Dynamic model of activity-type choice and scheduling

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Abstract This paper presents a model for the choice of activity-type and timing, incorporating the dynamics of scheduling, estimated on a six-week travel diary. The main focus of the study is the inclusion of past history of activity involvement and its influence on current activity choice. The econometric formulation adopted, explicitly accounts both for correlation across alternatives and for state dependency. The results indicate that behavioral variables are superior to socio-economic variables and that consideration of the correlation pattern over alternatives clearly improves the fit of the model. This is a first but significant contribution to changing the current static demand models into dynamic activity based ones. The availability of other multi-week travel surveys and the progress made recently on advanced econometric techniques should encourage the transferability of this study to different regions or model scale.

Keywords Activity involvement · Past history · Multi-week travel diary

Introduction

Transportation research has seen, in recent years, an increased analysis of the complexity of travel decisions. The need for a more behaviorally sound framework has led to a new generation of transport model systems called activity-based models (Bowman and Ben-Akiva 2000; Nagel and Rickert 2001; Bhat et al. 2004; Arentze and Timmermans 2004; Pendyala 2004; Salvini 2005). The general concept is that household and individual travel patterns are generated from the desire to pursue activities in time and space. However, their desires are limited by personal/physical constraints and by the institutional, economic and cultural context. Moreover, activity based models are not limited to one single purpose trip or tour, but often account for all trips/tours on a daily basis or on longer periods.

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An activity-based model consists of a multi-facet choice process; amongst its various components activity participation has certainly a core position but has seen relatively modest attention (Arentze 2000). Activity type preferences are, explicitly, modeled by Hamed and Mannering (1993) as part of a modeling system estimated for post-work activity patterns. The assumption they made is that a traveler is likely to choose the activity providing the maximum utility. The choice set consists of four alternatives: shopping, free-time, personal business and chain of activities; a number of socio-demographics, working schedule variables and previous activity indicator are estimated using a multinomial logit. Bhat has published (i.e. 2000, 2005) several papers on activity participation both for weekdays and weekends. These studies also develop advanced econometric techniques with flexible structures to accommodate complex error structures among different alternatives. In particular Bhat and Srinivasan (2005) propose an interesting framework, now accepted by the activity-based community; they assume that: "the decisions regarding pattern-level attributes (note by the authorsnumber of tours) are driven by the basic activity needs of the individual (and the household of which the individual is a part). Consequently, and consistent with the derived demand philosophy of the activity-based approach, the pattern-level decisions are considered to be at the highest level of the analysis hierarchy".

Other approaches using Poisson related methods (Ma 1999), multivariate negative binomial (Kockelman 1999) and structural equations (Lu and Pas 1999; Golob 2000) have been used to predict activity frequencies; all the studies cited here use multi-day travel surveys, although the number of observations from the same individual is often limited to 2 days.

It is often argued that past history of activity involvement may have potential influence on current activity choice. However, because the data available to transport modelers are often limited to a single day trip diary, the activity type choice models existing in the literature are estimated with the assumption of independence from past involvement (Hamed and Mannering 1993). Furthermore, in the 1980s and 1990s several researchers have pointed out that there is a significant day-to-day variability in travel behavior. A 1988 special issue of *Transportation* provides excellent insight into the analysis of variability in multi-day travel activity behavior. Jones and Clark focus on the policy implications of variability analysis and encourage the collection of more multi-day travel surveys. Eric Pas distinguishes long-term patterns from daily behavior and finds that the latter is independent of individual characteristics, when accounting for both. Hanson and Huff, by using a specific measure of repetition and variability, conclude that a 1-week record of travel does not capture longer-term travel behavior. In a more recent work, Pas and Sundar (1995) assess that this variability should have important implications for data collection, model estimation and model interpretation. Data collected over a period of time are still rare, but two recent studies have made available to researchers two continuous surveys on a 6-week period (Axhausen et al. 2002, 2007) (see Schlich 2004 for an analysis of the similarity of the activity patterns over time). Longitudinal observations of individual activity behavior allow the estimation of dynamic models. There are very few example of dynamic, disaggregate choice models in the existing activity-based literature. Hirsh et al. (1986) estimate a parametric model of dynamic decision-making process for weekly shopping activity behavior. The individual is assumed to proceed from period to period and the observed weekly activity pattern is the outcome of successive decisions. Action plans are then modified on the basis of actual behavior and of the additional information acquired in previous periods. Mahmassani and Chang (1986) investigate the dynamics of departure

¹ The longer-term is referred to a five-week period over which the travel survey used was conducted (Uppsala household travel survey, 1971).



time of urban commuters in a series of simulation experiments; user departure time decisions are modeled by means of heuristics, which also incorporate the experience cumulated by repeated use of the same facility. Very recently, Habib and Miller (2008) have presented an econometric model for daily activity program; the estimations are done using CHASE, a one-week travel survey collected in Toronto. The model accommodates dynamics in activity travel behavior; in particular, within-day dynamics in time-use is represented by the number of working hours, while day-to-day dynamics is explained by the number of activities executed the previous day. Dynamics and variability in within day mode choice has been studied by Ramadurai and Srinivasan (2006); variation over time of explanatory attributes, influence of state-dependence, role of habit persistence and the effect of unobserved preference heterogeneity were modeled by using mixed logit models. Dynamics in mode choice was also studied by Srinivasan and Bhargavi (2007) to account for rapid and substantial changes in fast growing Indian economy; data used for this model were recorded on two time points over a period of 5 years (2000–2005).

