

# Estimating the Number of Active Devices Within a Fixed area Using Wi-Fi Monitoring

Hai Li

Under the guidance of Prof. Chamberland

Electrical and Computer Engineering Department  
Texas A&M University

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# Outline

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- Problem Description
- System Model
- Estimation Scheme
- Simulation and Result
- Experimental Implementation and Result
- Conclusion

# Introduction

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Occupancy estimation has several applications

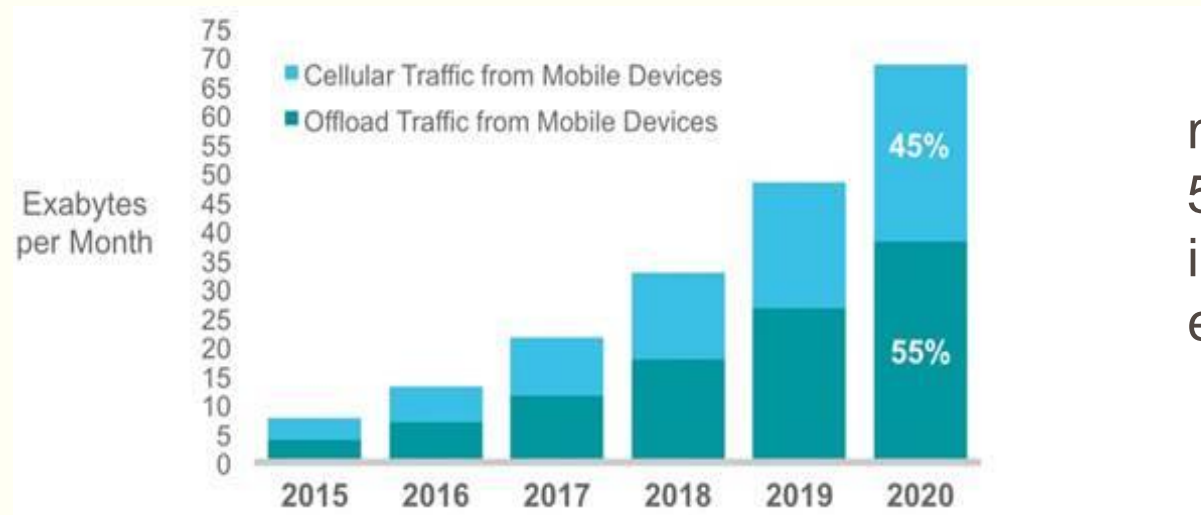
- Smart home management
- Emergency situation

Camera based approach is costly and limited by resolution

RF signal based approach consumes less power

# Wi-Fi Traffic Growth

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mobile offload increases from 51 percent (3.9 exabytes/month) in 2015 to 55 percent (38.1 exabytes/month) by 2020

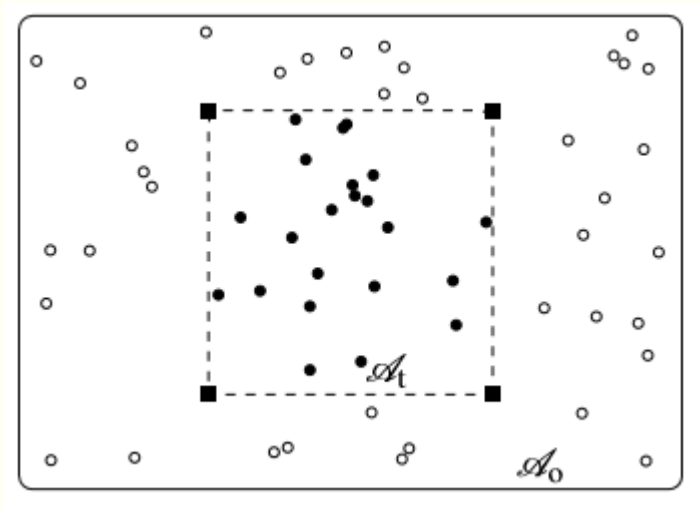
# Problem Description

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- Estimating number of active wireless units using Wi-Fi monitoring with statistical inference
- Increase the estimation accuracy by deploying proper directional antennas

# System Model

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$A_t$  is the area of target region.

$A_o$  is the area of complement region.

Location vector:

$$\underline{\mathbf{U}} = (\mathbf{U}_1, \dots, \mathbf{U}_{n_a})$$

Received signal vector:

$$\underline{\mathbf{P}} = (\mathbf{P}_1, \dots, \mathbf{P}_{n_a})$$

# Channel Model

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Log-normal channel model:

$$P_d[\text{dBm}] = A + B \log_{10}(d) + L_s + G_a$$

Log-normal distribution:

$$f_{L_s}(\ell) = \frac{1}{\sqrt{2\pi}\sigma_s} \exp\left(-\frac{\ell^2}{2\sigma_s^2}\right)$$

# Find Parameter Values

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Using least squares method:

$$\operatorname{argmin}_{a,b} \left\| \underbrace{\begin{bmatrix} \vdots \\ p_{ij} - G_i(\phi_{ij}) \\ \vdots \end{bmatrix}}_{\mathbf{y}} - \underbrace{\begin{bmatrix} \vdots & \vdots \\ 1 & \log_{10}(d_{ij}) \\ \vdots & \vdots \end{bmatrix}}_M \begin{bmatrix} a \\ b \end{bmatrix} \right\|^2.$$

$$[A \ B]^T = (M^T M)^{-1} M^T \mathbf{y}. \quad \sigma_s^2 = \frac{1}{N-1} \mathbf{y}^T (I - M(M^T M)^{-1} M^T) \mathbf{y}.$$



# System Model

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We assume the number and locations of wireless units form a Poisson point process

$$\Pr(R_t = r_t) = \frac{(\lambda_t A_t)^{r_t}}{r_t!} e^{-A_t \lambda_t} \quad r_t = 0, 1, \dots$$

$$\Pr(R_o = r_o) = \frac{(\lambda_o A_o)^{r_o}}{r_o!} e^{-A_o \lambda_o} \quad r_o = 0, 1, \dots$$

# Estimation Scheme

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Bayes Estimation:

- Applied when we know the Poisson parameters  $\lambda_t$  and  $\lambda_o$
- Prior probabilities are given by Poisson distributions

Maximum Estimation:

- Applied when we Poisson parameters  $\lambda_t$  and  $\lambda_o$  are unknown
- Likelihood function:  $\mathcal{L}(\lambda_t, \lambda_o; \underline{\mathbf{p}}, \underline{\mathbf{u}}) = f_{\underline{\mathbf{P}}, \underline{\mathbf{U}}}(\underline{\mathbf{p}}, \underline{\mathbf{u}}; \lambda_t, \lambda_o)$

# Bayes Estimation

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Distribution of  $\underline{\mathbf{U}}$

$$f_{\underline{\mathbf{U}}}(\underline{\mathbf{u}}) = \frac{1}{A_t^{R_t(\underline{\mathbf{u}})}} \frac{(\lambda_t A_t)^{R_t(\underline{\mathbf{u}})}}{(R_t(\underline{\mathbf{u}}))!} e^{-A_t \lambda_t} \frac{1}{A_o^{R_o(\underline{\mathbf{u}})}} \frac{(\lambda_o A_o)^{R_o(\underline{\mathbf{u}})}}{(R_o(\underline{\mathbf{u}}))!} e^{-A_o \lambda_o} = \frac{\lambda_t^{R_t(\underline{\mathbf{u}})}}{(R_t(\underline{\mathbf{u}}))!} \frac{\lambda_o^{R_o(\underline{\mathbf{u}})}}{(R_o(\underline{\mathbf{u}}))!} e^{-A_t \lambda_t - A_o \lambda_o}.$$

