

Trump Rallies and Hate Crimes: A Comment on Feinberg et al. (2019)

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Abstract

This note replicates and critiques the analysis of Feinberg et al (2019), which claims that Donald Trump election rallies led to a 226% increase in county hate incidents. After correcting their analysis for various problems, we find no effect of rallies on hate incidents.

1 Introduction

Feinberg et al (2019) claim evidence of an association between Donald Trump election rallies and county hate incidents. In particular, they argue that Donald Trump election rallies led to a heightened frequency of hate incidents. We collect the same data used by Feinberg et al., and then show that their conclusion is driven by the assumptions underlying their regression framework. Despite the language used in their paper, the analysis they present does not actually analyze whether counties that hosted Trump rallies saw increases in hate crimes following the rallies. Instead, the claim of a 226% increase relied on comparing counties with rallies to other counties without them. Yet politicians tend to hold political rallies near where large numbers of people live. And in places with more people, the raw number of crimes is generally mechanically higher. Appropriately controlling for population eliminates the effect they report. Further, as an example of the flaws in their model, we show that applying their methodology to Hillary Clinton election rallies yields the ostensible result that Clinton rallies led to an increase in hate incidents as large as the supposed Trump effect, if not larger.

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2 Data

We obtain the exact data used by Feinberg et al to the best of our abilities, given the description of the data they use in their paper. Table 1 reports the range and mean of each of the relevant variables in the data we obtained. Table 2 analogously reports the range and mean that Feinberg et al report in their data. They do not report the range and mean for every variable in their paper, so we list those for which they do. This allows us to assess the extent to which we have been able to collate the same or highly similar data, in order to replicate their study.

| Variable | Mean | Minimum | Maximum |
|--------------------------|---------|---------|---------|
| # Hate Incidents | 0.03 | 0 | 16 |
| # Hate Groups | 24.6 | 0 | 79 |
| Jewish Pop. Per Capita | 0.00297 | 0 | 0.29292 |
| % Rep 2012 | 59.6% | 7.3% | 95.9% |
| % College | 20.4% | 1.9% | 78.8% |
| Violent Crime (Per 10k) | 28.19 | 0 | 920.88 |
| Property Crime (Per 10k) | 200.89 | 0 | 3480.01 |
| Population | 98318 | 82 | 9818605 |

Table 1 – Means and Ranges of Control Variables Used in Analysis

| Variable | Mean | Minimum | Maximum |
|--------------------------|-------|---------|---------|
| # Hate Incidents | 0.03 | 0 | 16 |
| # Hate Groups | 24.6 | 0 | 79 |
| "Jewish Pop. Per Capita" | 315.7 | 0 | 28903.7 |
| % Rep 2012 | 59.7% | 7.1% | 95.9% |
| % College | 20.8% | 3% | 80.2% |

Table 2 – Means and Ranges Reported by Feinberg et al.

We obtain the hate incident data from the Anti-Defamation League (ADL) website, the link to which is provided in Feinberg et al.'s paper.¹ We match exactly the range and mean (the latter within determinable rounding error) of this variable with the range and mean reported by Feinberg et al. We obtain the hate group data from the Southern Poverty Law Center (SPLC) website.² Again, we match exactly the range and mean of this variable with those reported by Feinberg et al.

Feinberg et al. state that they obtain the data on county-level Jewish population "culled from the 2011 North American Jewish Data Bank and the 2011 U.S. Census American Com-

¹<https://www.adl.org/education-and-resources/resource-knowledge-base/adl-hate-map>

²<https://www.splcenter.org/hate-map?year=2016>

munity Survey by Comenetz (2012).” We interpret that they used the data from the North American Jewish Data Bank,^{3 4} which itself draws on the Comenetz data. Feinberg et al report that “per capita Jewish population ranges from 0 to 28,903.7 with a mean of 315.7”. By definition of per capita, this statement is impossible – so we assume that they are multiplying the per-capita numbers by 100,000. (Regardless, because this is a control variable, multiplication by a constant has no effect on any other coefficients in regression analysis.) If this is the case, then our numbers closely, but not exactly, match those reported by Feinberg et al.

