexity bound. We describe an algorithm for finding find an m-optimal approximation of X in steps.

The first step is to show that it is enough to limit our search to approximations X's such that

support $(X') \subseteq \operatorname{support}(X)$.

Lemma 2. There is an m-optimal-approximation X' of X such that $support(X') \subseteq support(X)$.

Proof. Let X'' be any m-optimal approximation of X. Let $\{x_1, \ldots, x_n\} = \operatorname{support}(X)$. The random variable X' whose probability mass function is $f_{X'}(x_i) =$.

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Next, note that every random variable X'' with support of size at most m that is contained in $\operatorname{support}(X)$ be described by first setting the (at most m) elements of the support of X''; then for

- every such option, determine X'' by setting probability values for the elements in the chosen support of X', and setting 0 for rest of the elements.
- Since from Lemma 2 we can assume without loss of generality that if X' is an m-optimal approx-
- imation variable for X then $\operatorname{support}(X') \subseteq \operatorname{support}(X)$, our search to find such X' takes two
- steps. Denote the set of random variables with support S by \mathbb{X}_S . In step 1, we find the m-optimal
- approximation random variable among all random variables in \mathbb{X}_S , and denote the m-optimal distance
- for X_S by $\varepsilon(X,S)$. Next, in Step 2, among all the possible supports we find the support setting S of
- size $\leq m$ for which $\varepsilon(X,S)$ is minimal: We describe an efficient way to do so.

70 **3.1 Step 1**

- 71 We first fix a set $S \subseteq \operatorname{support}(X)$ of size at most m, and among all the random variables in \mathbb{X}_S
- find one with a minimal distance from X. To that, set $S = \{x_1 < \dots < x_m\} \subseteq \operatorname{support}(X)$. To
- simplify the proofs set $x_0 = -\infty$, and $x_{m+1} = \infty$. Then $x_0 < x_1$ and $x_m < x_{m+1}$. In addition
- recall that for every random variable X'', $F_{X''}(-\infty) = 0$ and $F_{X''}(\infty) = 1$. Finally, for every
- 75 $1 \le i \le m$ let \hat{x}_i be the maximal element of support (X) that is smaller than x_i . For the rest of this
- section we assume S is fixed and therefore is not necessarily included in the notation.
- 77 Next, as the elements of S are also elements of $\operatorname{support}(X)$, we can define the following weight
- function that we use to find the m-optimal distance $\varepsilon(X,S)$.
- 79 **Definition 3.** For $0 \le i \le m$ let

$$w(x_i, x_{i+1}) = \begin{cases} P(x_i < X < x_{i+1}) & \text{if } i = 0 \text{ or } i = m; \\ P(x_i < X < x_{i+1})/2 & \text{otherwise.} \end{cases}$$

