

Lecture 1: Logistic Regression

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1.1 Two-Class Logistic Regression

$Prob(event\ happens) = p.$

$Prob(event\ not\ happens) = 1 - p = q.$

The **Odds** of winning event happens is defined as

$$Odds \triangleq \frac{p}{q}. \quad (1.1)$$

In logistic regression, we model the log odds as a linear equation

$$\log\left(\frac{p}{q}\right) = \beta^T x + b \quad (1.2)$$

From this definition, we obtain

$$\begin{aligned} \log\left(\frac{p}{q}\right) &= \beta^T x + b \\ \log\left(\frac{p}{1-p}\right) &= \beta^T x + b \\ \frac{p}{1-p} &= e^{\beta^T x + b} \\ p &= (1-p)e^{\beta^T x + b} \\ p(1 + e^{\beta^T x + b}) &= e^{\beta^T x + b} \\ p &= \frac{e^{\beta^T x + b}}{1 + e^{\beta^T x + b}} = \frac{1}{1 + e^{-(\beta^T x + b)}} \\ q &= 1 - p = 1 - \frac{1}{1 + e^{-(\beta^T x + b)}} = \frac{e^{-(\beta^T x + b)}}{1 + e^{-(\beta^T x + b)}} = \frac{1}{1 + e^{\beta^T x + b}}. \end{aligned} \quad (1.3)$$

Now, consider 2-class classification. If we assign class label as $y_i = \pm 1$, we can express the probability as

$$p(y_i) = \frac{1}{1 + e^{-y_i(\beta^T x + b)}}. \quad (1.4)$$

Maximum Likelihood Estimation (MLE) In the following, we set $y_i = 1$ if the event happens, otherwise $y_i = 0$. We obtain the model parameters from the maximization of the likelihood of the events x_1, \dots, x_n , which is defined as

$$L = \prod_{i=1}^n p(x_i) = \prod_{i=1}^n p^{y_i} (1-p)^{1-y_i}, \quad (1.5)$$

where $p = P(y_i = +1)$. This is **binomial distribution**.

Rather than maximizing the likelihood function, equivalently, we maximize the **log-likelihood** loss function

$$l = \log L = \sum_{i=1}^n [y_i \cdot \log(p) + (1 - y_i) \cdot \log(1 - p)], \quad (1.6)$$

to obtain the model parameters:

$$\max_{\theta} l(\theta). \quad (1.7)$$

Model 1. Binomial distribution

For binomial distribution, the model parameter is $\theta = p$. Then we obtain the optimal model parameter by maximizing the log-likelihood loss. The derivative of l with respect to p is computed as

$$\begin{aligned} \frac{\partial l}{\partial p} &= \sum_{i=1}^n \left(\frac{y_i}{p} - \frac{1-y_i}{1-p} \right) \\ &= \sum_{i=1}^n \frac{y_i(1-p) - p(1-y_i)}{p(1-p)} \\ &= \sum_{i=1}^n \frac{y_i - p}{p(1-p)} \\ &= \frac{n_+ - np}{p(1-p)}. \end{aligned} \quad (1.8)$$

Set $\frac{\partial l}{\partial p} = 0$, we have

$$\frac{n_+ - np}{p(1-p)} = 0. \quad (1.9)$$

With the assumption that $p \neq 0$ and $1-p \neq 0$, optimal parameter is obtained as

$$p = \frac{n_+}{n}. \quad (1.10)$$

Model 2. Logistic regression

For logistic regression, the model parameters are $\theta = (\beta, b)$ in the $p_i = \frac{1}{1 + e^{-(\beta^T x_i + b)}}$. After padding, $p_i = \frac{1}{1 + e^{-\tilde{\beta}^T \tilde{x}_i}}$, where $\tilde{\beta} = [\beta, b]$ and $\tilde{x}_i = [x_i, 1]$. The log-likelihood is now

$$l = \log L = \sum_{i=1}^n [y_i \cdot \log(p_i) + (1 - y_i) \cdot \log(1 - p_i)], \quad (1.11)$$

