
Kolmogorov Approximation

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1 Introduction

Many different approaches to approximation of probability distributions are studied in the literature [9, 12, 13]. The papers vary in the types random variables involved, how they are represented, and in the criteria used for evaluation of the quality of the approximations. This paper is on approximating discrete distributions represented as explicit probability mass functions with ones that are simpler to store and to manipulate. This is needed, for example, when a discrete distribution is given as a large data-set, obtained, e.g., by sampling, and we want to represent it approximately with a small table.

The main contribution of this paper is an efficient algorithm for computing the best possible approximation of a given random variable with a random variable whose complexity is not above a prescribed threshold, where the measures of the quality of the approximation and the complexity of the random variable are as specified in the following two paragraphs.

We measure the quality of an approximation by the distance between the original variable and the approximate one. Specifically, we use the Kolmogorov distance which is one of the most used in statistical practice and literature. Given two random variables X and X' whose cumulative distribution functions (cdfs) are F_X and $F_{X'}$, respectively, the Kolmogorov distance between X and X' is $d_K(X, X') = \sup_t |F_X(t) - F_{X'}(t)|$ (see, e.g., [8]). We say that X' is a good approximation of X if $d_K(X, X')$ is small.

The complexity of a random variable is measured by the size of its support, the number of values that it can take, $|\text{support}(X)| = |\{x: \Pr(X = x) \neq 0\}|$. When distributions are maintained as explicit tables, as done in many implementations of statistical software, the size of the support of a variable is proportional to the amount of memory needed to store it and to the complexity of the computations around it.

In summary, the exact notion of optimality of the approximation targeted in this paper is:

Definition 1. A random variable X' is an optimal m -approximation of a random variable X if $|\text{support}(X')| \leq m$ and there is no random variable X'' such that $|\text{support}(X'')| \leq m$ and $d_K(X, X'') < d_K(X, X')$.

The main contribution of the paper is an efficient algorithm that takes X and m as parameters and constructs an optimal m -approximation of X .

The rest of the paper is organized as follows. In Section 2 we describe how our work relates to other algorithms and problems studied in the literature. In Section 3 we detail the proposed algorithm,

31 analyze its properties, and prove Theorem ?? . In Section 4 we demonstrate how the proposed
 32 approach performs on the problem of estimating the probability of hitting deadlines is plans and
 33 compare it to alternatives approximation approaches from the literature. We also demonstrate the
 34 performance of our approximation algorithm on some randomly generated random variables. The
 35 paper is concluded with a discussion in Section 5.

36 2 Related Work

37 The problem studied in this paper is related to the theory of Sparse Approximation (aka Sparse
 38 Representation) that deals with sparse solutions for systems of linear equations, as follows.

Given a matrix $D \in \mathbb{R}^{n \times p}$ and a vector $x \in \mathbb{R}^n$, the most studied sparse representation problem is finding the sparsest possible representation $\alpha \in \mathbb{R}^p$ satisfying $x = D\alpha$:

$$\min_{\alpha \in \mathbb{R}^p} \|\alpha\|_0 \text{ subject to } x = D\alpha.$$

39 where $\|\alpha\|_0 = |\{i : \alpha_i \neq 0, i = 1, \dots, p\}|$ is the ℓ_0 pseudo-norm, counting the number of non-zero
 40 coordinates of α . This problem is known to be NP-Hard with a reduction to NP-complete subset
 41 selection problems.

In these terms, using also the ℓ_∞ norm that represents the maximal coordinate and the ℓ_1 norm that represents the sum of the coordinates, our problem can be phrased as:

$$\min_{\alpha \in [0, \infty)^p} \|x - D\alpha\|_\infty \text{ subject to } \|\alpha\|_0 = m \text{ and } \|\alpha\|_1 = 1.$$

42 where D is the all-ones triangular matrix (the entry at row i and column j is one if $i \leq j$ and zero
 43 otherwise), x is related to X such that the i th coordinate of x is $F_X(x_i)$ where $\text{support}(X) = \{x_1 <$
 44 $x_2 < \dots < x_n\}$ and α is related to X' such that the i th coordinate of α is $f_{X'}(x_i)$. The functions F_X
 45 and $f_{X'}$ represent, respectively, the cumulative distribution function of X and the mass distribution
 46 function of X' . This, of course, means that the coordinates of x are assumed to be positive and
 47 monotonically increasing and that the last coordinate of x is assumed to be one. We demonstrate an
 48 application for this specific sparse representation problem and show that it can be solve in $O(n^2m)$
 49 time and $O(m^2)$ memory.