- Note that $x_i = -\infty$ for i = 0 and $x_i = \infty$ for i = m + 1. Also note that $P(x_i < X < x_{i+1}) = 0$
- $F_X(\hat{x}_{i+1}) F_X(x_i)$, a fact that we will use throughout this section.
- Definition 4. Let $\varepsilon(X,S) = \max_{i=0,\dots,m} w(x_i,x_{i+1})$.
- We first show that $\varepsilon(X,S)$ is a lower bound. That is, every random variable in \mathbb{X}_S has a distance at
- least $\varepsilon(X,S)$. Then, we present a random variable $X' \in \mathbb{X}_S$ with distance $\varepsilon(X,S)$. It then follows
- that such X' is an m-optimal approximation random variable among all random variables in \mathbb{X}_S .
- The intuition behind choosing these specific weights and $\varepsilon(X,S)$ being a lower bound is as follows.
- Since for every $X' \in \mathbb{X}_S$ the probability values of X' for the elements not in S are set to 0, we have
- that $F_{X'}(\hat{x}_{i+1}) = F_{X'}(x_i)$. Therefore the distance between X' and X at points x_i and \hat{x}_{i+1} that we
- have to take into additional account is increased by $F_X(\hat{x}_{i+1}) F_X(x_i) = P(x_i < X < x_{i+1})$.
- 90 Formally we have the following.
- Proposition 5. If $X' \in \mathbb{X}_S$ then $d_k(X, X') \geq \varepsilon(X, S)$.
- 92 *Proof.* By definition, for every $0 \le i \le m$, $d_k(X, X') \ge \max\{|F_X(\hat{x}_{i+1}) F_{X'}(\hat{x}_{i+1})|, |F_X(x_i) F_{X'}(\hat{x}_{i+1})|\}$
- 93 $F_{X'}(x_i)$ | }. Note that $F_{X'}(\hat{x}_{i+1}) = F_{X'}(x_i)$ since the probability values for all the elements not in S
- 94 are set to 0.
- 95 If i=0, that is $x_i=-\infty$, we have that $F_X(x_i)=F_{X'}(x_i)=F_{X'}(\hat{x}_{i+1})=0$ and therefore
- 96 $d_k(X, X') \ge |F_X(\hat{x}_{i+1})| = |F_X(\hat{x}_{i+1}) F_X(x_i)| = P(x_i < X < x_{i+1}) = w(x_i, x_{i+1}).$
- 97 If i=m, that is $x_{i+1}=\infty$, we have that $F_X(\hat{x}_{i+1})=F_{X'}(\hat{x}_{i+1})=F_{X'}(x_i)=1$. and therefore
- 98 $d_k(X, X') \ge |1 F_X(\hat{x}_i)| = |F_X(\hat{x}_{i+1}) F_X(x_i)| = P(x_i < X < x_{i+1}) = w(x_i, x_{i+1}).$

