# **RBFAR-OLS**

This repository contains code for my research on a modified method of OLS (orthogonal least squares), applied to RBF-AR (radial basis function - autoregressive) models.

## Relevant Papers

- Original OLS algorithm for RBF networks
- One of the articles that introduced RBF-AR
- RBF-AR parameter optimization
- RBF-AR applications
- Adaptive learning for RBF-AR methods

## Formulation of the problem

$$\begin{split} X &\in \mathbb{R}^{l \times n} \\ d &\in \mathbb{R}^{l \times 1} \\ \Psi : \mathbb{R}^n &\to \left[0, 1\right]^m, \\ \Phi_i &= \Psi(X_i) \\ \mathbf{N} &\in \mathbb{R}^{m \times n} \end{split}$$

**Note:** the target variable d should be normalized, so it has a mean of 0. These models only work well for stationary data, but for data with a trend, this normalization should be done via a trend removal method.

## First approach (locally pre-trained coefficients)

$$\begin{aligned} y_i &= \Psi(X_i) \cdot (X_i \cdot \mathbf{N}^T) \\ \Longrightarrow d \approx y &= \left(\Phi \odot (X \cdot \mathbf{N}^T)\right) \cdot \underline{1} \\ \Longrightarrow \mathbf{N} &= ? \end{aligned}$$

**Solution:** first fit an AR model to each candidate centre, and then select from them and assign weights to each selected one.

Let  $\Xi_i$  be the set of training sample indicies associated with the *i*th candidate centre. We define  $\Xi_i$  as follows:

$$\Xi_i = \! \{j \in [1,l] \mid \! \Psi_i(\boldsymbol{X}_i) \geq \rho \},$$

where  $\rho$  is a threshold for the candidate centre selection, and  $\Psi_i(X_j)$  is the ith element of  $\Psi(X_j)$ 

Let  $\nu_i$  be the coefficient-vector associated with the ith candidate centre.

We fit each linear AR model separately, using a standard least squares approach:

$$\hat{\nu}_i = \arg\min_{\nu_i} \sum_{j \in \Xi_i} (d_j - X_j \cdot \nu_i)^2$$

Let N be the matrix of these coefficients:

$$\mathbf{N} = \begin{bmatrix} \hat{\nu}_1^T \\ \hat{\nu}_2^T \\ \vdots \\ \hat{\nu}_m^T \end{bmatrix}$$

Now we can calculate the activation matrix  $\Sigma$ . This matrix represents the output for each training point i and each linear AR model, associated with the jth candidate centre:

$$\begin{split} \Sigma &= \Phi \odot (X \cdot \mathbf{N}^T) \\ \Sigma_{i,j} &= \Psi_j(X_i) \cdot (X_i \cdot \hat{\nu}_j^T) \\ \Sigma &= \begin{bmatrix} \Psi_1(X_1) \cdot (X_1 \cdot \hat{\nu}_j^T) & \Psi_2(X_1) \cdot (X_1 \cdot \hat{\nu}_2^T) & \dots & \Psi_m(X_1) \cdot (X_1 \cdot \hat{\nu}_m^T) \\ \Psi_1(X_2) \cdot (X_2 \cdot \hat{\nu}_1^T) & \Psi_2(X_2) \cdot (X_2 \cdot \hat{\nu}_2^T) & \dots & \Psi_m(X_2) \cdot (X_2 \cdot \hat{\nu}_m^T) \\ \vdots & \vdots & \ddots & \vdots \\ \Psi_1(X_l) \cdot (X_l \cdot \hat{\nu}_1^T) & \Psi_2(X_l) \cdot (X_l \cdot \hat{\nu}_2^T) & \dots & \Psi_m(X_l) \cdot (X_l \cdot \hat{\nu}_m^T) \end{bmatrix} \end{split}$$

From now, we can continue with the OLS approach, such that the activations, normally denoted by P are given by the  $\Sigma$  matrix  $(P = \Sigma)$ .