The ADL, SPLC, and rally data are at the *place* level, whereas Feinberg et al. run their regressions at the county level. We use the Missouri Census Data Center’s geographic correspondence dataset,⁵ which is based on GIS, to assign places to the corresponding counties.

Feinberg et al. do not report from where they obtained their 2012 county-level presidential election data, so we obtain the county-level 2012 election returns data from Dave Leip’s Election Atlas,⁶ a standard source in the political science literature for election returns data. The range and mean of our % Republican vote share variable very closely matches theirs.

Feinberg et al. state that they obtained data on county-level college education rates for individuals aged 25+ from the 2015 American Community Survey. We also obtained data on this variable from the 2015 American Community Survey.⁷ The range and mean of this variable is close – but not identical – to that reported by Feinberg et al.

We obtain the list of rallies for the 2016 Donald Trump presidential campaign from Wikipedia, which is the source listed by Feinberg et al.⁸ We obtain the list of Hillary Clinton rallies for the 2016 general election campaign from FairVote.⁹

Feinberg et al. state that they obtain county-level violent crime and property crime rates from “the FBI’s 2015 Unified Crime Report”. The Unified Crime Report (UCR) encompasses many different datasets, one of which is the county-level crime data, available on ICPSR.¹⁰

³<https://www.jewishdatabank.org/databank/search-results/study/602>

⁴https://www.jewishdatabank.org/content/upload/bjdb/602/N-JewishMapUS_2011_Methodology.pdf

⁵<http://mcdc.missouri.edu/applications/geocorr2014.html>

⁶<https://uselectionatlas.org/>

⁷<https://data2.nhgis.org/main-Series B15003>

⁸https://en.wikipedia.org/wiki/List_of_rallies_for_the_2016_Donald_Trump_presidential_campaign

⁹<https://docs.google.com/spreadsheets/d/14Lxw0vc4YBUwQ8cZouyewZvOGg6PyzS2mArWNe3iJcY/edit#gid=0>

¹⁰<https://www.openicpsr.org/openicpsr/project/108164/version/V1/view>

Within this, there is data on Arrests and, separately, Offenses Known. Because not all offenses result in arrests but an offense known constitutes a crime nonetheless, we use the UCR County-Level Offenses Known to obtain our violent crime and property crime rate data.¹¹ Feinberg et al. do not report summary statistics for these variables. Also from the UCR, we obtain official data on hate crimes as an alternative measure to the ADL's data on hate incidents.¹²

We obtain county population data from the 2010 U.S. Census.¹³ Note that using intercensal estimates¹⁴ – approximations of county population made by the Census Bureau for non-Census years – corresponding to 2015 does not change our findings.

Finally, to run placebo tests, we obtain data on each Taco Bell location in the US.¹⁵

3 Analysis

3.1 Replication of Headline Result - Population Controls & Clinton Rally Effects

We run negative binomial regressions directly analogous to those of Feinberg et al. Table 1 reports the coefficients resulting from these regressions. In particular, column (1) repeats the main regression specification presented by Feinberg et al, the results of which we match rather closely, with a coefficient of 1.12 (compared to 1.18 in Feinberg et al) on the Trump rallies variable. Column (2) repeats this same regression, except with Clinton rallies, instead of Trump rallies, as the key variable of interest. The coefficient here is 1.21, an even larger number than the coefficient on Trump rallies in (1) (although the difference between the two is not statistically significant).

