

3.4 多维 r.v. 函数的分布

回忆一维 r.v. 函数的分布

多维例子: 已知平面上点 (X, Y) 分布, 求其到原点距离 $\sqrt{X^2 + Y^2}$ 的分布

3.4.1 多维离散情形.

二项分布可加性

例. 若 X 与 Y 独立, $X \sim B(m, p)$, $Y \sim B(n, p)$, 证 $Z = X + Y \sim B(m+n, p)$.

证. $\forall 0 \leq k \leq m+n$, 直观解释

$$\begin{aligned} P(Z=k) &= \sum_{l=0}^k P(X=l, Y=k-l) \\ &\stackrel{\text{独立}}{=} \sum_{l=0}^k P(X=l) P(Y=k-l) \\ &= \sum_{l=0}^k C_m^l p^l (1-p)^{m-l} C_n^{k-l} p^{k-l} (1-p)^{n-(k-l)} \\ &= \sum_{l=0}^k C_m^l C_n^{k-l} p^k (1-p)^{n+m-k} \\ &= C_{m+n}^k p^k (1-p)^{n+m-k} \quad \text{证毕} \quad \square \end{aligned}$$

例. X, Y 独立, $X \sim P(\lambda_1)$, $Y \sim P(\lambda_2)$ 证 $X+Y \sim P(\lambda_1 + \lambda_2)$.

泊松分布可加性

证: $\forall k \geq 0$,

$$\begin{aligned} P(X+Y=k) &= \sum_{l=0}^k P(X=l, Y=k-l) \\ &= \sum_{l=0}^k P(X=l) P(Y=k-l) \\ &= \sum_{l=0}^k \frac{\lambda_1^l}{l!} e^{-\lambda_1} \frac{\lambda_2^{k-l}}{(k-l)!} e^{-\lambda_2} \\ &= \frac{e^{-(\lambda_1 + \lambda_2)}}{k!} \sum_{l=0}^k C_k^l \lambda_1^l \lambda_2^{k-l} \\ &= \frac{(\lambda_1 + \lambda_2)^k}{k!} e^{-(\lambda_1 + \lambda_2)} \quad \text{证毕} \quad \square \end{aligned}$$

例. 设 X_1, X_2, X_3, X_4 独立同分布于 $B(1, p)$. 求

$$X = \begin{vmatrix} X_1 & X_2 \\ X_3 & X_4 \end{vmatrix}$$

的概率分布

解: $X = X_1 X_4 - X_2 X_3 \in \{-1, 0, 1\}$

$$\begin{aligned} P(X=1) &= P(X_1=1, X_4=1, X_2=0) + P(X_1=1, X_4=1, X_2=1, X_3=0) \\ &= p^2(1-p) + p^3(1-p) = p^2(1-p^2) \end{aligned}$$

同理, $P(X=-1) = p^2(1-p^2)$.

$$\therefore P(X=0) = 1 - 2p^2(1-p^2)$$

X 分布列为

X	-1	0	1
P	$p^2(1-p^2)$	$1-2p^2(1-p^2)$	$p^2(1-p^2)$

□

例. 设 X 与 Y 独立同分布, 且服从几何分布, 即

$$P(X=k) = p(1-p)^{k-1}, \quad k=1, 2, \dots$$

证明: 1) $Z = \min\{X, Y\}$ 与 $W = X - Y$ 独立.

2) Z 与 $U = \max\{X, Y\} - Z$ 独立.

证: 1) $\forall k \geq 0, \forall l, \quad (X \geq Y)$

$$P(Z=l, W=k) = P(Y=l, X=k+l)$$

$$\stackrel{\text{独立}}{=} P(X=k+l) P(Y=l) = p(1-p)^{k+l-1} p(1-p)^{l-1}$$

$$= p^2 (1-p)^k (1-p)^{2(l-1)}$$

同理, $\forall k \geq 0, \forall l$

$$P(Z=l, W=-k) = P(X=l, Y=l+k)$$

$$= p^2 (1-p)^{2(l-1)} (1-p)^k$$

$$\text{又} \because p(Z > l) = p(X > l, Y > l) = (1-p)^{2l}, \quad \forall l \geq 0$$

$$\therefore p(Z = l) = p(Z > l-1) - p(Z > l) = [1 - (1-p)^2] (1-p)^{2(l-1)}$$

$\forall k \geq 0$, 由对称性,

$$\begin{aligned} p(W = k) &= p(W = -k) = \sum_{l=1}^{\infty} p(Y = l, X = k+l) \\ &= \sum_{l=1}^{\infty} p^2 (1-p)^{l-1+k+l-1} = \frac{p^2 (1-p)^k}{1 - (1-p)^2} \end{aligned}$$

$$\therefore p(Z = l, W = k) = p(Z = l) p(W = k), \quad l \geq 1, k \in \mathbb{Z}. \quad \text{证毕.}$$

$$2) \quad U = \max\{X, Y\} - \min\{X, Y\} = |X - Y| = |W|.$$

$\therefore Z$ 与 W 独立

$\therefore Z$ 与 U 独立.

