

- ch 1
- exponential growth
- probability simplex / regular polygon formed by unit vector.
- HW - principle-Zallele 3/4 ODE of stable is a steady state
- With mutation

- Intensity of selection How does selection favors A?
- Fixation Prob. selection opposes invasion.
- vs law
- fish dominance
- ch 9. spatial Moran.
- how to jump from $i \rightarrow j$ cell
- Eder growth is a band cell same-based.

- ch 2
- what is Quasispecies. is a quasispecies
- sequence space
- 3/4 Quasispecies equation
- $$\dot{x} = xW - \phi x$$
- Adaptive. / Def / Empt.
- mutation meltdown / antiviral strategy.
- ch 3
- Some MC properties.
- Moran Def. / neutral / general case.
- Fixation Probability. comparing exponential to 1 - E[x]
- Mean fixation time. Poisson process - IAT to 1/2

- ch 10 G.W Process $\{Z_n, i\}$ Branching process
- Prob genotypic function
- $$f_{n+1} = f^{n+1}(s)$$
- its relation to Variance / Expectation
- Extinction Prob. and the theorem
- fixation Prob. $P_n(k, j)$

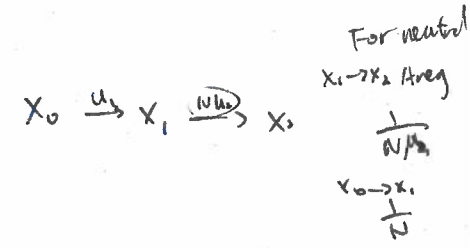
- ch 3
- Some MC properties.
- Moran Def. / neutral / general case.
- Fixation Probability. comparing exponential to 1 - E[x]
- Mean fixation time. Poisson process - IAT to 1/2

- ch 22
- Poset: set with "order operator"
- Distributed later
- genotype lattice is a chain
- Fitness / mutational neighbourhood.
- Mutational arch and res relationship are to the
- Multiple Branching / fixation time is a Poisson dist.

- ch 4
- Def cancer. How were
- TSG. Oncogenes. is a fixation
- Prevent of cancer / linear compartmentalization. $P = \frac{1}{N}$
- 2-stage model for TSG why high / tumorigen.

- ch 12
- Def of coalescent
- what genealogy tree comes from
- How does the $T(i)$ work
- How is $T(i)$ distributed / Derivation $N \rightarrow \infty$
- How the Expectation
- Infinite Sites Model / Assumption,
- Tajima's D. what does it do? null hypothesis

- CIN model
- each stage $WT \sim Exp(\lambda)$
- Multi-stage theory (Mutation Pathway) to the number
- Wg. Fisher Proves Binomial accuracy to the number
- Fixed
- Example / length
- Extension to 3/4 binary sequence of d P[sample to i]



- ch 6. Diffusion Theory
- Kolmogorov forward/backward.
- From prob.
- Mean Fix time

Rate of mutation

- ch 7 Pay-off matrix. / 3/4 fixation
- ESS unbeatable, strict Nash.
- Determine linear stability by $\det(A)$ why?
- TFT.
- $$f_A(1-\epsilon) > f_B(\epsilon)$$
- $$a(1-\epsilon) + b\epsilon > c(1-\epsilon) + d\epsilon$$

Is WF Process a death-birth process? No analytic sol.

- $P_k = P[Z=k]$ set by distribution
- P_{ij} transition probability. $P_{ik} = P_k$
- $P_{2j} = P_0 P_{j1} \dots P_{j0}$
- $$P_{ij} = \sum_{k: k_i=j} P_k \cdot P_{ki} \cdot \forall i$$

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五福 拼手气

—

集齐五福，拼手气

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集齐五福，拼手气分5亿

$$\frac{1-r^{-1}}{1-r^{-N}}$$

五福，拼手气

chap 2 evolution

Def. changes in the frequency of types from 1 gen. to the next.
 ↳ from mutation, replication and selection.

Def (Malthusian Law) Exponential Growth

$$x' = rx \Rightarrow x(t) = x_0 e^{rt} \quad r: \text{rate of cell division} \quad \text{average gen. time: } \frac{1}{r}$$

with death $\frac{1}{d}$ on average

$$x' = (r-d)x \Rightarrow x(t) = x_0 e^{(r-d)t}$$

$R_0 = r/d := \text{basic reproductive ratio.}$

> 1 ↑
 < 1 extinct
 $= 1$

Logistic growth.

population has a carrying capacity K

$$x' = rx(1 - \frac{x}{K}) \Rightarrow x(t) = \frac{K x_0 e^{rt}}{K + x_0(e^{rt} - 1)} \quad \lim_{t \rightarrow \infty} x(t) = K$$

Rescaling $x' = x(1-x)$ as growth rate.

Bifurcation

Selection: 2 exp. growing types.

type A: a $x(t)$

$$x' = ax$$

if $a > 0$ not extinct

B: b $y(t)$

$$y' = by$$

with cap. let $x(t) + y(t) = 1 \quad \forall t \geq 0$

有限

$$\begin{cases} x' = x(a - \phi) \\ y' = y(b - \phi) \end{cases} \quad \phi = ax + by: \text{average fitness of the population.}$$

就凭他们的
relative fitness

$$\Leftrightarrow x' = x(1-x)(a-b)$$

if $a > b \rightarrow$ All $a \quad x=1$
 All $b \quad x=0$

Probability Simplex

$$S_n = \{ (x_1, \dots, x_n) \mid \sum_{i=1}^n x_i = 1, x_i \geq 0 \}$$

every pt. in S_i is a prob. dist.

0.1

S_2

S_3

S_4

Regular

最可能的点会变成
(0, ..., 1, ..., 0)

Core type

$$x_1(t) + \dots + x_n(t) = 1 \quad \text{in } S_n.$$

$$\text{average fitness: } \phi = x_1 f_1 + \dots + x_n f_n$$

$$x_i' = x_i(f_i - \phi) \Rightarrow \text{a single equilibrium pt.}$$

sub/super exp. growth.

$$\begin{cases} x' = ax^c - \phi x \\ y' = by^c - \phi y \end{cases} \quad \phi = ax^c + by^c$$

$$\Rightarrow \begin{cases} c=0 & \text{linear (immigration)} \\ c=1 & \text{exp.} \\ c<1 & \text{sub.} \\ c>1 & \text{super.} \end{cases}$$

$$\Leftrightarrow x' = x(1-x)f(x), \quad f(x) = ax^{c-1} - b(1-x)^{c-1}$$

$$f(x)=0 \Rightarrow$$

$$\text{Fixed pt: } \begin{cases} x=0 \\ x=1 \\ x^* = 1/(1+(a/b)^{1/(c-1)}) \text{ if } c \neq 1 \end{cases} \quad \text{why}$$

$< 0, 1$ not stable.

when $c < 1$: survival of all

x^* is globally stable.

[Can invade to stable pt, 即使 $a > b$, 也可以到 stable]

$c > 1$: unstable x^* :

if $x > x^*$
 $x < x^*$

A outcomplete B
 B outcomplete A

Can NOT invade.

$\{0, 1\}$ stable

ation

$u_1 \text{ Prob}(A \rightarrow B)$

$u_2 \text{ Prob}(B \rightarrow A)$ during reproduction

$$\begin{cases} x' = x(1-u_1) + y u_2 - \phi x \\ y' = y(1-u_2) + x u_1 - \phi y \end{cases}$$

$$\phi = 1 \quad x+y=1$$

因为 growth difference 是一样的.

why

$$\Rightarrow x' = u_2 - x(u_1 + u_2)$$

$$\Rightarrow \text{stable equilibrium } x^* = \frac{u_2}{u_1 + u_2}$$

$$\text{If } u_2 = 0, \quad x' = x u_1 \Rightarrow A \text{ extinct}$$

station of a types.