- 99 Otherwise for every $1 \le i < m$, we use the fact that $max\{|a|,|b|\} \ge |a-b|/2$ for every $a,b \in$
- 100 \mathbb{R} , to have $d_k(X,X') \geq 1/2|F_X(\hat{x}_{i+1}) F_X(x_i) + F_{X'}(x_i) F_{X'}(\hat{x}_{i+1})|$. So $d_k(X,X') \geq 1/2|F_X(\hat{x}_{i+1}) F_X(x_i)|$
- 101 $1/2|F_X(\hat{x}_{i+1}) F_X(x_i)| = P(x_1 < X < x_2)/2 == w(x_i, x_{i+1}).$
- Therefore since $d_k(X, X') \geq w(x_i, x_{i+1})$ for every $0 \leq i \leq m$, by definition of $\varepsilon(X, S)$ proof
- 103 follows.
- Next we show a random variable $X' \in \mathbb{X}_S$ with a distance of $\varepsilon(X,S)$ from X. Thus X' is an
- noptimal approximation among the set X_S . We define X' as follows:
- 106 **Definition 6.** Let $f_{X'}(x_i) = w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)$ for i = 1, ..., m and $f_{X'}(x) = 0$
- 107 for $x \notin S$.
- We first see that X' is a properly defined random variable. We then discuss the properties of X'.
- 109 **Lemma 7.** $f_{X'}$ is a probability mass function.
- 110 *Proof.* From definition $f_{X'}(x_i) \ge 0$ for every i. To see that $\sum_i f_{X'}(x_i) = 1$, we have $\sum_i f_{X'}(x_i) = 1$
- $\sum_{i} (w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)) = \sum_{x_i \in S} f_X(x_i) + w(x_0, x_1) + \sum_{0 < i < m} 2w(x_i, x_{i+1}) + \sum_{0$
- 112 $w(x_m, x_{m+1}) = \sum_{x_i \in S} P(X = x_i) + P(x_0 < X < X_1) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x_i < X < X_{i+1}) + \sum_{0 < i < m} P(x$
- 113 $P(x_m < X < X_{m+1}) = 1$ since this sum is the entire cpt of X.
- Note that X' can be constructed in linear time to the size of the cdf of X. Intuitively the setting of
- 115 X' allows to take an "advantage" of distance from X at the elements of support (X'), to avoid the
- overall increased distance of X from X' at the elements that are not at $\mathrm{support}(X)$ and in which
- $f_{X'}$ is set to 0. Formally we have the following.
- **Lemma 8.** Let $x \in \text{support}(X)$ and $0 \le i \le m$ be such that $x_i \le x \le x_{i+1}$ then $-w(x_i, x_{i+1}) \le x_i$
- 119 $F_X(x) F_{X'}(x) \le w(x_i, x_{i+1}).$
- 120 *Proof.* We prove by induction on $0 \le i < m$.
- First see that $F_{X'}(j) = 0$ for every $x_0 < j < x_1$ and therefore $F_X(j) F_{X'}(j) = F_X(j) 0 \le 0$
- 122 $F_X(\hat{x}_1) = F_X(\hat{x}_1) F_X(x_0) = w(x_0, x_1)$. For $j = x_1$ we have $F_X(x_1) F_{X'}(x_1) = F_X(\hat{x}_1) + F_{X'}(x_1) = F_X(\hat{x}_1) + F_X(\hat{x}_1) = F_X(\hat{x$
- 123 $f_X(x_1) (w(x_0,x_1) + w(x_1,x_2) + f_X(x_1) = w(x_0,x_1) + f_X(x_1) (w(x_0,x_1) + w(x_1,x_2) + f_X(x_1) + f_X($
- 124 $f_X(x_1) = -w(x_1, x_2).$
- Next assume that $F_X(\hat{x}_i) F_{X'}(\hat{x}_i) = w(x_{i-1}, x_i)$. Then $F_X(x_i) F_{X'}(x_i) = F_X(\hat{x}_i) + f_X(x_i) F_{X'}(x_i)$
- 126 $(w(x_{i-1},x_i)+w(x_i,x_{i+1})+f_X(x_i))=w(x_{i-1},x_i)+f_X(x_i)-(w(x_{i-1},x_i)+w(x_i,x_{i+1})+f_X(x_i))$
- 127 $f_X(x_i) = -w(x_i, x_{i+1}).$
- As before we have that for all $x_i < j < x_{i+1}$, we have $F_X(j) F_{X'}(j) = F_X(j) F_{X'}(\hat{x}_{i+1}) \le$
- 129 $F_X(\hat{x}_{i+1}) F_{X'}(\hat{x}_{i+1})$. Then $F_X(\hat{x}_{i+1}) F_{X'}(\hat{x}_{i+1}) = (F_X(x_i) + P(x_i < x < x_{i+1})) F_{X'}(\hat{x}_{i+1})$
- 130 $F_{X'}(x_i) = -w(x_i, x_{i+1}) + 2w(x_i, x_{i+1}) = w(x_i, x_{i+1}).$
- Finally for $x_m \leq j \leq x_{m+1}$ we have that $F_{X'}(x_m) = 1$ therefore $F_X(x_m) F_{X'}(x_m) = (1 1)$
- 133 x_{m+1} we have $F_X(j) F_{X'}(j) < (1 P(x_m < X < x_{m+1})) 1 < -P(x_m < X < x_{m+1})) = 0$
- $-w(x_m,x_{m+1})$ as required.
- Finally the m-optimality of X' w.r.t. X_S is straightforward from Lemma 8 and the definition of
- 136 $\varepsilon(X,S)$. Therefore we have.
- 137 Corrolary 9. $d_k(X, X') = \varepsilon(X, S)$.

3.2 Step 2

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Chakravarty, Orlin, and Rothblum [2] proposed a polynomial-time method that, given a certain objective functions (additive), finds an optimal consecutive partition. Their method involves the construction of a graph such that the (consecutive) set partitioning problem is reduced to the problem 141 of finding the shortest path in that graph.