The derivative of l with respect to $\tilde{\beta}$ is computed as

$$\begin{aligned} \frac{\partial l}{\partial \tilde{\beta}} &= \frac{\partial l}{\partial p_i} \cdot \frac{\partial p_i}{\partial \tilde{\beta}} \\ &= \sum_{i=1}^n \left(\frac{y_i}{p_i} - \frac{1-y_i}{1-p_i} \right) \cdot \frac{\partial p_i}{\partial \tilde{\beta}}. \end{aligned} \quad (1.12)$$

In logistic regression, we use the activation function and its derivative

$$\sigma(z) = \frac{1}{1 + e^{(-z)}}, \quad \frac{d\sigma(z)}{dz} = \sigma \cdot (1 - \sigma), \quad (1.13)$$

where $z_i = \tilde{\beta}^T \tilde{x}_i$ and $p_i = \sigma(z_i)$. Therefore, we have

$$\begin{aligned} \frac{\partial p_i}{\partial \tilde{\beta}} &= \frac{\partial p_i}{\partial z} \cdot \frac{\partial z}{\partial \tilde{\beta}} \\ &= \sigma \cdot (1 - \sigma) \cdot \tilde{x}_i \\ &= p_i \cdot (1 - p_i) \cdot \tilde{x}_i, \end{aligned} \quad (1.14)$$

Finally, we have

$$\begin{aligned}\frac{\partial l}{\partial \beta} &= \sum_{i=1}^n \left(\frac{y_i}{p_i} - \frac{1-y_i}{1-p_i} \right) \cdot p_i \cdot (1-p_i) \cdot \tilde{x}_i \\ &= \sum_{i=1}^n (y_i(1-p_i) - (1-y_i)p_i) \cdot \tilde{x}_i \\ &= \sum_{i=1}^n (y_i - p_i) \cdot \tilde{x}_i.\end{aligned}\tag{1.15}$$

Here, because of the complicated expression, we cannot find an expression for the parameter $\tilde{\beta} = \dots$ as in Equ. 1.10. $\tilde{\beta}$ is computed numerically using an optimization algorithm such as the gradient descent algorithm.

1.2 Multi-Class Logistic Regression

For K class, we have probabilities $\{p_1, p_2, \dots, p_K\}$ with $\sum_{i=1}^K p_i = 1$. Then we define

$$\begin{cases} \log \frac{p_1}{p_K} = \beta_1^T x + b_1 \\ \log \frac{p_2}{p_K} = \beta_2^T x + b_2 \\ \dots \\ \log \frac{p_{K-1}}{p_K} = \beta_{K-1}^T x + b_{K-1} \end{cases}\tag{1.16}$$

$$\begin{cases} \frac{p_1}{p_K} = e^{\beta_1^T x + b_1} \\ \frac{p_2}{p_K} = e^{\beta_2^T x + b_2} \\ \dots \\ \frac{p_{K-1}}{p_K} = e^{\beta_{K-1}^T x + b_{K-1}} \end{cases}\tag{1.17}$$

$$\begin{cases} p_1 = p_K \cdot e^{\beta_1^T x + b_1} = [1 - (p_1 + \dots + p_{K-1})] e^{\beta_1^T x + b_1} \\ p_2 = p_K \cdot e^{\beta_2^T x + b_2} = [1 - (p_1 + \dots + p_{K-1})] e^{\beta_2^T x + b_2} \\ \dots \\ p_{K-1} = p_K \cdot e^{\beta_{K-1}^T x + b_{K-1}} = [1 - (p_1 + \dots + p_{K-1})] e^{\beta_{K-1}^T x + b_{K-1}} \end{cases}\tag{1.18}$$