Dynamic models have been applied more extensively to explain car ownership behavior; Goodwin and Mogridge (1981) note that "the resistance to change" is one of the aspect that model based on cross sectional data fail to account for. More technical questions arising in dynamic models, like state dependency versus heterogeneity, have also been addressed in household car ownership models by Kitamura (1988) and by Kitamura and Bunch (1990).

The present paper intends to introduce dynamics in travel behavior by explicitly accounting past activity involvement when modeling activity choices and their scheduling. Exogenous variables describing present and past activity involvement are calculated on short-term period (day), medium-term period (week) and long-term period (multi weeks). The survey, from which data have been derived, is Mobidrive, a six-week travel survey conducted in Germany; all activities and the related tours are accounted in the framework applied. The correlation across alternatives and the state dependency across choices made on the same period of time (day or week) or by the same person are modeled via a mixed logit structure with error components and repeated choices. The models presented are then validated and applied to reproduce aggregate choices on the 7 days of the week and during the six weeks.

The model formulation

The development of the dynamic activity type choice model proposed in this paper is based on discrete choice methods. We make use of mixed logit models and in particular we refer to the error component interpretation on panel data as described by Train (2003). This formulation is well adapted to our problem because it is able to deal with correlation over alternatives in the stochastic portion of utility, and it allows efficient estimation in presence of repeated choices by the same respondent. We assume that in each of the T time periods observed each person faces a choice among the alternatives in set J. The choice set and the number of choice situations can vary over people. The utility that person n obtains from alternative j in choice situation t is:

$$U_{nit} = \beta x_{nit} + \mu z_{nit} + \varepsilon_{nit} \tag{1}$$

where β is a vector of unobserved coefficients, x_{njt} is a vector of observed variables relating to person n, alternative j and choice situation t; μ is a vector of random terms with density $f(\mu)$, z_{njt} is a vector of error components, ε_{njt} is an unobserved random term i.i.d. extreme value distributed.



In our specification the coefficients β are assumed to be constant over individuals and over choice situations for a given period; in other terms we are not interested here in measuring individual taste variations. This assumption is appropriate if the decision-maker's tastes are stable over the observed time period. The systematic part of the utility function is additive. The terms in z_{njt} along with ε_{njt} , define the stochastic portion of the utility. The correlation pattern over alternatives depends on the specification of the vector z_{njt} . Various correlation structures can be, in fact, obtained by an appropriate choice of variables to enter as error components. In this model, we specify a dummy variable for each group k that equals 1 for each alternative in the group and zero for alternatives outside the group. The random terms μ that are multiplied by the error components are assumed to be independently normally distributed with zero mean and variance $\sigma_k N$ $(0, \sigma_k)$; the variance σ_k measures the magnitude of the correlation.

Considering a sequence of specific activity episodes, one for each time period $t = \{1,..., T\}$, the probability that the person makes this sequence of choices is the product of logit formulas (Revelt and Train 1998):

$$L_{nj}(\beta,\mu) = \prod_{t=1}^{T} \left[\frac{e^{\beta x_{nj_t t} + \mu z_{nj_t t}}}{\sum_{j \in J} e^{\beta x_{nj_t} + \mu z_{nj_t}}} \right]$$
(2)

where j_t is the alternative choice made by the individual n on time period t. We consider here four time periods: (1) the single activity episode (T = 1), (2) the day, (3) the week, (4) and the entire set of individual activity episodes.

The unconditional probability is the integral of this product over all values of μ :

$$P_{nj} = \int L_{nj}(\beta, \mu) f(\mu) d\mu. \tag{3}$$

The integrand is a product of logit formulas, one for each time period or group of choice situations.

The vector of unknown parameters is then estimated by maximizing the log-likelihood function, i.e., by solving the equation:

$$\max_{\beta,\mu} LL(\beta,\mu) = \max_{\beta,\mu} \sum_{1}^{N} \ln P_{nj}(\beta,\mu). \tag{4}$$

This involves the computation of P_{nj} (β , μ) for each individual n, n=1,...,N, which is impractical since it requires the evaluation of one multidimensional integral per individual. The value of P_{nj} (β , μ) is therefore replaced by a Monte-Carlo estimate obtained by sampling over μ , and given by

$$SP_{nj}^{R} = \frac{1}{R} \sum_{r=1}^{R} L_{nj}(\beta, \mu)$$
 (5)

where R is the number of random draws μ_r , taken from the normal distribution. As a result, μ is now computed as the solution of the simulated log-likelihood problem.

$$\max_{\beta,\mu} SLL^{R}(\beta,\mu) = \max_{\beta,\mu} \frac{1}{N} \sum_{n=1}^{N} \ln SP_{nj}^{R}(\beta,\mu). \tag{6}$$



We will denote by a $(\beta, \mu)_R^*$ solution of this last approximation (often called Sample Average Approximation, or SAA), while $(\beta, \mu)^*$ denote the solution of the true problem (4).

Our model has been calibrated using the software AMLET, which uses an adaptive stochastic programming algorithm to estimate the objective function (Bastin et al. 2006).