In addition to closely matching the estimated Trump rally effect in Feinberg et al., there are several other ways to demonstrate that we have closely replicated the full results of their model. Looking at the other control variables, these also typically have very similar coefficients to the estimates provided in Feinberg et al. Formally, denoting their point estimates as b and

¹¹Note that, in the county-level UCR numbers, data is missing for the counties corresponding to 4 out of 5 of New York's boroughs. This is because each borough is incorporated as a separate county, but the New York Police Department encompasses all five boroughs and is headquartered in Manhattan, so the data for the city as a whole is given for Manhattan alone. As such, we fill in the missing observations for the remaining 4 boroughs with the data for New York City as a whole.

¹²<https://www.icpsr.umich.edu/icpsrweb/ICPSR/studies/37060>

¹³This variable is included in the 2011 North American Jewish Data Bank, so we do not separately acquire it.

¹⁴<https://data.nber.org/data/census-intercensal-county-population-age-sex-race-hispanic.html>

¹⁵<https://www.arcgis.com/home/item.html?id=9206a8e05b29496991f69441c15bc6ed>

Table 3 – Replication of Base Specification - Trump and Clinton Rally Effects

| | (1) Hate Incidents | (2) Hate Incidents | (3) Hate Incidents | (4) Hate Incidents |
|--------------------|------------------------|------------------------|-------------------------|-------------------------|
| Trump Rally | 1.1152*** (0.1759) | | 0.1772 (0.1328) | |
| Clinton Rally | | 1.2128*** (0.3700) | | 0.1608 (0.2246) |
| Population (Log) | | | 1.1748*** (0.0727) | 1.1902*** (0.0704) |
| Jewish Pop. (p.c.) | 29.6317*** (5.8375) | 29.5561*** (5.4755) | 9.7918*** (2.9117) | 9.7363*** (2.9242) |
| Hate Groups | 0.0220*** (0.0051) | 0.0221*** (0.0052) | -0.0054 (0.0035) | -0.0057 (0.0035) |
| Violent Crime | 0.0043 (0.0029) | 0.0049 (0.0039) | 0.0025 (0.0032) | 0.0023 (0.0031) |
| Property Crime | 0.0007 (0.0004) | 0.0010** (0.0005) | 0.0008 (0.0008) | 0.0009 (0.0007) |
| % Rep. 2012 | -4.0908*** (0.7139) | -3.8977*** (0.7746) | -1.3461** (0.5709) | -1.2439** (0.5694) |
| % College | 5.5976*** (0.8383) | 6.6066*** (0.8663) | 3.0044*** (0.7814) | 3.1048*** (0.7929) |
| Northeast | 0.7020** (0.2759) | 0.7640*** (0.2717) | 0.7473*** (0.2308) | 0.7540*** (0.2288) |
| Midwest | -0.4082 (0.2673) | -0.3294 (0.2788) | -0.3835 (0.2491) | -0.3761 (0.2510) |
| South | -0.6969** (0.2864) | -0.6859** (0.2879) | -0.2221 (0.2086) | -0.2240 (0.2094) |
| January | -0.9767*** (0.1852) | -1.1787*** (0.1819) | -1.1383*** (0.1760) | -1.1824*** (0.1683) |
| February | -1.0075*** (0.1643) | -1.1788*** (0.1720) | -1.1961*** (0.1652) | -1.2346*** (0.1689) |
| March | -0.3967** (0.1553) | -0.5310*** (0.1577) | -0.5508*** (0.1485) | -0.5743*** (0.1546) |
| April | -0.7467*** (0.1692) | -0.8187*** (0.1694) | -0.8893*** (0.1491) | -0.9035*** (0.1540) |
| May | -0.7223*** (0.1733) | -0.7921*** (0.1751) | -0.8569*** (0.1586) | -0.8705*** (0.1614) |
| June | -0.7611*** (0.1759) | -0.7980*** (0.1796) | -0.9730*** (0.1569) | -0.9799*** (0.1643) |
| July | -1.1752*** (0.2148) | -1.2059*** (0.2180) | -1.3050*** (0.1891) | -1.3109*** (0.1932) |
| August | -0.9521*** (0.1690) | -1.0125*** (0.1666) | -1.0666*** (0.1586) | -1.0765*** (0.1613) |
| September | -0.9285*** (0.1671) | -0.9449*** (0.1647) | -0.9731*** (0.1596) | -0.9767*** (0.1603) |
| October | -0.6520*** (0.1434) | -0.6619*** (0.1417) | -0.6411*** (0.1389) | -0.6436*** (0.1382) |
| December | -0.4354** (0.1718) | -0.4234** (0.1654) | -0.4561*** (0.1438) | -0.4547*** (0.1438) |
| Constant | -3.8758*** (0.6148) | -4.1682*** (0.7073) | -17.5205*** (0.9791) | -17.7651*** (0.9631) |
| ln(α) | 1.4618*** (0.1717) | 1.5591*** (0.1592) | 0.3038 (0.2107) | 0.3194 (0.2069) |
| Pseudo R-squared | 0.3203 | 0.3130 | 0.4152 | 0.4149 |
| Log Likelihood | -2916.22 | -2947.36 | -2509.19 | -2510.46 |
| Observations | 37668 | 37668 | 37668 | 37668 |