We then select the centres and assign coefficients  $(\theta_i)$  to each selected centre using the OLS approach.

$$d = \Sigma \cdot \theta + e$$

The details are left out here, as they are similar to the OLS approach described below.

Having completed this step, we get:

$$y = (\Phi \odot (X \cdot \mathbf{N}^T)) \cdot \theta$$

This can be rewritten as the original equation, with the weights modified:

$$y = (\Phi \odot (X \cdot \tilde{\mathbf{N}}^T)) \cdot \underline{1},$$
 where  $\tilde{\mathbf{N}} = \mathbf{N} \odot \theta$ 

## Second approach (no pre-training)

We flatten the N matrix into a vector  $\nu$  and rewrite the equation as follows:

$$\begin{split} P \in \mathbb{R}^{l \times (n \cdot m)} \\ P_{i,(j-1) \cdot m+k} = \Phi_{i,j} \cdot X_{i,k} \\ P = \begin{bmatrix} \Phi_{1,1} X_{1,1} & \Phi_{1,2} X_{1,1} & \dots & \Phi_{1,m} X_{1,n} \\ \Phi_{2,1} X_{2,1} & \Phi_{2,2} X_{2,1} & \dots & \Phi_{2,m} X_{2,n} \\ \vdots & \vdots & \ddots & \vdots \\ \Phi_{l,1} X_{l,1} & \Phi_{l,2} X_{l,1} & \dots & \Phi_{l,m} X_{l,n} \end{bmatrix} \end{split}$$

let  $p_i$  be the ith column of P

$$\begin{split} \nu &= [\mathbf{N}_{1,1}, \mathbf{N}_{1,2}, \dots, \mathbf{N}_{1,n}, \mathbf{N}_{2,1}, \dots, \mathbf{N}_{m,n}]^T \\ \Longrightarrow \ d \approx y = P \cdot \nu \\ A &\in \mathbb{R}^{M_S \times M_S}, \end{split}$$

where  ${\cal M}_S$  is the number of selected centres

$$A = \begin{bmatrix} 1 & \alpha_{1,1} & \dots & \alpha_{1,M_S} \\ 0 & 1 & \dots & \alpha_{2,M_S} \\ \vdots & \ddots & \ddots & \vdots \\ 0 & 0 & \dots & 1 \end{bmatrix}$$

OLS criteria:

- $P = W \cdot A$ , where A is an upper triangular matrix, W has orthogonal columns
- $d = W \cdot g + e$

Thus:

$$d\approx W\cdot g=P\cdot \nu=W\cdot A\cdot \nu$$

## Centre selection

#### Initialization:

• define candidate centres C, and calculate  $\Phi$  accordingly

- $\bullet$  calculate P as defined above
- initialize the residuals:  $d^{(0)} = d$
- initialize the selected centres:  $C^{(0)} = \emptyset$
- initialize W to an empty matrix:  $W^{(0)} = \emptyset$
- initialize g to an empty vector:  $g^{(0)} = \emptyset$  initialize the selected indicies  $I_{\text{selected}}^{(0)} = \emptyset$

# At each iteration ( $k \in [1, |C_{candidates}|]$ ):

 $\forall i \in [1, |C_{\textbf{candidates}}|] \setminus I_{\textbf{selected}}^{(k-1)}:$ 

$$\begin{split} w_i &= p_i - \sum_{j=1}^{k-1} \alpha_{i,j} w_j, \\ \text{where } \alpha_{i,j} &= \frac{p_i^T \cdot w_j}{||w_j||^2} \\ g_i &= \frac{w_i^T \cdot d^{(k-1)}}{||w_i||^2} \\ err_i &= g_i^2 \cdot \frac{||w_i||^2}{||d^{(k-1)}||^2} = \frac{\left(w_i^T \cdot d^{(k-1)}\right)^2}{||w_i||^2 \cdot ||d^{(k-1)}||^2} \end{split}$$

