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# Kolmogorov Approximation

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## 1 Introduction

Many different approaches to approximation of probability distributions are studied in the literature [9, 12, 13]. The papers vary in the types random variables involved, how they are represented, and in the criteria used for evaluation of the quality of the approximations. This paper is on approximating discrete distributions represented as explicit probability mass functions with ones that are simpler to store and to manipulate. This is needed, for example, when a discrete distribution is given as a large data-set, obtained, e.g., by sampling, and we want to represent it approximately with a small table.

The main contribution of this paper is an efficient algorithm for computing the best possible approximation of a given random variable with a random variable whose complexity is not above a prescribed threshold, where the measures of the quality of the approximation and the complexity of the random variable are as specified in the following two paragraphs.

We measure the quality of an approximation by the distance between the original variable and the approximate one. Specifically, we use the Kolmogorov distance which is one of the most used in statistical practice and literature. Given two random variables  $X$  and  $X'$  whose cumulative distribution functions (cdfs) are  $F_X$  and  $F_{X'}$ , respectively, the Kolmogorov distance between  $X$  and  $X'$  is  $d_K(X, X') = \sup_t |F_X(t) - F_{X'}(t)|$  (see, e.g., [8]). We say that  $X'$  is a good approximation of  $X$  if  $d_K(X, X')$  is small.

The complexity of a random variable is measured by the size of its support, the number of values that it can take,  $|\text{support}(X)| = |\{x: \Pr(X = x) \neq 0\}|$ . When distributions are maintained as explicit tables, as done in many implementations of statistical software, the size of the support of a variable is proportional to the amount of memory needed to store it and to the complexity of the computations around it.

In summary, the exact notion of optimality of the approximation targeted in this paper is:

**Definition 1.** A random variable  $X'$  is an optimal  $m$ -approximation of a random variable  $X$  if  $|\text{support}(X')| \leq m$  and there is no random variable  $X''$  such that  $|\text{support}(X'')| \leq m$  and  $d_K(X, X'') < d_K(X, X')$ .

The main contribution of the paper is an efficient algorithm that takes  $X$  and  $m$  as parameters and constructs an optimal  $m$ -approximation of  $X$ .

The rest of the paper is organized as follows. In Section 2 we describe how our work relates to other algorithms and problems studied in the literature. In Section 3 we detail the proposed algorithm,

analyze its properties, and prove Theorem ???. In Section 4 we demonstrate how the proposed approach performs on the problem of estimating the probability of hitting deadlines is plans and compare it to alternatives approximation approaches from the literature. We also demonstrate the performance of our approximation algorithm on some randomly generated random variables. The paper is concluded with a discussion in Section 5.

## 2 Related Work

The problem studied in this paper is related to the theory of Sparse Approximation (aka Sparse Representation) that deals with sparse solutions for systems of linear equations, as follows.

Given a matrix  $D \in \mathbb{R}^{n \times p}$  and a vector  $x \in \mathbb{R}^n$ , the most studied sparse representation problem is finding the sparsest possible representation  $\alpha \in \mathbb{R}^p$  satisfying  $x = D\alpha$ :

$$\min_{\alpha \in \mathbb{R}^p} \|\alpha\|_0 \text{ subject to } x = D\alpha.$$

where  $\|\alpha\|_0 = |\{i : \alpha_i \neq 0, i = 1, \dots, p\}|$  is the  $\ell_0$  pseudo-norm, counting the number of non-zero coordinates of  $\alpha$ . This problem is known to be NP-Hard with a reduction to NP-complete subset selection problems.

In these terms, using also the  $\ell_\infty$  norm that represents the maximal coordinate and the  $\ell_1$  norm that represents the sum of the coordinates, our problem can be phrased as:

$$\min_{\alpha \in [0, \infty)^p} \|x - D\alpha\|_\infty \text{ subject to } \|\alpha\|_0 = m \text{ and } \|\alpha\|_1 = 1.$$

where  $D$  is the all-ones triangular matrix (the entry at row  $i$  and column  $j$  is one if  $i \leq j$  and zero otherwise),  $x$  is related to  $X$  such that the  $i$ th coordinate of  $x$  is  $F_X(x_i)$  where  $\text{support}(X) = \{x_1 < x_2 < \dots < x_n\}$  and  $\alpha$  is related to  $X'$  such that the  $i$ th coordinate of  $\alpha$  is  $f_{X'}(x_i)$ . The functions  $F_X$  and  $f_{X'}$  represent, respectively, the cumulative distribution function of  $X$  and the mass distribution function of  $X'$ . This, of course, means that the coordinates of  $x$  are assumed to be positive and monotonically increasing and that the last coordinate of  $x$  is assumed to be one. We demonstrate an application for this specific sparse representation problem and show that it can be solve in  $O(n^2m)$  time and  $O(m^2)$  memory.