* p<0.10, ** p<0.05, *** p<0.01

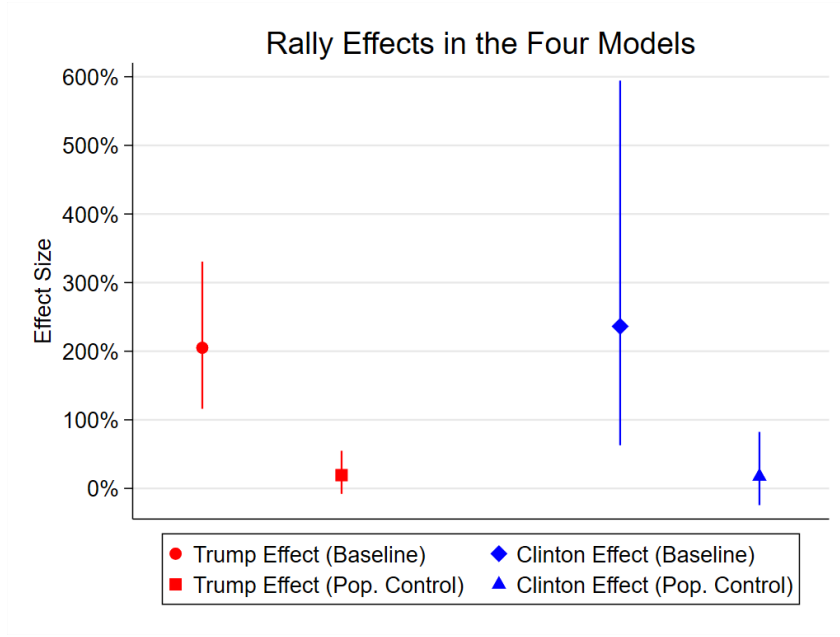


Figure 1 – Trump and Clinton Rally Effects in Base Specification, and with Population Control

ours as $\hat{\beta}$, we can measure the aggregate similarity of the predicted values of the two models. To do this, we use our data to construct the *index* for the two sets of estimates (respectively Xb and $X\hat{\beta}$). Across our regression sample, the resulting correlation between the two indices is 0.9732, and this would likely be even higher except for rounding error as the Feinberg et al. coefficients are often known to only two or fewer significant figures. This demonstrates that we have closely replicated the model in the original paper.

Crucially, we next add controls for the logarithm of population. In an exponential regression model such as this one, including the logarithm of population is the only correct way of controlling for population.¹⁶ Intuitively, negative binomial regressions estimate proportional rather than additive effects. As such, to properly control for the fact that the count of crime grows proportionally with the count of population, the correct approach is to include the log of population as a control. (Using the level of population as a control will not capture this relationship.) After doing this, the effects of both Trump rallies (column 3) and Clinton rallies (column 4) are no longer statistically significant. The model fit as measured by the pseudo R^2 and log likelihood, is also substantially improved. Figure 1 plots the four coefficients side-by-side, transformed to be represented in terms of a percentage increase (the way Feinberg et al reported their headline result).