$q_{ij} : \text{Prob}(i \rightarrow j)$

Q stochastic matrix.

$$x_i' = \sum_{j=1}^n x_j q_{ji} - \phi x_i \quad x' = xQ - \phi x$$

y-Weinberg Principle

$$aa \quad aa \quad AA: x \ y \ z$$

$$a: p \quad A: q$$

$$x+y+z=p+q=1$$

$$\text{Random mating} \Rightarrow \begin{cases} x = p^2 \\ y = 2pq \\ z = q^2 \end{cases}$$

$$p' = p \quad q' = q \quad \text{constant!}$$

$$x_2 = f(x_1) - \beta = 0$$

HIV Infection

stats:

Some biological bias on HIV

short genome

high mutation rate

short gen. time

large population.

Sequence space

Def space

$$\Lambda^L = \{ (a_1, \dots, a_L) \mid a_i \in \Lambda \}$$

$$\Lambda^0 = \cup_{L \geq 0} \Lambda^L$$

Evolution is a trajectory of a population in sequence space

Def. Hamming dist.

Def. Mismatch dist. no. of mismatches

Def. Fitness

$$f: G \rightarrow \mathbb{R} \quad \text{or} \quad f: G \rightarrow P \rightarrow \mathbb{R}.$$

hard to measure.

Fitness is a MAP from genotype to

Quasispecies equation

Ex. $i=0, 1, 2, \dots, n$

$x(t) = (x_0(t), \dots, x_n(t))$ genotypes frequencies

$Q = q_{ij} := \text{mutation matrix} = I$ (error-free)

$f = (f_0, \dots, f_n)$ fitness $\phi = x \cdot f$ average fitness

$$\dot{x}_i = \sum_{j=0}^n x_j f_j q_{ji} \quad i=0, \dots, n \Leftrightarrow \dot{x} = x f Q - \phi x$$

If Q is irreducible, global x^* exist.

Solving

$$x_i(t) = x_i(t) \in P(t)$$

$$\phi(t) = \int_0^t \phi(s) ds$$

$$\Rightarrow \dot{x} = \sum_{j=0}^n x_j f_j q_{ji} \Rightarrow x = \sum_{j=0}^n x_j = e^{\psi} \text{ thus } x_i = \frac{x_i}{x}$$

$$\Rightarrow \dot{x} = \psi e^{\psi} = \phi x \quad (x \text{ grow exponentially})$$

$x W = \phi x$ ϕ is the largest eigenvalue of W

This can be rewritten as

$$\dot{x} = x W - \phi x \quad \text{with } W = (w_{ij}) = (f_j q_{ji})$$

x^* is the solution of $x W = \phi x$ [Eigenvector problem]

$\Rightarrow \phi$ is the largest eigenvalue of W and x^* is left eigenvector.

option

Def. Location of maximal fitness in sequence space.

{ mutation: generates "peaks" > balance x^*
selection: drive to max peak

If mutation removes "fit" types adaptation is NOT possible.

Ex. Binary seq. of L $0: f_0$ $1: f_1$ $\rightarrow \phi$

Type 0 copy without error $q_0 = (1-u)^L$

$$\begin{cases} \dot{x}_0 = x_0 f_0 q_0 - \phi x_0 \\ \dot{x}_1 = x_0 f_0 (1-q_0) + x_1 - \phi x_1 \end{cases} \Leftrightarrow \dot{x}_0 = x_0 [f_0 q_0 - 1 - x_0 (f_0 - 1)]$$

What's mutation meltdown?

$$x_0^* = \begin{cases} \frac{f_0 q_0 - 1}{f_0 - 1} & f_0 q_0 > 1 \\ 0 & \text{else.} \end{cases}$$

To get the bound for mutation rate

$$\log(f_0 q_0) = 1 + L \log(1-u) \approx 1 - Lu > \log 1 \Leftrightarrow$$

$$u < \frac{1}{L}$$

$\Rightarrow u_c = \frac{1}{L}$ is the error threshold.

$u > \frac{1}{L}$ "mutation meltdown"
strategy to anti-viral

It's like the learning rate in some way

Ch 5. stochastic model in finite population

Basis exponential. pdf $f(x) = \lambda e^{-\lambda x}$
 $F(x) = 1 - e^{-\lambda x}$ $E[X] = \frac{1}{\lambda}$ $Var[X] = \frac{1}{\lambda^2}$

$X \sim \text{Exp}(\lambda)$ $Y \sim \text{Exp}(\mu)$ $\min(X, Y) \sim \text{Exp}(\lambda + \mu)$

MC transition matrix, homogeneous, absorbing / e.
 Ergodic \downarrow aperiodic ii) irreducible. iii) positive recurrent
 \downarrow π_i stationary dist.

fixation
 从某一个等位基因一直
 [N]

Moran simplest A.B 2-type fixed N population. state from A $i = 0, 1, 2, \dots, N$

$P_{i,i+1} = p(1-p)$ $P_{i,i-1} = (1-p)p$ $P_{i,i} = p^2 + (1-p)^2$ $p = i/N$ frequency of A

Each step only change 1 subjects inside population.

Birth-death process is a process with P (transition) and 1 subjects change in one step.

每次只能
改变

Fixation Probability x_i : prob. of getting to N starting from i

$x_i = \frac{i}{N}$ since each allele has the same chance of being fixed

More general cases

$P_{i,i+1} = \alpha_i$ $P_{i,i-1} = \beta_i$ $\alpha_0 = \beta_N = 0$

let $\gamma_i = \frac{\beta_i}{\alpha_i}$

$x_i = \frac{1 + \sum_{k=1}^{i-1} \prod_{k=1}^i \gamma_k}{1 + \sum_{k=0}^{N-1} \prod_{k=1}^i \gamma_k}$

Prob. of ending in N starting from i

absorption

Mean fixation time

$-N^2 [(1-p) \log(1-p) + p \log p]$ generation

Moran w/ constant selection

A waiting time $\sim \text{Exp}(\lambda_A = r)$ $B \sim \text{Exp}(1)$

r is the relative fitness of A over B

T_A T_B be the waiting times till birth.

$T_A \sim \min\{\text{Exp}(\lambda_A), \dots, \text{Exp}(\lambda_A)\} = \text{Exp}(\lambda_A i)$ $T_B \sim \text{Exp}((N-i)\lambda_B)$

$P(T_A < T_B) = \frac{r^i}{r^i + N - i}$

$P(T_A > T_B) = \frac{N-i}{r^i + N - i}$

competing exponential

Transit. Prob $\left\{ \begin{aligned} P_{i,i+1} &= \frac{r^i}{r^i + N - i} \cdot \frac{N-i}{N} \\ P_{i,i-1} &= \frac{N-i}{r^i + N - i} \cdot \frac{i}{N} \end{aligned} \right.$

$P_{i,i} = 1 - \dots$

Fixation Prob

$x_i = \frac{1 - 1/r^i}{1 - 1/r^N}$

Cancer

- imbalance of cellular cooperation
- somatic (Non-germ cells) evolution: selfish, uncontrolled replication
- accumulation of mutations [cellular cooperation]

Some hallmarks of cancer: disruption in i) integrity of genome ii) correct division iii) cell status iv) mobility.