## 3 An Algorithm for Optimal Approximation

Let, in the scope of this section,  $X$  be a given random variable with a finite support of size  $n$ , and let  $0 < m \leq n$  be a given complexity bound. We will now develop notations and collect facts towards an algorithm for finding an optimal  $m$ -approximation of  $X$ .

The first useful fact is that it is enough to limit our search to approximations  $X'$ 's such that  $\text{support}(X') \subseteq \text{support}(X)$ :

**Lemma 2.** *There is an optimal  $m$ -approximation  $X'$  of  $X$  such that  $\text{support}(X') \subseteq \text{support}(X)$ .*

*Proof.* Since minimizing  $d_K(X, X')$  subject to the constraints that  $\text{support}(X') \subseteq \text{support}(X)$  and  $|\text{support}(X')| \leq m$  is a convex optimization problem, it has a solution  $\varepsilon$ . Let  $X''$  be any random variable whose support is of size smaller or equal to  $m$ , let  $\{x_1, \dots, x_n\} = \text{support}(X)$ , and let  $x_{n+1} = \infty$ . Consider the the random variable  $X'$  whose probability mass function is  $f_{X'}(x_i) = P(x_i \leq X'' < x_{i+1})$  for  $i = 1, \dots, n$  and  $F_{X'}(x) = 0$  if  $x \notin \text{support}(X)$ . By definition  $d_K(X, X') = d_K(X, X'')$ . Since  $\text{support}(X') \subseteq \text{support}(X)$  and  $|\text{support}(X')| <$

63  $|\text{support}(X'')| \leq m$ , we get that  $d_K(X, X'') \geq \varepsilon$ . Thus,  $\varepsilon$  is a global minimum and any  
 64  $X' \in \text{argmin}\{d_K(X, X') : \text{support}(X') \subseteq \text{support}(X), |\text{support}(X')| \leq m\}$  is an  $m$ -optimal  
 65 approximation of  $X$ .  $\square$

66 Next, note that every random variable  $X''$  with support of size at most  $m$  that is contained in  
 67  $\text{support}(X)$  be described by first setting the (at most  $m$ ) elements of the support of  $X''$ ; then for  
 68 every such option, determine  $X''$  by setting probability values for the elements in the chosen support  
 69 of  $X'$ , and setting 0 for rest of the elements.

70 Since from Lemma 2 we can assume without loss of generality that if  $X'$  is an optimal  $m$ -  
 71 approximation variable for  $X$  then  $\text{support}(X') \subseteq \text{support}(X)$ , our search to find such  $X'$  takes  
 72 two steps. Denote the set of random variables with support  $S$  by  $\mathbb{X}_S$ . In step 1, we find the optimal  
 73  $m$ -approximation random variable among all random variables in  $\mathbb{X}_S$ , and denote the  $m$ -optimal  
 74 distance for  $\mathbb{X}_S$  by  $\varepsilon(X, S)$ . Next, in Step 2, among all the possible supports we find the support  
 75 setting  $S$  of size  $\leq m$  for which  $\varepsilon(X, S)$  is minimal: We describe an efficient way to do so.

### 76 3.1 Step 1

77 We first fix a set  $S \subseteq \text{support}(X)$  of size at most  $m$ , and among all the random variables in  $\mathbb{X}_S$   
 78 find one with a minimal distance from  $X$ . To that, set  $S = \{x_1 < \dots < x_m\} \subseteq \text{support}(X)$ . To  
 79 simplify the proofs set  $x_0 = -\infty$ , and  $x_{m+1} = \infty$ . Then  $x_0 < x_1$  and  $x_m < x_{m+1}$ . In addition  
 80 recall that for every random variable  $X''$ ,  $F_{X''}(-\infty) = 0$  and  $F_{X''}(\infty) = 1$ . Finally, for every  
 81  $1 \leq i \leq m$  let  $\hat{x}_i$  be the maximal element of  $\text{support}(X)$  that is smaller than  $x_i$ . For the rest of this  
 82 section we assume  $S$  is fixed and therefore is not necessarily included in the notation.

83 Next, as the elements of  $S$  are also elements of  $\text{support}(X)$ , we can define the following weight  
 84 function that we use to find the  $m$ -optimal distance  $\varepsilon(X, S)$ .