The data and the framework

The multi-day database used for the dynamic activity type model is derived from a six-week travel diary held in Karlsruhe (Germany) in 1999, part of the *Mobidrive* survey, which involved about 160 households and 360 individuals in the main survey. Further details on data collection techniques and descriptive results on rhythm of daily life can be found in Axhausen et al. (2002) and PTV AG et al. (2000).² The recorded days were structured according to the framework proposed by Bhat and Singh (2000) and extended by Cirillo (2001). We briefly describe the principal elements of the schema adopted here to define the activity chains on a daily basis. The schema is also presented in Fig. 1, where activities considered and their possible schedules are schematically represented. We distinguish working days, i.e., days including a commute or school trip, and non-working days. Given the multi-week nature of the survey, the days of one individual can belong to either category. All trips are grouped into tours. A tour is the sequence of trips performed by an individual, starting from a given base (usually home or workplace) until the individual returns to this base. Each tour has a main activity defined by duration, purpose and main mode.

For each daily chain we define: the *main activity* of the day as work/education for working days and the *principal activity* of the day as the longest duration out-of-home activity for non-working days. All daily activity chains are represented in relation with this pivotal activity.

The work tour is divided into outbound and return legs, which are called the *morning* and *evening commute*. All activities that take place before the morning commute will be referred to as *morning activities* and the associated displacements grouped into one or more *morning tours*; they constitute the *morning pattern*. Similarly, all activities taking place after the return from work to home (the evening commute) will be referred to as *evening activities* and the associated displacements grouped into one or more *evening tours*; which together constitute the *evening pattern*. Additionally, all activities taking place outside the work location after the morning, but before the evening commute will be called *midday activities* and the associated displacements, whose origin and destination are at work, are grouped into one or more *midday tours*, in turn aggregated into the *midday pattern*.

We organize the daily pattern of the non-working days in a similar manner. The morning pattern represents the activities and travel undertaken before leaving home to perform the principal activity of the day. The *principal activity pattern* represents the activities and travel performed within the tour comprising the principal activity of the day; the principal activity tour is divided into outbound and return legs. Note that this definition implies that the principal activity pattern always consists of a single tour. The (afternoon and) evening pattern comprises the activities and travel of individuals after their return home from their principal activity. By applying this framework to Mobidrive we obtained

² See also http://www.ivt.ethz.ch/vpl/research/mobidrive for a list of the papers employing the Mobidrive data.



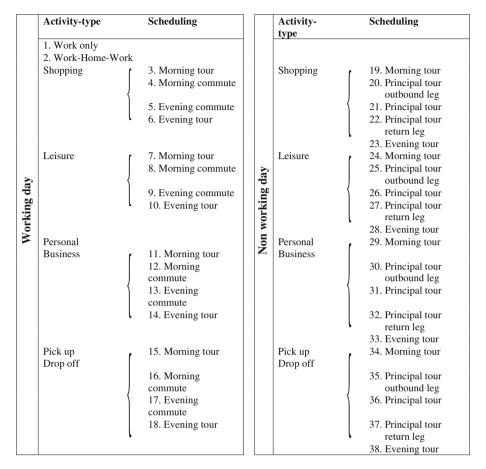


Fig. 1 Activity-scheduling choice tree

4,952 activity episodes, 3,212 daily schedules, 773 weekly schedules and 144 individual schedules. Table 1 shows the number of activity episodes for each purpose and the relative scheduling for working and non-working days.

Choice set, model structure and variables' definition

The model developed to generate the activity programs is disaggregate and based on individual choices. The alternatives are discrete and are exogenously defined by the analyst. A number of assumptions/hypothesis needed to be made in order to define the choice set. In fact, one of the main difficulties in modeling activity patterns is that analysts usually observe just the chosen pattern and it is difficult to reconstruct the entire set of alternatives available to the individuals. In this study, we consider only out-of home activities and the decision process is modeled at the tour level³; however, multiple tours on the same day are

³ This choice model is based on single tour decisions and therefore differs from the one proposed by Bowman and Ben-Akiva (2000), where the full activity pattern for the given day was modeled.



Table 1 Number of activity episodes

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Type	Shopping	Leisure	Personal business	Pick up/drop off	Shopping Leisure Personal business Pick up/drop off No extra activities Home from work	Home from work	All
Working day							
Morning tour (MT)	41	14	41	5			101
Morning commute (MC)	44	24	92	26			159
Evening commute (EC)	207	103	105	40			455
Evening tour (ET)	143	376	06	54			663
Work as the only activity (W)					547		547
Work-home-work (WHW)						159	159
Non working day							
Morning tour (MT)	372	126	163	72			733
Principal tour outbound leg (PT out-leg)	88	35	06	54			267
Principal tour (PT)	404	859	251	64			1,377
Principal tour return leg (PT ret-leg)	73	59	40	14			186
Evening tour (ET)	76	93	62	53			305
Total	1,469	1,488	206	382			4,952



same as tour-based mode choice

possible. Tours with multiple activities are included in our model; in this case we just model the choice of the principal activity. In a comparative study, the first author has analyzed activity patterns in four European countries (*including Mobidrive*) and concluded that secondary activities are executed at the most in 10% of the total number of (morning, midday and evening) tours. Our approximations seem perfectly reasonable in this context.