¹⁶See <https://www.mathematica-journal.com/2013/06/negative-binomial-regression/> for a detailed explanation of this.

3.2 Replication Using FBI Hate Crimes Data

Next, in Table 4, we reproduce the same regressions in columns 1-4 of Table 3, except using the FBI UCR Official Hate Crime Data instead of the ADL data on Hate Incidents. Similarly, we get a significant coefficient on the Trump and Clinton rally variables if we do not control for population; once we do, both coefficients again fall to insignificant levels very close to zero. Figure 2 again plots these four transformed coefficients side-by-side.

Table 4 – Base Specification with FBI Hate Crime Data

| | (1) Hate Crimes | (2) Hate Crimes | (3) Hate Crimes | (4) Hate Crimes |
|------------------|-----------------------|-----------------------|-----------------------|-----------------------|
| Trump Rally | 1.0168*** (0.1891) | | -0.0440 (0.1079) | |
| Clinton Rally | | 1.3833*** (0.2935) | | 0.0724 (0.2181) |
| Population (Log) | | | 1.0688*** (0.0349) | 1.0617*** (0.0325) |
| Pseudo R-squared | 0.1936 | 0.1886 | 0.3035 | 0.3034 |
| Log Likelihood | -10824.11 | -10890.90 | -9349.54 | -9349.68 |
| Observations | 37668 | 37668 | 37668 | 37668 |

Control variables in these regressions are identical to those in Table 3; results are suppressed for brevity. * p<0.10, ** p<0.05, *** p<0.01

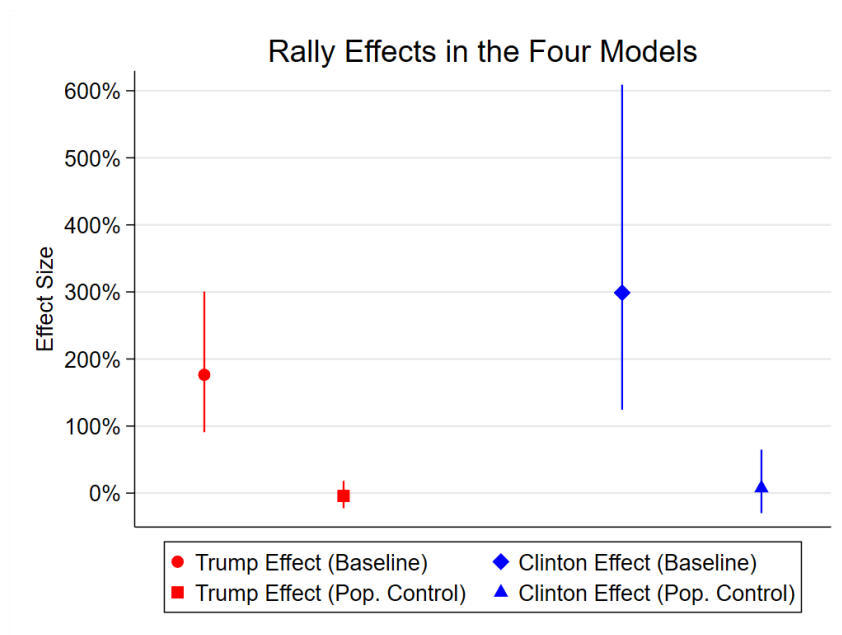


Figure 2 – Trump and Clinton Rally Effects Using FBI Hate Crime Data

3.3 Back to the Future

Feinberg et al. aim to investigate whether “following a Trump rally, the hosting county is more likely to report an increase in incidents targeting minorities”. However, the analysis they present did not actually analyze whether counties that hosted Trump rallies saw an increase in hate crimes following the rally relative to their prior level. Instead, the headline claim of a 226% increase relied on comparing counties with rallies to other counties without them.