Conditional distribution of  $P_j$  given  $U_j$

$$\begin{aligned} & f_{\mathbf{P}_j | \mathbf{U}_j}(\mathbf{p}_j | \mathbf{u}_j) \\ &= \prod_{i=1}^{n_s} f_{L_{ij}}(p_{ij} - A - B \log_{10}(d_{ij}) - G_i(\phi_{ij})) \\ &= \frac{1}{(2\pi\sigma_s^2)^{\frac{n_s}{2}}} \prod_{i=1}^{n_s} e^{-\frac{(p_{ij} - A - B \log_{10}(d_{ij}) - G_i(\phi_{ij}))^2}{2\sigma_s^2}} \\ &= (2\pi\sigma_s^2)^{-\frac{n_s}{2}} e^{-\frac{\sum_{i=1}^{n_s} (p_{ij} - A - B \log_{10}(d_{ij}) - G_i(\phi_{ij}))^2}{2\sigma_s^2}}. \end{aligned}$$

# Bayes Estimation

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Conditional distribution of  $\underline{\mathbf{P}}$  given  $\underline{\mathbf{U}}$

$$f_{\underline{\mathbf{P}}|\underline{\mathbf{U}}}(\underline{\mathbf{p}}|\underline{\mathbf{u}}) = \prod_{j=1}^{n_a} f_{\mathbf{P}_j|\mathbf{U}_j}(\mathbf{p}_j|\mathbf{u}_j)$$

Marginal distribution of  $\underline{\mathbf{P}}$

$$\begin{aligned} f_{\underline{\mathbf{P}}}(\underline{\mathbf{p}}) &= \int_{\{\underline{\mathbf{u}}: R_t(\underline{\mathbf{u}}) + R_o(\underline{\mathbf{u}}) = n_a\}} f_{\underline{\mathbf{P}}|\underline{\mathbf{U}}}(\underline{\mathbf{p}}|\underline{\mathbf{u}}) f_{\underline{\mathbf{U}}}(\underline{\mathbf{u}}) d\underline{\mathbf{u}} \\ &= \sum_{(r_t, r_o): r_t + r_o = n_a} \sum_{\{\mathbb{I} \subset [n_a]: |\mathbb{I}| = r_t\}} \frac{\lambda_t^{r_t} \lambda_o^{r_o}}{r_t! r_o!} e^{-A_t \lambda_t - A_o \lambda_o} \\ &\quad \times \prod_{j \in \mathbb{I}} \mathcal{I}_{\mathcal{A}_t}(j) \prod_{j \in \mathbb{I}^c} \mathcal{I}_{\mathcal{A}_o}(j). \end{aligned}$$

The integral components

$$\mathcal{I}_{\mathcal{A}_t}(j) = \int_{\mathcal{A}_t} f_{\mathbf{P}_j|\mathbf{U}_j}(\mathbf{p}_j|\mathbf{u}_j) d\mathbf{u}_j$$

$$\mathcal{I}_{\mathcal{A}_o}(j) = \int_{\mathcal{A}_o} f_{\mathbf{P}_j|\mathbf{U}_j}(\mathbf{p}_j|\mathbf{u}_j) d\mathbf{u}_j.$$

# Bayes Estimation

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The posterior distribution of  $R_t$  given  $\underline{\mathbf{P}}$

$$\begin{aligned} \Pr(R_t = r_t | \underline{\mathbf{P}} = \underline{\mathbf{p}}) \\ = \sum_{\{\mathbb{I} \subset [n_a] : |\mathbb{I}| = r_t\}} \frac{\lambda_t^{r_t} \lambda_o^{r_o} e^{-A_t \lambda_t - A_o \lambda_o}}{r_t! r_o! f_{\underline{\mathbf{P}}}(\underline{\mathbf{p}})} \\ \times \prod_{j \in \mathbb{I}} \mathcal{J}_{\mathcal{A}_t}(j) \prod_{j \in \mathbb{I}^c} \mathcal{J}_{\mathcal{A}_o}(j) \end{aligned}$$

Bayes estimator

$$\hat{R}_t(\underline{\mathbf{p}}) = \mathbb{E}[R_t | \underline{\mathbf{P}} = \underline{\mathbf{p}}] = \sum_{r_t=0}^{n_a} r_t \Pr(R_t = r_t | \underline{\mathbf{P}} = \underline{\mathbf{p}}).$$

Performance criterion

$$\text{BMSE}[\hat{R}_t] = \mathbb{E}\left[\left(\hat{R}_t(\underline{\mathbf{P}}) - R_t\right)^2\right] \quad \text{BMSE}[\hat{R}_t] \approx \frac{1}{M} \sum_{m=1}^M \left(\hat{R}_t^{(m)}(\underline{\mathbf{P}}^{(m)}) - R_t^{(m)}\right)^2$$

# Maximum Likelihood Estimation

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Likelihood function

$$\begin{aligned} & \mathcal{L}(\lambda_t, \lambda_o; \underline{\mathbf{p}}, \underline{\mathbf{u}}) \\ &= f_{\underline{\mathbf{P}}, \underline{\mathbf{U}}}(\underline{\mathbf{p}}, \underline{\mathbf{u}}; \lambda_t, \lambda_o) \\ &= f_{\underline{\mathbf{P}}|\underline{\mathbf{U}}}(\underline{\mathbf{p}}|\underline{\mathbf{u}}) f_{\underline{\mathbf{U}}}(\underline{\mathbf{u}}; \lambda_t, \lambda_o). \end{aligned}$$

Marginal likelihood function of observed data  $\underline{\mathbf{P}}$

$$\begin{aligned} \mathcal{L}(\lambda_t, \lambda_o; \underline{\mathbf{p}}) &= \int_{\{\underline{\mathbf{u}}: R_t(\underline{\mathbf{u}}) + R_o(\underline{\mathbf{u}}) = n_a\}} f_{\underline{\mathbf{P}}|\underline{\mathbf{U}}}(\underline{\mathbf{p}}|\underline{\mathbf{u}}) f_{\underline{\mathbf{U}}}(\underline{\mathbf{u}}; \lambda_t, \lambda_o) d\underline{\mathbf{u}} \\ &= e^{-\Lambda_t \lambda_t - \Lambda_o \lambda_o} \sum_{(r_t, r_o): r_t + r_o = n_a} \frac{\lambda_t^{r_t} \lambda_o^{r_o}}{r_t! r_o!} \sum_{\{\mathbb{I} \subset [n_a]: |\mathbb{I}| = r_t\}} \prod_{j \in \mathbb{I}} \mathcal{J}_{\mathcal{A}_t}(j) \prod_{j \in \mathbb{I}^c} \mathcal{J}_{\mathcal{A}_o}(j) \end{aligned}$$

# Maximum Likelihood Estimation

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Property of likelihood function

$$\mathcal{L}(c\lambda_t, c\lambda_o; \underline{\mathbf{P}}) = c^{n_a} e^{-(A_t\lambda_t + A_o\lambda_o)(c-1)} \mathcal{L}(\lambda_t, \lambda_o; \underline{\mathbf{P}}).$$

Set c

$$c = \frac{n_a}{A_t\lambda_t + A_o\lambda_o}.$$

$$\max_{\lambda_t, \lambda_o} \mathcal{L}(\lambda_t, \lambda_o; \underline{\mathbf{p}}) = \max_{\alpha} \mathcal{L}\left(\frac{n_a}{A_t}\alpha, \frac{n_a}{A_o}(1 - \alpha); \underline{\mathbf{p}}\right)$$