□

3.4.2. 多维连续情形

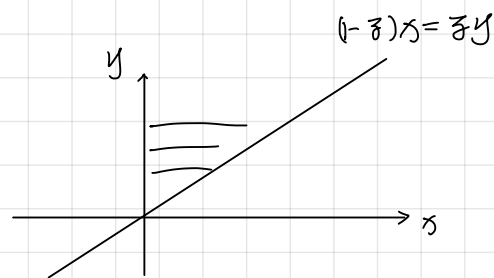
例. 若 X, Y 独立同分布于 $E(\lambda)$, $Z = \frac{X}{X+Y}$, 求 Z 的分布.

解. 显然, $0 \leq Z \leq 1$. $\forall 0 < z < 1$,

$$p(Z \leq z) = p\left(\frac{X}{X+Y} \leq z\right)$$

确定积分区域

$$= \int \int_{\substack{\frac{x}{x+y} \leq z, \\ x > 0, y > 0}} \lambda^2 e^{-\lambda x} e^{-\lambda y} dx dy$$



$$= \int_0^{+\infty} dy \int_0^{\frac{z}{1-z}y} dx \lambda^2 e^{-\lambda x} e^{-\lambda y}$$

$$= \int_0^{+\infty} dy \lambda e^{-\lambda y} (1 - e^{-\lambda \frac{z}{1-z}y})$$

$$= \int_0^{+\infty} dy \lambda e^{-\lambda y} - \lambda e^{-\lambda \frac{1}{1-z}y}$$

$$= 1 - (1-z) = z$$

$$\therefore Z \sim U(0, 1).$$

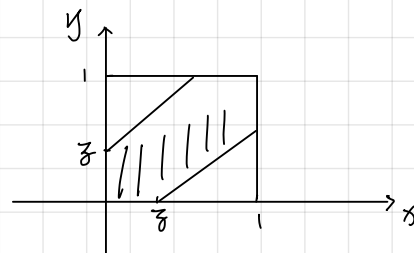
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例. 在 $(0,1)$ 内任取两点, 求两点间距离的分布函数与概率密度

解. $X, Y \sim U(0,1)$ 且独立. 令 $Z = |X - Y|$. 显然, $0 \leq Z \leq 1$.

$\forall z \in (0,1),$

$$P(Z \leq z) = \iint_{\substack{|x-y| \leq z \\ 0 < x < 1 \\ 0 < y < 1}} dx dy$$



$$= 1 - (1-z)^2 = 2z - z^2$$

$$\therefore F_Z(z) = \begin{cases} 0, & z < 0 \\ 2z - z^2, & 0 \leq z < 1 \\ 1, & z \geq 1 \end{cases}$$

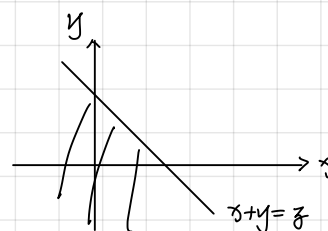
$$f_Z(z) = \begin{cases} 2 - 2z, & 0 \leq z < 1 \\ 0, & \text{其它.} \end{cases}$$

□

和的分布

设 (X, Y) 概率密度为 $f(x, y)$. $Z = X + Y$, 则

$$\begin{aligned} P(Z \leq z) &= \iint_{x+y \leq z} f(x, y) dx dy \\ &= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{z-x} f(x, y) dy \end{aligned}$$



$$\therefore f_Z(z) = \int_{-\infty}^{+\infty} f(x, z-x) dx = \int_{-\infty}^{+\infty} f(z-y, y) dy.$$

若 X 与 Y 独立, 则

$$f_Z(z) = \int_{-\infty}^{+\infty} f_X(x) f_Y(z-x) dx =: f_X * f_Y(z)$$

卷积

例. $X, Y \sim N(0, 1)$ 且独立, 证 $X+Y \sim N(0, 2)$.