To demonstrate the pitfalls of using cross-sectional analysis to verify a temporal claim, and more specifically the flaws underlying the statistical model used by Feinberg et al., we present an additional falsification test by regressing *past* hate incidents on *future* Trump rallies. In other words, we ask whether Trump rallies in, say, November cause hate incidents the previous February. This would be an absurdity. To perform this analysis, we define a county as having a future rally for candidate x if it subsequently held a rally for that candidate in a later month, and had not already held a rally for that candidate in our sample.¹⁷ Aside from this temporal switch, we retain the primary regression specification estimated by Feinberg et al. Despite the absurd nature of this hypothesis, we do indeed find that future Trump and Clinton rallies cause hate incidents in the past. Furthermore, the magnitude of the coefficients on both the Trump (1) and Clinton (2) specifications are almost indistinguishable from the base-line estimates in Table 3. This is an example of the hazards of running static regression specifications to study a dynamic question (i.e., do Trump rallies cause an increase in hate crimes in the county where the rally is held relative to the numbers of hate crimes beforehand?).

Table 5 – Falsification Test - Future Rally Effects on Past Hate Incidents

| | (1) Hate Incidents | (2) Hate Incidents |
|----------------------|-----------------------|-----------------------|
| Future Trump Rally | 1.0561*** (0.2282) | |
| Future Clinton Rally | | 1.1946*** (0.3177) |
| Pseudo R-squared | 0.3147 | 0.3137 |
| Log Likelihood | -2940.34 | -2944.30 |
| Observations | 37668 | 37668 |

Control variables in these regressions are identical to those in Table 3; results are suppressed for brevity. * $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$

¹⁷This ensures that effects of past rallies cannot be attributed to the Future Rally variables.

3.4 Yo Quiero Taco Bell

By studying a count variable without controlling for population appropriately, Feinberg et al.'s method is prone to producing *curious* and implausible findings. As political rallies typically occur in high population areas, their method is likely to claim an effect on any outcome which is increasing in population. To illustrate this, we look at the alleged effect of Trump rallies on the frequency of Taco Bell restaurants, where there is a very strong *a priori* reason for believing the true effect is zero.

We again estimate the baseline specification of Feinberg et al., changing only the outcome of interest to the number of Taco Bell stores. Column (1) shows that this implausibly yields an estimated effect from Trump rallies of 160%. Subsequent to their failure to control for population being pointed out, Feinberg et al. have stated that they still find a (much smaller) effect when including a control for the level of population. As above, this is the incorrect way of controlling for population in a negative binomial regression. Column (2) documents how adding such a control still yields erroneous results; the estimated effect of Trump rallies on Taco Bell restaurants remains a highly statistically significant 41%. Only when controlling for the logarithm of population in column (3) does the effect fall to a tightly estimated zero.

This again demonstrates that invalidity of the statistical methodology upon which the 226% hate crime effect claim is based. Further, it shows that spurious results are still generated if population is incorrectly controlled for. In a cross-sectional analysis such as this, it is necessary to control for the log of population.

Table 6 – Falsification Test - Rally Effects on Taco Bell Locations

| | (1) Restaurants | (2) Restaurants | (3) Restaurants |
|-------------------|-----------------------|-----------------------|-----------------------|
| Trump Rally | 0.9567*** (0.1160) | 0.3439*** (0.0893) | 0.0507 (0.0350) |
| Population ('000) | | 0.0027*** (0.0004) | |
| Population (Log) | | | 1.0471*** (0.0172) |
| Pseudo R-squared | 0.1601 | 0.2085 | 0.4203 |
| Log Likelihood | -55502.30 | -52304.78 | -38312.59 |
| Observations | 37668 | 37668 | 37668 |

* p<0.10, ** p<0.05, *** p<0.01

3.5 Dynamic Analysis

Despite these implementation issues, Feinberg et al. are correct in one regard - the natural thought experiment to test is whether hate crimes rose in counties hosting a Trump rally following the rally, relative to their previous level. Such dynamic analysis also makes it possible to bypass arguments about the correct way of controlling for differences (such as population) between locations, as identification comes from within-county variation.