# Maximum Likelihood Estimation

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Maximum likelihood estimator

$$\begin{aligned} & \hat{R}_t(\underline{\mathbf{p}}) \\ &= E_{\hat{\lambda}_t, \hat{\lambda}_o} [R_t | \underline{\mathbf{P}} = \underline{\mathbf{p}}] \\ &= \sum_{r_t=0}^{n_a} r_t \Pr(R_t = r_t | \underline{\mathbf{P}} = \underline{\mathbf{p}}; \hat{\lambda}_t, \hat{\lambda}_o). \end{aligned}$$

Performance criterion

$$\text{MSE} [\hat{R}_t] \approx \frac{1}{M} \sum_{m=1}^M \left( \hat{R}_t^{(m)}(\underline{\mathbf{P}}^{(m)}) - R_t^{(m)} \right)^2$$

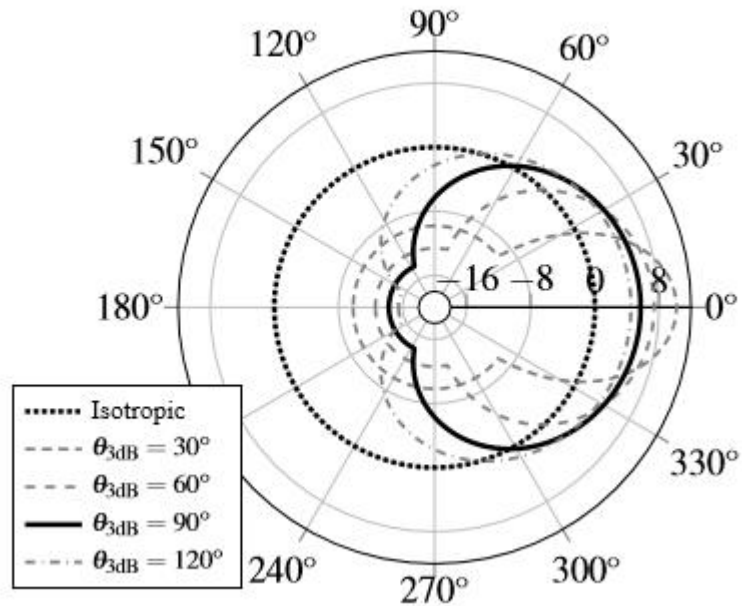


# Numerical Simulation

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3GPP antenna model

$$G_i(\phi_{ij}) = -\min \left\{ 12 \left( \frac{\phi_{ij} - \theta_i}{\theta_{3\text{dB}}} \right)^2, G_{\text{floor}} \right\} - G_{\text{avg}}$$



# Channel Parameters

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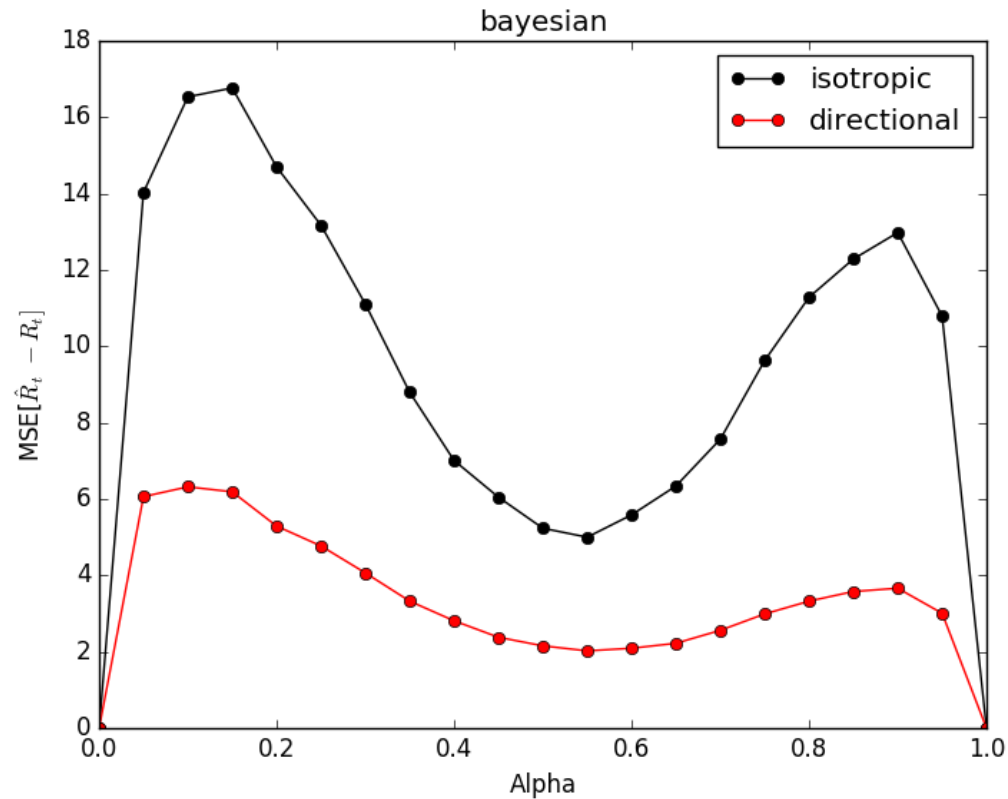
$$P[\text{dBm}] = A + B \log_{10}(d) + L + G(\phi)$$

$$A = P_t + 20 \log_{10} \left( \frac{3 \times 10^8}{f_{\text{carrier}}} \right) - 20 \log_{10}(4\pi)$$

$$B = -20 \text{dBm}$$

$$\sigma_s = 2.0 \text{dBm}$$

# Simulation Results



Bayes estimator

Intensity  $\lambda = 32$

21 points on each curve

50000 trials for each point

# Simulation Results

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$$\alpha = 1 - (\text{confidence level}/100)$$

$$p^* = 1 - \alpha/2$$

$$df = N - 1$$

$$SE = \frac{\sigma}{\sqrt{N}}$$

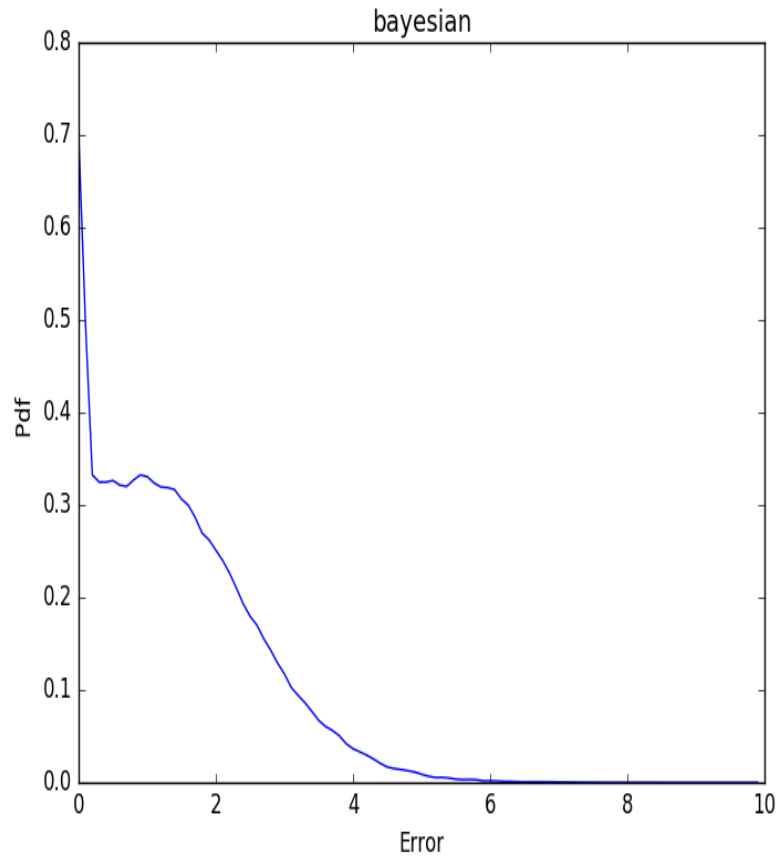
$$\text{Confidence interval} = \mu \pm \text{Margin of error}$$