For each tour generated by an out of home activity episode, individuals are supposed to make a two-steps choice: first they choose the type of activity to be pursued and then they schedule the chosen activity within a specific time frame. Four activity types are possible: shopping, leisure, personal business and pick up/drop off. In the original data set more activity categories were available (i.e., long term shopping vs. short term shopping); however, in order to reduce the total number of discrete alternatives and to avoid a low number of observations for each category, we grouped them into the above main activities. Moreover, the activity "other-purposes" was excluded given the small number of observations reported for this category. The scheduling is intended to be the decision regarding the specific time frame when the activity is executed; in this model time frames are defined with respect to the main activity of the day. As follows from the framework described in Sect. 3, the main activity of the day is by definition the work activity for working days and it is the activity with the longest duration for non-working days. For working days the possible times frame are: morning tour, morning commute, evening commute and evening tour; work based tour to home during lunch time are also considered. Similar is the choice set for non-working days; the main difference consists in the possibility to execute secondary activities on the outbound and return leg of the principal tour. The combinations of activity types and of time frames define the alternatives in the choice set. A total number of thirty-eight alternatives are modeled; the relative choice set is reported in Fig. 1.

The availability of each alternative for a given individual was then determined on the basis of the activity patterns reported during the entire survey period. In particular, if an activity type was never reported by a respondent during the 6-week period, this activity type is considered not available for the individual. Shopping alternatives are available from 8:00 am to 6:00 pm; they are not available on Sundays except if they have been chosen.

We now turn our attention to the variables used in model estimation. Variables types and their categories are listed in Table 2; all the variables enter the utility function in an additive form. We estimate (1) socio-economic variables, (2) level of service variables and (3) individual activity involvement variables. Socio-economic variables are added as constant and are activity specific. Age (by categories), number of young children in the household, number of working hours per week, employment status and sex enter the final specification. Level-of-service variables are included through the logsum of a mode choice model. This model is a simplified version of an existing mode choice model (Cirillo and Axhausen 2006), which is estimated on five alternatives (car as driver, car as passenger, transit, walk and bike) and contains a full set of alternative specific constants, travel time and travel cost. To account for the variability in value of travel time savings, four mode choice models were estimated, one for each activity-type considered in our final activity programs. For the non-chosen activity-type alternatives time and cost variables are calculated as the average on all the locations visited by each individual for the same purpose. Each individual is supposed to know the average durations of the possible activities to be performed and the travel times and travel costs to reach the relative locations. Patternrelated variables are defined on three different time horizons: the day, the week and the

⁴ Sunday trading is still the exception in Germany, and it was more so in 1999, when the survey was conducted. Still, a limited number of outlets at train stations, gas stations and bakeries were available.



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Level	Variables	Description/categories	Type
Day	Activity duration (min)	It is randomly drawn from the vector of activity (with the same purpose) durations, reported by the same individual, over the entire survey period.	Continuous
Day	Time budget (min)	It is calculated as 24 h minus the time spent on previous activities (time at home included) and previous travel.	Continuous
Day	Available time before work (min)	It is the time available between the shop opening hour (8:00 am) and the arrival time to work.	Continuous
Day	Available time after work (min)	It is the time available after the departure from work and the assumed shop closing time (6:00 pm).	Continuous
Week	High week episode	Dummy variable, which is 1 if, for the week considered, the number of activity episodes with a specific purpose is greater than two.	Dummy
Six-week	Last time	Number of days since the day when the same activity was undertaken last	Discrete
Six-week	Immobile days	Number of days that the individual spent at home since the first day of the reporting period.	Discrete
Level of service	Logsum	Logsum of the mode choice model.	Continuous
Individual socio-economic	Age	Age 6–18/age 26–35/age 51–65	Dummy
variables	Presence of children	Number of children under 12	Dummy
	Professional status	Number of working hours per week Full time worker Female and employed part-time	Continuous Dummy Dummy



Table 3 Model fit (final log-likelihood/rho-squared adjusted)

Model	Degrees of freedom	Logit	Mixed—err comps	Mixed—day	Mixed—week	Mixed—individuals
Socio-economic variables	16 + 4 e.c.	-11952.6 0.024	-11845.1 0.033	-11850.8 0.032	-11685.3 0.046	-11351.8 0.073
Day behavioral variables	32 + 4 e.c.	-10777.7 0.119	Convergence not achieved	-10643.9 0.129	-10495.3 0.142	-10154.3 0.169
Week behavioral variables	36 + 4 e.c.	-9868.7 0.193	-9797.0 0.198	-9847.3 0.194	-9862.6 0.193	-9758.1 0.201
Long-term behavioral variables	46 + 4 e.c.	-9194.9 0.247	_9134.9 0.251	-9179.7 0.248	-9189.3 0.247	-9149.0 0.250

e.c. = error component



entire survey period. Durations are estimated across activities and different coefficients are estimated per different time frames; the remaining pattern-related variables are estimated as activity specific. For non-chosen alternatives the activity durations are drawn from the vector of activity (with the same purpose) durations, reported by the same individual, over the entire survey period. Again this is a strong assumption, but it is reasonable to assume that individuals are aware of the possible durations and of the locations of their activities and that they consider the time to be spent at the destinations when deciding their activity agenda. In discrete-continuous choice model of activity-type and duration time is allocated just if the activity is chosen (Munizaga et al. 2008). The implicit assumption of hazard joint model of activity type choice and duration is that the day is planned coherently in one step (Bhat 1996). We make here the assumption of independence of the forward looking choices. The variables then are only the result of past choices, which are then assumed to be taken as given by the traveler. Time availability for shopping activities and weekly frequency were specified as in Hirsh et al. (1986). The definition of policy related variables responds to the need of more accurate forecasting in activity-based models. In fact, the model proposed in this paper is sensitive to both work duration and time availability before and after working activities. Day-to-day dynamics are explained by the variable that counts the number of days passed since the same activity was performed last. Individual past history is also represented by the number of days spent at home since the beginning of the survey period.

