

# 20200929 Homework

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1. Prove that  $Z = \max\{H(X|Y), H(Y|X)\}$  is a distance of discrete random variables  $X$  and  $Y$ .

$$\textcircled{1} Z(X, Y) = \max\{H(X|Y), H(Y|X)\} = \max\{H(X), H(Y)\} - I(X; Y) = Z(Y, X)$$

$$\textcircled{2} \text{ since } H(X) \geq 0, \text{ so } Z(X, Y) \geq 0$$

$$\textcircled{3} Z(X, Y) = 0 \text{ only if and only if } H(X|Y) = H(Y|X) = 0$$

$$\therefore H(X) - I(X; Y) = H(Y) - I(X; Y) \Leftrightarrow H(X) = H(Y)$$

$\therefore X = Y$  or there is a one-to-one function mapping from  $X$  to  $Y$ .

$$\textcircled{4} Z(X, Y) + Z(Y, T) = \max\{H(X), H(Y)\} + \max\{H(Y), H(T)\} - I(X; Y) - I(Y; T)$$

$$Z(X, T) = \max\{H(X), H(T)\} - I(X; T)$$

$$\text{if } H(X) \geq H(Y) \geq H(T): Z(X, Y) + Z(Y, T) = H(X|Y) + H(Y|T) \geq H(X|Y, T) + I(X; Y|T) = H(X|T) = Z(X, T)$$

$$\text{if } H(X) \geq H(T) \geq H(Y): Z(X, Y) + Z(Y, T) = H(X|Y) + H(T|Y) \geq H(X|Y, T) + H(T|Y) \geq H(X|T) + I(X; Y|T) = H(X|T) = Z(X, T)$$

$$\text{if } H(T) \geq H(X) \geq H(Y): Z(X, Y) + Z(Y, T) = H(X|Y) + H(T|Y) \geq H(X|Y, T) + H(T|Y) \geq H(X|T) + I(X; Y|T) = H(X|T) = Z(X, T)$$

The other cases where  $H(X) \leq H(Y)$  are similar.

$\therefore Z = \max\{H(X|Y), H(Y|X)\}$  is a metric for all  $X, Y$ .

2. Prove chain rules.

$$\textcircled{1} I(X_1, X_2, \dots, X_n; Y) = H(X_1, \dots, X_n) - H(X_1, X_2, \dots, X_n|Y)$$

$$= \sum_{i=1}^n H(X_i|X_1, \dots, X_{i-1}) - \sum_{i=1}^n H(X_i|X_1, X_2, \dots, X_{i-1}, Y)$$

$$= \sum_{i=1}^n I(X_i; Y|X_1, X_2, \dots, X_{i-1})$$

$$\textcircled{2} I(X_1, X_2, \dots, X_n; Y|Z) = H(X_1, X_2, \dots, X_n|Z) - H(X_1, X_2, \dots, X_n|Y, Z)$$

$$= \sum_{i=1}^n H(X_i|X_1, \dots, X_{i-1}, Z) - \sum_{i=1}^n H(X_i|X_1, X_2, \dots, X_{i-1}, Y, Z)$$

$$= \sum_{i=1}^n I(X_i; Y|X_1, X_2, \dots, X_{i-1}, Z)$$

3.  $a_i > 0, b_i > 0$ , prove log-sum inequality.

$$\text{let } a_i' = \frac{a_i}{\sum_{j=1}^n a_j}, b_i' = \frac{b_i}{\sum_{j=1}^n b_j}. \text{ thus } 0 \leq D(a' \| b') = \sum_{i=1}^n a_i' \log \frac{a_i'}{b_i'} = \sum_{i=1}^n \frac{a_i}{\sum_{j=1}^n a_j} \log \frac{a_i}{b_i} \cdot \frac{\sum_{j=1}^n b_j}{\sum_{j=1}^n a_j}$$

$$= \frac{1}{\sum_{j=1}^n a_j} \left[ \sum_{i=1}^n a_i \log \frac{a_i}{b_i} - \sum_{i=1}^n a_i \cdot \log \frac{\sum_{j=1}^n a_j}{\sum_{j=1}^n b_j} \right]$$

$$\therefore \sum_{i=1}^n a_i \log \frac{a_i}{b_i} \geq \left( \sum_{i=1}^n a_i \right) \log \frac{\sum_{i=1}^n a_i}{\sum_{i=1}^n b_i}$$

4. prove  $D(p \| q)$  is convex function of the pair  $(p, q)$ .

$$\text{let } a_1 = \lambda p_1, a_2 = (1-\lambda)p_2, b_1 = \lambda q_1, b_2 = (1-\lambda)q_2$$

$$\text{then } D(a_1 + a_2 \| b_1 + b_2) = \sum_x (a_1 + a_2) \log \frac{a_1 + a_2}{b_1 + b_2} \leq \sum_x a_1 \log \frac{a_1}{b_1} + \sum_x a_2 \log \frac{a_2}{b_2} = \lambda \sum_x p_1 \log \frac{p_1}{q_1} + (1-\lambda) \sum_x p_2 \log \frac{p_2}{q_2}$$

$$= \lambda D(p_1 \| q_1) + (1-\lambda) D(p_2 \| q_2)$$

$$\therefore D(\lambda p_1 + (1-\lambda)p_2 \| \lambda q_1 + (1-\lambda)q_2) \leq \lambda D(p_1 \| q_1) + (1-\lambda) D(p_2 \| q_2)$$