95% Confidence interval of  $|r_t - \hat{r}_t|$

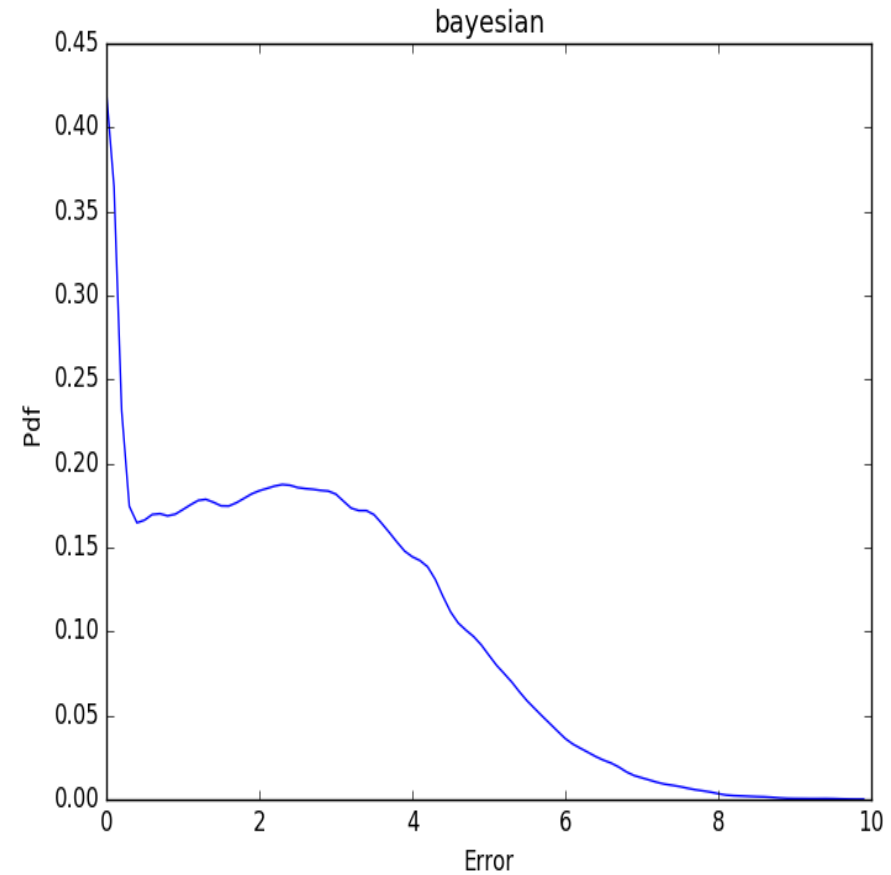
Antenna type	Confidence interval
Directional	$1.412411 \pm 0.002166$
Isotropic	$2.454623 \pm 0.003450$

# Simulation Results

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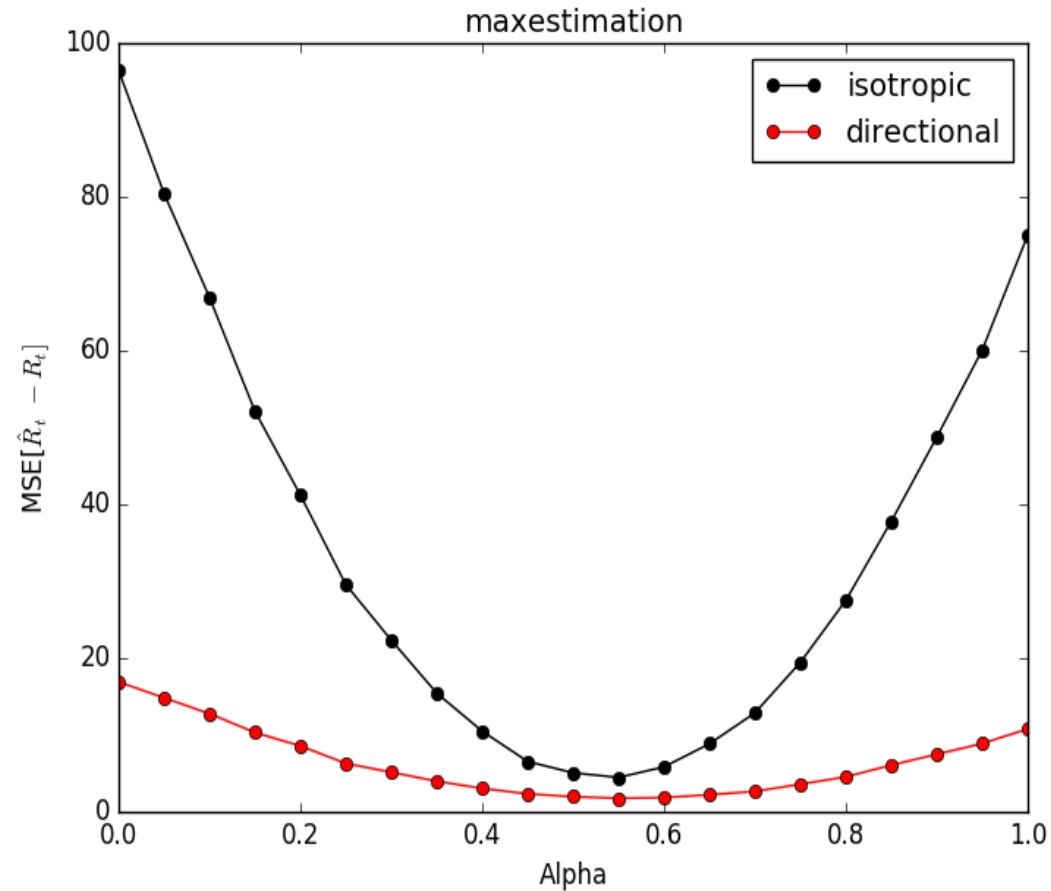


