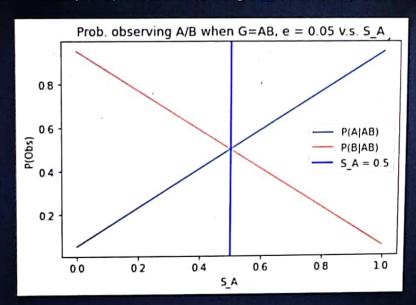
O SNP Cally & Sampley Brobablish P(E:=1)=P(E:=1 [G) =0, & PC SI=DIG=AB)= SA a) Ploi: A/G=AB): Plo: A/E: 0, S:=A, G=AB) + Ploi: A/E:=1, S:=B. (=AB) ·P(E:=0) ~ Same as P(E::0/6:M) · P(S:=B) = 1.de). Sa + 1.e. (1-SA) = SA(1-e) + (1-SA)e = Sa te - Ze Sa. b Ploi. B1 (n=AB)= Ploi= B1€=0, Si= B, G=AB). PC€=0). P(Si=B) +PCO;=BIE:=1, Si=A, G=AB) P(e:=1)-PCSi=A) = (1-e)(1-SA) + e-SA = 1-e +2eSA -SA o > abbrevialed as P(AS) 1) P(ABIO, -- ON) = P(O, -ON) (F-AB) P(G-AB) PCO: On (AB) P(AB) + P(O, and BB) PLBB HPCO, - ON (AB) PCAA) From a), b), Ulcelihoods-are: P(0:--On IAB) = TT RO:=AIAB) TJ P(0:=BIAB) Na: # of & observed, = (SA te-ZeSA) NA (1-e+ZeSH-SA) NB NB: # of B observed=N-IVA and RO, and AA)= TTP(O;=A/AA) = TT (O:=B/AA) = eneches P(U1... UMBB)= eMA (1-e) MB Thus P(AB10, -On) = [(SA+e-ZeSA) NA (1-e+ZeSA-SO)NB] P(AB) T(Sate-2026) Ma(1-e+2e-So-So) No]P(AB)+ e1/8/1-e) MAPLE MA(+e) P(B) [(Sate-ZeSa) NA (1-etZeSa-Sa) N-AA] P(AB) (Sate-Lesa) Macl-etzesa-Sa) N-Ma JP(AB) + eN-Na(1-e) Map(AA) + en (1-e) N-Nap(BB) (), e) see codes

```
def PobsA_AB(S_A,err):
      return S_A + err - 2*err*S_A
 def PobsB_AB(S_A,err):
      return 1 - err + 2*err*S_A - S_A #or 1 - PobsA_AB
 S = np.linspace(0,1,100)
  e = 0.05
  ya = [PobsA_AB(sa,e) for sa in S]
 yb = [PobsB_AB(sa,e) for sa in S]
  plt.plot(S,ya,label = "P(A|AB)")
  plt.plot(S,yb,label = "P(B|AB)")
  plt.axvline(x = 0.5, color = 'b', label = 'S_A = 0.5')
  plt.xlabel("S_A")
  plt.legend()
  plt.ylabel("P(Obs)")
  plt.title("Prob. observing A/B when G=AB, e = 0.05 v.s. S_A")
✓ 0.1s
```

Text(0.5, 1.0, 'Prob. observing A/B when G=AB, e = 0.05 v.s. S_A ')



- When sampling rates are the same for A and B strands (=0.5), the probabilities of observing A and B are the same given the true genotype is AB
- As the chance of sampling the A strand goes up, the probability of observing A in reads, given the true genotype is AB and the error rates are the same for A&B strands, goes up; while the probability of observing B in the reads goes down.

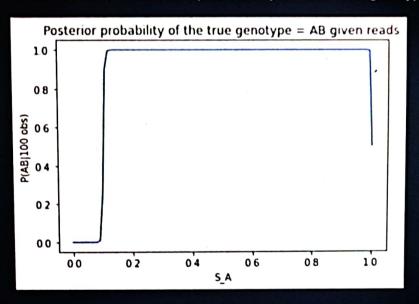
Pyth

0.107 17 0.114 (0. 5] 0 0

Given the assumptions and the reads, the prosterior probability of the true genotype G = AB is:

$$\mathsf{P}(\mathsf{AB}|O_1...O_{100}) = \frac{(S_A + 0.05 - 2*0.05*S_A)^{60} * (1 - 0.05 + 2*0.05*S_A - S_A)^{40}}{(S_A + 0.05 - 2*0.05*S_A)^{60} * (1 - 0.05 + 2*0.05*S_A - S_A)^{40} + 0.05^{40} * 0.95^{40} * 0.95^{40}}$$

Text(0.5, 1.0, 'Posterior probability of the true genotype = AB given reads')



(/>

- Given 60 reads are A, 40 reads are B, and the error rates = 0.05 for both A and B strands sampled, the posterior probability of the true genotype = AB rises shapping when P(S) > 0.1, then remains pretty much at 1 until the P(S) goes over 0.9.
- Intuitively, because the reads are pretty much half-and-half, and the error rates are very low, we could infer (which is what the posterior probability is doing) that the true genotype shouldn't be homozygous, unless the sampling rates are very biased. For example, if the chance of sampling B is very high if the true genotype is AB (S_A closes to 0), but we are still observing so many A in our reads tells us that the true genotype is more likely to be AA.