To conclude this Section, it is important to say that we are observing the extension of daily models into multi-day time horizons (Buliung 2005); in order move on in that direction it is necessary to use appropriate data but also to understand how activity and scheduling variables should enter our model specifications.

Model estimations results

In Table 4 we report the results of the model estimation. Estimated coefficients and relevant statistics are given for five model formulations: logit model, mixed logit model with error components, mixed logit model with error components and correlation across repeated observations from the same day, the same week and the same individual. The final specifications include fifteen socio-economic variables, thirty activity/scheduling related variables, the logsum of the mode choice model and four error components. Error components are specific to activity types, which induces correlations across activities of the same type performed at different time frames. Given the high number of alternatives specified the model is certainly identified (Walker 2001). We have tried to include a number of different socio-economic variables and their combinations, but they turned out to be not significant and were excluded from the final specification. Behavioral variables are estimated as specific to activity types (where it was possible). Activity durations including work duration, estimated as specific to tour scheduling, are found to be negative and significant.⁵ Negative value for shopping activity durations have been also found by Hirsh et al. (1986) in their dynamic model of weekly activity pattern. In terms of behavioral interpretation, it can be said that longer duration are more negatively perceived in the morning and commuting tours (which is natural, given the individual time

⁵ As pointed out by one of the reviewers, negative values for activity durations could be determined by endogeneity effect between duration and frequency.



5.32

0.5548 4.6440 0.2093 0.0029

9.6

0.5536 4.4205 0.2111 0.0025

0.003 1.540 0.234

7.22

7.87

0.5463 4.2962 0.2111

Shopping

Last time

Leisure

2.81

1.0072

12.84

38.72

2.76

60:1

5.44

0.0042

0.0020

Personal business Pick up/drop off

8.47

t-Stat 11.80 14.85 12.24 16.28 Mixed—individuals 14.11 12.23 99.6 90.6 5.09 7.30 7.45 3.48 2.15 6.25 4.42 7.82 3.01 Estimates -0.0198-0.0013-0.0208-0.0159-0.0068-0.0044-0.0200-0.0026-0.7019-0.3122-0.4830-0.0164-0.0081-0.2721-0.00200.0042 1.4461 0.4623 1.1124 .8323 t-Stat 18.58 31.97 23.59 21.45 22.58 29.99 20.82 10.90 2.19 5.64 2.92 9.21 5.41 7.73 8.68 4.82 5.86 Mixed—week Estimates -0.0044-0.0012-0.0200-0.0158-0.0064-0.0195-0.0080-0.0025 -0.0207-0.6679-0.3029-0.4592-0.0018-0.2647-0.01630.0040.5057 0.46992.0947 1.1091t-Stat 10.34 17.66 32.87 21.59 22.83 26.75 17.77 21.91 4.56 2.19 6.40 7.23 9.60 8.44 8.84 Mixed—day Estimates -0.016-0.020-0.004-0.008-0.020-0.003-0.021-0.016-0.006-0.277-0.682-0.313-0.474-0.002-0.0010.004 .605 0.509 1.142 2.207 t-Stat 10.65 12.38 12.10 13.64 11.65 10.10 12.35 11.80 2.95 0.89 Mixed—err comps 11.61 9.70 4.40 3.90 7.52 9.80 5.43 Estimates -0.0196-0.0204-0.0166-0.0040-0.0193-0.0044-0.0013-0.0077-0.0034-0.0155-0.0052-0.9463-0.7697-0.4001-0.47270.0078 2.6685 0.7568 1.7625 3.7895 19.19 25.00 24.46 18.43 15.03 15.10 t-Stat 13.19 10.45 12.03 35.46 10.67 2.14 29.9 4.88 9.00 8.62 5.93 3.03 3.97 5.94 Estimates -0.0162-0.0195-0.0044-0.0012-0.0200-0.0207-0.3014-0.0018-0.0080-0.0025-0.0159-0.0063-0.6555-0.4375-0.26210.0039 1.4918 1.0966 0.4874 2.0591 Logit PT out-leg NWday PT ret-leg NWday Personal business Personal business Pick up/drop off Pick up/drop off MT NWday PT NWday ET NWday Alternative MT Wday MC Wday ET Wday Shopping Shopping EC Wday Shopping Shopping W-H-W Leisure Leisure Available time before work Available time post work High week episode Activity duration Time budget Variable