85 **Definition 3.** For  $0 \leq i \leq m$  let

$$w(x_i, x_{i+1}) = \begin{cases} P(x_i < X < x_{i+1}) & \text{if } i = 0 \text{ or } i = m; \\ P(x_i < X < x_{i+1})/2 & \text{otherwise.} \end{cases}$$

86 Note that  $x_i = -\infty$  for  $i = 0$  and  $x_i = \infty$  for  $i = m + 1$ . Also note that  $P(x_i < X < x_{i+1}) =$   
 87  $F_X(\hat{x}_{i+1}) - F_X(x_i)$ , a fact that we will use throughout this section.

88 **Definition 4.** Let  $\varepsilon(X, S) = \max_{i=0, \dots, m} w(x_i, x_{i+1})$ .

89 We first show that  $\varepsilon(X, S)$  is a lower bound. That is, every random variable in  $\mathbb{X}_S$  has a distance at  
 90 least  $\varepsilon(X, S)$ . Then, we present a random variable  $X' \in \mathbb{X}_S$  with distance  $\varepsilon(X, S)$ . It then follows  
 91 that such  $X'$  is an optimal  $m$ -approximation random variable among all random variables in  $\mathbb{X}_S$ .

92 The intuition behind choosing these specific weights and  $\varepsilon(X, S)$  being a lower bound is as follows.  
 93 Since for every  $X' \in \mathbb{X}_S$  the probability values of  $X'$  for the elements not in  $S$  are set to 0, we have  
 94 that  $F_{X'}(\hat{x}_{i+1}) = F_{X'}(x_i)$ . Therefore the distance between  $X'$  and  $X$  at points  $x_i$  and  $\hat{x}_{i+1}$  that we  
 95 have to take into additional account is increased by  $F_X(\hat{x}_{i+1}) - F_X(x_i) = P(x_i < X < x_{i+1})$ .

96 Formally we have the following.

97 **Proposition 5.** If  $X' \in \mathbb{X}_S$  then  $d_k(X, X') \geq \varepsilon(X, S)$ .

98 *Proof.* By definition, for every  $0 \leq i \leq m$ ,  $d_k(X, X') \geq \max\{|F_X(\hat{x}_{i+1}) - F_{X'}(\hat{x}_{i+1})|, |F_X(x_i) -$   
99  $F_{X'}(x_i)|\}$ . Note that  $F_{X'}(\hat{x}_{i+1}) = F_{X'}(x_i)$  since the probability values for all the elements not in  $S$   
100 are set to 0.

101 If  $i = 0$ , that is  $x_i = -\infty$ , we have that  $F_X(x_i) = F_{X'}(x_i) = F_{X'}(\hat{x}_{i+1}) = 0$  and therefore  
102  $d_k(X, X') \geq |F_X(\hat{x}_{i+1})| = |F_X(\hat{x}_{i+1}) - F_X(x_i)| = P(x_i < X < x_{i+1}) = w(x_i, x_{i+1})$ .

103 If  $i = m$ , that is  $x_{i+1} = \infty$ , we have that  $F_X(\hat{x}_{i+1}) = F_{X'}(\hat{x}_{i+1}) = F_{X'}(x_i) = 1$ . and therefore  
104  $d_k(X, X') \geq |1 - F_X(\hat{x}_i)| = |F_X(\hat{x}_{i+1}) - F_X(x_i)| = P(x_i < X < x_{i+1}) = w(x_i, x_{i+1})$ .

105 Otherwise for every  $1 \leq i < m$ , we use the fact that  $\max\{|a|, |b|\} \geq |a - b|/2$  for every  $a, b \in$   
106  $\mathbb{R}$ , to have  $d_k(X, X') \geq 1/2|F_X(\hat{x}_{i+1}) - F_X(x_i) + F_{X'}(x_i) - F_{X'}(\hat{x}_{i+1})|$ . So  $d_k(X, X') \geq$   
107  $1/2|F_X(\hat{x}_{i+1}) - F_X(x_i)| = P(x_1 < X < x_2)/2 = w(x_i, x_{i+1})$ .

108 Therefore since  $d_k(X, X') \geq w(x_i, x_{i+1})$  for every  $0 \leq i \leq m$ , by definition of  $\varepsilon(X, S)$  proof  
109 follows.  $\square$

110 Next we show a random variable  $X' \in \mathbb{X}_S$  with a distance of  $\varepsilon(X, S)$  from  $X$ . Thus  $X'$  is an optimal  
111  $m$ -approximation among the set  $\mathbb{X}_S$ . We define  $X'$  as follows:

112 **Definition 6.** Let  $f_{X'}(x_i) = w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)$  for  $i = 1, \dots, m$  and  $f_{X'}(x) = 0$   
113 for  $x \notin S$ .

114 We first show that  $X'$  is a properly defined random variable:

115 **Lemma 7.**  $f_{X'}$  is a probability mass function.