50 3 An Algorithm for Optimal Approximation

51 We next describe in details the proof of theorem ??.

52 In the following we set X as a random variable with a finite support of size n , and we set $0 < m \leq n$.
 53 We need to find an m -optimal approximation random variable X' .

54 Our first step is to show that it is enough to limit our search to X' 's such that $\text{support}(X') \subseteq$
 55 $\text{support}(X)$.

56 **Lemma 2.** *There is an m -optimal-approximation X' of X such that $\text{support}(X') \subseteq \text{support}(X)$.*

57 [DF: This proof is unclear to me, please clean.] Assume for contradiction is a random variable X''
 58 with support size m such that $d_K(X, X'')$ is minimal but $\text{support}(X'') \not\subseteq \text{support}(X)$. We will
 59 show how to transform X'' support such that it will be contained in $\text{support}(X)$. Let v' be the first
 60 $v' \in \text{support}(X'')$ and $v' \notin \text{support}(X)$. Let $v = \max\{i : i < v' \wedge i \in \text{support}(X)\}$. Every v'
 61 we will replace with v and name the new random variable X' , we will show that $d_K(X, X') =$

62 $d_K(X, X')$. First, note that: $F_{X''}(v) = F_{X'}(v)$, $F_X(v) = F_X(v)$. Second, $F_{X'}(v) - F_X(v) =$
63 $F_{X'}(v) - F_X(v)$. Therefore, $d_K(X, X'') = d_K(X, X')$ and X' is also an optimal approximation of
64 X . \square

65 Next, note that every random variable X'' with support of size at most m that is contained in
66 $\text{support}(X)$ be described by first setting the (at most m) elements of the support of X'' ; then for
67 every such option, determine X'' by setting probability values for the elements in the chosen support
68 of X' , and setting 0 for rest of the elements.

69 Since from Lemma 2 we can assume wlog that if X' is an m -optimal approximation variable for
70 X then $\text{support}(X') \subseteq \text{support}(X)$, our search to find such X' takes two steps. Denote the set of
71 random variables with support S by \mathbb{X}_S . In step 1, we find the m -optimal approximation random
72 variable among all random variables in \mathbb{X}_S , and denote the m -optimal distance for \mathbb{X}_S by $\varepsilon(X, S)$.
73 Next, in Step 2, among all the possible supports we find the support setting S of size $\leq m$ for which
74 $\varepsilon(X, S)$ is minimal: We describe an efficient way to do so.

75 3.1 Step 1

76 We first fix a set $S \subseteq \text{support}(X)$ of size at most m , and among all the random variables in \mathbb{X}_S
77 find one with a minimal distance from X . To that, set $S = \{x_1 < \dots < x_m\} \subseteq \text{support}(X)$. To
78 simplify the proofs set $x_0 = -\infty$, and $x_{m+1} = \infty$. Then $x_0 < x_1$ and $x_m < x_{m+1}$. In addition
79 recall that for every random variable X'' $F_{X''}(-\infty) = 0$ and $F_{X''}(\infty) = 1$. For the rest of this
80 section we assume S is fixed and therefore is not necessarily included in the notation.

81 Next, as the elements of S are also elements of $\text{support}(X)$, we can define the following weight
82 function that we use to find the m -optimal distance $\varepsilon(X, S)$.

83 **Definition 3.** For $0 \leq i < m$ let

$$w(x_i, x_{i+1}) = \begin{cases} P(x_i < X < x_{i+1}) & \text{if } i = 0 \text{ or } i = m; \\ P(x_i < X < x_{i+1})/2 & \text{otherwise.} \end{cases}$$

84 Note that when $i = 0$ (resp. $i = m + 1$) then $x_i = -\infty$ (resp. $x_i = \infty$).

85 Finally define:

$$\varepsilon(X, S) = \max_{i=0, \dots, m} w(x_i, x_{i+1}) \quad (1)$$

86 We first show that $\varepsilon(X, S)$ is a lower bound. That is, every random variable in \mathbb{X}_S has a distance at
87 least $\varepsilon(X, S)$. Then, we present a random variable $X' \in \mathbb{X}_S$ with distance $\varepsilon(X, S)$. It then follows
88 that such X' is an m -optimal approximation random variable among all random variables in \mathbb{X}_S .