5. prove.  $H(p)$  is a concave function of  $p$ .

$$H(\lambda p_1 + (1-\lambda)p_2) = \lambda \sum_{i=1}^n p_{1i} \log \frac{1}{\lambda p_{1i} + (1-\lambda)p_{2i}} + (1-\lambda) \sum_{i=1}^n p_{2i} \log \frac{1}{\lambda p_{1i} + (1-\lambda)p_{2i}}$$

$$\text{let } q_i = \lambda p_{1i} + (1-\lambda)p_{2i}, \quad \sum_{i=1}^n p_{1i} = 1, \quad \sum_{i=1}^n q_i \leq 1$$

$$\therefore H(\lambda p_1 + (1-\lambda)p_2) = \lambda \sum_{i=1}^n p_{1i} \log \frac{1}{q_i} + (1-\lambda) \sum_{i=1}^n p_{2i} \log \frac{1}{q_i}$$

$$\because \ln x \leq x-1 \quad (x > 0)$$

$$\therefore \ln \frac{q_i}{p_{1i}} \leq \frac{q_i}{p_{1i}} - 1 \Leftrightarrow p_{1i} \ln \frac{1}{p_{1i}} \leq p_{1i} \ln \frac{1}{q_i} + q_i - p_{1i}, \text{ summing on } i \text{ gives:}$$

$$\sum_{i=1}^n p_{1i} \ln \frac{1}{p_{1i}} \leq \sum_{i=1}^n p_{1i} \ln \frac{1}{q_i} + \sum_{i=1}^n q_i - \sum_{i=1}^n p_{1i} \leq \sum_{i=1}^n p_{1i} \ln \frac{1}{q_i}$$

$$\therefore H(\lambda p_1 + (1-\lambda)p_2) \geq \lambda \sum_{i=1}^n p_{1i} \log \frac{1}{p_{1i}} + (1-\lambda) \sum_{i=1}^n p_{2i} \log \frac{1}{p_{2i}} = \lambda H(p_1) + (1-\lambda) H(p_2)$$

6. Prove the Fano's inequality.

$$\text{def: } P_e = P_r\{X \neq \hat{X}\}, \quad Y = \begin{cases} 0, & \text{if } X = \hat{X} \\ 1, & \text{if } X \neq \hat{X} \end{cases} \quad P(Y=1) = P_e, \quad H(Y|X, \hat{X}) = 0, \quad H(X|\hat{X}=\hat{x}, Y=0) = 0$$

$$H(X|\hat{X}) = H(X|\hat{X}) + H(Y|X, \hat{X}) = H(X, Y|\hat{X}) = H(Y|\hat{X}) + H(X|Y, \hat{X}) \leq H(Y) + \sum_{\hat{x} \in \mathcal{X}} [P(Y=0, \hat{X}=\hat{x}) \cdot H(X|\hat{X}=\hat{x}, Y=0) + P(Y=1, \hat{X}=\hat{x}) \cdot H(X|\hat{X}=\hat{x}, Y=1)]$$

$$\leq H(Y) + P(Y=1) \log(|\mathcal{X}| - 1) = h_b(P_e) + P_e \log(|\mathcal{X}| - 1)$$

7. If  $X_1 \rightarrow \dots \rightarrow X_6$  is a Markov chain, judge and show whether  $(X_3, X_6) \rightarrow (X_2, X_5) \rightarrow (X_1, X_4)$  is also a Markov chain.

$$\text{Only need to prove: } p(x_1, x_2, \dots, x_6) p(x_2, x_5) = p(x_2, x_3, x_5, x_6) p(x_1, x_2, x_4, x_5) \quad \text{--- (1)}$$

$$X_1 \rightarrow X_2 \rightarrow \dots \rightarrow X_6 \Rightarrow p(x_1, x_2, \dots, x_6) p(x_2, x_5) = p(x_1, x_2) p(x_2, x_3) p(x_3, x_4) p(x_4, x_5) p(x_5, x_6) \quad \text{--- (2)}$$

According to the proposition of Markov subchains,  $X_2 \rightarrow X_3 \rightarrow X_5 \rightarrow X_6$  and  $X_1 \rightarrow X_2 \rightarrow X_4 \rightarrow X_5$  form a Markov chain, thus:

$$p(x_2, x_3, x_5, x_6) p(x_2, x_5) = p(x_2, x_3) p(x_3, x_5) p(x_5, x_6) \quad \text{--- (3)}$$

$$p(x_1, x_2, x_4, x_5) p(x_2, x_5) = p(x_1, x_2) p(x_2, x_4) p(x_4, x_5) \quad \text{--- (4)}$$

Based on (1)-(4), we only need to prove (5):

$$p(x_3, x_5) \cdot p(x_2, x_4) = p(x_2, x_5) p(x_3, x_4) \quad \text{--- (5)}$$

which is equivalent to prove (6):

$$p(x_5|x_3) p(x_4|x_2) = p(x_4|x_2) p(x_5|x_2) \quad \text{when } p(x_3) > 0, p(x_2) > 0 \quad \text{--- (6)}$$

### 1. Distance of discrete random variables.

Prove that  $\max \{H(X|Y), H(Y|X)\}$  is a distance of discrete random variables  $X$  and  $Y$ .

