2017年概率 A 期末试卷答案

一填空题(3分)

1.
$$\Omega = \{(x,y) | x+y = 1, \exists 1 \ 0 < x < 1, 0 < y < 1 \}$$
; 2. 0.4; 3. 1/9;

4.
$$f(x,y) = \begin{cases} 6, & (0 < x < 1, x^2 < y < x) \\ 0, & \text{ } \sharp \dot{\Xi} \end{cases}$$
; 5. 1/2

二选择题(3分)

B; C; C; D; D.

 \equiv

11. $(10 \, \text{分})$ 解: 设A表示考生会解这道题,B表示考生选出正确答案,则有

(1) 根据全概率公式可得

$$P(A)=0.8,\; P(\overline{A})=0.2, P(B\Big|A)=1, P(B\Big|\overline{A})=\frac{1}{4}=0.25$$
 ,

$$P(B) = P(A) \cdot P(B \, \middle| \, A) + P(\overline{A}) \cdot P(B \, \middle| \, \overline{A}) = 0.8 \times 1 + 0.2 \times 0.25 = 0.85 \ .$$

(2) 根据条件概率公式可得

$$P(A|B) = \frac{P(AB)}{P(B)} = \frac{P(A) \cdot P(B|A)}{P(B)} = \frac{0.8 \times 1}{0.85} = \frac{16}{17} \approx 0.941$$

12. (6分)

13.解: (1) (6分)

$$f_{X}(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_{0}^{2} \left(x^{2} + \frac{1}{3}xy\right) dy, & 0 \le x \le 1 \\ 0, & \text{!! } \stackrel{\triangle}{\text{!!}} \end{cases} = \begin{cases} 2x^{2} + \frac{2}{3}x, & 0 \le x \le 1 \\ 0, & \text{!! } \stackrel{\triangle}{\text{!!}} \end{cases}$$

$$f_{Y}(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_{0}^{1} \left(x^{2} + \frac{1}{3}xy\right) dx, & 0 \le y \le 2 \\ 0, & \text{!!} \dot{\Xi} \end{cases} = \begin{cases} \frac{y}{6} + \frac{1}{3}, & 0 \le y \le 2 \\ 0, & \text{!!} \dot{\Xi} \end{cases}$$

由于 $f(x,y) \neq f_X(x)f_Y(y)$, 所以X 和Y不独立.

(2)
$$(4 \%) P(X + Y \ge 1) = \iint_D f(x, y) dxdy = 1 - \int_0^1 dx \int_0^{-x+1} \left(x^2 + \frac{1}{3}xy\right) dy = \frac{65}{72}$$

14. (10分) 解: 由独立性可得

(X,Y)	(1, 2)	(1, 4)	(3, 2)	(3, 4)
P(X=x,Y=y)	0.18	0.12	0.42	0.28
X + Y	3	5	5	7
X - Y	- 1	- 3	1	- 1

所以
$$Z = X + Y$$
 的分布律为 $\begin{pmatrix} 3 & 5 & 7 \\ 0.18 & 0.54 & 0.28 \end{pmatrix}$, $W = X - Y$ 的分布律为 $\begin{pmatrix} -3 & -1 & 1 \\ 0.12 & 0.46 & 0.42 \end{pmatrix}$

15. (10 \(\frac{1}{2}\))
$$E(X) = \int_0^1 dx \int_0^1 x(2-x-y)dy = \frac{5}{12}, \quad E(X^2) = \int_0^1 dx \int_0^1 x^2(2-x-y)dy = \frac{1}{4}$$

$$D(X) = \frac{1}{4} - \left(\frac{5}{12}\right)^2 = \frac{11}{144}, \quad \text{in which } E(Y) = \frac{5}{12}, \quad D(Y) = \frac{11}{144}$$

$$E(XY) = \int_0^1 dx \int_0^1 xy(2-x-y)dy = \frac{1}{6}, \quad Cov(X,Y) = E(XY) - E(X)E(Y) = -\frac{1}{144}$$

所以
$$X$$
和 Y 的相关系数为: $\rho_{XY} = \frac{Cov(X,Y)}{\sqrt{DX}\sqrt{DY}} = -\frac{1}{11}$.

16. (8分)解: (1)(3分) X 服从二项分布,参数: n = 100, p = 0.2, 即 $X \sim B(100, 0.2)$,

其概率分布为

$$P(X = k) = C_{100}^{k} 0.2^{k} 0.8^{100-k}, \quad k = 0, 1, \dots, 100;$$

(2) (5 分)
$$E(X) = np = 20$$
, $D(X) = np(1-p) = 16$, 根据德莫弗—拉普拉斯定理

$$P\left\{14 \le X \le 30\right\} = P\left\{\frac{14 - 20}{4} \le \frac{X - 20}{4} \le \frac{30 - 20}{4}\right\} = P\left\{-1.5 \le \frac{X - 20}{4} \le 2.5\right\}$$

$$\approx \Phi(2.5) - \Phi(-1.5) = \Phi(2.5) - [1 - \Phi(1.5)]$$

$$=\Phi(2.5)+\Phi(1.5)-1=0.994+0.933-1=0.927$$
.

17. (8分) **解**: 似然函数为
$$L(x_1,\dots,x_n,\lambda) = \prod_{i=1}^n \lambda \alpha x_i^{\alpha-1} e^{-\lambda x_i^{\alpha}}$$
,

$$(x_i > 0, i = 1, \dots, n) \implies L(x_1, \dots, x_n, \lambda) = \lambda^n \alpha^n e^{-\lambda \sum_{i=1}^n x_i^{\alpha}} \prod_{i=1}^n x_i^{\alpha-1},$$

取对数,有
$$\ln L = n \ln \lambda + \ln \alpha^n - \lambda \sum_{i=1}^n x_i^{\alpha} + (\alpha - 1) \sum_{i=1}^n \ln x_i$$

令
$$\frac{d \ln L}{d \lambda} = \frac{n}{\lambda} - \sum_{i=1}^{n} x_i^{\alpha} = 0 \,, \ \text{ 便得 } \lambda \text{ 的最大似然估计量为: } \hat{\lambda} = \frac{n}{\sum_{i=1}^{n} x_i^{\alpha}} \,.$$

18. $(8 \%) \bar{x} = 403$, s = 6.16, $X \sim N(400, \sigma^2)$, $(\alpha = 0.05)$,

因
$$\left| \frac{\overline{X} - \mu}{s / \sqrt{n}} \right| = \left| \frac{403 - 400}{6.16 / \sqrt{16}} \right| \approx 1.948 < t_{0.025} (16 - 1) = 2.1315$$
,所以合格。