HW10

04/07/2024

Q1

(a) The kernel of posterior density for $\beta|y,X$ is not easily recognisable as any know distribution, so analytic solution may not be easily obtained and MH algorithm is then suitable for sampling from the posterior.

Let $q(\beta_{new} \mid \beta_{old})$ be the density of given proposal distribution $N_p(\beta_{old}, \widehat{Cov}(\hat{\beta}_{ML}))$. The algorithm in pseudo code is as follows:

- (i) Set the staring value of $\beta^{(0)}$, e.g. using prior mean b_0 .
- (ii) For t = 1, ..., 1000:
 - 1. Draw β^* from $q(\beta^* \mid \beta^{(t-1)})$.
 - 2. Compute the acceptance probability: $\alpha(\boldsymbol{\beta^*} \mid \boldsymbol{\beta^{(t-1)}}) = min\left(1, \frac{p(\boldsymbol{\beta^*} \mid \boldsymbol{y}, \boldsymbol{X})q(\boldsymbol{\beta^{(t-1)}} \mid \boldsymbol{\beta^*})}{p(\boldsymbol{\beta^{(t-1)}} \mid \boldsymbol{y}, \boldsymbol{X})q(\boldsymbol{\beta^*} \mid \boldsymbol{\beta^{(t-1)}})}\right)$
 - 3. With probability $\alpha(\beta^* \mid \beta^{(t-1)})$ set $\beta^{(t)} = \beta^*$. Else, set $\beta^{(t)} = \beta^{(t-1)}$.
- (iii) Optionally, discard first B iterations as burn-in.
- (iv) Output samples $\beta^{(B+1)}, \ldots, \beta^{(1000)}$
- (b) Point estimates of individual parameters can be approximated from the generated random draws. The posterior mean is approximated by the sample average:

$$\mathbb{E}(\boldsymbol{\beta} \mid \boldsymbol{y}, \boldsymbol{X}) \approx \frac{1}{1000 - B} \sum_{t=B+1}^{1000} \boldsymbol{\beta}^{(t)}$$

(c) The credible interval can be defined via empirical quantiles of the respective random draws, e.g. 90% credible interval:

$$\hat{Q}_{0.05}(\beta \mid y, X) \approx q_{0.05}(\beta^{(B+1)}, \dots, \beta^{(1000)})$$

$$\hat{Q}_{0.95}(\boldsymbol{\beta} \mid \boldsymbol{y}, \boldsymbol{X}) \approx q_{0.95}(\boldsymbol{\beta^{(B+1)}}, \dots, \boldsymbol{\beta^{(1000)}})$$