Pdf of  $|r_t - \hat{r}_t|$   
System with directional antennas



Pdf of  $|r_t - \hat{r}_t|$   
System with isotropic antennas

# Simulation Results



Maximum likelihood estimator

Intensity  $\lambda = 32$

21 points on each curve

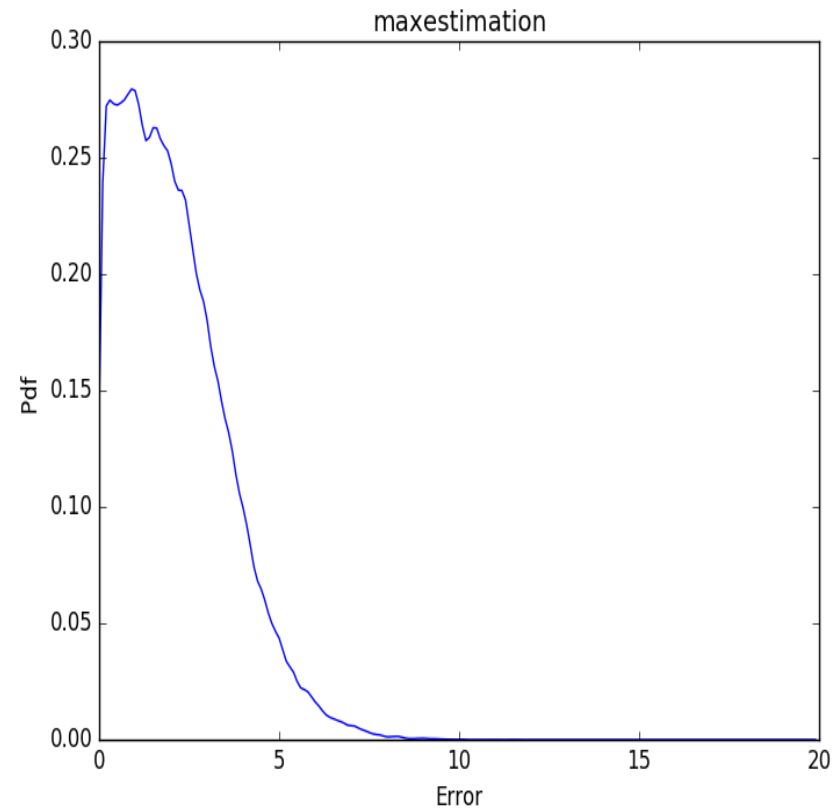
50000 trials for each point

95% Confidence interval of  $|r_t - \hat{r}_t|$

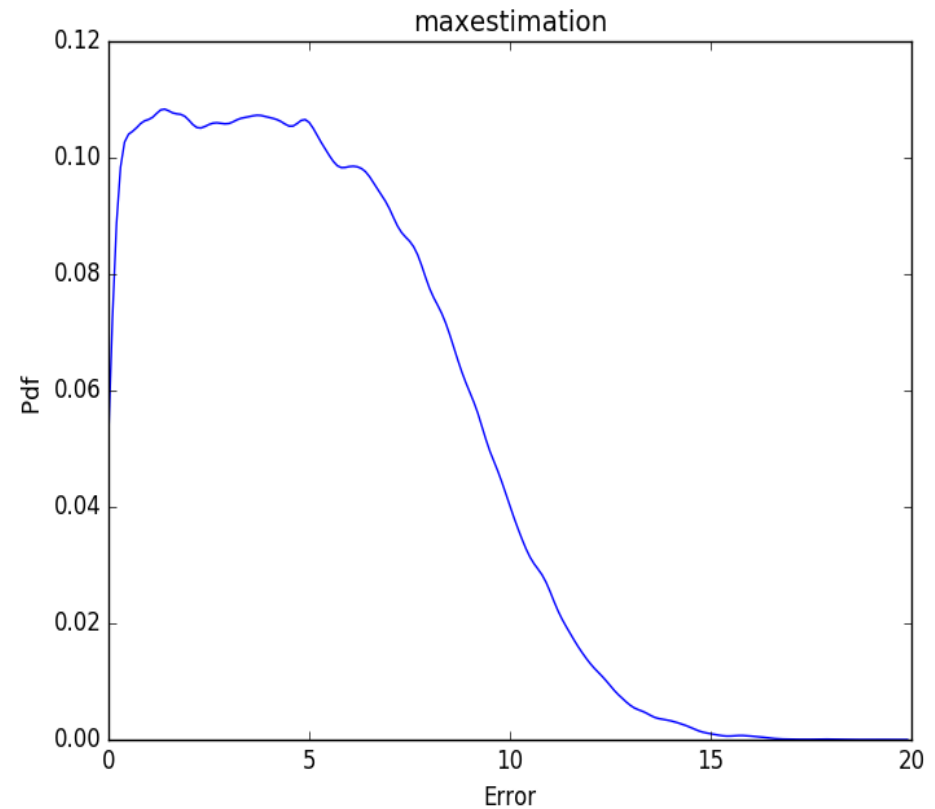
Antenna type	Confidence interval
Directional	$2.080484 \pm 0.002809$
Isotropic	$4.972427 \pm 0.006020$

# Simulation Results

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Pdf of  $|r_t - \hat{r}_t|$   
System with directional antennas



Pdf of  $|r_t - \hat{r}_t|$   
System with isotropic antennas

# Simulation results

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In both results, BMSE and MSE of directional antennas systems are smaller.  
The error is of system with directional antennas distributed more closely to 0.  
System with directional antennas perform better.



# Experimental Implement

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Experimental setup consists of three components

- Alpha network card AWUS036NHA
- Local processing units: Intel Next Unit of Computing
- Cloud server

NIC operates in monitor mode to capture Wi-Fi packets

Local processing units parses information: MAC addresses, RSSI values, time stamps

