

HW7

13/06/2024

Q1(a) Likelihood:

$$\begin{aligned}
 L(\lambda, \mu) &= f(\mathbf{x}, \mathbf{y} \mid \lambda, \mu) \\
 &= f(x_1, \dots, x_m \mid \lambda) f(y_1, \dots, y_n \mid \mu) \quad (\mathbf{X}, \mathbf{Y} \text{ independent}) \\
 &= \prod_{i=1}^m f(x_i \mid \lambda) \prod_{j=1}^n f(y_j \mid \mu) \quad \left(X_i \stackrel{i.i.d.}{\sim} \text{Exp}(\lambda), Y_j \stackrel{i.i.d.}{\sim} \text{Exp}(\mu) \right)
 \end{aligned}$$

Log-likelihood:

$$\begin{aligned}
 l(\lambda, \mu) &= \log \left(\prod_{i=1}^m f(x_i \mid \lambda) \prod_{j=1}^n f(y_j \mid \mu) \right) \\
 &= \sum_{i=1}^m \log f(x_i \mid \lambda) + \sum_{j=1}^n \log f(y_j \mid \mu) \\
 &= \sum_{i=1}^m \log(\lambda \exp(-\lambda x_i)) + \sum_{j=1}^n \log(\mu \exp(-\mu y_j)) \\
 &= m \log \lambda - \lambda m \bar{x} + n \log \mu - \mu n \bar{y}
 \end{aligned}$$

Score function:

$$\begin{aligned}
 s(\lambda, \mu \mid \mathbf{x}, \mathbf{y}) &= \left(\frac{\partial l(\lambda, \mu)}{\partial \lambda}, \frac{\partial l(\lambda, \mu)}{\partial \mu} \right)^T \\
 &= \left(\frac{m}{\lambda} - m \bar{x}, \frac{n}{\mu} - n \bar{y} \right)^T \stackrel{!}{=} (0, 0)^T \\
 &\Leftrightarrow \begin{cases} \frac{m}{\lambda} = m \bar{x} \\ \frac{n}{\mu} = n \bar{y} \end{cases} \\
 \Rightarrow (\hat{\lambda}_{ML}, \hat{\mu}_{ML})^T &= \left(\frac{1}{\bar{X}}, \frac{1}{\bar{Y}} \right)^T
 \end{aligned}$$

(b)

$$\begin{aligned}
\mathbb{E}[V(\mathbf{X}, \mathbf{Y})] &= \left(\mathbb{E} \left(\frac{1}{\hat{\lambda}_{ML}} \right), \mathbb{E} \left(\frac{1}{\hat{\mu}_{ML}} \right) \right)^T \\
&= (\mathbb{E}(\bar{X}), \mathbb{E}(\bar{Y}))^T \\
&= \left(\frac{1}{m} \sum_{i=1}^m \mathbb{E}(X_i), \frac{1}{n} \sum_{j=1}^n \mathbb{E}(Y_j) \right)^T \\
&= \left(\frac{1}{\bar{\lambda}}, \frac{1}{\bar{\mu}} \right)^T \quad \left(X_i \stackrel{i.i.d.}{\sim} \text{Exp}(\lambda), Y_j \stackrel{i.i.d.}{\sim} \text{Exp}(\mu) \right) \\
\text{Cov}(V(\mathbf{X}, \mathbf{Y})) &= \text{Cov} \left(\left(\frac{1}{\hat{\lambda}_{ML}}, \frac{1}{\hat{\mu}_{ML}} \right)^T \right) \\
&= \text{Cov}((\bar{X}, \bar{Y})^T) \\
&= \begin{pmatrix} \text{Var}(\bar{X}) & 0 \\ 0 & \text{Var}(\bar{Y}) \end{pmatrix} \quad (\mathbf{X}, \mathbf{Y} \text{ independent}) \\
&= \begin{pmatrix} \frac{1}{m^2} \text{Var}(\sum_{i=1}^m X_i) & 0 \\ 0 & \frac{1}{n^2} \text{Var}(\sum_{j=1}^n Y_j) \end{pmatrix} \\
&= \begin{pmatrix} \frac{1}{m\lambda^2} & 0 \\ 0 & \frac{1}{n\mu^2} \end{pmatrix} \quad \left(X_i \stackrel{i.i.d.}{\sim} \text{Exp}(\lambda), Y_j \stackrel{i.i.d.}{\sim} \text{Exp}(\mu) \right)
\end{aligned}$$