The KolmogorovApprox algorithm (Algorithm 1) starts by constructing a directed weighted graph 143 G similar to the method of Chakravarty, Orlin, and Rothblum [2]. The nodes V consist of the 144 support of X together with an extra two nodes, $-\infty$ and ∞ for technical reasons, whereas the 145 edges E connect every pair of nodes in one direction (lines 1-2). The weight w of each edge 146 $e = (x, y) \in E$ is determined by one of two cases as in Definition 3. The first is where nodes 147 x or y are the source or target nodes respectively. In this case, the weight is the probability of X 148 to get a value between x and y, non inclusive, i.e., w(e) = Pr(x < X < y). The second case 149 is where x and y are not a source or target nodes, here the weight is the probability of X to get a 150 value between x and y, non inclusive, divided by two i.e., w(e) = Pr(x < X < y)/2. The values 151 taken are non inclusive, since we are interested only in the error value. The source node of the 152 shortest path problem at hand corresponds to the $-\infty$ node added to G in the construction phase, 153 and the target node is the extra node ∞ . The set of all solution paths in G, i.e., those starting at 154 $-\infty$ and ending in ∞ with at most m edges, is called $paths(G, -\infty, \infty)$. The goal is to find the 155 path l in $paths(G, -\infty, \infty)$ with the lightest bottleneck (line 3). This can be achieved by using the 156 Bellman - Ford algorithm with two tweaks. The first is to iterate the graph G in order to find only 157 paths with length of at most m edges. The second is to find the lightest bottleneck as opposed to 158 the traditional objective of finding the shortest path. This is performed by modifying the manner of 159 "relaxation" to bottleneck(x) = min[max(bottleneck(v), w(e))], done also in [14]. Consequently, 160 we find the lightest maximal edge in a path of length $\leq m$, which represents the minimal error, 161 $\varepsilon(X,S)$, defined in Definition ?? where the nodes in path l represent the elements in set S. The 162 approximated random variable X' is then derived from the resulting path l (lines 4-5). Every node 163 $x \in l$ represent a value in the new calculated random variable X', we than iterate the path l to find 164 the probability of the event $f_{X'}(x)$ as described in Definition 6. For every edge $(x_i, x_j) \in l$ we 165 determine: if (x_i, x_j) is the first edge in the path l (i.e. $x_i = -\infty$), then node x_j gets the full weight 166 $w(x_i, x_j)$ and it's own weight in X such that $f_{X'}(x_j) = f_X(x_j) + w(x_i, x_j)$. If (x_i, x_j) in not the 167 first nor the last edge in path l then we divide it's weight between nodes x_i and x_j in addition to their 168 own original weight in X and the probability that already accumulated. If (x_i, x_j) is the last edge 169 in the path l (i.e. $i = \infty$) then node i gets the full weight $w(x_i, x_i)$ in addition to what was already 170 accumulated such that $f_{X'}(x_i) = f_{X'}(x_i) + w(x_i, x_i)$. 171

Algorithm 1: KolmogorovApprox(X, m)

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1 S = \operatorname{support}(X) \cup \{\infty, -\infty\}
2 G = (V, E) = (S, \{(x, y) : x < y\})
3 (x_0, \dots, x_{m+1}) = l = \operatorname{argmin}_{l \in paths(G, -\infty, \infty), |l| \le m} \max \{ w(e) : e \in l \}
4 for 0 < i < m + 1 do
    f_{X'}(x_i) = w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)
6 return X'
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Theorem 10. KolmogorovApprox(X, m) is an m-optimal-approximation of X.

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Proof. If we consider the vertexes S = l \setminus \{-\infty, \infty\} for a path l \in paths(G, -\infty, \infty), |l| \leq m we
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      have that \max\{w(e): e \in l\} = \varepsilon(X, S).
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Theorem 11. The KolmogorovApprox(X, m) algorithm runs in time $O(mn^2)$, using $O(n^2)$ mem-176 ory where $n = |\operatorname{support}(X)|$.

