Exercises Chapter 6

1. An alternative, equivalent representation of the *F*-test statistic is the following:

$$F = \frac{\left(\sum_{i=1}^{n} \hat{\varepsilon}_{iR}^{2} - \sum_{i=1}^{n} \hat{\varepsilon}_{iU}^{2}\right)/q}{\left(\sum_{i=1}^{n} \hat{\varepsilon}_{iU}^{2}\right)/(n-K)} = \frac{\left(SS_{R} - SS_{U}\right)/q}{SS_{U}/(n-K)},$$

where $\hat{\varepsilon}_{iU}$ are the residuals from the *un*restricted (i.e., the usual) regression of Y on X, and where $\hat{\varepsilon}_{iR}$ are the residuals from the *re*stricted ordinary least squares regression which minimizes the following *restricted* version of the OLS-objective function

$$\min_{\tilde{\beta}} S_n(\tilde{\beta}) = (Y - X\tilde{\beta})'(Y - X\tilde{\beta}) \quad \text{such that} \quad R\tilde{\beta} - r = 0$$

where the restriction is just the null hypothesis.

The standard F-**test.** The standard F-test for a linear regression tests the hypothesis that all coefficients except the intercept are equal to zero. In this case, $\hat{\varepsilon}_{iR}$ are simply the residuals from regressing Y on only the intercept. In this standard case we have,

$$F_{1} = \frac{\left(\sum_{i=1}^{n} (Y_{i} - \bar{Y})^{2} - \sum_{i=1}^{n} \hat{\varepsilon}_{iU}^{2}\right) / (K - 1)}{\left(\sum_{i=1}^{n} \hat{\varepsilon}_{iU}^{2}\right) / (n - K)}$$

since here $\hat{\varepsilon}_{iR}^2 = (Y_i - \bar{Y})^2$.

Show that F_1 is equal to F_2 with

$$F_2 = \frac{R_U^2/(K-1)}{(1-R_U^2)/(n-K)},$$

where R_U^2 denotes the coefficient of determination of the unrestricted regression model.

- 2. Install the R package AER and load the package. The ARE-package contains the data set Journals. Check ?Journals to learn more about the data. Create the variables citeprice (journal price per citations) and age (journal age) as following:
 - > # install.packages("AER")
 - > suppressMessages(library("AER"))
 - > ## attach the data-set Journals to the current R-session
 - > data("Journals", package = "AER")
 - > ## ?Journals # Check the help file
 - > ##
 - > ## Select variables "subs" and "price"
 - > journals <- Journals[, c("subs", "price")]
 - > ## Define variable 'journal-price per citation'
 - > journals\$citeprice <- Journals\$price/Journals\$citations
 - > ## Define variable 'journal-age'
 - > journals\$age <- 2020 Journals\$foundingyear</pre>
 - > ## Check variable names in 'journals'
 - > names(journals)

Estimate the coefficients β_1 and β_2 of the following linear regression model

$$log(Y_i) = \beta_1 + \beta_2 log(X_i) + \varepsilon_i, \quad i = 1, \dots, n$$

with $\log(Y) = \log(\text{subs})$ (i.e., logarithm of the number of library subscriptions) and $\log(X) = \log(\text{citeprice})$ (i.e., logarithm of the journal price per citations).

- (a) Do you have heteroscedastic error-term variances? Explain your answer by discussing a diagnostic plot showing the residuals against the fitted values.
- (b) Estimate the standard error of the OLS estimator $\hat{\beta}_2$ using an appropriate variance estimator.
- 3. Consider the following multiple linear regression model:

$$Y_i=\beta_1+\beta_2X_{2i}+\beta_3X_{3i}+\varepsilon_i,\quad i=1,\dots,n$$
 (in matrix notation)
$$Y=X\beta+\varepsilon$$

where $\beta = (1, -5, 5)'$, ε_i is a heteroscedastic error term

$$\varepsilon_i \sim N(0, \sigma_i^2)$$
 with $\sigma_i = |X_{3i}|$,

and where for all $i = 1, \ldots, n = 100$:

- $X_{2i} \sim N(10, 1.5^2)$
- $X_{3i} \sim U[0.2, 8]$

You're given the following data generated from this regression model:

```
> set.seed(109) # Sets the "seed" of the random number generators:
> n <- 100  # Number of observations
> ## Generate two explanatory variables plus an intercept-variable:
> X_1 <- rep(1, n)  # Intercept
> X_2 <- rnorm(n, mean=10, sd=1.5)  # Draw realizations form a normal distr.
> X_3 <- runif(n, min = 0.2, max = 8) # Draw realizations form a t-distr.
> X <- cbind(X_1, X_2, X_3)  # Save as a Nx3-dimensional data matrix.
> beta <- c(1, -5, 5)
> ## Generate realizations from the heteroscadastic error term
> eps <- rnorm(n, mean=0, sd=abs(X_3))
> ## Dependent variable:
> Y <- X %*% beta + eps</pre>
```