Mutation type: i) point ii) chromosomal iii) indel iv) mitotic

Mutations accumulates if the rate of replication is high \rightarrow ex. colon cancer.

Tumor suppressor genes

Oncogenes: \uparrow fitness if one allele is mutated or inappropriately expressed

Fixation of oncogenes: Moran w/ N being the number of cells. $p = X_i = \frac{1-r^{-i}}{1-r^{-N}}$ $i=1$ if one mutation start

$P(t) = 1 - e^{-N\mu t}$: Prob. of a mutant is fixed by time t

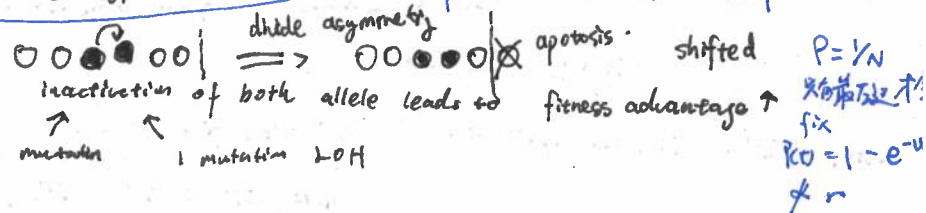
$P(t) \uparrow$ in N if $r > 1$ else decreasing.

large N increases the ~~rate~~ fixation probability so \Rightarrow slow it down by small cellular compartment.

Linear process \uparrow Have to look at video. Eg. cryptos in colon. Compartmentalization helps

1-D linear dimension process.

TSG: w/o it, you would get cancer \Rightarrow inactivation of both allele leads to fitness advantage \uparrow shifted $P = 1/N$



type 0 $\xrightarrow{u_1}$ type 1 $\xrightarrow{u_2}$ type 2 fitness advantage.

Fixation: $\frac{1}{P} = N \frac{1}{Nu_2}$. So type 1 cells reach fixation before type 2. if $N \ll \frac{1}{Nu_2}$.

Let $X_i(t)$ be probability of staying in $\begin{cases} 0 & \text{type 0} \\ 1 & \text{type 1} \\ 2 & \text{type 2} \end{cases}$

ODE $\begin{aligned} \dot{x}_0 &= -u_1 x_0 \\ \dot{x}_1 &= u_1 x_0 - Nu_2 x_1 \\ \dot{x}_2 &= Nu_2 x_1 \end{aligned}$ $X(0) = (1, 0, 0)$ at first "all healthy" \rightarrow $t \rightarrow 1$ $X(t) = (0, 0, 1)$

$P(t) = X_2(t) = 1 - Nu_1 e^{-u_1 t} - u_1 e^{-Nu_2 t}$

感兴趣: $P[\text{到一细胞被 hit 了两次}]$ 为略有两波

Time scales

short:

long:

very long:

$Nu_2 - u_1$

$t \ll 1/Nu_2$

$1/Nu_2 < t < 1/u_1$

$t \gg 1/u_1$

No rate limit

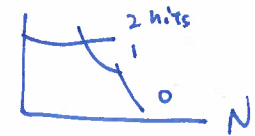
$P(t) \approx Nu_1 u_2 t^2 / 2$

$P(t) \approx 1 - \exp(-u_1 t)$

[2 rate limiting events]

[first hit is rate limiting]

"Tunnelling" of type 1 $1/Nu_2 \leq t \leq 1/u_1$



$P(t) = 1 - e^{-Nu_1 \int_0^t x_1(t') dt'}$

(the probability at least one cell with two hits has arisen before t)

Large population size ($N \gg 1/u_1$)

Type 1 are generated as soon as possible

$x_1(t) = Nu_1 t$

Why

Probly of Type 2 growth during type 1 growth

$P(t) = 1 - \exp \left\{ -u_2 \int_0^t x_1(t') dt' \right\} = 1 - \exp \left\{ -\frac{1}{2} Nu_1 u_2 t^2 \right\}$

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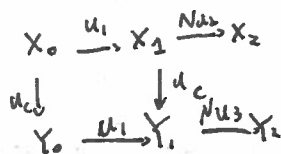
Genetic instability

MIN
microsatelliteCIN
Chromosomal
 $\begin{cases} \text{I} \\ \text{II} \\ \text{III} \end{cases}$

one is mutated or lost
one is mutated
both alleles

only
suppressor

Neutral CIN and $A^{+/+}$ are neutral $P = 1/N$. A^{-} will be fixed immediately.



ODE

$$\begin{cases} \dot{X}_0 = -(u_1 + u_c)X_0 \\ \dot{X}_1 = u_1 X_0 - (u_c + Nu_2)X_1 \\ \dot{X}_2 = Nu_2 X_1 \end{cases}$$

$$\dot{Y}_0 = u_c X_0 - u_1 Y_0$$

$$\dot{Y}_1 = u_c X_1 + u_1 Y_0 - Nu_3 Y_1$$

$$\dot{Y}_2 = Nu_3 Y_1$$

$$X_2(t) \approx Nu_2 u_1 t^2/2$$

$$Y_2(t) \approx u_1 u_c t^2$$

$Y_1 \approx Y_2$ if $u_3 = 10^{-2}$ is much higher

$Y_1 \rightarrow Y_2$ is negligible.

Costly CIN. in small apertures.

CIN with $r < 1$

$$P = \frac{1-r^{-1}}{1-r^{-N}}$$

note: CIN \rightarrow CIN with $Np u_c$

large aperture complicated.

$$X_2(t) = Nu_2 u_1 t^2/2$$

$$X_2(t) = Np u_1 u_c t^2$$

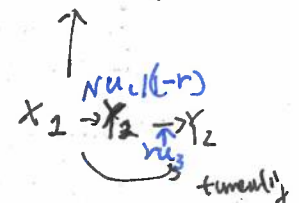
$$X_2(t) = Nu_2 u_1 t^2/2$$

$$Y_2(t) = R u_1 t^2/2$$

$$R = Nu_c u_3 / (1-r)$$

Why the difference b/w large and small aperture in CIN?

Np become so small such that $A^{+/+}$ CIN (Y_0) will NOT reach fixation.



ch 5. Cancer progression.

In the last chapter, we know that small compartment ($N \downarrow$) helps decrease the fixation probability? why.
In colon, epithelial cells replicates in a rapid rate. Cypto combined with linear evolution

Progression A few specific mutation \rightarrow adenoma ~~tumor~~ \rightarrow cancer

Multistage theory

cancer progress in a discrete state space in a linear fashion.
 $0 \xrightarrow{u_1} 1 \xrightarrow{u_2} 2 \xrightarrow{u_3} \dots \xrightarrow{u_k} k$ u_i are small \Leftarrow transition are rare.

$T_{01}, T_{ij} \sim \text{Exp}(u_i)$ $\gamma_k \sim \frac{\text{time to get to final state } k}{\text{Exp}(u_1) + \dots + \text{Exp}(u_k)}$ independent

$$E[\gamma_k] = \sum_{j=1}^k E[\text{Exp}(u_j)] = \sum_{j=1}^k \frac{1}{u_j}$$

why? since the independence of each transition and $\gamma_k = \min\{T_{01}, \dots, T_{(k-1)k}\}$ waiting the for k out of d .