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 Model results

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Variable	Alternative	Logit		Mixed—err comps	comps	Mixed—day	8	Mixed—week	sek	Mixed—individuals	viduals
		Estimates	t-Stat	Estimates	t-Stat	Estimates	t-Stat	Estimates	t-Stat	Estimates	t-Stat
Last time leisure	Work only	-0.2431	3.06	0.1558	1.09	-0.191	2.24	-0.2451	4.82	-0.2555	2.85
Last time PB	Work only	-0.1359	2.17	-0.1170	1.61	-0.136	2.19	-0.1357	2.96	-0.1300	1.75
Immobile days	Shopping	0.0063	1.35	0.0097	1.46	900.0	1.11	0.0065	1.09	9900.0	0.74
	Leisure	0.0063	1.38	0.0101	1.52	0.007	1.14	0.0065	1.10	0.0068	0.77
	Personal business	0.0067	1.46	0.0107	1.59	0.007	1.21	0.0069	1.18	0.0070	0.81
	Pick up/drop off	0900.0	1.30	0.0096	1.42	900.0	1.07	0.0063	1.06	0.0065	0.73
Logsum (Mode choice)	All	0.1157	1.95	0.1965	2.11	0.128	2.05	0.1473	2.35	0.2027	1.91
Age 36–50	Leisure	-0.4089	1.96	-0.5514	1.97	-0.415	2.78	-0.4567	3.17	-0.6149	2.68
Age 51–65	Pick up/drop off	0.4383	4.20	0.7369	6.87	0.443	4.76	0.4952	6.62	0.5898	7.31
Age 6–18	W-H-W	-0.3353	1.92	-0.0664	0.40	-0.305	1.78	-0.3231	3.32	-0.2489	1.40
Number of children under 12	Leisure	0.1502	2.46	0.3471	3.97	0.169	2.75	0.1532	2.63	0.2023	2.22
	Pick up/drop off	0.4722	5.87	0.8920	6.97	0.515	6.79	0.5244	7.71	0.7264	8.66
Number of working hours per week	Work only	0.0240	9.02	0.0318	8.52	0.025	9.21	0.0240	8.20	0.0234	5.04
Full time worker	Shopping	-0.6940	3.35	-1.4388	4.88	-0.795	3.81	-0.7093	2.90	-0.5629	1.07
	Leisure	-1.0277	4.88	-1.7898	6.34	-1.128	5.46	-1.0418	4.16	-0.8856	1.55
	Personal business	-0.5937	2.80	-1.1476	4.06	-0.670	3.17	-0.6093	2.47	-0.4726	0.88
	Pick up/drop off	-0.4466	1.97	-0.9747	3.25	-0.497	2.13	-0.4829	1.75	-0.3146	0.54
	Work only	-1.6649	7.99	-2.6170	8.85	-1.780	8.79	-1.6817	7.27	-1.6275	3.00
Female and part time	Shopping	0.5714	3.93	0.6456	3.83	0.577	4.11	0.5725	4.51	0.6851	11.60
	Leisure	0.4363	2.98	0.6079	4.03	0.441	3.21	0.4274	3.36	0.4625	4.55
	Personal business	0.6801	4.67	0.8423	5.17	0.677	4.82	0.6802	5.37	0.7736	8.91
	Pick up/drop off	0.5211	2.49	0.7080	4.24	0.588	3.13	0.5709	2.85	0.7632	6.38



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Variable	Alternative	Logit		Mixed—err comps	comps	Mixed—day	y	Mixed—week	sek	Mixed—individuals	viduals
		Estimates <i>t</i> -Stat	t-Stat	Estimates <i>t</i> -Stat	t-Stat	Estimates t-Stat	t-Stat	Estimates <i>t</i> -Stat	t-Stat	Estimates	t-Stat
Error components	Shopping	ı	I	2.4667	11.78	0.693	11.37	-0.0013	0.03	0.2413	3.13
	Leisure	ı	ı	0.6685	96.0	-0.077	0.74	0.2569	3.21	0.5161	5.62
	Personal business	ı	ı	2.2663	10.31	0.464	4.25	0.0068	0.23	-0.1685	1.75
	Pick up/drop off	ı	ı	2.8491	8.43	0.774	8.84	0.5549	6.50	0.5597	4.80
Number of observations		4,952		4,952		4,952		4,952		4,952	
No. of observations with repetitions		4,952		4,952		3,212		773		144	
Log likelihood at zero		-12268.27		-12268.27		-12268.27		-12268.27		-12268.27	
Log likelihood final		9194.9		9134.9		9179.7		9189.3		9149.02	
Rho squared adjusted		0.2468		0.2514		0.2477		0.2469		0.2502	

Wday working day, NWday non-working day, MT morning tour, MC morning commute, EC evening commute, ET evening tour, PT principal tour, PT out-leg principal tour return leg, W-H-W work-home-work



constraints). The same negative effect has been found for secondary activities during the principal tour and in the evening tour for non-working days.

The time budget variables are highly significant, highlighting that more mandatory activities, such as private business and pick up/drop off have a different status from the others. The same is true for period since the last activity performance, which captures in a rough way the rhythms reported in detail by Schönfelder (2006). The time available for shopping before work is found to be negative due to the low number of shopping activity episodes in the morning tours. This means that temporal constraints remain significant for workers. Time available for shopping after work is instead positively evaluated by our sample. The high week episode variable is found to be highly significant as already outlined in a previous dynamic model proposed by Hirsh et al. (1986). Moreover, the variable specific for pick up/drop off and shopping have the highest values given their fixed commitment nature. The number of days passed since an activity type was done last time is also affecting positively the probability of doing again the same activity. We also find that the same variable specified for leisure and pick up/drop off purposes affect negatively the probability of working all day. Immobile days increase the likelihood of out-of-home activities during the following days. The logsum of the mode choice model is positive and less than one, which is consistent with random utility maximization theory (Ben-Akiva and Lerman 1985); this variable is significant in all the model specifications shown in that paper. A number of socio-economics attributes are also part of the explanatory set of variables. Young people have a low probability to go home for lunch; individuals between 36 and 50 year old are less likely to go out for leisure, while adults over 50 are more likely to be engaged in picking up/dropping off other members of the family. Turning to individuals' working status, it is interesting to see that full time workers are not frequently engaged in activities other than work; the increasing number of working hours per week also affects the probability of work only alternative. In contrast, being female and working part-time highly affect the complexity of daily patterns.