- (a) Compute the theoretical covariance matrix variance $Var(\hat{\beta})$ of the OLS estimator $\hat{\beta}$ for the given data generating process and the given data.
- (b) Use a Monte-Carlo simulation to generate 10000 variance estimates

$$\widehat{\operatorname{Var}}_{\mathsf{HC3},1}(\hat{\beta}_2),\ldots,\widehat{\operatorname{Var}}_{\mathsf{HC3},10000}(\hat{\beta}_2)$$

and 10000 variance estimates

$$\widehat{\operatorname{Var}}_{\mathsf{HC3},1}(\hat{\beta}_3),\ldots,\widehat{\operatorname{Var}}_{\mathsf{HC3},10000}(\hat{\beta}_3).$$

These estimates represent typical estimation results. (Of course, in practice you observe only one variance estimation result $\widehat{\mathrm{Var}}_{\mathsf{HC3}}(\hat{\beta}_2)$ for $\mathrm{Var}(\hat{\beta}_2)$ and one $\widehat{\mathrm{Var}}_{\mathsf{HC3}}(\hat{\beta}_3)$ for $\mathrm{Var}(\hat{\beta}_3)$.)

- (i) Visualize the Monte Carlo realizations for the variance estimates. Add points displaying the sample mean of the Monte Carlo realizations and points displaying the true variance values.
- (ii) Do the Monte Carlo realizations $\widehat{\mathrm{Var}}_{\mathsf{HC3},r}(\hat{\beta}_2)$ and $\widehat{\mathrm{Var}}_{\mathsf{HC3},r}(\hat{\beta}_3)$, $r=1,\ldots,10000$ estimate the true variances $\mathrm{Var}(\hat{\beta}_2)$ and $\mathrm{Var}(\hat{\beta}_3)$ well on average?
- (iii) Are there large estimation uncertainties?

4. The Boston housing data set (contained in the R package MASS) contains observations on housing values in suburbs of Boston. Let's consider the following regression model

```
\mathsf{medv}_i = \beta_1 + \beta_2 \mathsf{ptratio}_i + \beta_3 \mathsf{lstat}_i + \beta_4 \mathsf{age}_i + \beta_5 \mathsf{crim}_i + \beta_6 \mathsf{nox}_i + \varepsilon_i
```

where $i=1,\dots,n$ indexes the suburbs. Check ?Boston in R to get an overview about the variables. You can assume that the assumptions of Chapter 6 hold. The following R code computes the regression estimates:

```
> library("lmtest") # for coeftest()
> library("sandwich")# for robust se
> library("MASS") # for Boston housing data
> data("Boston") # Check: ?Boston; names(Boston)
> lm_obj <- lm(medv ~ ptratio + lstat + age + crim + nox, data = Boston)
> vcovHC3_mat <- vcovHC(lm_obj, type = "HC3")</pre>
> round(coeftest(lm_obj, vcov = vcovHC3_mat), 3)
t test of coefficients:
            Estimate Std. Error t value Pr(>|t|)
(Intercept) 55.746 3.013 18.504 <2e-16 ***
ptratio -1.181
lstat -0.868
                        0.147 -8.046 <2e-16 ***
lstat
                        0.081 -10.662 <2e-16 ***
           0.060 0.017 3.527 <2e-16 ***
-0.024 0.036 -0.674 0.501
-8.059 3.318 -2.429 0.016 *
age
crim
nox
Signif. codes: 0 '*** 0.001 '** 0.01 '* 0.05 '.' 0.1 ' 1
```

- (a) Use R to test $H_0: \beta_6=0$ versus $H_A: \beta_6<0$ by means of an t-test. What is the correct p-value and what is the test decision when using a significance level of $\alpha=0.01$?
- (b) Use R to test $H_0: \beta_5 = \beta_6 = 0$ versus $H_A: \beta_5 \neq 0$ and/or $\beta_6 \neq 0$ by means of an F-test. What is the marginal significance value in this case?
- (c) What is the maximal probability of a type I error if you test the null hypothesis in (b) by means of two separate t-tests instead of one F-test? How does this compare to the probability of a type I error for the F test in (b).