Add all equations in Equ 1.18 together

$$\begin{aligned}p_1 + \dots + p_{K-1} &= [1 - (p_1 + \dots + p_{K-1})] \left(e^{\beta_1^T x + b_1} + e^{\beta_2^T x + b_2} + \dots + e^{\beta_{K-1}^T x + b_{K-1}} \right) \\ 1 - p_K &= p_K \cdot \left(\sum_{i=1}^{K-1} e^{\beta_i^T x + b_i} \right) \\ 1 &= p_K \cdot \left(1 + \sum_{i=1}^{K-1} e^{\beta_i^T x + b_i} \right) \\ p_K &= \frac{1}{1 + \sum_{i=1}^{K-1} e^{\beta_i^T x + b_i}}.\end{aligned}\tag{1.19}$$

Put p_K in Equ 1.19 back into Equ 1.18

$$p_k = p_K \cdot e^{\beta_k^T x + b_k} = \frac{e^{\beta_k^T x + b_k}}{1 + \sum_{i=1}^{K-1} e^{\beta_i^T x + b_i}},\tag{1.20}$$

where k varies in $\{1, 2, \dots, K-1\}$.

Using padding $\tilde{\beta}_k = [\beta_k, b_k]$ and $\tilde{x} = [x, 1]$, Equ 1.20 is transformed as

$$\begin{cases} p_1 = \frac{e^{\tilde{\beta}_1^T \tilde{x}}}{1 + \sum_{i=1}^{K-1} e^{\tilde{\beta}_i^T \tilde{x}}} \\ p_2 = \frac{e^{\tilde{\beta}_2^T \tilde{x}}}{1 + \sum_{i=1}^{K-1} e^{\tilde{\beta}_i^T \tilde{x}}} \\ \dots \\ p_{K-1} = \frac{e^{\tilde{\beta}_{K-1}^T \tilde{x}}}{1 + \sum_{i=1}^{K-1} e^{\tilde{\beta}_i^T \tilde{x}}} \\ p_K = \frac{e^{\tilde{\beta}_K^T \tilde{x}}}{1 + \sum_{i=1}^{K-1} e^{\tilde{\beta}_i^T \tilde{x}}} \end{cases} \quad (1.21)$$

Maximum Likelihood Estimation (MLE)

$$L \propto p_1^{n_1} p_2^{n_2} \dots p_K^{n_K} \quad (1.22)$$

where $\sum_{k=1}^K n_k = n$ and $\sum_{k=1}^K p_k = 1$. This is **multinomial distribution**.

The log-likelihood loss function

$$l = \log L \propto \sum_{i=1}^K n_i \cdot \log p_i. \quad (1.23)$$

Model 3. Multinomial distribution

If we do multinomial distribution, model parameters are $\{\theta_1 = p_1, \dots, \theta_K = p_K\}$. Then we obtain optimal model parameters by maximizing the log-likelihood loss

$$\begin{aligned} & \max_{\{p_1, \dots, p_K\}} \sum_{i=1}^K n_i \cdot \log p_i, \\ & \text{subject to } \sum_{i=1}^K p_i = 1. \end{aligned} \quad (1.24)$$

Using **Augmented Lagrangian Multiplier (ALM)**, we have the following equivalent optimization form

$$\max_{\{p_1, \dots, p_K\}} \mathcal{L}(\lambda) = \left(\sum_{i=1}^K n_i \cdot \log p_i \right) - \lambda \cdot \left(\sum_{i=1}^K p_i - 1 \right) \quad (1.25)$$

The derivative of \mathcal{L} with respect to p_i and λ is computed respectively as

$$\begin{aligned} \frac{\partial \mathcal{L}}{\partial p_i} &= \frac{n_i}{p_i} - \lambda, \\ \frac{\partial \mathcal{L}}{\partial \lambda} &= \sum_{i=1}^K p_i - 1. \end{aligned} \quad (1.26)$$

Set $\frac{\partial \mathcal{L}}{\partial p_i} = 0$, we have

$$n_i = \lambda p_i \quad (i=1, \dots, K). \quad (1.27)$$

Add all the K equations in Equ 1.27 together,

$$\begin{aligned} n_1 + n_2 + \dots + n_K &= \lambda \cdot (p_1 + p_2 + \dots + p_K) \\ n &= \lambda \cdot 1 \\ \lambda &= n \end{aligned} \quad (1.28)$$