89 The intuition behind choosing these specific weights and $\varepsilon(X, S)$ being a lower bound is as follows.
90 For every $1 \leq i \leq m$ let \hat{x}_i be the maximal element of $\text{support}(X)$ that is smaller than x_i . Then
91 since for every $X' \in \mathbb{X}_S$ the probability values of X' for the elements not in S are set to 0, we
92 have that $F_{X'}(\hat{x}_{i+1}) = F_{X'}(x_i)$. Therefore the distance between X' and X at points x_i and \hat{x}_{i+1} is
93 increased by $F_X(\hat{x}_{i+1}) - F_X(x_i) = P(x_i < X < x_{i+1})$.

94 Formally we have the following.

95 **Proposition 4.** For every random variable X' with $\text{support}(X') = S$ we have $d_K(X, X') \geq$
96 $\varepsilon(X, S)$.

97 *Proof.* Let X' be a random variable with support S . Then by definition, for every $0 \leq i \leq m$,
 98 $d_k(X, X') \geq \max\{|F_X(x_i) - F_{X'}(x_i)|, |F_X(\hat{x}_{i+1}) - F_{X'}(\hat{x}_{i+1})|\}$. Note that $F_{X'}(x_i) = F_{X'}(\hat{x}_{i+1})$
 99 since the probability value for all the elements not in S is set to 0.

100 If $i = 0$, that is $x_i = -\infty$, we have that $F_X(x_i) = F_{X'}(x_i) = F_{X'}(\hat{x}_{i+1}) = 0$ and therefore
 101 $d_k(X, X') \geq |F_X(\hat{x}_{i+1}) - F_{X'}(\hat{x}_{i+1})| = P(x_i < X < x_{i+1}) = w(x_i, x_{i+1})$.

102 If $i = m$, that is $x_{i+1} = \infty$, note that $F_X(\hat{x}_{i+1}) = F_{X'}(\hat{x}_{i+1}) = 1$. Therefore $F_{X'}(x_i) = 1$ as well.
 103 Therefore $d_k(X, X') \geq |F_X(\hat{x}_i) - F_{X'}(\hat{x}_i)| = P(x_i < X < x_{i+1}) = w(x_i, x_{i+1})$.
 104 [[DF: fix]]

105 Otherwise for every $1 \leq i < m$, we use the fact that $\max\{|a|, |b|\} \geq |a - b|/2$ for every $a, b \in \mathfrak{R}$, to
 106 have $d_k(X, X') \geq \max\{|F_X(x_i) - F_{X'}(x_i)|, |F_X(\hat{x}_{i+1}) - F_{X'}(\hat{x}_{i+1})|\}$, and therefore $d_k(X, X') \geq$
 107 $1/2|F_X(x_i) - F_X(\hat{x}_{i+1}) + F_{X'}(\hat{x}_{i+1}) - F_{X'}(x_i)|$. Since it is given that $F_{X'}(\hat{x}_{i+1}) - F_{X'}(x_i) =$
 108 $P(x_i < X' < x_{i+1}) = 0$, we have that $d_k(X, X') \geq 1/2|F_X(x_i) - F_X(\hat{x}_{i+1})| = P(x_1 < X <$
 109 $x_2)/2 = w(x_i, x_{i+1})$.

110 We saw that $d_k(X, X') \geq w(x_i, x_{i+1})$ for every $0 \leq i \leq m$. Therefore by definition of $\varepsilon(X, S)$,
 111 proof follows. \square

112 [[DF: here I stopped]]

113 Let X' be defined by $f_{X'}(x_i) = w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)$ for $i = 1, \dots, m$ and $f_{X'}(x) =$
 114 0 for $x \notin S$.