证.

$$\begin{aligned} f_{X+Y}(z) &= \int_{-\infty}^{+\infty} f_X(x) f_Y(z-x) dx \\ &= \frac{1}{2\pi} \int_{-\infty}^{+\infty} e^{-\frac{x^2+(z-x)^2}{2}} dx \\ &= \frac{1}{2\pi} \int_{-\infty}^{+\infty} e^{-(x-\frac{z}{2})^2} e^{-\frac{z^2}{4}} dx \\ &= \frac{1}{2\pi} e^{-\frac{z^2}{4}} \quad \text{证毕.} \end{aligned}$$

一般地, 若 $X_i \sim N(\mu_i, \sigma_i^2)$, $1 \leq i \leq n$ 且独立, 则

$$\sum_{i=1}^n a_i X_i \sim N\left(\sum_{i=1}^n a_i \mu_i, \sum_{i=1}^n a_i^2 \sigma_i^2\right).$$

例. X, Y 独立, $X \sim U(0, 1)$, $f_Y(y) = \begin{cases} 2y, & 0 \leq y \leq 1 \\ 0, & \text{其它.} \end{cases}$

求 $Z = X+Y$ 的概率密度.

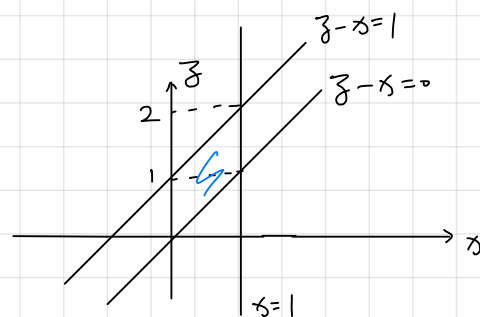
解.

$$\begin{aligned} f_Z(z) &= \int_{-\infty}^{+\infty} f_X(x) f_Y(z-x) dx \\ &= \int_0^1 f_Y(z-x) dx \end{aligned}$$

$$0 < z < 1, \quad f_Z(z) = \int_0^z 2y dy = z^2$$

$$1 < z < 2, \quad f_Z(z) = \int_{z-1}^1 2y dy = 1 - (z-1)^2 = 2z - z^2$$

$$f_Z(z) = 0 \quad \text{其它.}$$

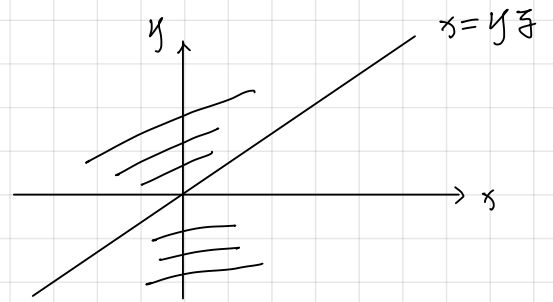


商的分布

设 (X, Y) 联合密度为 $f(x, y)$, $Z = \frac{X}{Y}$, 则

$$F_Z(z) = P\left(\frac{X}{Y} \leq z\right)$$

$$= \iint_{\frac{x}{y} \leq z} f(x, y) dx dy$$



$$= \iint_{y>0, x \leq yz} f(x, y) dx dy$$

$$+ \iint_{y<0, x \geq yz} f(x, y) dx dy$$

$$= \int_0^{+\infty} dy \int_{-\infty}^{yz} f(x, y) dx + \int_{-\infty}^0 dy \int_{yz}^{+\infty} f(x, y) dx$$

$$\Rightarrow f_Z(z) = F'_Z(z) = \int_0^{+\infty} y f(yz, y) dy + \int_{-\infty}^0 (-y) f(yz, y) dy$$

$$= \int_{-\infty}^{+\infty} |y| f(yz, y) dy.$$

例 设 X, Y 独立同分布于 $E(1)$. 求 $Z = \frac{X}{Y}$ 概率密度.

解: $f(x, y) = e^{-(x+y)}, x, y > 0$

则 $z > 0$ 时,

$$f_Z(z) = \int_{-\infty}^{+\infty} f(yz, y) |y| dy$$

$$= \int_0^{+\infty} e^{-(yz+y)} y dy = (1+z)^{-2}.$$

$f_Z(z) = 0$, 其它.