Looking at the final log-likelihood values and adjusted rho-squared (on the basis of the degrees of freedom) we observe that the inclusion of error components definitely improves the fit of the model. However, when accounting for correlation across observations (from the same day, the same week and the same respondent) the improvements are less remarkable; the best model is the one obtained by including error components and without accounting for panel effects. This result differs from what was recently found by Cherchi (2007), who used the same panel data to estimate a mode choice model with different temporal correlation patterns. Their results show that for mode choice model the correlation due to panel effect was very significant and that the best model was the one accounting for correlation across the observations from the same individual. In Table 3 we also report the rho-squared adjusted obtained by calibrating a series of models of increasing richness. The basic model includes just socio-economic variables; the followings gradually include daily behavioral variables, weekly behavioral variables and long-term variables. The low fit obtained by including just socio-economic variables clearly shows the superiority of behavioral variables in the context of activity-type choice models and scheduling.

In one case (mixed logit model with error components specified on socio-economic and day variables) the convergence was not achieved after 1,500 iterations, which is the maximum number of iterations set for our runs. The main reason of that failure is the high value assumed by the error component related to the leisure activity.

The best model in terms of goodness of fit has been applied to test the ability of the model to forecast observed choices. Figures 2 and 3 compare observed and predicted choices for Monday and Friday, Figs. 4 and 5 report the same statistics for the weekend