Put $\lambda = n$ back in Equ 1.27,

$$n_i = n p_i \quad (i=1, \dots, K). \quad (1.29)$$

Finally, optimal parameters are obtained

$$p_i = \frac{n_i}{n} \quad (i=1, \dots, K). \quad (1.30)$$

Model 4. Multinomial logistic regression

Now we do multinomial logistic regression, model parameters $\theta = \tilde{\beta}_k$ are in $p_{k,i} = \frac{e^{\tilde{\beta}_k^T \tilde{x}_i}}{1 + \sum_{h=1}^{K-1} e^{\tilde{\beta}_h^T \tilde{x}_i}}$ ($k = 1, \dots, K-1$) and $p_{K,i} = \frac{1}{1 + \sum_{h=1}^{K-1} e^{\tilde{\beta}_h^T \tilde{x}_i}}$.

For generality, set $\tilde{\beta}_K = 0$. Then we have

$$p_{k,i} = \frac{e^{\tilde{\beta}_k^T \tilde{x}_i}}{\sum_{h=1}^K e^{\tilde{\beta}_h^T \tilde{x}_i}}, \quad k = \{1, \dots, K\}. \quad (1.31)$$

Let $Y_i^k = I(Y_i = k)$ be one-of-K encoding of Y_i , also known as indicator vector with $\sum_{i=1}^K Y_i^k = 1$. Define the log-likelihood loss function as

$$\begin{aligned} l &= \log L \\ &= \log \left[\prod_{i=1}^n \left(p_{1,i}^{Y_i^1} \cdots p_{K,i}^{Y_i^K} \right) \right] \\ &= \sum_{i=1}^n \log \left(p_{1,i}^{Y_i^1} \cdots p_{K,i}^{Y_i^K} \right) \\ &= \sum_{i=1}^n \sum_{k=1}^K Y_i^k \log(p_{k,i}) \\ &= \sum_{i=1}^n \left\{ \sum_{k=1}^K Y_i^k \log \frac{e^{\tilde{\beta}_k^T \tilde{x}_i}}{\sum_{h=1}^K e^{\tilde{\beta}_h^T \tilde{x}_i}} \right\} \\ &= \sum_{i=1}^n \left\{ \sum_{k=1}^K Y_i^k (\beta_k^T \tilde{x}_i) - \log \left(\sum_{h=1}^K e^{\tilde{\beta}_h^T \tilde{x}_i} \right) \right\} \end{aligned} \quad (1.32)$$

The derivative of l with respect to $\tilde{\beta}_k$ ($k = 1, \dots, K-1$) is computed as

$$\begin{aligned} \frac{\partial l}{\partial \beta_k} &= \sum_{i=1}^n Y_i^k \cdot \tilde{x}_i - \left(\frac{e^{\tilde{\beta}_k^T \tilde{x}_i}}{\sum_{h=1}^K e^{\tilde{\beta}_h^T \tilde{x}_i}} \right) \cdot \tilde{x}_i \\ &= \sum_{i=1}^n (Y_i^k - p_{k,i}) \cdot \tilde{x}_i, \end{aligned} \quad (1.33)$$

which is same as the derivative result of binomial logistic regression in Equ 1.15 .

1.3 Softmax Regression

Probability is defined as

$$\begin{cases} p_1 &= \frac{e^{\tilde{\beta}_1^T \tilde{x}}}{\sum_{k=1}^K e^{\tilde{\beta}_k^T \tilde{x}}}, \\ \dots & \\ p_K &= \frac{e^{\tilde{\beta}_K^T \tilde{x}}}{\sum_{k=1}^K e^{\tilde{\beta}_k^T \tilde{x}}}. \end{cases} \quad (1.34)$$

Actually, **Softmax Regression** is the general form of **Multi-class Logistic Regression** without specifically setting $\tilde{\beta}_K$ to 0.

Therefore, the derivative of log-likelihood loss with respect to $\tilde{\beta}_k$ is same as the results in Equ 1.33.