115 **Lemma 5.** For $i > 1$, if $F_{X'}(x_i) - F_X(x_i) = w(x_i, x_{i+1})$ then $F_{X'}(x_{i+1}) - F_X(x_{i+1}) =$
 116 $w(x_{i+1}, x_{i+2})$.

Proof.

$$F_{X'}(x_{i+1}) - F_X(x_{i+1}) = \quad (2)$$

$$\begin{aligned} &= f_{X'}(x_{i+1}) - f_X(x_{i+1}) - P(X < x_{i+1}) + P(X' < x_{i+1}) \\ &= f_{X'}(x_{i+1}) - f_X(x_{i+1}) - F_X(x_i) - P(x_i < X < x_{i+1}) + F_{X'}(x_i) \end{aligned}$$

$$= f_{X'}(x_{i+1}) - f_X(x_{i+1}) - F_X(x_i) - 2w(x_i, x_{i+1}) + F_{X'}(x_i) \quad (3)$$

$$= f_{X'}(x_{i+1}) - f_X(x_{i+1}) - 2w(x_i, x_{i+1}) + w(x_i, x_{i+1}) \quad (4)$$

$$= w(x_i, x_{i+1}) + w(x_{i+1}, x_{i+2}) - 2w(x_i, x_{i+1}) + w(x_i, x_{i+1}) \quad (5)$$

$$= w(x_{i+1}, x_{i+2})$$

117 By Definition 3 the probability $P(x_{i-1} < X < x_i) = 2w(x_{i-1}, x_i)$ as in Equation (3). Equation (4)
 118 is deduced by the induction hypothesis and Equation (5) where $f_{X'}(x_i) - f_X(x_i) = w(x_{i-1}, x_i) +$
 119 $w(x_i, x_{i+1})$ is true by construction, see Definition?? \square

120 **Lemma 6.** Base case: $i = 1, F_{X'}(x_1) - F_X(x_1) = w(x_1, x_2)$.

Proof.

$$\begin{aligned} F_{X'}(x_1) - F_X(x_1) &= \\ &= f_{X'}(x_1) - f_X(x_1) - w(x_0, x_1) \\ &= w(x_0, x_1) + w(x_1, x_2) - w(x_0, x_1) \\ &= w(x_1, x_2) \end{aligned}$$

122 **Proposition 7.** *There exists X' such that $\text{support}(X') = S$ and $d_k(X, X') = \varepsilon(X, S)$.*

123 3.2 Step 2

124 Chakravarty, Orlin, and Rothblum [2] proposed a polynomial-time method that, given a certain
 125 objective functions (additive), finds an optimal consecutive partition. Their method involves the
 126 construction of a graph such that the (consecutive) set partitioning problem is reduced to the problem
 127 of finding the shortest path in that graph.

128 The KolmogorovApprox algorithm (Algorithm 1) starts by constructing a directed weighted graph
 129 G similar to the method of Chakravarty, Orlin, and Rothblum [2]. The nodes V consist of the
 130 support of X together with an extra two nodes, $-\infty$ and ∞ for technical reasons, whereas the
 131 edges E connect every pair of nodes in one direction (lines 1-2). The weight w of each edge
 132 $e = (x, y) \in E$ is determined by one of two cases as in Definition 3. The first is where nodes
 133 x or y are the source or target nodes respectively. In this case, the weight is the probability of X
 134 to get a value between x and y , non inclusive, i.e., $w(e) = \Pr(x < X < y)$. The second case
 135 is where x and y are not a source or target nodes, here the weight is the probability of X to get a
 136 value between x and y , non inclusive, divided by two i.e., $w(e) = \Pr(x < X < y)/2$. The values
 137 taken are non inclusive, since we are interested only in the error value. The source node of the
 138 shortest path problem at hand corresponds to the $-\infty$ node added to G in the construction phase,
 139 and the target node is the extra node ∞ . The set of all solution paths in G , i.e., those starting at
 140 $-\infty$ and ending in ∞ with at most m edges, is called $\text{paths}(G, -\infty, \infty)$. The goal is to find the
 141 path l in $\text{paths}(G, -\infty, \infty)$ with the lightest bottleneck (line 3). This can be achieved by using the
 142 *Bellman – Ford* algorithm with two tweaks. The first is to iterate the graph G in order to find only
 143 paths with length of at most m edges. The second is to find the lightest bottleneck as opposed to
 144 the traditional objective of finding the shortest path. This is performed by modifying the manner of
 145 “relaxation” to $\text{bottleneck}(x) = \min[\max(\text{bottleneck}(v), w(e))]$, done also in [14]. Consequently,
 146 we find the lightest maximal edge in a path of length $\leq m$, which represents the minimal error,
 147 $\varepsilon(X, S)$, defined in Definition ?? where the nodes in path l represent the elements in set S . The
 148 approximated random variable X' is then derived from the resulting path l (lines 4-5). Every node
 149 $x \in l$ represent a value in the new calculated random variable X' , we than iterate the path l to find
 150 the probability of the event $f_{X'}(x)$ as described in Definition ???. For every edge $(x_i, x_j) \in l$ we
 151 determine: if (x_i, x_j) is the first edge in the path l (i.e. $x_i = -\infty$), then node x_j gets the full weight
 152 $w(x_i, x_j)$ and it's own weight in X such that $f_{X'}(x_j) = f_X(x_j) + w(x_i, x_j)$. If (x_i, x_j) in not the
 153 first nor the last edge in path l then we divide it's weight between nodes x_i and x_j in addition to their
 154 own original weight in X and the probability that already accumulated. If (x_i, x_j) is the last edge
 155 in the path l (i.e. $i = \infty$) then node i gets the full weight $w(x_i, x_j)$ in addition to what was already
 156 accumulated such that $f_{X'}(x_j) = f_{X'}(x_j) + w(x_i, x_j)$.