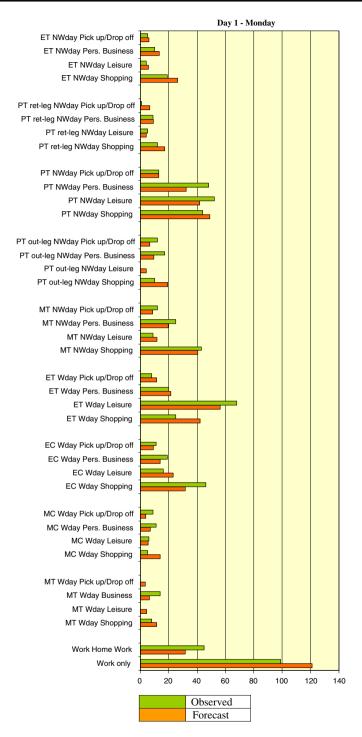


Fig. 2 Model application: Monday



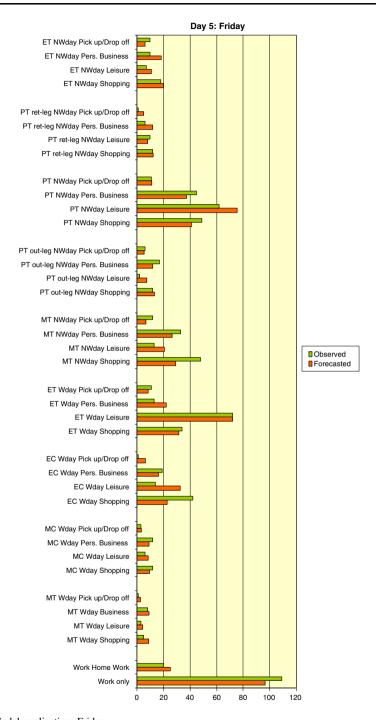


Fig. 3 Model application: Friday

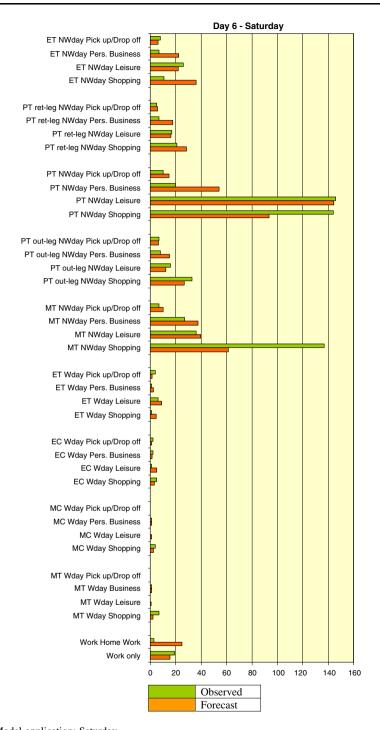


Fig. 4 Model application: Saturday



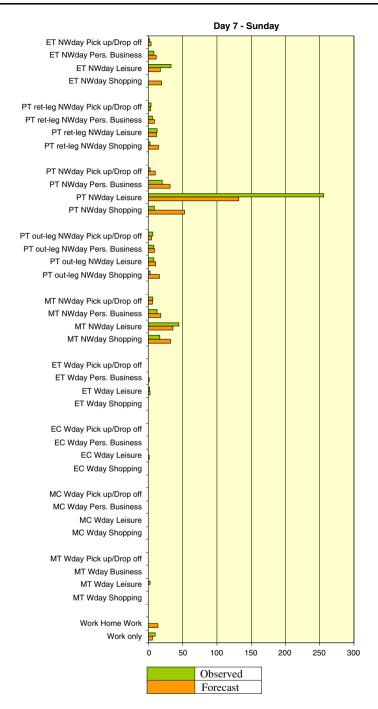


Fig. 5 Model application: Sunday



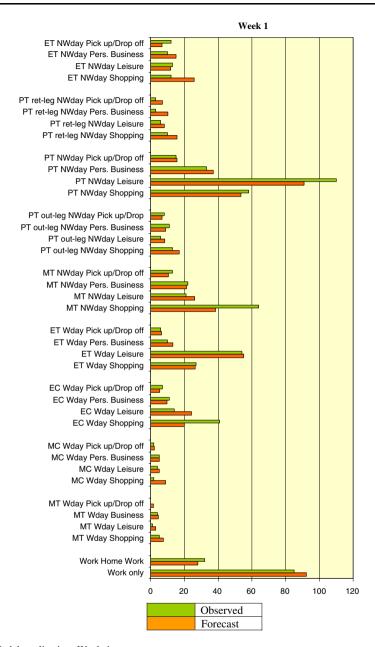


Fig. 6 Model application: Week 1

days, while in Figs. 6 and 7 the model application to the first and last week of the survey is given. This exercise aims at testing if one dynamic model can reproduce activity choices and scheduling over different days and weeks. Models estimated on panel data are usually estimated for each single day to capture day-to-day variability (Hirsh et al. 1986; Habib and Miller 2008). The model does a quite good job in reproducing aggregate choices for



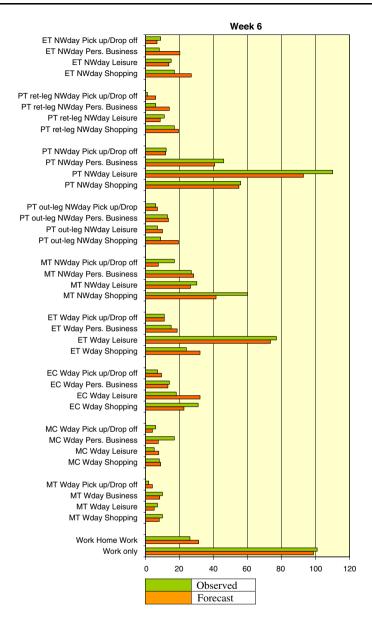


Fig. 7 Model application: Week 6

weekdays and for the different weeks of the survey. There are clearly some effects that are not captured. On Saturdays the number of shopping tours is underestimated during the morning and the principal tour, while on Sundays shopping tours are overestimated and leisure tours underestimated. More detailed information about the shopping opening hours on weekend days could perhaps help to set appropriate alternative availabilities and to reduce the gap between choices observed and predicted. One could also propose to estimate different models for the weekend days, but that will destroy the temporal continuity of our model structure. Overall, we can conclude that the dynamic model proposed is able



to reproduce pattern choices and that it also adapts well to the variability of rhythms of daily life.

Conclusions

The paper has implemented a model of activity type and timing choice, which extends the argument of Cirillo and Axhausen (2006), that one should and can improve the formulation of the systematic utility function by including variables that describe the dynamics of the day. Here again, the bulk of the explanatory power rests in those dynamic variables, which give a real insight into the constraints and usefulness of the activities. We model the sequential choice of the tour elements, i.e., the dynamic choices over the course of the day: the estimates confirm that previous choices impact through the restrictions on the degrees of freedom, particular vis-a-vis time. This is intuitively obvious, but shown here for the first time in a discrete choice framework. Previous work used independent hazard models to look at these sequential choices, or only at the rhythms of identical activities/tour elements. The socio-demographic variables, in other models dominant, loose their prominence. The personal preferences here are captured exclusively with the activity type specific error components.

There are various avenues for further research. Our model starts to count past history variable on day one of the survey; this causes the so called "initial conditions problem" of dynamic choice models. There are several ways to overcome this limitation; one could simply start to model from week two of the survey or adopt methods from the stochastic process theory. The latter approach has been proposed by Heckman (1981a, b) and Heckman and Singer (1986); however, those procedures are quite complex and their application in an activity based context could be problematic. It would be interesting to test if non-linear formulations of the variables, in particular of prior activity durations and budgets, would improve the fit. Equally, one could try to capture the known regularities of activity performance through the introduction of relevant interaction terms with the variables describing the interval to the previous activity performance. Within the framework of the current model, it would make sense to differentiate leisure further, as the large differences between leisure activities are well known, or to integrate effects of the weather. One could improve the estimates of the variables for the non-observed choices further, for example using models, which account for sample selectivity, one could estimate the duration, modes and travel times more appropriately. The same is true for shopping on Sundays, which is mostly grocery shopping. In the same vein, one could analyze the impact of the censoring affecting the variables describing the first observations of each respondent. It would be also interesting to compute elasticities of the different types of variables considered. Relevant policy simulations using simpler models and the full model can be conducted to assess the potential of dynamic models over static models.

The availability of the second half of the Mobidrive data collected in Halle and the more recent 2003 Thurgau survey (Axhausen et al. 2007), raise the question of how stable the parameters and model structure are. We hope to address this question of transferability, or better stability in the future.

The main application of the work undertaken here is the simulation of the dynamics of the activity program over time. We can now undertake to model weeks or even longer time periods, while accounting for the impacts of shocks to transport system.



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