$k=1 \quad \gamma_1 = \min\{T_{01}, \dots, T_{0d}\} \sim \text{Exp}(\lambda_1 + \dots + \lambda_d) = \text{Exp}(d\lambda)$ independent

Assume $\lambda_i = \lambda \forall i$ $\gamma_j \sim \gamma_{j+1} + \text{Exp}((d-j+1)\lambda) \Rightarrow E[\gamma_j] = \frac{1}{\lambda} \sum_{j=1}^k \frac{1}{d-j+1}$

Mutation pathway

Exit i : set of all possible mutations in step i

Exit 1 = $\{1, 2, 3\}$

Probability of a mutation pathway: $P = j_1 \rightarrow \dots \rightarrow j_k$

$\text{Prob}(P) = \prod_{i=1}^k \frac{\lambda_{j_i}}{\sum_{j \in \text{Exit } i} \lambda_j}$

competing exponentials

Expected time for P :

$E[\gamma_P] = \sum_{i=1}^k \frac{1}{\sum_{j \in \text{Exit } i} \lambda_j}$

The Wright-Fisher Process

- Haploid population of A/B of size N
- Reproduction occurs in discrete, non-overlapping generations. ?
- $X(t) :=$ number of A in size $N \in \{0, 1, \dots, N\}$

$X(t)$ is a MC. with $X(t+1) | X(t) \sim \text{Bin}(N, p/N)$ Markov chain

$P_{ij} = \text{Prob}[X(t+1)=j | X(t)=i] = \binom{N}{j} \left(\frac{i}{N}\right)^j \left(1 - \frac{i}{N}\right)^{N-j}$

Properties $E[X(t+1)] = E[X(t)]$

Variance $Ni(1-p) [1 - (1-p)^t]$

Absorbing state $X(t)=0$
 $X(t)=N$

let $x_i = \lim_{t \rightarrow \infty} P[X(t)=N | X(0)=i]$ Probability of absorbed starting from i

$i = \lim_{t \rightarrow \infty} E[X(t)] = 0(1-x_i) + N x_i \Rightarrow x_i = \frac{i}{N}$

Expected Fixation Time

$k_i :=$ no. of generation

Prob. of not fixed. Prob. of fix

$k_2 \approx 2(\log N - 1 + \gamma)$

Extension of Wright-Fisher

accumulating mutation

genome binary of length d $\{0, 1\}^d$

$X_j(t)$ be the number of cells with j^{th} mutation at time t .

$$x_j(t) = X_j(t)/N$$

At first $X_{00} = N$ ~~the first is no mutation~~

Assumed fitness $f(t, s)$

Ignore backward.

Probab

Probability of getting a j -cell [with j mutation]

$$\begin{aligned} \theta_j(t) &= \sum_{i=0}^j P[i\text{-cell} \rightarrow j\text{-cell}] = \sum_{i=0}^j \underbrace{P[i\text{-cell}]}_{\downarrow} P[i \rightarrow j] \\ &= \sum_{i=0}^j \frac{(1+s)^i x_i(t)}{\sum_{k=0}^d (1+s)^k x_k(t)} \Rightarrow \underbrace{(1-u)^{d-j}}_{\text{Binomial}} \underbrace{\binom{d-i}{j-i} u^{j-i}}_{\text{Binomial}} \end{aligned}$$

另, 此同是递推关系

Transition from $X_{t+1} \leftarrow X_t$

Prob of getting a i

$$P_{m,n} = P[X_{t+1} = n | X_t = m], \text{ multi-normal dist.}$$

generalization of binomial

mutation waves travel at constant speed

waiting time for j -mutated cells are fixed in the population.

Assumed: wave is Gaussian.

$$(1+s)^i \approx 1 + si \quad s \ll 1$$

$$X_{(j)}(t) = A \exp\left\{-\left(\frac{j - vt}{2\sigma^2}\right)^2\right\} \quad \text{Normal.}$$

$$A \approx \frac{1}{\sqrt{2\pi\sigma^2}}$$

v determined by decoupling clonal expansion
6 generations of new types.

$$\dot{x}_j = s x_j \left[j - \sum_{i=0}^d i x_i(t) \right] \Rightarrow v = s\sigma^2$$

Adaptation

generating a new mutant
 $v = 1/\tau$ is solved by

τ : average time it takes to produce a new mutant

$$X_{j+1}(\tau) = u d \int_0^\tau x_j(t) dt$$

speed of adaptation

$$v \approx \frac{2s \log N}{\log [s/(ud)]^2} \propto s$$

approxim. of asexual reproduction

waiting time to cancer [Time for the first cells w/ k mutations]

$$\begin{aligned} \tau_k &\propto \frac{1}{s} \\ &\approx \frac{k \log [s/(ud)]^2}{2s \log N} \end{aligned}$$

Selection \rightarrow dominates in large pop

Drift in small pop.

Probability of a population over time / instead of paths.

$\psi(p, t) :=$ a certain allele is at frequency p at time t

Evolutionary Process

Directional process $M(p)$: Non-zero expected change ^{in allele frequency.} _{selection/migration}

Non-directional process $V(p)$: Zero expected change. _{random}

$\psi(p, t)$ is



Let g : Probability that frequency changes from $p \rightarrow p+\epsilon$ in time interval dt _{drift} ^{Expected variance in the next gen. distr.}

$g(p, \epsilon; dt)$

$$\psi(p, t+dt) = \int \psi(p-\epsilon, t) g(p+\epsilon, \epsilon; dt) d\epsilon$$

$$(\text{Taylor expansion}) = \int \left[\psi(p, t) g(p, \epsilon; dt) - \epsilon \frac{\partial (\psi g)}{\partial p} + \frac{\epsilon^2}{2} \frac{\partial^2 (\psi g)}{\partial p^2} \right] d\epsilon$$

$$\begin{aligned} E\epsilon. \\ V(p) &= E[\text{Var } p(t+1) | p(t)] \\ &= \frac{2p(1-p)}{N^2} \end{aligned}$$

Approximation

$$\begin{aligned} \psi(p, t+dt) &= \underbrace{\psi(p, t) \int g(p, \epsilon; dt) d\epsilon}_1 - \frac{\partial}{\partial p} \psi \underbrace{\int \epsilon g(p, \epsilon; dt) d\epsilon}_0 + \frac{1}{2} \frac{\partial^2}{\partial p^2} \psi \underbrace{\int \epsilon^2 g(p, \epsilon; dt) d\epsilon}_{E(\epsilon^2) \propto \text{Var}(\epsilon)} \\ &= \psi - \frac{\partial}{\partial p} \psi \underbrace{\int \epsilon g(p, \epsilon; dt) d\epsilon}_0 + \frac{1}{2} \frac{\partial^2}{\partial p^2} \psi \underbrace{\int \epsilon^2 g(p, \epsilon; dt) d\epsilon}_{E(\epsilon^2) \propto \text{Var}(\epsilon)} \end{aligned}$$

$$\Rightarrow \psi(p, t+dt) = \psi(p, t)$$

$$- \frac{\partial}{\partial p} [\psi(p, t) M(p)] dt + \frac{1}{2} \frac{\partial^2}{\partial p^2} [\psi(p, t) V(p)] dt$$

$$\Rightarrow \frac{\partial \psi(p, t)}{\partial t} = - \frac{\partial}{\partial p} [\psi M] + \frac{1}{2} \frac{\partial^2}{\partial p^2} [\psi V]$$

change in allele frequency in terms of $\left\{ \begin{array}{l} \text{shape of dist. } M \\ V \end{array} \right.$

Equilibrium

$$\frac{\partial}{\partial t} = 0 = - \frac{\partial}{\partial p} \psi M + \frac{1}{2} \frac{\partial^2}{\partial p^2} \psi V$$

ODE

$$\frac{1}{2} \frac{\partial}{\partial p} \psi V - \psi M = 0$$

after integrate over p

$$\Rightarrow \psi^*(p) = \frac{C}{V(p)} \exp \left[\int_0^p \frac{2M(q)}{V(q)} dq \right]$$

\hookrightarrow accounts for Non-directional forces

Selective competition b/w 2 alleles.