Proof. Constructing the graph G takes $O(n^2)$. The number of edges is $O(E) \approx O(n^2)$ and for every 177 edge the weight is at most the sum of all probabilities between the source node $-\infty$ and the target 178 node ∞ , which can be done efficiently by aggregating the weights of already calculated edges. The 179 construction is also the only stage that requires memory allocation, specifically $O(E+V) = O(n^2)$. 180 Finding the shortest path takes $O(m(E+V)) \approx O(mn^2)$. Since G is DAG (directed acyclic graph) 181 finding shortest path takes O(E+V). We only need to find paths of length $\leq m$, which takes 182 O(m(E+V)). Deriving the new random variable X' from the computed path l takes O(mn). For 183 every node in l (at most m nodes), calculating the probability $P(s < X < \infty)$ takes at most n. 184 To conclude, the worst case run-time complexity is $O(n^2 + mn^2 + mn) = O(mn^2)$ and memory 185 complexity is $O(E+V) = O(n^2)$. 186

187 4 A case study and experimental results

The case study examined in our experiments is the problem of task trees with deadlines [4, 3].
Hierarchical planning is a well-established field in AI [5, 6, 7], and is still relevant nowadays [1, 15].
A hierarchical plan is a method for representing problems of automated planning in which the
dependency among tasks can be given in the form of networks, here we focus on hierarchical plans
represented by task trees. The leaves in a task tree are *primitive* actions (or tasks), and the internal
nodes are either *sequence* or *parallel* actions. The plans we deal with are of stochastic nature, where
the duration of a primitive action is given by a random variable.

A sequence node denotes a series of tasks that should be performed consecutively, whereas a parallel 195 node denotes a set of tasks that begin at the same time. A valid plan is one that is fulfilled before some 196 given deadline, i.e., its makespan is less than or equal to the deadline. The objective in this context 197 is to compute the probability that a given plan is valid, or more formally computing P(X < T), 198 where X is a random variable representing the makespan of the plan and T is the deadline. As said 199 above, resource consumption (task duration) is uncertain, and described as probability distributions 200 in the leaf nodes. We assume that the distributions are independent but not necessarily identically 201 distributed and that the random variables are discrete and have a finite support. 202

The problem of finding the probability that a task tree satisfies a deadline is known to be NP-hard. In fact, even the problem of summing a set of random variables is NP-hard [10]. This is an example of an explicitly given random variable that we need to estimate deadline meeting probabilities for.

In the first experiment we focus on is the problem of task trees with deadlines, and consider three 206 types of task trees. The first type includes logistic problems of transporting packages by trucks and 207 airplanes (from IPC2 http://ipc.icaps-conference.org/). Hierarchical plans of those logistic problems 208 were generated by the JSHOP2 planner [11] (see example problem, Figure 1, one parallel node with 209 all descendant task nodes being in sequence). The second type consists of task trees used as execution 210 plans for the ROBIL team entry in the DARPA robotics challenge (DRC simulation phase), and the 211 third type is of linear plans (sequential task trees). The primitive tasks in all the trees are modeled as 212 discrete random variables with support of size M obtained by discretization of uniform distributions 213 over various intervals. The number of tasks in a tree is denoted by N. 214

We implemented the approximation algorithm for solving the deadline problem with four different methods of approximation. The first two are for achieving a one-sided Kolmogorov approximation – the OptTrim [3] and the Trim [4] operators, and the third is a simple sampling scheme. We used those

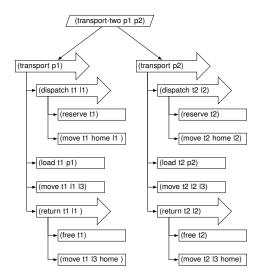


Figure 1: A plan generated by the JSHOP2 algorithm. Arrow shapes represent sequence nodes, parallelograms represent parallel nodes, and rectangles represent primitive nodes.

methods as a comparison to the Kolmogorov approximation with the suggested Kolmogorov Approx algorithm. The parameter m of $\operatorname{OptTrim}$ and $\operatorname{KolmogorovApprox}$ corresponds to the inverse of ε given to the Trim operator. Note that in order to obtain some error ε , one must take into consideration the size of the task tree N, therefore, $m/N=1/(\varepsilon\cdot N)$. We ran also an exact computation as a reference to the approximated one in order to calculate the error. The experiments conducted with the following operators and their parameters: Kolmogorov Approx operator with $m=10\cdot N$, the $\operatorname{OptTrim}$ operator with $m=10\cdot N$, the Trim as operator with $\varepsilon=0.1/N$, and two simple simulations, with a different samples number $s=10^4$ and $s=10^6$.