116 *Proof.* From definition  $f_{X'}(x_i) \geq 0$  for every  $i$ . To see that  $\sum_i f_{X'}(x_i) = 1$ , we have  $\sum_i f_{X'}(x_i) =$   
117  $\sum_i (w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)) = \sum_{x_i \in S} f_X(x_i) + w(x_0, x_1) + \sum_{0 < i < m} 2w(x_i, x_{i+1}) +$   
118  $w(x_m, x_{m+1}) = \sum_{x_i \in S} P(X = x_i) + P(x_0 < X < x_1) + \sum_{0 < i < m} P(x_i < X < x_{i+1}) +$   
119  $P(x_m < X < x_{m+1}) = 1$  since this sum is the entire cpt of  $X$ .  $\square$

120 Note that  $X'$  can be constructed in linear time to the size of the cdf of  $X$ . Intuitively the setting of  
121  $X'$  allows to take an "advantage" of distance from  $X$  at the elements of  $\text{support}(X')$ , to avoid the  
122 overall increased distance of  $X$  from  $X'$  at the elements that are not at  $\text{support}(X)$  and in which  
123  $f_{X'}$  is set to 0. Formally we have the following.

124 **Lemma 8.** Let  $x \in \text{support}(X)$  and  $0 \leq i \leq m$  be such that  $x_i \leq x \leq x_{i+1}$  then  $-w(x_i, x_{i+1}) \leq$   
125  $F_X(x) - F_{X'}(x) \leq w(x_i, x_{i+1})$ .

126 *Proof.* We prove by induction on  $0 \leq i < m$ .

127 First see that  $F_{X'}(j) = 0$  for every  $x_0 < j < x_1$  and therefore  $F_X(j) - F_{X'}(j) = F_X(j) - 0 \leq$   
128  $F_X(\hat{x}_1) = F_X(\hat{x}_1) - F_X(x_0) = w(x_0, x_1)$ . For  $j = x_1$  we have  $F_X(x_1) - F_{X'}(x_1) = F_X(\hat{x}_1) +$   
129  $f_X(x_1) - (w(x_0, x_1) + w(x_1, x_2) + f_X(x_1)) = w(x_0, x_1) + f_X(x_1) - (w(x_0, x_1) + w(x_1, x_2) +$   
130  $f_X(x_1)) = -w(x_1, x_2)$ .

131 Next assume that  $F_X(\hat{x}_i) - F_{X'}(\hat{x}_i) = w(x_{i-1}, x_i)$ . Then  $F_X(x_i) - F_{X'}(x_i) = F_X(\hat{x}_i) + f_X(x_i) -$   
132  $(w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)) = w(x_{i-1}, x_i) + f_X(x_i) - (w(x_{i-1}, x_i) + w(x_i, x_{i+1}) +$   
133  $f_X(x_i)) = -w(x_i, x_{i+1})$ .

134 As before we have that for all  $x_i < j < x_{i+1}$ , we have  $F_X(j) - F_{X'}(j) = F_X(j) - F_{X'}(\hat{x}_{i+1}) \leq$   
135  $F_X(\hat{x}_{i+1}) - F_{X'}(\hat{x}_{i+1})$ . Then  $F_X(\hat{x}_{i+1}) - F_{X'}(\hat{x}_{i+1}) = (F_X(x_i) + P(x_i < x < x_{i+1})) -$   
136  $F_{X'}(x_i) = -w(x_i, x_{i+1}) + 2w(x_i, x_{i+1}) = w(x_i, x_{i+1})$ .

137 Finally for  $x_m \leq j \leq x_{m+1}$  we have that  $F_{X'}(x_m) = 1$  therefore  $F_X(x_m) - F_{X'}(x_m) = (1 -$   
138  $P(x_m < X < x_{m+1})) - 1 = P(x_m < X < x_{m+1}) = w(x_m, x_{m+1})$ , and for every  $x_m < j <$   
139  $x_{m+1}$  we have  $F_X(j) - F_{X'}(j) < (1 - P(x_m < X < x_{m+1})) - 1 < -P(x_m < X < x_{m+1}) =$   
140  $-w(x_m, x_{m+1})$  as required.  $\square$

141 From Lemma 8, by the definition of  $\varepsilon(X, S)$ , we then have:

142 **Corollary 9.**  $d_k(X, X') = \varepsilon(X, S)$ .

### 143 3.2 Step 2

144 Chakravarty, Orlin, and Rothblum [2] proposed a polynomial-time method that, given a certain  
145 objective functions (additive), finds an optimal consecutive partition. Their method involves the  
146 construction of a graph such that the (consecutive) set partitioning problem is reduced to the problem  
147 of finding the shortest path in that graph.