(c) Based on the given assumption on $V(\mathbf{X}, \mathbf{Y})$, we let $\boldsymbol{\theta} = (\frac{1}{\lambda}, \frac{1}{\mu})^T$, $\hat{\boldsymbol{\theta}}_n = V(\mathbf{X}, \mathbf{Y})$, $V(\boldsymbol{\theta}) = \boldsymbol{\Sigma}_V$ for Delta method notation.

$$\begin{pmatrix} \sqrt{m} & 0 \\ 0 & \sqrt{n} \end{pmatrix} (\hat{\boldsymbol{\theta}}_n - \boldsymbol{\theta}) \xrightarrow{d} N_2(\mathbf{0}, V(\boldsymbol{\theta}))$$

Let $h : \mathbb{R}^2 \rightarrow \mathbb{R}^2$, $h(\boldsymbol{\theta}) = h((\frac{1}{\lambda}, \frac{1}{\mu})^T) = (\lambda, \mu)^T$

Now check conditions to apply the Delta method are fulfilled:

$$\begin{aligned}
H(\boldsymbol{\theta}) &= \left(\frac{\partial h(\boldsymbol{\theta})}{\partial \lambda}, \frac{\partial h(\boldsymbol{\theta})}{\partial \mu} \right) \\
&= \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}
\end{aligned}$$

so, H is of full rank $r = 2$ while every row not equal to zero vector. h is component-wise continuously partially differentiable.

Apply Delta method,

$$\begin{aligned}
h(\hat{\boldsymbol{\theta}}_n) &= T(\mathbf{X}, \mathbf{Y}) \\
\begin{pmatrix} \sqrt{m} & 0 \\ 0 & \sqrt{n} \end{pmatrix} (h(\hat{\boldsymbol{\theta}}_n) - h(\boldsymbol{\theta})) &\xrightarrow{d} N_2(0, H(\boldsymbol{\theta})V(\boldsymbol{\theta})H(\boldsymbol{\theta})^T) \\
\begin{pmatrix} \sqrt{m} & 0 \\ 0 & \sqrt{n} \end{pmatrix} (T(\mathbf{X}, \mathbf{Y}) - (\lambda, \mu)^T) &\xrightarrow{d} N_2\left(0, \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix} \boldsymbol{\Sigma}_V \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}\right) \\
\Rightarrow T(\mathbf{X}, \mathbf{Y}) &\stackrel{a}{\approx} N_2\left((\lambda, \mu)^T, \begin{pmatrix} \sqrt{m} & 0 \\ 0 & \sqrt{n} \end{pmatrix}^{-1} \boldsymbol{\Sigma}_V \begin{pmatrix} \sqrt{m} & 0 \\ 0 & \sqrt{n} \end{pmatrix}^{-T}\right)
\end{aligned}$$

(d) Since Cramer-Rao bound for $T(\mathbf{X}, \mathbf{Y})$ is $I^{-1}((\lambda, \mu)^T)$,

$$\begin{aligned}
I((\lambda, \mu)^T) &= \mathbb{E}_{\lambda, \mu} \left[-\frac{\partial s(\lambda, \mu \mid \mathbf{X}, \mathbf{Y})}{\partial(\lambda, \mu)} \right] \\
&= \mathbb{E}_{\lambda, \mu} \left[-\begin{pmatrix} \frac{-m}{\lambda^2} & 0 \\ 0 & \frac{-n}{\mu^2} \end{pmatrix} \right] \\
&= \begin{pmatrix} \frac{m}{\lambda^2} & 0 \\ 0 & \frac{n}{\mu^2} \end{pmatrix} \\
I^{-1}((\lambda, \mu)^T) &= \frac{\lambda^2 \mu^2}{mn} \begin{pmatrix} \frac{n}{\mu^2} & 0 \\ 0 & \frac{m}{\lambda^2} \end{pmatrix} \\
&= \begin{pmatrix} \frac{\lambda^2}{m} & 0 \\ 0 & \frac{\mu^2}{n} \end{pmatrix}
\end{aligned}$$

For single observation (X_i, Y_j) ,

$$i^{-1}((\lambda, \mu)^T) = \begin{pmatrix} \lambda^2 & 0 \\ 0 & \mu^2 \end{pmatrix}$$

Cramer-Rao bound is achieved if

$$\begin{aligned}
\boldsymbol{\Sigma}_V &= i^{-1}((\lambda, \mu)^T) \\
\Leftrightarrow \begin{cases} \frac{1}{\lambda^2} = \lambda^2 \\ \frac{1}{\mu^2} = \mu^2 \end{cases} \\
\Leftrightarrow \begin{cases} \lambda = 1 \\ \mu = 1 \end{cases} &\quad (\text{consider real solution only and given } \lambda, \mu > 0)
\end{aligned}$$