Select the centre  $c_i$  with the highest  $err_i$ .

$$\begin{split} I_{\text{selected}}^{(k)} &= I_{\text{selected}}^{(k-1)} \cup \{i\} \\ C^{(k)} &= C^{(k-1)} \cup \{c_i\} \\ W^{(k)} &= W^{(k-1)} \oplus w_i \\ g^{(k)} &= g^{(k-1)} \oplus g_i \\ d^{(k)} &= d^{(k-1)} - w_i \cdot g_i \end{split}$$

Stopping criteria:

- 1. Maximum number of iterations reached:  $k = K_{max}$
- 2. Convergence:  $||d^{(k)}||^2 < \epsilon$
- 3. No significant improvement:  $\frac{||d^{(k)}||^2}{||d^{(k-1)}||^2} < \delta$

If any of the stopping criteria is met, the algorithm terminates with  ${\cal M}_S=k.$ 

## Finding the coefficients

$$d = W \cdot g + e = P \cdot \nu = W \cdot A \cdot \nu + e \approx W \cdot \hat{g}$$

W has orthogonal columns  $\implies W^T \cdot W = H$ ,

where H is a diagonal matrix, and thus invertible

We can calculate 
$$\hat{g}$$
 by  $H^{-1} \cdot W^T \cdot d = \hat{g}$ 

$$\implies A \cdot \hat{\nu} = H^{-1} \cdot W^T \cdot d$$
$$A \cdot \hat{\nu} = \hat{q}$$

which is easily and efficiently solvable, due to A being an upper triangular matrix.

## SVD-based center selection

Let us assume, we have calculated the activation matrix P, using either approach. Let us then compute its SVD:

$$P = Q \cdot \Sigma \cdot V^T,$$

where Q and V are orthogonal matrices, and  $\Sigma$  is a diagonal matrix containing the singular values of P.

We then get the following equation for the optimal weights:

$$\hat{\nu} = \operatorname{argmin}_{\nu}\{\left\|d - P \cdot \nu\right\|_{2}^{2} + \alpha \left\|\nu\right\|_{2}^{2}\},$$

where  $\alpha$  is a regularization parameter.

We can transform this as follows:

$$\hat{\nu} = \operatorname{argmin}_{\nu} \{ \|d - P \cdot \nu\|_{2}^{2} + \alpha \|\nu\|_{2}^{2} \}$$

$$= \operatorname{argmin}_{\nu} \{ d^{T}d - 2\nu^{T}P^{T}d + \nu^{T}P^{T}P\nu + \alpha\nu^{T}\nu \}$$

$$\Rightarrow \frac{\delta}{\delta\nu} = -2P^{T}d + 2P^{T}P\nu + 2\alpha\nu = 0$$

$$\Rightarrow \hat{\nu} = (P^{T}P + \alpha I)^{-1}P^{T}d$$

$$= (V\Sigma^{T}Q^{T}Q\Sigma V^{T} + \alpha VV^{T})^{-1}V\Sigma^{T}Q^{T}d$$

$$Q \text{ is orthonormal } \Rightarrow Q^{T}Q = I \Rightarrow \hat{\nu} = (V\Sigma^{T}I\Sigma V^{T} + \alpha VV^{T})^{-1}V\Sigma^{T}Q^{T}d$$

$$= (V\Sigma^{T}\Sigma V^{T} + \alpha VV^{T})^{-1}V\Sigma^{T}Q^{T}d$$

$$= (V(\Sigma^{T}\Sigma + \alpha I)V^{T})^{-1}V\Sigma^{T}Q^{T}d$$

$$V \text{ is orthogonal } \Rightarrow V^{-1} = V^{T} \Rightarrow \hat{\nu} = V(\Sigma^{T}\Sigma + \alpha I)^{-1}V^{T}V\Sigma^{T}Q^{T}d$$