148 The KolmogorovApprox algorithm (Algorithm 1) starts by constructing a directed weighted graph  
149  $G$  similar to the method of Chakravarty, Orlin, and Rothblum [2]. The nodes  $V$  consist of the  
150 support of  $X$  together with an extra two nodes,  $-\infty$  and  $\infty$  for technical reasons, whereas the  
151 edges  $E$  connect every pair of nodes in one direction (lines 1-2). The weight  $w$  of each edge  
152  $e = (x, y) \in E$  is determined by one of two cases as in Definition 3. The first is where nodes  
153  $x$  or  $y$  are the source or target nodes respectively. In this case, the weight is the probability of  $X$   
154 to get a value between  $x$  and  $y$ , non inclusive, i.e.,  $w(e) = Pr(x < X < y)$ . The second case  
155 is where  $x$  and  $y$  are not a source or target nodes, here the weight is the probability of  $X$  to get a  
156 value between  $x$  and  $y$ , non inclusive, divided by two i.e.,  $w(e) = Pr(x < X < y)/2$ . The values  
157 taken are non inclusive, since we are interested only in the error value. The source node of the  
158 shortest path problem at hand corresponds to the  $-\infty$  node added to  $G$  in the construction phase,  
159 and the target node is the extra node  $\infty$ . The set of all solution paths in  $G$ , i.e., those starting at  
160  $-\infty$  and ending in  $\infty$  with at most  $m$  edges, is called  $paths(G, -\infty, \infty)$ . The goal is to find the  
161 path  $l$  in  $paths(G, -\infty, \infty)$  with the lightest bottleneck (line 3). This can be achieved by using the  
162 *Bellman – Ford* algorithm with two tweaks. The first is to iterate the graph  $G$  in order to find only  
163 paths with length of at most  $m$  edges. The second is to find the lightest bottleneck as opposed to  
164 the traditional objective of finding the shortest path. This is performed by modifying the manner of  
165 “relaxation” to  $bottleneck(x) = \min[\max(bottleneck(v), w(e))]$ , done also in [14]. Consequently,  
166 we find the lightest maximal edge in a path of length  $\leq m$ , which represents the minimal error,  
167  $\varepsilon(X, S)$ , defined in Definition ?? where the nodes in path  $l$  represent the elements in set  $S$ . The  
168 approximated random variable  $X'$  is then derived from the resulting path  $l$  (lines 4-5). Every node  
169  $x \in l$  represent a value in the new calculated random variable  $X'$ , we than iterate the path  $l$  to find  
170 the probability of the event  $f_{X'}(x)$  as described in Definition 6. For every edge  $(x_i, x_j) \in l$  we  
171 determine: if  $(x_i, x_j)$  is the first edge in the path  $l$  (i.e.  $x_i = -\infty$ ), then node  $x_j$  gets the full weight  
172  $w(x_i, x_j)$  and it's own weight in  $X$  such that  $f_{X'}(x_j) = f_X(x_j) + w(x_i, x_j)$ . If  $(x_i, x_j)$  in not the  
173 first nor the last edge in path  $l$  then we divide it's weight between nodes  $x_i$  and  $x_j$  in addition to their  
174 own original weight in  $X$  and the probability that already accumulated. If  $(x_i, x_j)$  is the last edge  
175 in the path  $l$  (i.e.  $x_j = \infty$ ) then node  $x_i$  gets the full weight  $w(x_i, x_j)$  in addition to what was already  
176 accumulated such that  $f_{X'}(x_j) = f_{X'}(x_i) + w(x_i, x_j)$ .

177 **Theorem 10.**  $KolmogorovApprox(X, m)$  is an  $m$ -optimal-approximation of  $X$ .

178 *Proof.* If we consider the vertexes  $S = l \setminus \{-\infty, \infty\}$  for a path  $l \in paths(G, -\infty, \infty)$  we have  
179 that  $\max\{w(e) : e \in l\} = \varepsilon(X, S)$ . Therefore, line 3 of the algorithm essentially computes a set

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**Algorithm 1:** KolmogorovApprox( $X, m$ )

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1  $S = \text{support}(X) \cup \{\infty, -\infty\}$ 
2  $G = (V, E) = (S, \{(x, y) : x < y\})$ 
3  $(x_0, \dots, x_{m+1}) = l \in \text{argmin}_{l \in \text{paths}(G, -\infty, \infty), |l| \leq m} \max\{w(e) : e \in l\}$ 
4 for  $0 < i < m + 1$  do
5    $f_{X'}(x_i) = w(x_{i-1}, x_i) + w(x_i, x_{i+1}) + f_X(x_i)$ 
6 return  $X'$ 
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180  $S \in \text{argmin}_{S \subseteq \text{support}(X), |S| \leq m} \varepsilon(X, S)$ . By Corollary 9, the variable  $X'$  constructed in lines 4 and  
181 5 satisfies  $d_K(X, X') = \varepsilon(X, S)$  and by the minimality of  $S$  and by Proposition 5, it is an optimal  
182 approximation.  $\square$