### 2. Prove the following two chain rules.

\* **Proposition 2.26 (Chain Rule for Mutual Information).**

$$I(X_1, X_2, \dots, X_n; Y) = \sum_{i=1}^n I(X_i; Y | X_1, X_2, \dots, X_{i-1})$$

\* **Proposition 2.27 (Chain Rule for Conditional Mutual Information)**

For random variables  $X_1, X_2, \dots, X_n, Y$  and  $Z$ ,

$$I(X_1, X_2, \dots, X_n; Y | Z) = \sum_{i=1}^n I(X_i; Y | X_1, X_2, \dots, X_{i-1}, Z)$$

### 3. Log-Sum inequality: For non-negative numbers $a_i$ and $b_i$ , prove

$$\sum_{i=1}^n a_i \log \frac{a_i}{b_i} \geq \left( \sum_{i=1}^n a_i \right) \log \frac{\sum_{i=1}^n a_i}{\sum_{i=1}^n b_i}$$

### 4. Convex Relative Entropy

If  $(p_1, q_1)$  and  $(p_2, q_2)$  are pairs of probability mass functions then

$$D(\lambda p_1 + (1 - \lambda)p_2 \| \lambda q_1 + (1 - \lambda)q_2) \leq \lambda D(p_1 \| q_1) + (1 - \lambda)D(p_2 \| q_2)$$

for all  $0 \leq \lambda \leq 1$ . That is,  $D(p \| q)$  is **convex function** of the pair  $(p, q)$ .

### 5. Concave Entropy

Let  $p$  be the probability mass function of discrete random variable  $X$ . Here  $H(X)$  is denoted by  $H(p)$ . Prove that

$$H(\lambda p_1 + (1 - \lambda)p_2) \geq \lambda H(p_1) + (1 - \lambda)H(p_2).$$

That is,  $H(p)$  is a **concave function** of  $p$ .

### 6. Prove the Fano's inequality.

7. If  $X_1 \rightarrow X_2 \rightarrow X_3 \rightarrow X_4 \rightarrow X_5 \rightarrow X_6$  is a Markov chain, **judge and show** whether  $(X_3, X_6) \rightarrow (X_2, X_5) \rightarrow (X_1, X_4)$  is also a Markov chain.

### 8. Remark. Convex and Concave mutual information

Mutual Information  $I(X; Y)$  can be expressed by a function of input distribution  $p(x)$  and transition distribution  $p(y|x)$ , i.e.,

$$I(X; Y) = f(p(x), p(y|x)).$$

A. For given input distribution  $p(x)$ , we say that  $I(X; Y)$  is **convex** of transition distribution  $p(y|x)$ .

B. For given transition distribution  $p(y|x)$ , we say that  $I(X; Y)$  is **concave** of input distribution  $p(x)$ .

#### Convex 8A can be explained by:

Let  $p_1(x, y) = p(x)p_1(y|x)$  and  $p_2(x, y) = p(x)p_2(y|x)$  be two joint distributions, we define

- $p(x, y) \triangleq \lambda p_1(x, y) + (1 - \lambda)p_2(x, y) = \lambda p(x)p_1(y|x) + (1 - \lambda)p(x)p_2(y|x) = p(x)\{\lambda p_1(y|x) + (1 - \lambda)p_2(y|x)\}$
- $q(x, y) \triangleq \lambda q_1(x, y) + (1 - \lambda)q_2(x, y) \triangleq \lambda p(x)p_1(y) + (1 - \lambda)p(x)p_2(y) = p(x)\{\lambda p_1(y) + (1 - \lambda)p_2(y)\}$
- $I_\lambda(X; Y) \triangleq D(p(x, y) \| q(x, y))$
- $I_1(X; Y) \triangleq D(p_1(x, y) \| q_1(x, y)) = D(p_1(x, y) \| p(x)p_1(y))$
- $I_2(X; Y) \triangleq D(p_2(x, y) \| q_2(x, y)) = D(p_2(x, y) \| p(x)p_2(y))$

Then, for given input distribution  $p(x)$ ,  $I(X; Y)$  is convex of transition distribution  $p(y|x)$ , i.e.,

$$I_{\lambda}(X; Y) \leq \lambda I_1(X; Y) + (1 - \lambda) I_2(X; Y) \text{ i.e.}$$

$$D(p(x, y) || q(x, y)) \leq \lambda D(p_1(x, y) || q_1(x, y)) + (1 - \lambda) D(p_2(x, y) || q_2(x, y))$$