最大最小值分布

设 X_1, \dots, X_n i.i.d. (独立同分布), 概率密度为 $f(x)$, 分布函数

$$\text{为 } F(x). \quad Y = \max_{1 \leq i \leq n} X_i, \quad Z = \min_{1 \leq i \leq n} X_i.$$

$$\text{则 } F_Y(y) = P(Y \leq y) = P(X_i \leq y, \forall 1 \leq i \leq n)$$

$$= P(X_i \leq y)^n = F_X(y)^n$$

$$\therefore f_Y(y) = F_Y'(y) = n F_X(y)^{n-1} f_X(y).$$

$$F_Z(z) = P(Z \leq z) = 1 - P(\min_{1 \leq i \leq n} X_i > z)$$

$$= 1 - P(X_i > z)^n = 1 - (1 - F_X(z))^n$$

$$\therefore f_Z(z) = n(1 - F_X(z))^{n-1} f_X(z).$$

例. n 个独立工作的电子元件寿命 $\sim E(\lambda)$. 就下列工作方式求系统寿命: 1) 串联 2) 并联

解: 1) X_1, \dots, X_n i.i.d. $\sim E(\lambda)$.

$$Y = \min_{1 \leq i \leq n} X_i$$

$$P(Y > y) = P(X_1 > y)^n = e^{-n\lambda y}, \quad y > 0$$

$$\text{即 } Y \sim E(n\lambda).$$

$$2) \quad Z = \max_{1 \leq i \leq n} X_i$$

$$\forall z > 0, \quad P(Z \leq z) = P(X_1 \leq z)^n = (1 - e^{-\lambda z})^n$$

$$\therefore f_Z(z) = n\lambda(1 - e^{-\lambda z})^{n-1} e^{-\lambda z}, \quad z > 0.$$

一般情形

例. 若 $X \sim N(\mu, \sigma^2)$, Y 分布列为

Y	-1	1
P	$\frac{1}{3}$	$\frac{2}{3}$

且 X, Y 独立, 求 $Z = XY$ 的分布

解: $\forall z \in \mathbb{R}$,

$$P(Z \leq z) = P(Y=1, X \leq z) + P(Y=-1, -X \leq z)$$

$$= \frac{2}{3} P(X \leq z) + \frac{1}{3} P(X \geq -z)$$

$$= \frac{2}{3} \Phi\left(\frac{z-\mu}{\sigma}\right) + \frac{1}{3} \left(1 - \Phi\left(\frac{-z-\mu}{\sigma}\right)\right)$$

$$= \frac{2}{3} \Phi\left(\frac{z-\mu}{\sigma}\right) + \frac{1}{3} \Phi\left(\frac{z+\mu}{\sigma}\right)$$

$$\therefore f_Z(z) = \frac{2}{3} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(z-\mu)^2}{2\sigma^2}} + \frac{1}{3} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(z+\mu)^2}{2\sigma^2}}$$

□

例. 设 X_1, X_2, \dots, X_n i.i.d., $F(x), f(x)$.

$$Y = \max_{1 \leq i \leq n} X_i, \quad Z = \min_{1 \leq i \leq n} X_i$$

求 (Y, Z) 联合密度.

解. $\forall y > z$,

$$F(y, z) = P(Y \leq y, Z \leq z) = P(Y \leq y) - P(Y \leq y, Z > z)$$

$$= F(y)^n - P(z < X_i \leq y, \forall 1 \leq i \leq n)$$

$$= F(y)^n - [F(y) - F(z)]^n$$

$$\therefore f(y, z) = \frac{\partial^2 F}{\partial y \partial z}(y, z)$$

$$= n(n-1) [F(y) - F(z)]^{n-2} f(y) f(z).$$

□

例. 设 X 与 Y 独立, $X \sim U(0,1)$, Y 分布列为

Y	-1	0	1
P	$\frac{1}{2}$	$\frac{1}{3}$	$\frac{1}{6}$

求 $Z = X + Y$ 的分布

解. $-1 \leq Z \leq 2$. $\forall -1 < z < 2$,

$$F_Z(z) = P(X + Y \leq z)$$

$$= \frac{1}{2}P(X \leq z+1) + \frac{1}{3}P(X \leq z) + \frac{1}{6}P(X \leq z-1).$$

$-1 < z < 0$ 时,

$$F_Z(z) = \frac{1}{2}(z+1)$$

$0 < z < 1$ 时,

$$F_Z(z) = \frac{1}{2} + \frac{1}{3}z$$

$1 < z < 2$ 时,

$$F_Z(z) = \frac{5}{6} + \frac{1}{6}(z-1) = \frac{2}{3} + \frac{z}{6}$$

$$\therefore F_Z(z) = \begin{cases} 0, & z < -1 \\ \frac{z+1}{2}, & -1 \leq z < 0 \\ \frac{1}{2} + \frac{z}{3}, & 0 \leq z < 1 \\ \frac{2}{3} + \frac{z}{6}, & 1 \leq z < 2 \\ 1, & z \geq 2. \end{cases}$$

四

总结: 1) 联合分布函数、分布列、概率密度

边缘分布函数、分布列、概率密度

2) 独立性

3) 求多维 r.v. 的函数的分布