183 **Theorem 11.** *The KolmogorovApprox( $X, m$ ) algorithm runs in time  $O(mn^2)$ , using  $O(n^2)$  mem-*  
184 *ory where  $n = |\text{support}(X)|$ .*

185 *Proof.* Constructing the graph  $G$  takes  $O(n^2)$ . The number of edges is  $O(E) \approx O(n^2)$  and for every  
186 edge the weight is at most the sum of all probabilities between the source node  $-\infty$  and the target  
187 node  $\infty$ , which can be done efficiently by aggregating the weights of already calculated edges. The  
188 construction is also the only stage that requires memory allocation, specifically  $O(E + V) = O(n^2)$ .  
189 Finding the shortest path takes  $O(m(E + V)) \approx O(mn^2)$ .

190 [[GW: put a reference to the work of the fellows from the Technion to avoid some of this?]]

191 Since  $G$  is DAG (directed acyclic graph) finding shortest path takes  $O(E + V)$ . We only need  
192 to find paths of length  $\leq m$ , which takes  $O(m(E + V))$ . Deriving the new random variable  $X'$   
193 from the computed path  $l$  takes  $O(mn)$ . For every node in  $l$  (at most  $m$  nodes), calculating the  
194 probability  $P(s < X < \infty)$  takes at most  $n$  steps. To conclude, the worst case run-time complexity  
195 is  $O(n^2 + mn^2 + mn) = O(mn^2)$  and memory complexity is  $O(E + V) = O(n^2)$ .  $\square$

## 196 4 A case study and experimental results

197 The case study examined in our experiments is the problem of task trees with deadlines [4, 3].  
198 Hierarchical planning is a well-established field in AI [5, 6, 7], and is still relevant nowadays [1, 15].  
199 A hierarchical plan is a method for representing problems of automated planning in which the  
200 dependency among tasks can be given in the form of networks, here we focus on hierarchical plans  
201 represented by task trees. The leaves in a task tree are *primitive* actions (or tasks), and the internal  
202 nodes are either *sequence* or *parallel* actions. The plans we deal with are of stochastic nature, where  
203 the duration of a primitive action is given by a random variable.

204 A sequence node denotes a series of tasks that should be performed consecutively, whereas a parallel  
205 node denotes a set of tasks that begin at the same time. A *valid* plan is one that is fulfilled before some  
206 given *deadline*, i.e., its *makespan* is less than or equal to the deadline. The objective in this context  
207 is to compute the probability that a given plan is valid, or more formally computing  $P(X < T)$ ,  
208 where  $X$  is a random variable representing the makespan of the plan and  $T$  is the deadline. As said  
209 above, resource consumption (task duration) is uncertain, and described as probability distributions  
210 in the leaf nodes. We assume that the distributions are independent but *not* necessarily identically  
211 distributed and that the random variables are discrete and have a finite support.

212 The problem of finding the probability that a task tree satisfies a deadline is known to be NP-hard. In  
 213 fact, even the problem of summing a set of random variables is NP-hard [10]. This is an example of  
 214 an explicitly given random variable that we need to estimate deadline meeting probabilities for.

215 In the first experiment we focus on is the problem of task trees with deadlines, and consider three  
 216 types of task trees. The first type includes logistic problems of transporting packages by trucks and  
 217 airplanes (from IPC2 <http://ipc.icaps-conference.org/>). Hierarchical plans of those logistic problems  
 218 were generated by the JSHOP2 planner [11] (see example problem, Figure 1, one parallel node with  
 219 all descendant task nodes being in sequence). The second type consists of task trees used as execution  
 220 plans for the ROBIL team entry in the DARPA robotics challenge (DRC simulation phase), and the  
 221 third type is of linear plans (sequential task trees). The primitive tasks in all the trees are modeled as  
 222 discrete random variables with support of size  $M$  obtained by discretization of uniform distributions  
 over various intervals. The number of tasks in a tree is denoted by  $N$ .

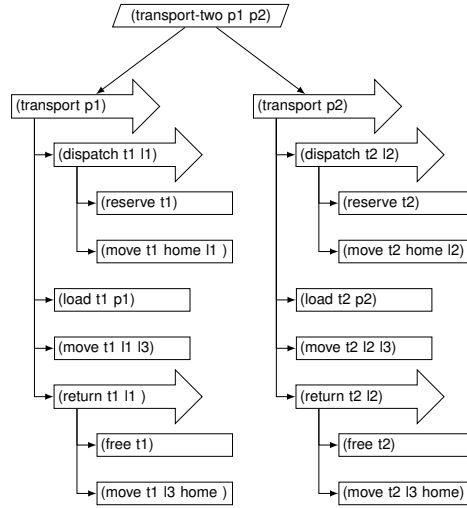


Figure 1: A plan generated by the JSHOP2 algorithm. Arrow shapes represent sequence nodes, parallelograms represent parallel nodes, and rectangles represent primitive nodes.