A_1 A_2 p $(1-p)$

fitness w_1, w_2 .

$$\bar{w} = p w_1 + (1-p) w_2$$

$$\frac{d\bar{w}}{dp} = w_1 - w_2$$

In the next gen. $p' = \frac{p w_1}{\bar{w}}$

Bright's equation for adaptive landscape

For selection

$$\Delta p_{\text{sel}} = p' - p = \frac{p(w_1 - \bar{w})}{\bar{w}} = \frac{p(1-p)(w_1 - w_2)}{\bar{w}} = p(1-p) \frac{d \log \bar{w}}{dp}$$

Mutation

u_1 : rate from A_1 to A_2

u_2 : rate from A_2 to A_1

per-generation change due to mutation

Ch 7. Evolutionary Game Theory

Fitness may change according to the dynamics of the population. [Frequency-dependent]

$$\begin{cases} \dot{x}_A = x_A [f_A(x) - \phi(x)] \\ \dot{x}_B = x_B [f_B(x) - \phi(x)] \end{cases} \Leftrightarrow \dot{x} = x(1-x) [f_A(x) - f_B(x)]$$

x is the frequency

$$x^* = 0 \text{ if } f_A(0) < f_B(0)$$

$$x^* = 1 \text{ if } f_A(1) > f_B(1)$$

a third equilibrium is possible if

$$\frac{\partial f_A}{\partial x}(x^*) < \frac{\partial f_B}{\partial x}(x^*)$$

Evolutionary Games

Pay-off matrix

	A	B
A	a	b
B	c	d

row: expected pay-off

against A

• fitness as expected pay-off

$$f_A(x_A, x_B) = ax_A + bx_B$$

$$f_B(x_A, x_B) = cx_A + dx_B$$

let $x_B = 1 - x \Rightarrow$

$$\dot{x} = x(1-x) [(a-b-c+d)x + b-d]$$

See page 11

Nash equilibrium

If two players play the same and no one increase by changing strategy.

\rightarrow Nash equilibrium.

[best reply to itself]

A is strict NE if $a > c$

[选A是严格]

Similarly B is NE if $d > b$

NE $c \geq c$

$d \geq b$

Evolutionary stable strategy

If ϵ of B invades a All-A population

selection will oppose

If $\begin{cases} f_A(1-\epsilon) > f_B(\epsilon) \\ a(1-\epsilon) + b\epsilon > c(1-\epsilon) + d\epsilon \end{cases} \rightarrow$

$\text{if } \epsilon \rightarrow 0$

① $a > c$ or ② $(a=c, b > d)$

ESS

Hierarchy of Nash

for all $i \neq k$.

Unbeatable

strict

NESS

Nash

$a_{kk} > a_{ik}$ $a_{ki} > a_{ii}$

$a_{kk} > a_{ik}$

$a_{kk} > a_{ik}$

or $a_{kk} = a_{ik}$

$a_{ki} > a_{ii}$

$a_{kk} > a_{ik}$

选k是严格纳什

$$\phi(x) = \sum_i x_i f_i(x)$$

$$f_i = \sum_j x_j a_{ij}$$

Replicator equation

$$\dot{x}_i = x_i [f_i(x) - \phi(x)]$$

kind of takes the R.P.C game

$$A = \begin{pmatrix} 0 & 1 & -1 \\ -1 & 0 & 1 \\ 1 & -1 & 0 \end{pmatrix}$$

$$\phi(x) = 0$$

Case 1: $\det(A) > 0 \Rightarrow$ unique stable equilibrium

Case 2: $\det(A) < 0 \Rightarrow$ unique unstable equilibrium

$= 0$ neutral oscillation

通过 A 的 determinant 判断是否稳定

The prisoner's dilemma

$$\begin{matrix} \text{silent} & S & C \\ \text{confess} & \begin{pmatrix} -5 & -1 \\ 0 & -7 \end{pmatrix} \end{matrix}$$

example $\begin{matrix} C & D \\ C & \begin{pmatrix} 3 & 0 \\ 5 & 1 \end{pmatrix} \end{matrix}$

$$f_C < f_D \Rightarrow D \text{ dominates}$$

$$\begin{matrix} C & D \\ C & \begin{pmatrix} R & S \\ D & \begin{pmatrix} T & P \end{pmatrix} \end{pmatrix} \end{matrix}$$

$$T > R > P > S \quad R > \frac{T+P}{2}$$

R reward P punishment

[FI] Repeat opponent last round: (m : number of Round)

TFT vs ALLD.

$$\begin{pmatrix} \text{TFT} & \\ \bar{m}R & S + (\bar{m}-1)P \\ T + (\bar{m}-1)P & \bar{m}P \end{pmatrix}$$

TFT = never first to defect
Not greedy.

If $\bar{m} > (T-P)/(R-P)$ resist invasion

GRIM :: Cooperating at first as long as opponent does not defect

$$\text{GRIM} \begin{pmatrix} mR & S + (m-1)P \\ T + (m-1)P & mP \end{pmatrix}$$

Noise in TFT

long run is as long as choosing randomly b/c C and D

TFT vs ALLC

$\begin{pmatrix} \bar{m}R & \bar{m}R \\ \bar{m}R & \bar{m}R \end{pmatrix}$ TFT Not evolutionarily stable \Rightarrow in the long run TFT can become ALLC

PD as MC

b/w S_1, S_2 is a Markov chain of CC, DD, DC, CD.

M is the transition matrix for it

$X(t)$ be the prob. of CC, ... DD at time t / games

$$X(t+1) = X(t) \cdot M$$

May reach stationary.

$$E(S_1, S_2) = R s_1 s_2 + S s_1 (1-s_2) + T (1-s_1) s_2 + P (1-s_1) (1-s_2)$$

CTFT $S(1, 1/3)$ $f_0 \begin{pmatrix} 3 & 0 \\ 5 & 1 \end{pmatrix}$

Memory one: affected by Opponent's and own last move

$$\text{TFT} = S(1, 0, 1, 0)$$

$$\text{WSLS} = S(1, 0, 0, 1) \quad [\text{cooperate after CC or DD, defect after CD or DC}]$$

ch 8 Evolutionary Games in finite populations.

Moran Process

$$A \ B$$

$$\begin{pmatrix} a & b \\ c & d \end{pmatrix}$$

A: i individuals.