$$V \text{ is orthogonal } \Rightarrow V^{T}V = I \Rightarrow \hat{\nu} = V(\Sigma^{T}\Sigma + \alpha I)^{-1}\Sigma^{T}Q^{T}d$$

$$= V\begin{bmatrix} \frac{\sigma_{1}}{\sigma_{1}^{2}+\alpha} & 0 & \cdots & 0\\ 0 & \frac{\sigma_{2}}{\sigma_{2}^{2}+\alpha} & \cdots & 0\\ \vdots & \vdots & \ddots & \vdots\\ 0 & 0 & \cdots & \frac{\sigma_{n}}{\sigma_{2}^{2}+\alpha} \end{bmatrix} Q^{T}d$$

 $\hat{\nu} = V \begin{bmatrix} 0 & \frac{\sigma_2}{\sigma_2^2 + \alpha} \\ \vdots & \vdots \\ 0 & 0 \end{bmatrix}$ 

This means that:

$$\begin{split} y &= P\hat{\nu} \\ &= Q\Sigma V^T\hat{\nu} \\ &= Q\Sigma V^TV \begin{bmatrix} \frac{\sigma_1}{\sigma_1^2 + \alpha} & 0 & \cdots & 0 \\ 0 & \frac{\sigma_2}{\sigma_2^2 + \alpha} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \frac{\sigma_n}{\sigma_n^2 + \alpha} \end{bmatrix} Q^Td \\ &= Q\Sigma \begin{bmatrix} \frac{\sigma_1}{\sigma_1^2 + \alpha} & 0 & \cdots & 0 \\ 0 & \frac{\sigma_2}{\sigma_2^2 + \alpha} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \frac{\sigma_n}{\sigma_n^2 + \alpha} \end{bmatrix} Q^Td \\ &= Q\begin{bmatrix} \frac{\sigma_1^2}{\sigma_1^2 + \alpha} & 0 & \cdots & 0 \\ 0 & \frac{\sigma_2^2}{\sigma_2^2 + \alpha} & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \frac{\sigma_n^2}{\sigma_2^2 + \alpha} \end{bmatrix} Q^Td \end{split}$$

Let  $\sigma_i' = \frac{\sigma_i^2}{\sigma_i^2 + \alpha}$ , and let  $\Sigma'$  be defined as follows:

$$\Sigma' = \begin{bmatrix} \frac{\sigma_1^2}{\sigma_1^2 + \alpha} & 0 & \cdots & 0\\ 0 & \frac{\sigma_2^2}{\sigma_2^2 + \alpha} & \cdots & 0\\ \vdots & \vdots & \ddots & \vdots\\ 0 & 0 & \cdots & \frac{\sigma_n^2}{\sigma_n^2 + \alpha} \end{bmatrix}$$
$$= \begin{bmatrix} \sigma_1' & 0 & \cdots & 0\\ 0 & \sigma_2' & \cdots & 0\\ \vdots & \vdots & \ddots & \vdots\\ 0 & 0 & \cdots & \sigma_n' \end{bmatrix}$$

Thus we get:

$$y = Q\Sigma'Q^Td$$

#### The selection algorithm

We iteratively select the first k columns of Q and  $\Sigma'$ , which correspond to the k largest singular values. This gives us a reduced representation of the data, focusing on the most important features. We then calculate the corresponding error and stop when the error is below a certain threshold.

Error = 
$$||d - P\hat{\nu}||_2^2$$
  
=  $||d - Q\Sigma'Q^Td||_2^2$ 

Let us define  $Q^{(k)}$  as the matrix formed by the first k columns of Q and zeros elsewhere, and  $\Sigma^{(k)}$  as the matrix formed by the first k rows and columns of  $\Sigma'$  and zeros in the rest. Then we can rewrite the error as follows:

$$\mathrm{Error}_k = \|d - Q^{(k)} \Sigma'^{(k)} Q^{(k)T} d\|_2^2$$

Our stopping criterion is based on the error:

$$\mathrm{Error}_k < \epsilon$$

where  $\epsilon$  is a predefined threshold. When this condition is met, we stop the selection process and use the current  $Q^{(k)}$  and  $\Sigma'^{(k)}$  for our reduced representation.