RSSI values are employed as basis for inference

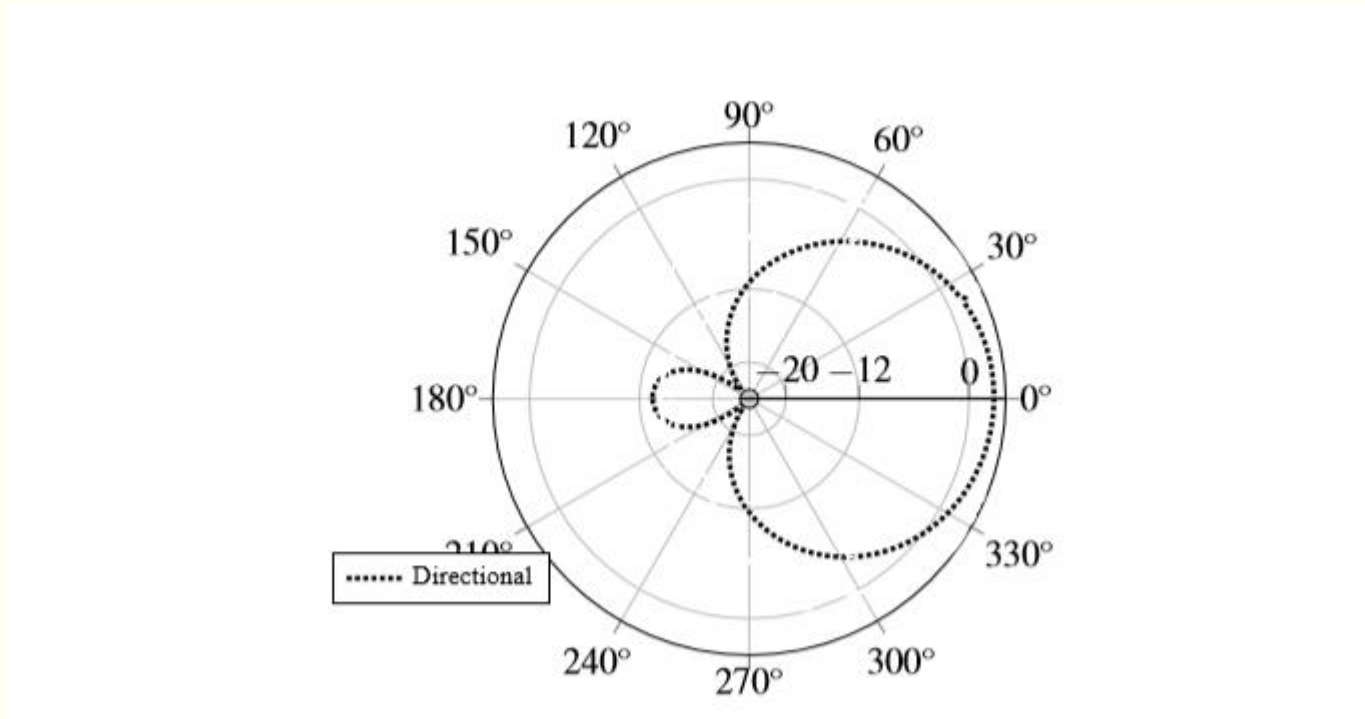


Fig. This graph depicts directional antenna radiation patterns

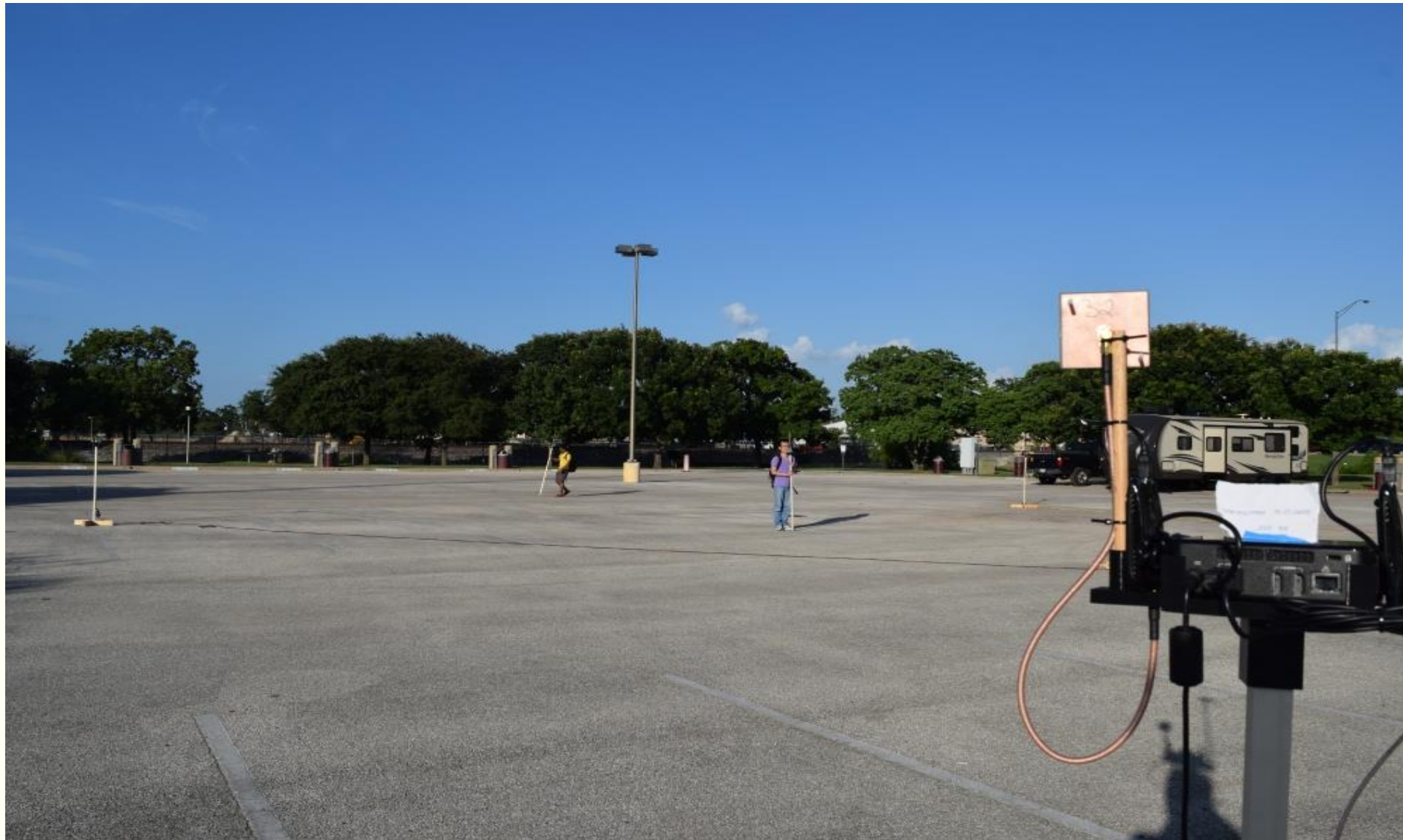
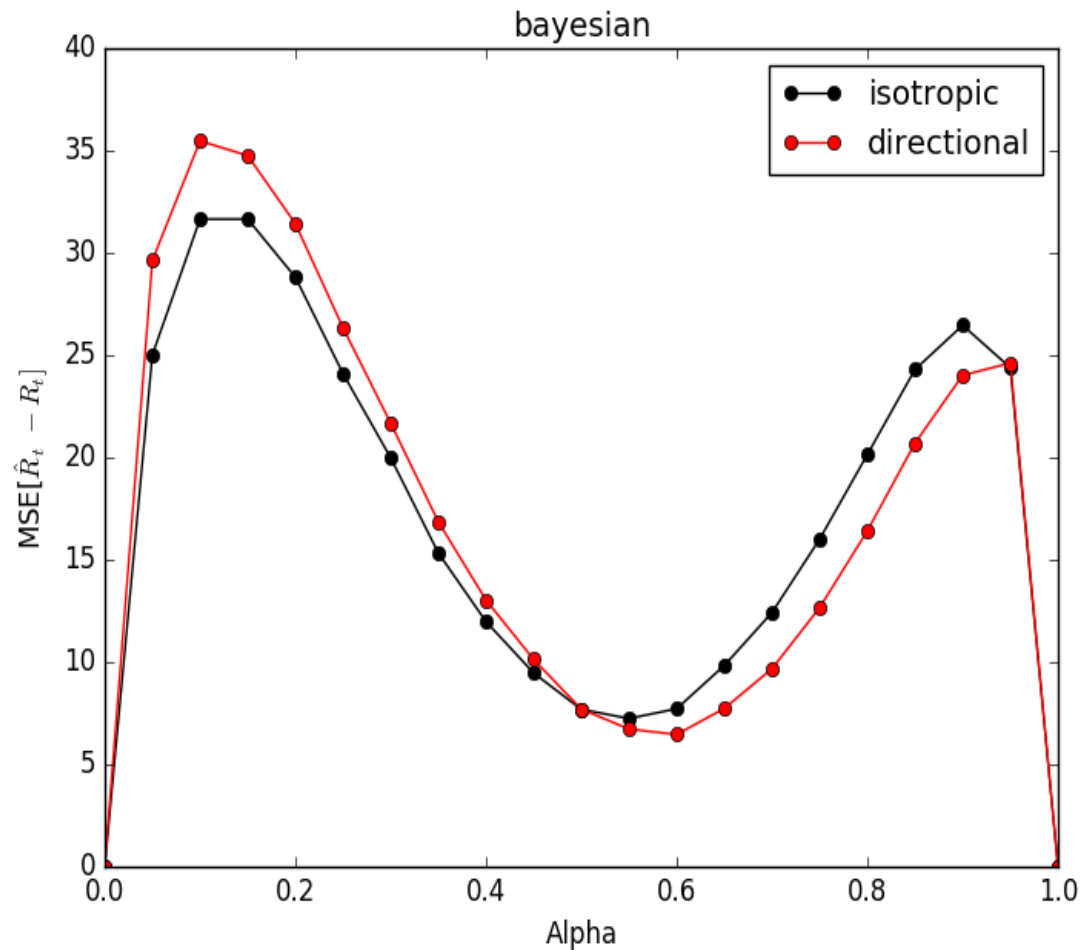


Figure: The place we conduct our experiment



Fig. This figure highlights the site used for the experiments and marks locations of experiment data.