To this end, we run difference-in-differences specifications – comparing the changes in hate incidents before/after Trump (Clinton) rallies in counties where rallies were held with the contemporaneous changes in counties where rallies were not held. If Trump (Clinton) rallies did indeed cause hate incidents, then we should see an increase in hate incidents in the counties where they were held relative to the counties where they were not. Estimating such a specification requires adding a fixed effect (dummy variable) for each county. Having included these, demographic controls that are constant across time (such as those controls included in previous regressions) become redundant. We do however also include dummy variables for month, to separate the effects of rallies from general temporal variation in the number of incidents. In fact, as seen in Table 6, we find no significant difference.¹⁸ This is true for both Trump rallies (column 1) and Clinton rallies (column 3). It remains true if we again use FBI UCR hate crimes as the outcome variable (columns 2 and 4), rather than ADL hate incidents. Indeed, not only are the point estimates for Trump rallies on hate crimes statistically insignificant, they are not even consistently signed.

Table 7 – Difference in Differences Estimates

| | (1) Hate Incidents | (2) Hate Crimes | (3) Hate Incidents | (4) Hate Crimes |
|----------------|-----------------------|---------------------|-----------------------|--------------------|
| Trump Rally | 0.1968 (0.1662) | -0.1357 (0.0861) | | |
| Clinton Rally | | | 0.0885 (0.1446) | 0.0130 (0.0968) |
| Month FE | Yes | Yes | Yes | Yes |
| Log Likelihood | -1481.51 | -5575.59 | -1482.56 | -5577.99 |
| Observations | 3000 | 9648 | 3000 | 9648 |

* p<0.10, ** p<0.05, *** p<0.01

¹⁸The regressions we run in Table 6 are Poisson regressions. We consider these to be closest in nature to the negative binomial regressions of Feinberg et al. while still being tractable in a difference-in-differences framework.

4 Conclusion

Do Trump rallies boost hate crimes? Feinberg et al. initially reported that Trump rallies led to an increase in hate incidents on the order of 226% in the counties wherein the rallies were held. This is an undeniably huge effect size. However, we find that the analogous purported effect of Clinton rallies on hate incidents is just as large, if not larger.

When we control for the logarithm of population – the correct way of controlling for population in a negative binomial model – both coefficients immediately fall to levels statistically indistinguishable from zero. This demonstrates that the headline result of Feinberg et al. was an artefact of their failure to account for the fact that higher population places will tend to have proportionally higher counts of crime. While other, flawed, ways of controlling for population retain nonzero Trump rally effects, these incorrect methods also make highly implausible claims; for example that Trump rallies increased the number of Taco Bell locations. This highlights the importance of controlling for population correctly.

In a further extension, we use FBI UCR Official Hate Crime data instead of ADL hate incident data. Again, controlling for population eliminates the result. Finally, we observe that, contrary to the claim that Trump rallies led to an increase in hate incidents, Feinberg et al. did not even analyze a dynamic statistical model, only comparing at the static level hate incidents in counties with Trump rallies to the level in counties without them. We show that regressing *past* hate crimes on *future* Trump rallies leads to a significant effect similar in magnitude to the headline result, demonstrating that hate incidents were already higher in counties with Trump rallies before the rallies were ever held (an unsurprising result given their higher population). Proceeding further in this vein, we run a difference-in-differences specification to more appropriately answer the main research question by comparing the *change* in the number of hate incidents in counties with Trump rallies before/after the rally to the change in counties without Trump rallies. In so doing, we again find no significant effect.

Reference

Feinberg, Branton, and Martinez-Ebers (2019). "The Trump Effect: How 2016 Campaign Rallies Explain Spikes in Hate." Working Paper.