Algorithm 1: KolmogorovApprox(X, m)

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1  $S = \text{support}(X) \cup \{\infty, -\infty\}$ 
2  $G = (V, E) = (S, \{(x, y) : x < y\})$ 
3  $(x_0, \dots, x_{m+1}) = l = \text{argmin}_{l \in \text{paths}(G, -\infty, \infty), |l| \leq m} \max\{w(e) : e \in l\}$ 
4 for  $0 < i < m + 1$  do
5    $\lfloor f_{X'}(x_i) = w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)$ 
6 return  $X'$ 

```

157 **Theorem 8.** $\text{KolmogorovApprox}(X, m)$ is an m -optimal-approximation of X .

158 **Theorem 9.** The $\text{KolmogorovApprox}(X, m)$ algorithm runs in time $O(mn^2)$, using $O(n^2)$ memory
 159 where $n = |\text{support}(X)|$.

160 *Proof.* Constructing the graph G takes $O(n^2)$. The number of edges is $O(E) \approx O(n^2)$ and for every
 161 edge the weight is at most the sum of all probabilities between the source node $-\infty$ and the target
 162 node ∞ , which can be done efficiently by aggregating the weights of already calculated edges. The
 163 construction is also the only stage that requires memory allocation, specifically $O(E + V) = O(n^2)$.
 164 Finding the shortest path takes $O(m(E + V)) \approx O(mn^2)$. Since G is DAG (directed acyclic graph)
 165 finding shortest path takes $O(E + V)$. We only need to find paths of length $\leq m$, which takes
 166 $O(m(E + V))$. Deriving the new random variable X' from the computed path l takes $O(mn)$. For
 167 every node in l (at most m nodes), calculating the probability $P(s < X < \infty)$ takes at most n .
 168 To conclude, the worst case run-time complexity is $O(n^2 + mn^2 + mn) = O(mn^2)$ and memory
 169 complexity is $O(E + V) = O(n^2)$. \square

170 4 A case study and experimental results

171 The case study examined in our experiments is the problem of task trees with deadlines [4, 3].
 172 Hierarchical planning is a well-established field in AI [5, 6, 7], and is still relevant nowadays [1, 15].
 173 A hierarchical plan is a method for representing problems of automated planning in which the
 174 dependency among tasks can be given in the form of networks, here we focus on hierarchical plans
 175 represented by task trees. The leaves in a task tree are *primitive* actions (or tasks), and the internal
 176 nodes are either *sequence* or *parallel* actions. The plans we deal with are of stochastic nature, where
 177 the duration of a primitive action is given by a random variable.

178 A sequence node denotes a series of tasks that should be performed consecutively, whereas a parallel
 179 node denotes a set of tasks that begin at the same time. A *valid* plan is one that is fulfilled before some
 180 given *deadline*, i.e., its *makespan* is less than or equal to the deadline. The objective in this context
 181 is to compute the probability that a given plan is valid, or more formally computing $P(X < T)$,
 182 where X is a random variable representing the makespan of the plan and T is the deadline. As said
 183 above, resource consumption (task duration) is uncertain, and described as probability distributions
 184 in the leaf nodes. We assume that the distributions are independent but *not* necessarily identically
 185 distributed and that the random variables are discrete and have a finite support.