# Experiment Results



Bayes estimator

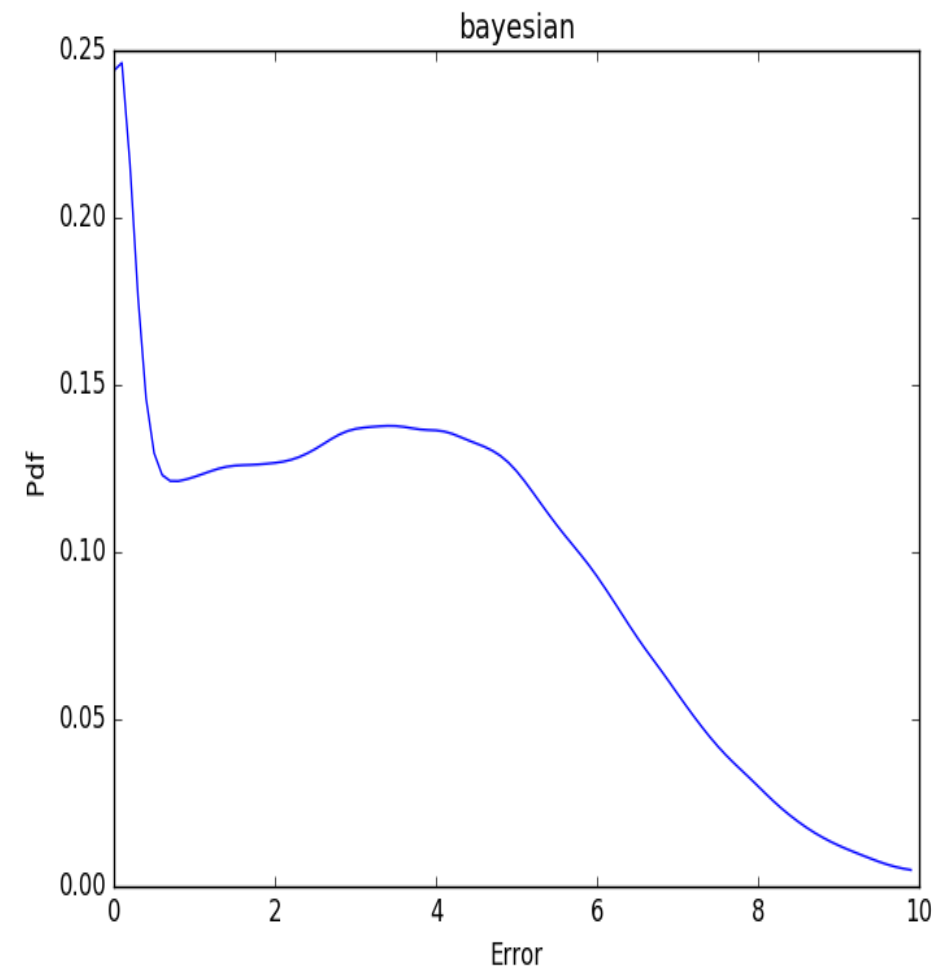
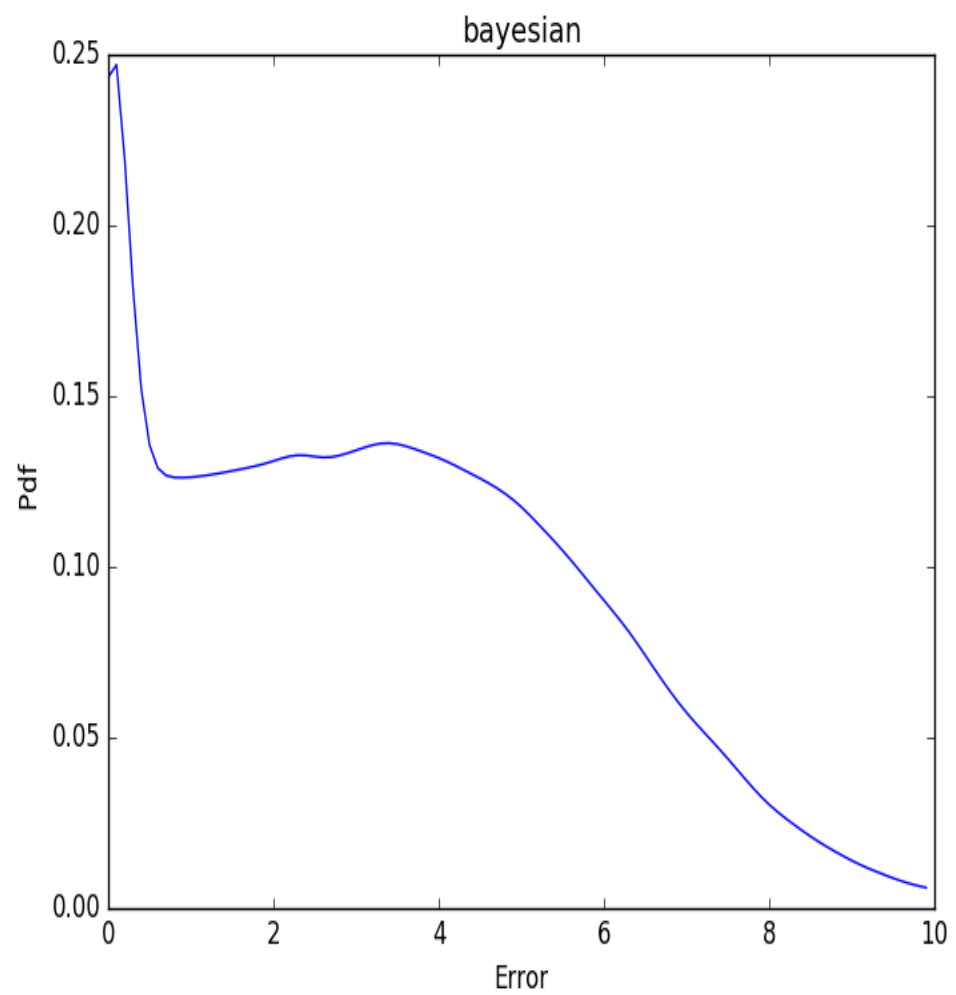
Intensity  $\lambda = 32$