Pay-off: $A: F_i = \frac{(i-1)a + (N-i)b}{N-1}$ $B: G_i = \frac{ic + (N-i-1)d}{N-1}$

selection opposes A invading B if

$$F_i < G_i \Leftrightarrow b(N-1) < c + d(N-2) \text{ why "c"}$$

if $b < c$ invades selection

Intensity of selection

$$f_i = 1 - w + w F_i \rightarrow \text{Expected pay-off}$$

$$g_i = 1 - w + w G_i$$

$w=0$ very weak selection
Game does not contribute.
 $w=1$, full game effect

Fixation Probability

For a moran

$$P_{i,i+1} = \frac{f_i}{f_i + (N-i)g_i} \cdot \frac{N-i}{N}$$

Fixation def

$$P_A = \frac{1}{1 + \sum_{k=1}^{N-1} \prod_{i=1}^k (g_i/f_i)}$$

$$P_{i,i+1} = \frac{g_i}{f_i}$$

weak selection limit.

$$P_A \approx \frac{1}{N} \frac{1}{1 - (2N-3)w/6}$$

$$\alpha = a + 2b - c - 2d$$

$$\beta = 2a + b - c - 4d$$

if $P_A > \frac{1}{N}$ selection favors the fixation of A

$$\alpha > \beta$$

$$\alpha(N-2) + b(2N-1) > c(N+1) + d(2N-4)$$

$$\frac{1}{2}N = 2 \quad b > c$$

$$N \rightarrow \infty \uparrow$$

$$a + 2b > c + 2d$$

Equilibrium

$$x^* = \frac{F_i = G_i}{a-b-c+d}$$

if $a > c$ $d > b$,

A, B, $\frac{a}{a-b-c+d}$ is the best strategy.

$P_A > 1/N$ Favors

$$\Leftrightarrow x^* < \frac{1}{3}$$

β is ESSN if 1 $b(N-1) < c + d(N-2)$ Expect $F_1 < G$

2. $P_A < 1/N$ for $w > 0$.

$$a(N-2) + b(2N-1) < c(N+1) + d(2N-4)$$

Evolutionary Graph Theory

Mark process $\frac{1}{2}$ Full matrix Every one can be another / die and reproduce
However it does not reflect the spatial connection b/w

$$P = \frac{1 - 1/r}{1 - 1/r^N}$$

Directed cycle



$$\begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix}$$

A: 1 B: r. Taking # of B

$$P_{m-1} = \frac{1}{N-m+rm}$$

A: 1/2 B: 1/2

$$P_{m+1} = \frac{r}{N+m+rm}$$

same as Mon

$$P = \frac{1 - 1/r}{1 - 1/r^N}$$

Breeding cycle - 1/2

Linear

$$P = \frac{1}{N}$$

independent of r.

G is amplifier suppress $P_G > P_{mon}$
 $P_G < P_{mon}$

Strongest suppress $\frac{1}{N}$ & r

Isotone Thm

$P_G = P_{mon}$ iff G is isotone

$$T_{comp}: T_j = w_{1j} + \dots + w_{nj}$$

$$T_i = T_j \quad \forall i, j$$

Prof $V = (v_1, \dots, v_N)$ $v_i = \begin{cases} 0 & \text{occupied by A} \\ 1 & \text{occupied by B} \end{cases}$

$$m = \sum v_i$$

$$P_{m+1} = \frac{r \sum_{ij} w_{ij} v_i (1-v_j)}{rm + N - m}$$



$$P_{m-1} = \frac{\sum w_{ij} (1-v_i) v_j}{rm + N - m}$$

iff

$$\frac{P_{m-1}}{P_{m+1}} = \frac{1}{r}$$

for all v. \Rightarrow

$$\sum w_{ij} (1-v_i) v_j = \sum w_{ij} v_i (1-v_j)$$

$$\Rightarrow \sum_j w_{kj} = \sum_i w_{ik} = 1$$

cellular automata

Gridspace with each grid can take one of the states

Eden growth

2 state.

S_0 , S_1 occupied.

Rule. a site in the neighbor of S_1 switched to S_1 .

variants

Available: 有 S_1 邻居的 S_0 中随机选一个

Bond: $S_0 \rightarrow S_1$ & no. of S_1

Cell-fan: 选一个 S_1 邻居的 S_0 , 再随机选择 S_0

P deme-based models

每个 Deme 有各自的细胞. W.F. Proc Moran process

cells can migrate from Deme to another

↓
non-spatial.

当 $N_{deme} = 1$

↓

stochastic

Spatial Moran

n = # of mutation vector

+1 $w_i^+(n) = \frac{\mu(1+s)}{N} (N - n_i) \left[n_i + \frac{m}{2} u_i'' \right]$ 这里不需要再除以 N 吗?

-1 $w_i^-(n) = \frac{\mu}{N} n_i \left[(N - n_i) - \frac{m}{2} n_i'' \right]$
 $u_i'' = (n_{i-1} + n_{i+1} - 2n_i)$

μ death rate

s rel. diff. in fitness

ch10. Branching process in biology.

Galton-Watson Process

After a time, produce Random no. of offspring. Z according to a fixed distribution.

$\{Z_n | n=0, \dots, \infty\}$ is a MC.

$$\square \begin{cases} Z_1 \\ Z_2 \end{cases}$$

$$Z_{n+1} = \sum_{j=1}^{Z_n} Z_n^{(j)}$$

Time-homogeneous

Transition Probability

$$P_k = \text{Prob}(Z=k)$$

P_{ij} transition from $i \rightarrow j$

$$P_{ij} = \sum_{k_1 + k_2 = j} P_{k_1} \dots P_{k_i}$$

$$P_k = P(1, k)$$

$$P(2, j) = P_0 P_j + P_1 P_{j-1} + \dots + P_j P_0$$

Probability Generating Function

Discrete RV $Z \sim \{p_k | k \geq 0\}$

$$f(s) = E[s^Z] = \sum_{k=0}^{\infty} p_k s^k \quad s \in [0, 1]$$

generally distribution on P

$$\frac{d^k f}{ds^k}(0) = k! p_k$$

Properties

$$E[Z] = f'(1)$$

$$\text{Var}[Z] = f'(1) + f''(1) - f'(1)^2$$

$$f(s) = \sum_j P(1, j) s^j$$

Iteration Def

$$f^{(0)}(s) = s$$

$$f^{(1)}(s) = f(s)$$

$$f^{(n+1)}(s) = f \circ f^{(n)}(s)$$

$$(f(s))^k = \sum_j P(k, j) s^j$$

f_n be the pgf of Z_n .

$P_n(i, j)$ be the n -step transition probabilities.

The CK equations

$$P_{n+m}(i, j) = \sum_{k=0}^{\infty} P_n(i, k) P_m(k, j)$$

$$\underline{f_n = f^{(n)}}$$

$$\underline{f_{n+1}(s) = \sum_j P_{n+1}(1, j) s^j = \sum_j \sum_k P_n(1, k) P(k, j) s^j = \sum_k P_n(1, k) \sum_j P(k, j) s^j = \sum_k P_n(1, k) f(s)^k}$$

$P_{n+1}(1, 1)$ means

$$= f_n(f(s)) \quad \text{deduction}$$

$$= f^{n+1}(s)$$

Moment of Z_n

Assume $p_0 + p_1 < 1$

See Exercise

$$E[Z_n] = m^n$$

$$\text{Var}[Z_n] = \begin{cases} 6^2 m^{n-1} (m^n - 1) / (m - 1) & m \neq 1 \\ n 6^2 & m = 1 \end{cases}$$

Extinction

$Z_n = 0$ Absorbing state

$$P = \text{Prob}(Z_i = 0 \text{ for } i \geq 0) = \lim_{n \rightarrow \infty} \text{Prob}(Z_i = 0) = \lim_{n \rightarrow \infty} f_n(0) = \lim_{n \rightarrow \infty} f^{(n)}(0)$$

properties of pof

Power series

non negative coefficients adding up to 1

i f is strictly convex

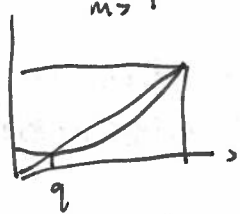
$$m = f'(1) < 1$$

ii $f(0) = p_0$ $f(1) = 1$

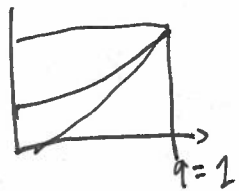
iii if $f'(1) = m \leq 1$ then $f(s) > s$ for $s \in [0, 1)$

iv if $f'(1) = m > 1$ $f(s) = s$ has a unique root in $[0, 1)$

$m > 1$



$m \leq 1$



Extinction Probability

Thm The extin prob. of GW process is the smallest non-negative root of $f(s) = s$
最小非负根为零