Python

(Psuedo)

0) 12=3

b) length = 4, K=3

Modeled:

AGGA, GGAT, GATA ATGA, TIGA, TGAT.

Unmedched: 44-6= 250

(if no skipping / Loub dracking are implemented when failed to match.

ATGA: ATG OTGA = = 42 AGGA; AGG O GGA

TT64: TT60 T64 = - +3 GGATA GGATA GATA

TGAT: TGANGAT: = STATA: GAT OATA = = = titots

- It changed Because tilts do not share a t-mer with the same sequence, while talets does (TGATA). . So the number of equivalen dasses would

reduce to 5 (including unmadded), with the It, to, til label and of the picture

The counts would also be different. All madebal's become: 1, while the unmadehed increases by 2.

d) k=4, length of read=4 AGGA GGAT GATA ATGA TGAT

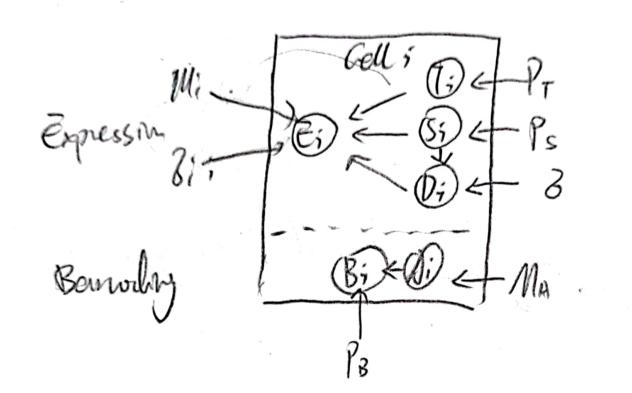
see the dicolumn

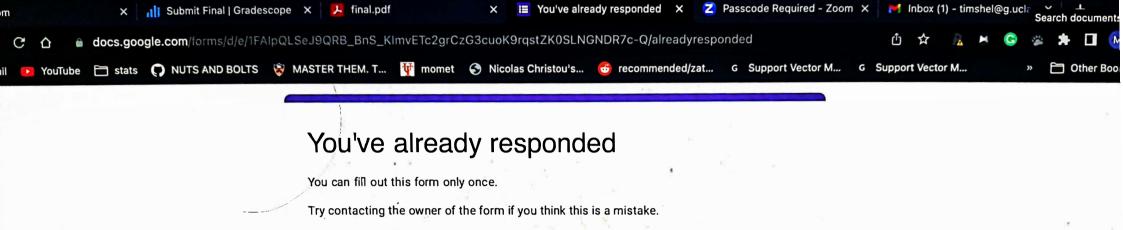
# counts					
Equialen classes	b/	0/3	d/		
- t.	2	1	2		
= tutitis.	1	0	1		
= tr	.1.	F	1	,	
= trits	1	1+	1		
- +3	1	- 11	1		
unmotched	250	1020	200	ho envi	
Todal possible	256	1024	256	رح	
	-		-	ļ	

C) Based on the Lable, a larger le for the some read length of 4 does not seem to change equivolence class counts/help disambiguate

@ Generative models			
61-5-62-M. cells	dades downtal by		
≥ 8, 8	death) => obs =0		
O Banodes of the cold	number of cell sample dasses	led. We devide for each	cell L be generaled:
& Cell states is in Car	degenical (Ps)	degunical (PT) probability, g	Peach cell type
1 Barcodes of the	S= G,	itely of cell being at a w	rdan cell state
banode #: Ain	Risson (Ma) Assum	e; cell state distribution is codes in a cell. L'barrodes get assigned for the formulal where hi is the total event modeled by Poisson, a probabilities of a specific barrobabilities of a specific barrobabilities of a specific barrobabilities.	3 independent of types.
Dell Death	Mutti (Ai, PB) what	Charcodes get assigned for binomial where Di is the top	Mores
Desume: By die end of the process, it roundarily die with a probability of a give no observations of gene expression	Mstarte cells of P 3. Pead cells	e tell modeled by Poisson, a specific bar	ind Pa is a rector
• 0	: ~ Bernoulli (3) Pro	bability of coll death of a	M-stage edt.
O (rene expression, nould be zero if the cell de other wise would fullow noun distributions depending on the cell states and the cell type expression b) P(E;) = P(EilPi, Si, 7 beaude [718:) - D(EilPi, Si, 7	15: a Bernoudling	(b), Si=M	dependent on cell states
(D) Gene expression,			
would be zero if the cell d	ies · EilDi ~ S	b, pi= 1	
other wise world fellow non	nal (NC 745 725 , Di= 0), 1.0
distribution depending on the	2		
expression and the cell type	3	offerted by	coll stude keed type.
b) P(Ei) = P(Ei)Pi, Si, 7	:)PCP: 15:2PCS	(;) P(T;)	
Pubi) = Pubilai)Pubi	Jan Jan All		
Ei: Expression, Di, cell death, s	i, rell state, Ti, rell t	type. Ai, total aveilable boun	des, Br, barrede Jo decell

c) Plate Mold





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