186 The problem of finding the probability that a task tree satisfies a deadline is known to be NP-hard. In
 187 fact, even the problem of summing a set of random variables is NP-hard [10]. This is an example of
 188 an explicitly given random variable that we need to estimate deadline meeting probabilities for.

189 In the first experiment we focus on is the problem of task trees with deadlines, and consider three
 190 types of task trees. The first type includes logistic problems of transporting packages by trucks and
 191 airplanes (from IPC2 <http://ipc.icaps-conference.org/>). Hierarchical plans of those logistic problems
 192 were generated by the JSHOP2 planner [11] (see example problem, Figure 1, one parallel node with
 193 all descendant task nodes being in sequence). The second type consists of task trees used as execution
 194 plans for the ROBIL team entry in the DARPA robotics challenge (DRC simulation phase), and the
 195 third type is of linear plans (sequential task trees). The primitive tasks in all the trees are modeled as
 196 discrete random variables with support of size M obtained by discretization of uniform distributions
 197 over various intervals. The number of tasks in a tree is denoted by N .

198 We implemented the approximation algorithm for solving the deadline problem with four different
 199 methods of approximation. The first two are for achieving a one-sided Kolmogorov approximation –

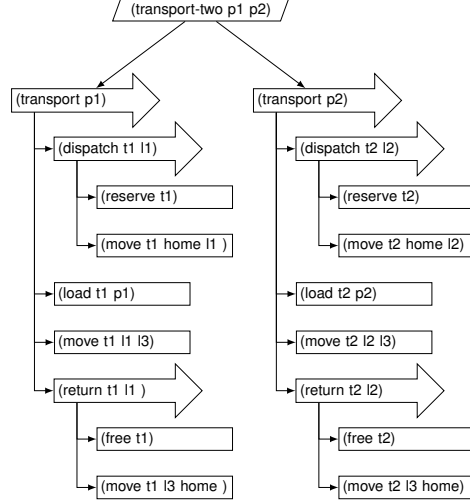


Figure 1: A plan generated by the JSHOP2 algorithm. Arrow shapes represent sequence nodes, parallelograms represent parallel nodes, and rectangles represent primitive nodes.

the OptTrim [3] and the Trim [4] operators, and the third is a simple sampling scheme. We used those methods as a comparison to the Kolmogorov approximation with the suggested KolmogorovApprox algorithm. The parameter m of OptTrim and KolmogorovApprox corresponds to the inverse of ε given to the Trim operator. Note that in order to obtain some error ε , one must take into consideration the size of the task tree N , therefore, $m/N = 1/(\varepsilon \cdot N)$. We ran also an exact computation as a reference to the approximated one in order to calculate the error. The experiments conducted with the following operators and their parameters: KolmogorovApprox operator with $m = 10 \cdot N$, the OptTrim operator with $m = 10 \cdot N$, the Trim as operator with $\varepsilon = 0.1/N$, and two simple simulations, with a different samples number $s = 10^4$ and $s = 10^6$.

Task Tree	M	KolmogorovApprox	OptTrim	Trim	Sampling	
		$m/N=10$	$m/N=10$	$\varepsilon \cdot N=0.1$	$s=10^4$	$s=10^6$
Logistics ($N=34$)	2	0	0	0.0019	0.007	0.0009
	4	0.0024	0.0046	0.0068	0.0057	0.0005
Logistics ($N=45$)	2	0.0002	0.0005	0.002	0.015	0.001
	4	0	0.003	0.004	0.008	0.0006
DRC-Drive ($N=47$)	2	0	0.004	0.009	0.0072	0.0009
	4	0.001	0.008	0.019	0.0075	0.0011
Sequential ($N=10$)	2	0.0093	0.015	0.024	0	0
	4	0	0.024	0.04	0.008	0.0016
	10	0	0.028	0.06	0.0117	0.001

Table 1: Comparison of estimated errors with respect to the reference exact computation on various task trees.