If $m \leq 1$ $q = 1$

if $m > 1$ $q < 1$

With Prob 1 $Z_1 \rightarrow 0$ or $Z_1 \rightarrow \infty$ as $n \rightarrow \infty$

Instability

$$\text{Prob}(\lim Z_n = 0) = q$$

$$\text{Prob}(\lim Z_n = \infty) = 1 - q$$

supercritical

$m > 1$

$$E[Z_n] > \infty$$

$q < 1$

critical

$m = 1$

$$E[Z_n] = 1$$

$q = 1$

sub-

$m < 1$

$$E[Z_n] < \infty$$

$q = 1$

Multiple Types

2 types 0 wt 1 mutant.

$$Z_0(t) \quad Z_1(t)$$

Type 0 $\xrightarrow{\alpha}$ Type 1

invisible

$$F = F_0, F_1$$

$$F_0(s_0, s_1; t) = E[s_0^{Z_0(t)} s_1^{Z_1(t)} | Z_0(0) = 1, Z_1(0) = 0]$$

$$F_1(s_0, s_1; t) = E[s_0^{Z_0(t)} s_1^{Z_1(t)} | Z_0(0) = 0, Z_1(0) = 1]$$

Recurrence equations

$$F_0(s; t) = [(1-\alpha) F_0(s; t-1) + \alpha F_1(s; t-1)]^2$$

$$F_1(s; t) = [F_1(s; t-1)]^2$$

Mutation

Differentiation

at $s=(1,1)$

$$E[Z_0(t) | Z_1(0) = \delta_{0i}] = [2(1-\alpha)]^t$$

$$E[Z_0(t) | Z_1(0) = \delta_{1i}]$$

$$= 2 E[Z_0(t-1) | Z_1(0) = \delta_{ii}] = 0$$

$$E[Z_0(t) | Z_1(0) = \delta_{0i}]$$

$$= 2(1-\alpha) E[Z_0(t-1) | \times]$$

$$\Rightarrow N(t) = E[Z_0(t) + Z_1(t) | Z_1(0) = \delta_{0i}] = 2^t$$

$$r(t) = E[Z_1(t) | Z_1(0) = \delta_{0i}] = 2^t - (2(1-\alpha))^t = 2^t [1 - (1-\alpha)^t]$$

Prob. of a mutant-free pop.

$$P_0(t) = F_0(1,0;t) = E[1_{0^{z_0(t)} 0^{z_1(t)}} \mid z_i(0) = \delta_{0,i}]$$

$$0^{z_1(t)} = \begin{cases} 1 & \text{if } z_1(t) = 0 \\ 0 & \text{if } z_1(t) > 0 \end{cases}$$

Ch 11 Evolutionary escape.

Reaching 1 before extinction

Simple ex in binary genome

1: mutant 0: WT

Partially ordered sets (Posets): a set \mathcal{E} together with a binary relation, denoted " \leq "

- properties
- reflexive: for all $e \in \mathcal{E}$, $e \leq e$
 - antisymmetric: for all $e_1, e_2 \in \mathcal{E}$ ($e_1 \leq e_2$ and $e_2 \leq e_1$) $\Rightarrow e_1 = e_2$
 - transitive: $e_1 \leq e_2$ and $e_2 \leq e_3 \Rightarrow e_1 \leq e_3$

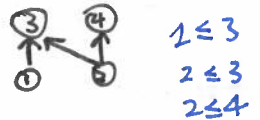
If $e_1 \neq e_2 \Rightarrow e_1 \leq e_2 \Rightarrow e_1 < e_2$

$e_1 < e_2$ is a cover relation.

Distributive lattice

A order ideal g in a poset \mathcal{E} is a subset of \mathcal{E} that is closed downward if $e_2 \in g$ and $e_1 \leq e_2$, then $e_1 \in g$

Example:



The distributive lattice is the set of all order ideal $J(\mathcal{E})$ under inclusion. $J(\mathcal{E})$ 都可以找到 sup, inf 对 J 都成立

- $J(\mathcal{E})$ forms a poset
- every pair of g, g_2 has a unique infimum and unique supremum.

The genotype lattice

- \mathcal{E} be a set of $n = |\mathcal{E}|$ irreversible genetic events (e.g. point mutation at different loci, indels, genomic)
- The poset indicates that the mutation can be accumulate.
- the order ideals g of $J(\mathcal{E})$ are precisely the subset of \mathcal{E} that are compatible with the partial order. [genotypes that can evolve subject to the order constraints]

Empty poset

- $\mathcal{E} = \{1 \dots n\}$ (no relations)
- $G = J(\mathcal{E})$ is the set of all subsets of \mathcal{E} / Boolean lattice of $\{0, 1\}^n$
- All genotypes can occur in any order; there are no order constraints.

Chain

A chain in $G = (J(\mathcal{E}), \leq)$ of length k is a totally ordered subsets $g_1 \subset g_2 \dots \subset g_k$. how to compare g_1, g_2

[Mutation pathways in cancer]

Fitness landscape

$$f: G \rightarrow \mathbb{R}$$

Mutational neighbourhood: The mutational neighbourhood of a genotype $g \in G$ is the set of genotypes that can be reached by mutation.

$$N(g) = \{h \in G \mid g \subset h\}$$

can be

$$[g \subset h] \text{ 表示 } g \rightarrow h$$

k-steps offspring

Set $F = \text{diag}(f)$ f is the fitness F is a diagonal matrix ($m \times m$)

Mutation matrix

$m = |G|$ mutation matrix $U = (U_{gh})_{g,h \in G}$ by $\pi_{e \in h, g} \mu_e$ if $h \in N(g)$
Entry i, j : $i \rightarrow j$ 多少步 mutation 概率 0 otherwise

The entry (g, h) of $U \cdot F$ is the probability of genotype h produced in one step. $U \cdot F = F \cdot \text{diag}(f)$
 (g, h) of $(U \cdot F)^k$ k step. producing h along the way

$$B = (I - U \cdot F)^{-1} - I = U \cdot F + (U \cdot F)^2 + \dots + (U \cdot F)^n$$

Properties

The entry b_{gh} of B is zero, unless $[g \subset h]$ $g \rightarrow h$

If $[g \subset h]$ then: $b_{gh} = \frac{\mu_{gh} f(h)}{P_{gh}(f)}$ where P_{gh} is a polynomial function of degree $(h'g) - 1$ on P

P_{gh} is the generating function for all chains from g to h in G

the risk polynomial $f(g) = f_g \quad g \in G$. case $g=0, h=1$

$P_{03}(f) = \sum_{0=g_0 < g_1 \dots g_k=1} f_{g_1} \dots f_{g_{k-1}}$ "the sum over all chains from 0 to 1 in the genotype lattice G "
 $0 \rightarrow 1$ all possible combinations.
 [Revise chain]

$R(g:f) = P_{03}(f)$ is the risk polynomial

assumption Let R_g be the basic reproductive ratio of an invading pathogen.