We can refine this process by optimizing the calculation of  $\mathrm{Error}_k$ . We can achieve this by only calculating the differences of the errors.

$$\begin{split} & \operatorname{Error}_{0} = \|d\|_{2}^{2} \\ & \operatorname{Error}_{k} - \operatorname{Error}_{k-1} = \|d - Q^{(k)} \Sigma'^{(k)} Q^{(k)T} d\|_{2}^{2} - \|d - Q^{(k-1)} \Sigma'^{(k-1)} Q^{(k-1)T} d\|_{2}^{2} \\ & = \left( \|d\|_{2}^{2} - 2 \sum_{i=1}^{k} d^{T} q_{i} \sigma'_{i} q_{i}^{T} d + \sum_{i=1}^{k} d^{T} q_{i}^{T} \sigma'_{i} q_{i} q_{i} \sigma'_{i} q_{i}^{T} d \right) \\ & - \left( \|d\|_{2}^{2} - 2 \sum_{i=1}^{k-1} d^{T} q_{i} \sigma'_{i} q_{i}^{T} d + \sum_{i=1}^{k-1} d^{T} q_{i}^{T} \sigma'_{i} q_{i} q_{i} \sigma'_{i} q_{i}^{T} d \right) \\ & = \left( \|d\|_{2}^{2} - 2 \sum_{i=1}^{k} \sigma'_{i} \left( q_{i}^{T} d \right)^{2} + \sum_{i=1}^{k} \sigma'_{i}^{2} \left( q_{i}^{T} d \right)^{2} \right) \\ & - \left( \|d\|_{2}^{2} - 2 \sum_{i=1}^{k-1} \sigma'_{i} \left( q_{i}^{T} d \right)^{2} + \sum_{i=1}^{k-1} \sigma'_{i}^{2} \left( q_{i}^{T} d \right)^{2} \right) \\ & = \left( \|d\|_{2}^{2} + \sum_{i=1}^{k} \sigma'_{i} \left( \sigma'_{i} - 2 \right) \left( q_{i}^{T} d \right)^{2} \right) \\ & - \left( \|d\|_{2}^{2} + \sum_{i=1}^{k-1} \sigma'_{i} \left( \sigma'_{i} - 2 \right) \left( q_{i}^{T} d \right)^{2} \right) \\ & = \sigma'_{k} \left( \sigma'_{k} - 2 \right) \left( q_{k}^{T} d \right)^{2} \\ \hline & \operatorname{Error}_{k} - \operatorname{Error}_{k-1} = \sigma'_{k} \left( \sigma'_{k} - 2 \right) \left( q_{k}^{T} d \right)^{2} \end{split}$$

To be able to consistently use the same sort of  $\epsilon$  values, we can normalize the error by dividing it by the initial error:

$$\text{Normalized Error}_k = \frac{\text{Error}_k}{\text{Error}_0} = \frac{\text{Error}_k}{\|d\|_2^2}$$

Thus, we get our optimized centre-selection algorithm:

- 1. Normalized  $Error_0 := 1$
- 2. While  $k < |C_{candidates}|$ :
  - 1. Normalized  $\operatorname{Error}_k := \operatorname{Normalized} \operatorname{Error}_{k-1} + \sigma_k' \left(\sigma_k' 2\right) \frac{\left(q_k^T d\right)^2}{\|d\|_2^2}$
  - 2. If Normalized  $\mathrm{Error}_k < \epsilon$  then:
    - 1. Stop

## **TODO**

- regularizations
  - L1 regularization
- non global sigma estimation
- multistep forecast