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Table 1 shows the results of the main experiment. The quality of the solutions provided by using the OptTrim operator are better (lower errors) than those provided by the Trim operator, following the optimality guarantees, but is interesting to see that the quality gaps happen in practice in each of the examined task trees. However, in some of the task trees the sampling method produced better results than the approximation algorithm with OptTrim. Nevertheless, the approximation algorithm comes with an inherent advantage of providing an exact quality guarantees, as opposed to the probabilistic guarantees provided by sampling.

216 In order to better understand the quality gaps in practice between KolmogorovApprox, OptTrim,
 217 and Trim, we investigate their relative errors when applied on single random variables with support
 218 size $n = 100$, and different support sizes of the resulting random variable approximation (m). In each
 219 instance of this experiment, a random variable is randomly generated by choosing the probabilities of
 220 each element in the support from a uniform distribution and then normalizing these probabilities so
 221 that they sum to one.

222 Figure 2 present the error produced by the above methods. The depicted results are averages over
 223 several instances (50 instances) of random variables. The curves in the figure show the average error
 224 of OptTrim and Trim operators with comparison to the average error of the optimal approximation
 225 provided by KolmogorovApprox as a function of m .

226 According to the depicted results it is evident that increasing the support size of the approxima-
 227 tion m reduces the error, as expected, in all three methods. However, errors produced by the
 228 KolmogorovApprox are significantly smaller, safe to say, a half of the error produced by OptTrim
 229 and Trim, it is clear both in the table (the relative error is mostly above 1) and in the graph.

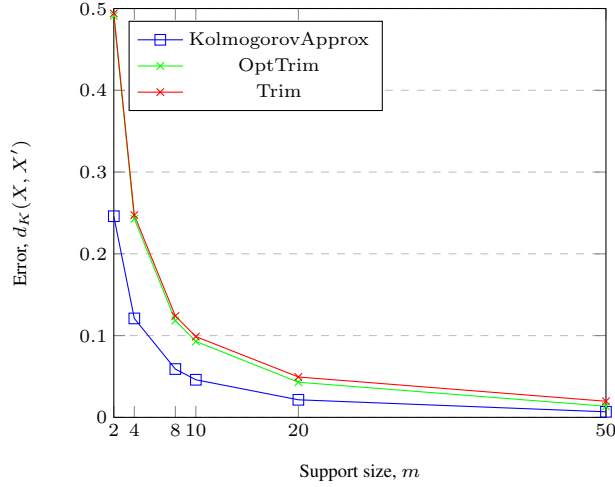


Figure 2: Error comparison between KolmogorovApprox, OptTrim, and Trim, on randomly generated random variables as function of m .

230 We also examined how our algorithm compares to linear programming as described and discussed, for
 231 example, in [12]. We ran an experiment to compare the run-time between the KolmogorovApprox
 232 algorithm with the run-time of a state-of-art implementation of linear programming. We used the
 233 “Minimize” function of Wolfram Mathematica and fed it with the equations $\min_{\alpha \in \mathbb{R}^n} \|x - \alpha\|_\infty$
 234 subject to $\|\alpha\|_0 \leq m$ and $\|\alpha\|_1 = 1$. The run-time comparison results were clear and persuasive,
 235 for a random variable with support size $n = 10$ and $m = 5$, the LP algorithm run-time was 850
 236 seconds, where the KolmogorovApprox algorithm run-time was less than a tenth of a second. For
 237 $n = 100$ and $m = 5$, the KolmogorovApprox algorithm run-time was 0.14 seconds and the LP
 238 algorithm took more than a day. Due to these timing results of the LP algorithm we did not proceed
 239 to examine it any further. Since it is not trivial to formally analyze the run-time of the LP algorithm,
 240 we conclude by the reported experiment that in this case the LP algorithm might not be as efficient as
 241 KolmogorovApprox algorithm whose complexity is proven to be polynomial in Theorem 9.

5 Discussion

References

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