Interested in the case where $R_1 > 1$ and $R_g < 1$ for all $g \neq 1$

Fitness landscape $f_g = \frac{R_g}{1-R_g} = R_g + R_g^2 + R_g^3 + \dots$ (useful)

Then $R_g = \frac{f_g}{1-f_g}$ $f_g \approx R_g$ for $g \neq 1$

Multi-type Branching Process

Eg. G w/ a Poisson offspring distribution.

$P_{gh}^k = \text{Pois}(k; u_{gh} R_g) = \frac{(u_{gh} R_g)^k}{k!} e^{-u_{gh} R_g}$

The risk of escape

Let ϵ_g be the probability of escape (reaching 1 before extinction) starting w/ one individual of type g .

$\epsilon_g :=$ Probab. reach 1 [escape] $1 - \epsilon_g :=$ Probability of extinction

$1 - \epsilon_g = \prod_{h \geq g} (1 - \epsilon_h)^k P_{gh}^k$. P_{gh} is defined as before $u_{gg}=1$
 what does k mean.

Recurrence equation

substituting the Poisson distr. $\epsilon_g < 1$ For $g \neq 1$ $\epsilon_g < 1$ and $R_g^2 \approx 0$

$\log(1 - \epsilon_g) = - \sum_{h \geq g} \epsilon_h u_{gh} R_g$. $\epsilon_g \approx R_g \sum_{h \geq g} \epsilon_h u_{gh}$

In particular, $\epsilon_0 = f_0 \sum_{h \in G} \epsilon_h u_{0h}$

$\Rightarrow \epsilon_0 \approx \epsilon_0 f_0 \prod_{g \in G} \mu_g R(g; f)$. How to solve this recursion

The risk of escape of N WT pathogens is

$1 - (1 - \epsilon_0)^N \approx 1 - e^{-N \epsilon_0}$ 假设 ϵ_0 足够小.

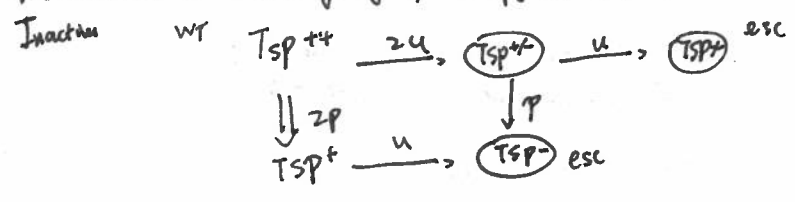
The critical population size

$N^* = 1/\epsilon_0$

$\gg N^* \Rightarrow$ escape is almost certain.
 $= N^* \Rightarrow$ risk of escape $1 - 1/e$
 $< N^* \Rightarrow$ $1/e :=$ prob. of successful intervention
 escape is almost impossible.

Example: Anti-cancer therapy.

Cancer cells can be x chemotherapy by TSG inactivation. We assume that at the beginning of therapy all cells have TSG.



the coalescent

WF process

To get into gene that gives rise to the final offspring.

Branching time (looking back into root)

Branch time is the time where lineages break off

[coalescent event]



All started at this

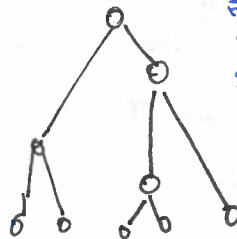
Coalescent time

Time for lineages for multiple to coale. / come together.

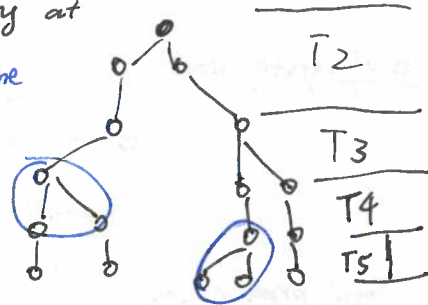
In principle you can have multiple of these events happening at the same time.

Interested in branch length

Every vector is a coalescent event.



then m.r.p can be simplified as such



1 coalescent event

The prob

of j genes has no common ancestor in previous generation.
= 一代不同色

$$j=2 \quad (1 - \frac{1}{N}) \quad \text{选了一个}$$

$$j=3 \quad (1 - \frac{1}{N}) (1 - \frac{1}{N} - \frac{1}{N}) \quad \text{选了两个}$$

$$\Rightarrow \prod_{i=1}^{j-1} (1 - \frac{i}{N}) = 1 - (\frac{j}{2})N^{-1} + O(N^{-2})$$

Coalescent time

let $T(j)$ be coalescent time b/w j and $j-1$ genes.

$$P[T(j) > t] = \left[\prod_{i=1}^{j-1} (1 - \frac{i}{N}) \right]^{Ne} \rightarrow \exp \left\{ -(\frac{j}{2})t \right\} \text{ as } N \rightarrow \infty$$

pairwise coalescent process

多线没有 coalescent

exponential dist

MRCA

$$T_{MRCA}(S) = T(S) + \dots + T(2) = \sum_{j=2}^n T(j)$$

How old is the tumor.

Expectation

$$E[T_{MRCA}(n)] = \sum_{j=2}^n E[T(j)] = \sum_{j=2}^n \frac{2}{j(j-1)} = 2 \sum_{j=1}^n \frac{1}{j-1} - \frac{1}{j} = 2(1 - \frac{1}{n})$$

2 generations?

$$E[T(2)] = 1$$

Variance

$$Var(T_{MRCA}(n)) = \sum_{j=2}^n Var(T(j)) = \sum_{j=2}^n \left(\frac{2}{j(j-1)} \right)^2 = 4 \sum_{j=1}^n \left(\frac{1}{j-1} - \frac{1}{j} \right)^2$$

why \rightarrow i.i.d.

$$Var(T(2)) = 1 \Rightarrow$$

T_{MCA} is dominated by $T(2)$

by 1st split 最重要

Huge Variation b/w any two coalescent event $T(2)$

Mutation

Poisson process puts down mutation at $\frac{\theta}{2} = \theta n$
 $\frac{1}{2}$ Branch.



Infinite sites model.

- 有 1 个位点 for gene
- infinite # of alleles.
- Each mutation will create a new mutation

每个 mutation 都在新位点!

of segregating sites.

大概有多, 不同

S equals the total # of mutations of the genealogy

Number of ~~the~~ column that are not identical.

Total branch len. Under Infinite Site Model

$$T_{tot}(n) = \sum_{j=2}^n j T(j)$$

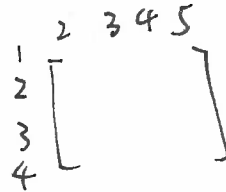
$$E[S] = \frac{\theta}{2} E[T_{tot}(n)] = \frac{\theta}{2} \sum_{j=2}^n j \frac{1}{\binom{n}{j}} = \theta \sum_{j=2}^n \frac{1}{j-1} = \theta (H_n) \leq \ln n + 2$$

Average pairwise nucleotide distance K

Expected value of

$$E[K] = \frac{\theta}{2} E[T(2)] = \theta$$

parallel lines



Hamming distance

Detecting selection

Under neutral infinite site model.

$$E[K] = \theta = n^{-1} E[S] \leftarrow \text{sensitive to low-allele frequencies}$$

strongly affected by allele frequencies

Tajima's D

H_0 : no selection

$$D = \frac{\hat{\pi} - n^{-1} \hat{S}}{\sqrt{\hat{\pi}}}$$

Variable

\Rightarrow To check whether it's outlined.

This is essentially a Random tree generator