Task Tree	M	KolmogorovApprox	OptTrim	Trim	Sampling	
		m/N=10	m/N=10	$\varepsilon \cdot N = 0.1$	$s=10^4$	$s=10^{6}$
Logistics $(N = 34)$	2	0	0	0.0019	0.007	0.0009
	4	0.0024	0.0046	0.0068	0.0057	0.0005
Logistics	2	0.0002	0.0005	0.002	0.015	0.001
(N=45)	4	0	0.003	0.004	0.008	0.0006
DRC-Drive	2	0	0.004	0.009	0.0072	0.0009
(N=47)	4	0.001	0.008	0.019	0.0075	0.0011
	2	0.0093	0.015	0.024	0	0
Sequential	4	0	0.024	0.04	0.008	0.0016
(N=10)	10	0	0.028	0.06	0.0117	0.001

Table 1: Comparison of estimated errors with respect to the reference exact computation on various task trees.

Table 1 shows the results of the main experiment. The quality of the solutions provided by using the OptTrim operator are better (lower errors) than those provided by the Trim operator, following the optimality guarantees, but is interesting to see that the quality gaps happen in practice in each of the examined task trees. However, in some of the task trees the sampling method produced better results than the approximation algorithm with OptTrim. Nevertheless, the approximation algorithm comes with an inherent advantage of providing an exact quality guarantees, as opposed to the probabilistic guarantees provided by sampling.

In order to better understand the quality gaps in practice between KolmogorovApprox, OptTrim, and Trim, we investigate their relative errors when applied on single random variables with support

size n=100, and different support sizes of the resulting random variable approximation (m). In each instance of this experiment, a random variable is randomly generated by choosing the probabilities of each element in the support from a uniform distribution and then normalizing these probabilities so that they sum to one.

Figure 2 present the error produced by the above methods. The depicted results are averages over several instances (50 instances) of random variables. The curves in the figure show the average error of OptTrim and Trim operators with comparison to the average error of the optimal approximation provided by KolmogorovApprox as a function of m.

According to the depicted results it is evident that increasing the support size of the approximation m reduces the error, as expected, in all three methods. However, errors produced by the KolmogorovApprox are significantly smaller, safe to say, a half of the error produced by OptTrim and Trim, it is clear both in the table (the relative error is mostly above 1) and in the graph.

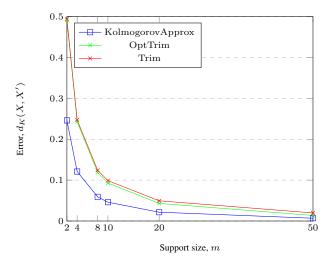


Figure 2: Error comparison between KolmogorovApprox, OptTrim, and Trim, on randomly generated random variables as function of m.

We also examined how our algorithm compares to linear programing as described and discussed, for example, in [12]. We ran an experiment to compare the run-time between the KolmogorovApprox algorithm with the run-time of a state-of-art implementation of linear programing. We used the "Minimize" function of Wolfram Mathematica and fed it with the equations $\min_{\alpha \in \mathbb{R}^n} \|x - \alpha\|_{\infty}$ subject to $\|\alpha\|_0 \le m$ and $\|\alpha\|_1 = 1$. The run-time comparison results were clear and persuasive, for a random variable with support size n = 10 and m = 5, the LP algorithm run-time was 850 seconds, where the KolmogorovApprox algorithm run-time was less than a tenth of a second. For n = 100 and m = 5, the KolmogorovApprox algorithm run-time was 0.14 seconds and the LP algorithm took more than a day. Due to these timing results of the LP algorithm we did not proceed to examine it any further. Since it is not trivial to formally analyze the run-time of the LP algorithm, we conclude by the reported experiment that in this case the LP algorithm might not be as efficient as KolmogorovApprox algorithm whose complexity is proven to be polynomial in Theorem 11.

59 5 Discussion

260 References

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