21 points on each curve

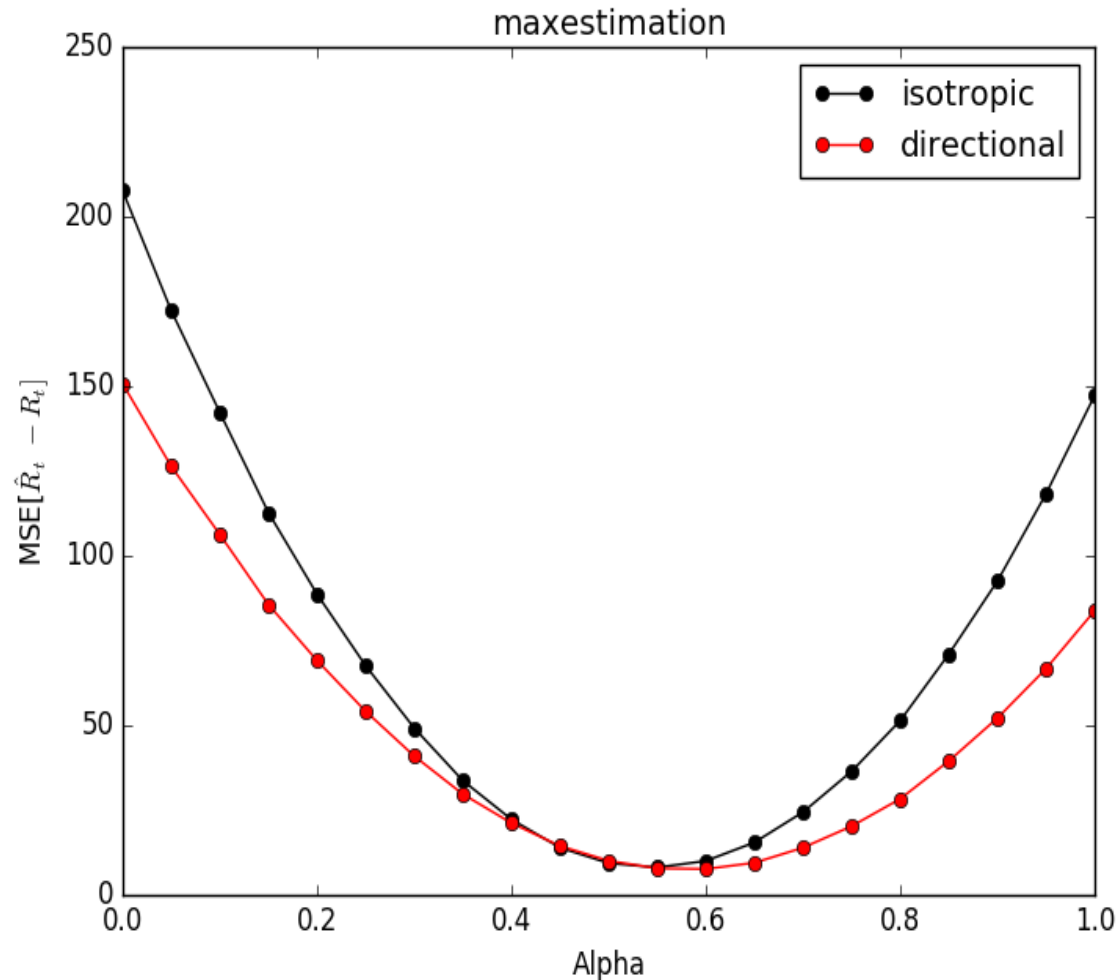
10000 trials for each point

95% Confidence interval of  $|r_t - \hat{r}_t|$

Antenna type	Confidence interval
Directional	$3.317535 \pm 0.010430$
Isotropic	$3.331094 \pm 0.010274$



# Experiment result



Maximum likelihood estimator  
Intensity  $\lambda = 32$   
21 points on each curve  
10000 trials for each point

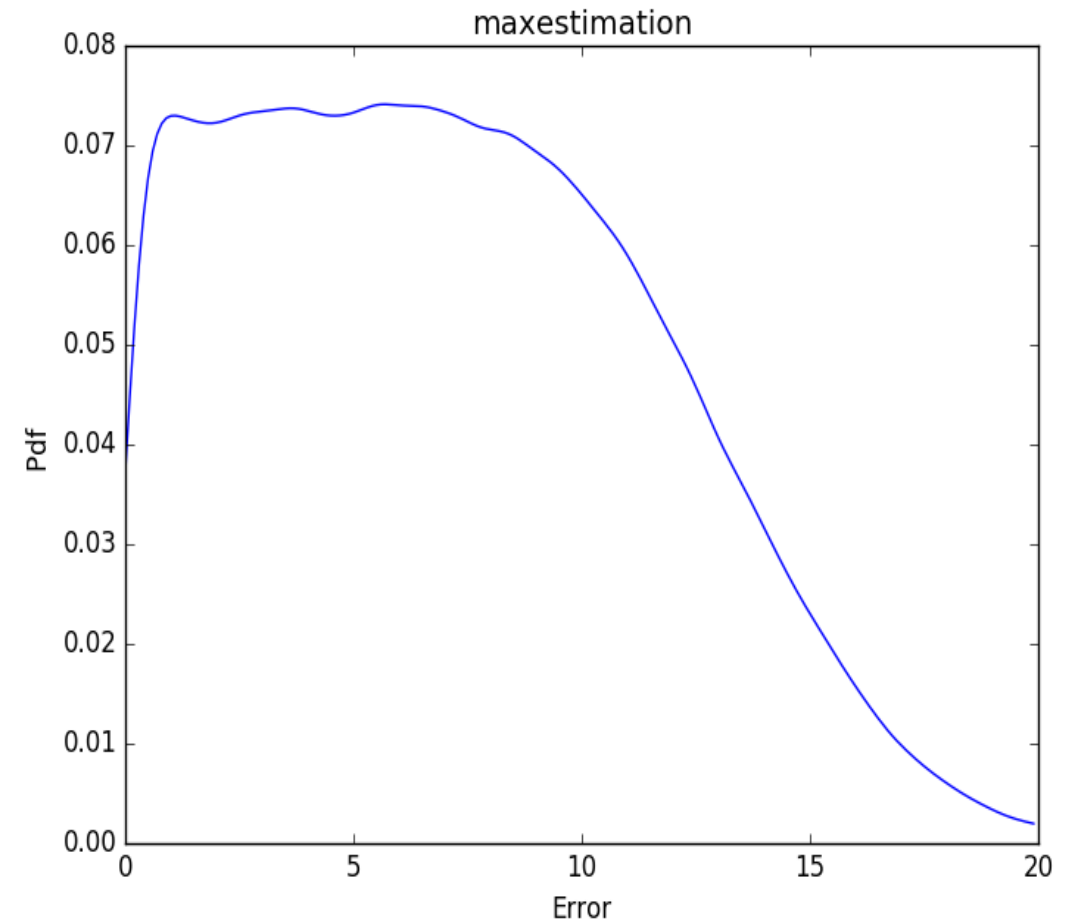
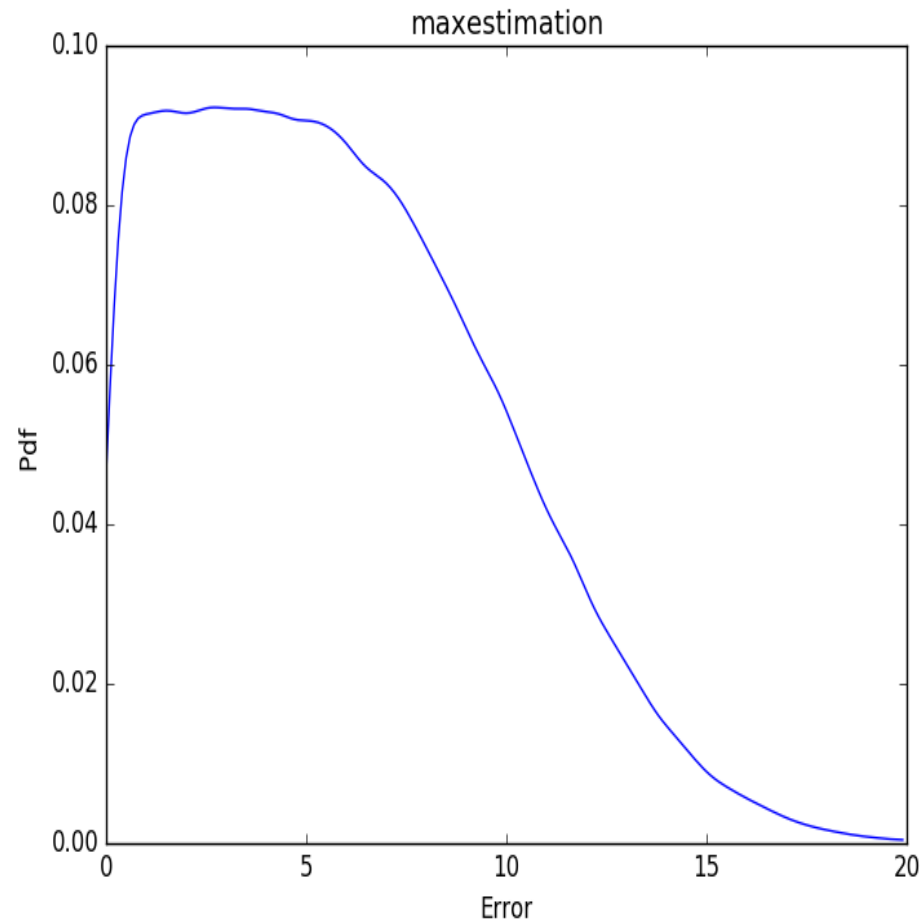
95% Confidence interval of  $|r_t - \hat{r}_t|$

Antenna type	Confidence interval
Directional	$5.881027 \pm 0.016484$
Isotropic	$7.144900 \pm 0.019182$



# Experiment Results

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# Conclusion

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The directional antenna do improve the performance of estimator accuracy.

This work may extend to be used in track specific device.

# Thank You

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