Chaper 3 homework

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2.

(a)

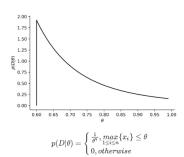
$$L(D| heta) = \left\{ egin{aligned} rac{1}{ heta^n}, \max_{1 \leq i \leq n} \{x_i\} \leq heta \ 0, otherwise \end{aligned}
ight.$$

Consider all heta that satisfy $heta \geq \max_{1 \leq i \leq n} \{x_i\}$, we have

$$L(\theta) = \frac{1}{\theta^n} \leq \frac{1}{\max\limits_{1 \leq i \leq n} \{x_i\}}$$

Therefore, $\hat{ heta} = \max_{1 \leq i \leq n} \{x_i\}$

(b)



 $p(D^n| heta)$ is merely relevant to $\max_{1 \le i \le n} \{x_i\}$ and the sample number n

4.

$$\begin{split} P(x_1, x_2, \cdots, x_n | \theta) &= \prod_{k=1}^n \prod_{i=1}^d \theta_i^{x_{ki}} (1 - \theta_i^{1-x_{ki}}) \\ L(\theta) &= \sum_{k=1}^n \sum_{i=1}^d [x_{ki} ln \theta_i + (1 - x_{ki}) ln (1 - \theta_i)] \\ \frac{\partial L}{\partial \theta_i} &= \sum_{k=1}^n \sum_{i=1}^d (\frac{x_{ki}}{\theta_i} - \frac{1 - x_{ki}}{1 - \theta_i}) \end{split}$$

Let $\frac{\partial L}{\partial \theta_i} = 0$

$$(1-\hat{ heta_i})\sum_{k=1}^n x_{ki} = n\hat{ heta_i} - \hat{ heta_i}\sum_{k=1}^n x_{ki}$$
 $\hat{ heta_i} = rac{1}{n}\sum_{k=1}^n x_k$

21.

When $p(\theta|D^n)$ converge, we have

$$\lim_{n\to\infty}p(\theta|D^n)=\lim_{n\to\infty}\frac{p(x_n|\theta)p(\theta|D^{n-1})}{\int p(x_n|\theta)p(\theta|D^{n-1})d\theta}=\lim_{n\to\infty}p(\theta|D^{n-1})$$

From equation above, we get

$$p(x_n|\theta) = \int p(x_n|\theta)(\theta|D^{n-1})d\theta \tag{1}$$

$$\lim_{n \to \infty} p(\theta|D^n) = p(\theta) \neq 0 \tag{2}$$

Also assume that series x_n converges to x^st . To formulate this sentence, we have

$$\lim_{n \to \infty} x_n = x^* \tag{3}$$

From equation(1) to (3), we get

$$\begin{split} p(x^*|\theta) &= \lim_{n \to \infty} p(x_n|\theta) \\ &= \lim_{n \to \infty} \int p(x^n|\theta) p(\theta|D^{n-1}) d\theta \\ &= \lim_{n \to \infty} \int p(x^n|\theta) p(\theta) d\theta \\ &= \int \lim_{n \to \infty} p(x^n|\theta) p(\theta) d\theta \\ &= \int p(x^*|\theta) p(\theta) d\theta \\ &= p(x^*) \end{split}$$

To guarantee the convergence, variables satisfy

•
$$\lim_{n\to\infty} p(\theta|D^n) = p(\theta) \neq 0$$

• $p(x|\theta) = p(x)$

•
$$p(x|\theta) = p(x)$$

37.

Recall that

$$\Sigma(\beta) = (1 - \beta)\Sigma + \beta I \tag{77}$$

We can see that

$$\lim_{\beta \to 1} \Sigma(\beta) = I$$

$$\lim_{\beta \to 1} (1 - \beta)\sigma_{ii} + \beta = 1$$

$$\sigma_{ii} = \lim_{\beta \to 0} \frac{1 - \beta}{1 - \beta} = 1$$

38.

(a)

$$J_1(w) = \frac{(\mu_1 - \mu_2)^2}{\sigma_1^2 + \sigma_2^2} = \frac{w^T (\boldsymbol{\mu_1} - \boldsymbol{\mu_2})^2 w}{w^T (\Sigma_1 + \Sigma_2) w} = \frac{w^T S_B w}{w^T S_W w}$$

We can get the maximun $J_1(w)$ when

$$S_B w = \lambda S_W w$$

Note that S_B is the cross product of two identical vector. S_B always directs to m_1-m_2 where m_1,m_2 refer to the different mean value of two class. In other words, using constant scalar c, we gain

 $S_B w = c(m_1 - m_2).$

Without scalar, we have

$$w = S_W^{-1}(m_1 - m_2) = (\Sigma_1 + \Sigma_2)^{-1}(oldsymbol{\mu_1} - oldsymbol{\mu_2})$$

(b)

$$J_2(w) = \frac{(\mu_1 - \mu_2)^2}{P(w_1)\sigma_1^2 + P(w_2)\sigma_2^2} = \frac{w^T(\mu_1 - \mu_2)^2 w}{w^T(P(w_1)\Sigma_1 + P(w_2)\Sigma_2)w} = \frac{w^TS_Bw}{w^TS_Ww}$$

Similar to (a), without scalar, we have

$$w = S_W^{-1}(P(w_1)m_1 - P(w_2)m_2) = [P(w_1)\Sigma_1 + P(w_2)\Sigma_2]^{-1}(\boldsymbol{\mu_1} - \boldsymbol{\mu_2})$$

(c)

$$\begin{split} J(w) &= \frac{|\tilde{m_1} - \tilde{m_2}|^2}{\tilde{s_1}^2 + \tilde{s_2}^2} \\ &= \frac{(\mu_1 - \mu_2)^2}{\sum\limits_{y \in D_1} (y - \mu_1)^2 + \sum\limits_{y \in D_2} (y - \mu_2)^2} \\ &= \frac{(\mu_1 - \mu_2)^2}{n(D)(\frac{n(D_1)}{n(D)}(y - \mu_1)^2 + \frac{n(D_2)}{n(D)}(y - \mu_2)^2)} \\ &\approx \frac{(\mu_1 - \mu_2)^2}{n(D)(P(w_1)\sigma_1^2 + P(w_2)\sigma_2^2)} \end{split}$$

where n(D) denotes the total number of the sample set. From the result above, we can conclude that (b) is more similar to J(w).

source code