223

224 We implemented the approximation algorithm for solving the deadline problem with four different  
 225 methods of approximation. The first two are for achieving a one-sided Kolmogorov approximation –  
 226 the OptTrim [3] and the Trim [4] operators, and the third is a simple sampling scheme. We used those  
 227 methods as a comparison to the Kolmogorov approximation with the suggested KolmogorovApprox  
 228 algorithm. The parameter  $m$  of OptTrim and KolmogorovApprox corresponds to the inverse of  $\varepsilon$   
 229 given to the Trim operator. Note that in order to obtain some error  $\varepsilon$ , one must take into consideration  
 230 the size of the task tree  $N$ , therefore,  $m/N = 1/(\varepsilon \cdot N)$ . We ran also an exact computation as  
 231 a reference to the approximated one in order to calculate the error. The experiments conducted  
 232 with the following operators and their parameters: KolmogorovApprox operator with  $m = 10 \cdot N$ ,  
 233 the OptTrim operator with  $m = 10 \cdot N$ , the Trim as operator with  $\varepsilon = 0.1/N$ , and two simple  
 234 simulations, with a different samples number  $s = 10^4$  and  $s = 10^6$ .

235 Table 1 shows the results of the main experiment. The quality of the solutions provided by using the  
 236 OptTrim operator are better (lower errors) than those provided by the Trim operator, following the  
 237 optimality guarantees, but is interesting to see that the quality gaps happen in practice in each of the  
 238 examined task trees. However, in some of the task trees the sampling method produced better results  
 239 than the approximation algorithm with OptTrim. Nevertheless, the approximation algorithm comes

Task Tree	$M$	KolmogorovApprox	OptTrim	Trim	Sampling	
		$m/N=10$	$m/N=10$	$\varepsilon \cdot N=0.1$	$s=10^4$	$s=10^6$
Logistics ( $N=34$ )	2	0	0	0.0019	0.007	0.0009
	4	0.0024	0.0046	0.0068	0.0057	0.0005
Logistics ( $N=45$ )	2	0.0002	0.0005	0.002	0.015	0.001
	4	0	0.003	0.004	0.008	0.0006
DRC-Drive ( $N=47$ )	2	0	0.004	0.009	0.0072	0.0009
	4	0.001	0.008	0.019	0.0075	0.0011
Sequential ( $N=10$ )	2	0.0093	0.015	0.024	0	0
	4	0	0.024	0.04	0.008	0.0016
	10	0	0.028	0.06	0.0117	0.001

Table 1: Comparison of estimated errors with respect to the reference exact computation on various task trees.

with an inherent advantage of providing an exact quality guarantees, as opposed to the probabilistic guarantees provided by sampling.

In order to better understand the quality gaps in practice between KolmogorovApprox, OptTrim, and Trim, we investigate their relative errors when applied on single random variables with support size  $n = 100$ , and different support sizes of the resulting random variable approximation ( $m$ ). In each instance of this experiment, a random variable is randomly generated by choosing the probabilities of each element in the support from a uniform distribution and then normalizing these probabilities so that they sum to one.

Figure 2 present the error produced by the above methods. The depicted results are averages over several instances (50 instances) of random variables. The curves in the figure show the average error of OptTrim and Trim operators with comparison to the average error of the optimal approximation provided by KolmogorovApprox as a function of  $m$ .

According to the depicted results it is evident that increasing the support size of the approximation  $m$  reduces the error, as expected, in all three methods. However, errors produced by the KolmogorovApprox are significantly smaller, safe to say, a half of the error produced by OptTrim and Trim, it is clear both in the table (the relative error is mostly above 1) and in the graph.

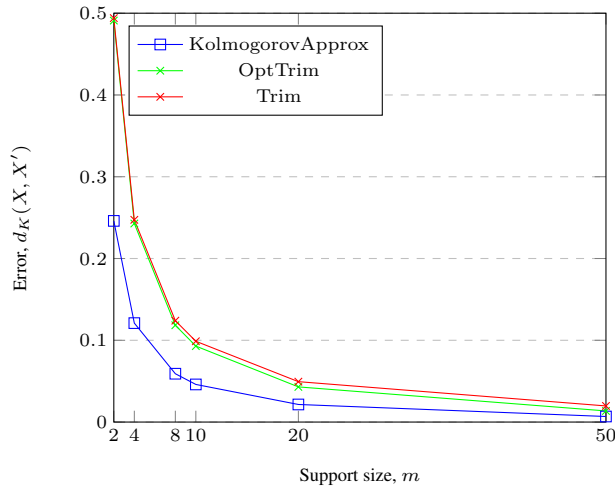


Figure 2: Error comparison between KolmogorovApprox, OptTrim, and Trim, on randomly generated random variables as function of  $m$ .



We also examined how our algorithm compares to linear programming as described and discussed, for example, in [12]. We ran an experiment to compare the run-time between the KolmogorovApprox algorithm with the run-time of a state-of-art implementation of linear programming. We used the “Minimize” function of Wolfram Mathematica and fed it with the equations  $\min_{\alpha \in \mathbb{R}^n} \|x - \alpha\|_\infty$  subject to  $\|\alpha\|_0 \leq m$  and  $\|\alpha\|_1 = 1$ . The run-time comparison results were clear and persuasive, for a random variable with support size  $n = 10$  and  $m = 5$ , the LP algorithm run-time was 850 seconds, where the KolmogorovApprox algorithm run-time was less than a tenth of a second. For  $n = 100$  and  $m = 5$ , the KolmogorovApprox algorithm run-time was 0.14 seconds and the LP algorithm took more than a day. Due to these timing results of the LP algorithm we did not proceed to examine it any further. Since it is not trivial to formally analyze the run-time of the LP algorithm, we conclude by the reported experiment that in this case the LP algorithm might not be as efficient as KolmogorovApprox algorithm whose complexity is proven to be polynomial in Theorem 11.

## 5 Discussion

## References

- [1] R. Alford, V. Shivashankar, M. Roberts, J. Frank, and D. W. Aha. Hierarchical planning: Relating task and goal decomposition with task sharing. In *IJCAI*, pages 3022–3029, 2016.
- [2] A. Chakravarty, J. Orlin, and U. Rothblum. A partitioning problem with additive objective with an application to optimal inventory groupings for joint replenishment. *Operations Research*, 30(5):1018–1022, 1982.
- [3] L. Cohen, T. Grinshpoun, and G. Weiss. Optimal approximation of random variables for estimating the probability of meeting a plan deadline. In *Proceedings of the Thirty-Second AAAI Conference on Artificial Intelligence, New Orleans, Louisiana, USA, February 2-7, 2018*, 2018.
- [4] L. Cohen, S. E. Shimony, and G. Weiss. Estimating the probability of meeting a deadline in hierarchical plans. In *IJCAI*, pages 1551–1557, 2015.
- [5] T. Dean, R. J. Firby, and D. Miller. Hierarchical planning involving deadlines, travel time, and resources. *Computational Intelligence*, 4(3):381–398, 1988.
- [6] K. Erol, J. Hendler, and D. S. Nau. HTN planning: Complexity and expressivity. In *AAAI*, volume 94, pages 1123–1128, 1994.
- [7] K. Erol, J. Hendler, and D. S. Nau. Complexity results for HTN planning. *Annals of Mathematics and Artificial Intelligence*, 18(1):69–93, 1996.
- [8] J. D. Gibbons and S. Chakraborti. Nonparametric statistical inference. In *International encyclopedia of statistical science*, pages 977–979. Springer, 2011.
- [9] A. C. Miller and T. R. Rice. Discrete approximations of probability distributions. *Management Science*, 29(3):352–362, 1983.
- [10] R. Möhring. Scheduling under uncertainty: Bounding the makespan distribution. *Computational Discrete Mathematics*, pages 79–97, 2001.
- [11] D. S. Nau, T.-C. Au, O. Ilghami, U. Kuter, J. W. Murdock, D. Wu, and F. Yaman. SHOP2: An HTN planning system. *Journal of Artificial Intelligence Research*, 20:379–404, 2003.

- 294 [12] K. Pavlikov and S. Uryasev. CVaR distance between univariate probability distributions and  
295 approximation problems. Technical Report 2015-6, University of Florida, 2016.
- 296 [13] A. N. Pettitt and M. A. Stephens. The kolmogorov-smirnov goodness-of-fit statistic with  
297 discrete and grouped data. *Technometrics*, 19(2):205–210, 1977.
- 298 [14] E. Shufan, H. Ilani, and T. Grinshpoun. A two-campus transport problem. In *MISTA*, pages  
299 173–184, 2011.
- 300 [15] Z. Xiao, A. Herzig, L. Perrussel, H. Wan, and X. Su. Hierarchical task network planning with  
301 task insertion and state constraints. In *IJCAI*, pages 4463–4469, 2017.