Introduction to R

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30/07/2020

Timetable and material

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Day	hour
Thursday	16:00-19:00 hours
Friday	10:00-13:00 hours
Saturday	11:00-14:00 hours

There will a 30-minute break during each session.

The material of this workshop is here

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Logit model:

- You should have defined a project (working directory) somewhere in your computer (e.g., Desktop).
- R is the actual programming language
- RStudio is an IDE (Integrated Development Environment) for R.
- R is case sensitive; e. g., Mean \neq mean
- R(Studio) may not work very well when files (or directory containing working files) have accented characters. If the locale language of your filesystem is not in English, then some errors may occur in those cases.

Basic calculations

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- R can be used as a calculator
- Mathematical constants:
 - pi = 3.142
 - $\exp(1) = 2.718$
- Logarithms:
 - log(e) = 1
 - $\log 10(100) = 2$
 - log(16, base = 4) = 2

Display Information

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- Getting help: in the console, put a question mark before the function name; RStudio will display the documentation:
- Set the display of decimal digits to 3, for a better output:

```
options(digits = 3)
```

Installing packages

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One of the strengths of R is the increasing number of available packages (more than 14,000 of them): CRAN Packages

To install and use a package we have

- firstly to install it;
- then, to load it in the current session.

To install a package (e.g., tseries), we can use the install.packages("tseries") function.

To load it in the current session, use library(tseries)

To install a package, we can also use the Tools option in the RStudio window and follow Install Packages...

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Data types

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The data types used by R are

- numeric (double precision): 2.718, 1.4, ...
- integer: 1, −13, . . .
- complex: 2 − 3*i*, . . .
- logical: TRUE, FALSE. Also NA is considered logical
- character: "one plus two", "Hello world!"

Data Structure

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- Vector: basic data structure in R. Its components have the same data type.
- Matrix: think of linear algebra.
 - A matrix as a collection of vectors.
- Dataframe: similar to a matrix, except that it is not necessarily homogeneous.
 - A collection of vectors (of possible different types) with the same length.
- List: generic data structure containing other objects (vectors, other lists), not necessarily of the same length.

Assignment operator

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- In R, we assign a value, value, to an object x by means of
 x <- value
- To type it, you can use the shortcut ALT+"-" (ALT key and minus sign); it works with Windows, MacOS.
- It is also possible to use =, but the equal sign has lower priority than <-.
 - Check the discussion at StackExchange.

Vectors

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The easiest way to create a vector in R is to use the c() (concatenate) function:

$$v \leftarrow c(1, 3, 5, 7, 9)$$

V

Vectors: coercion

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If we define a vector with components of different data types, the result will be coerced to the same data type

```
w \leftarrow c(1.56, "Hello World", 4, TRUE)
```

typeof(w)

```
## [1] "character"
```

Notice that

```
u <- c(1.56, 4, TRUE) typeof(u)
```

```
## [1] "double"
```

Vectors by sequences

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• Create a vector with the first 30 integer numbers:

$$(x < -1:30)$$

In reverse order:

$$(x < -30:1)$$

```
## [1] 30 29 28 27 26 25 24 23 22 21 20 19 18 17 16
## [26] 5 4 3 2 1
```

Vectors by sequences

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```
• Create a sequence of the first odd numbers, up to 11:
```

$$(y \leftarrow seq(1, 11, 2))$$

Repeat the character "Hello" 5 times:

```
## [1] "Hello" "Hello" "Hello" "Hello"
```

• Repeat the vector y, defined above, 2 times:

Length of a vector

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```
The length of a vector (number of its components) is obtained with the function length()
```

```
x <- 1:30 length(x)
```

[1] 30

```
y <- seq(1, 11, 2)
length(y)
```

[1] 6

```
z <- rep("Hello", 5)
length(z)</pre>
```

[1] 5

Subsetting

```
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```

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```
Let y \leftarrow seq(1,11,2): 1, 3, 5, 7, 9, 11.
```

- Select the third component: y[3] = 5
- Exclude the fourth component: y[-4]: 1, 3, 5, 9, 11.
- Select the first four components: y[c(1:4)]: 1, 3, 5, 7
- Select the first, fifth and last element:

```
y[c(1, 5, length(y))]
```

```
## [1] 1 9 11
```

or

```
y[c(1, length(y) - 1, length(y))]
```

[1] 1 9 11

Subsetting

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- Remove the first four components: y[-c(1:4)]: 9, 11.
- Select the second, fifth and sixth components: y[c(2, 5, 6)]: 3, 9, 11.
- Select components by using a vector of logic type:

```
s <- c(TRUE, TRUE, FALSE, FALSE, TRUE, FALSE)
y[c(s)]
```

```
## [1] 1 3 9
```

Vectorization

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Consider the vector

$$x < -1:100$$

Then,

- $x^2 = 1$, 4, 9, 16, 25, 36 ...
- 2*x+3 = 5, 7, 9, 11, 13, 15 ...
- sqrt(x) = 1, 1.414, 1.732, 2, 2.236, 2.449 ...

Vectorization

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We can add up two vectors even when they have different lengths, provided that the length of one is an integer multiple of the other:

```
x < - seq(1,11,2)
y < -1:12
x + y
```

[1] 2 5 8 11 14 17 8 11 14 17 20 23 ##

```
Check for identity
```

```
all(x + y == rep(x,2) + y)
## [1] TRUE
identical(x + y, rep(x,2) + y)
## [1] TRUE
```

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Each of the following operators returns either TRUE or FALSE.

- == (equality)
- != (not equal to)
- > (greater than)
- (less than)
- >= (greater than or equal to)
- (less than or equal to)
- !x (not x)
- x | y (x OR y)
- x & y (x AND y)

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Examples

Let $x \leftarrow c(1,2,3,4,5,6)$

- x > 4: FALSE, FALSE, FALSE, FALSE, TRUE, TRUE
- x == 4: FALSE, FALSE, FALSE, TRUE, FALSE, FALSE
- x != 3: TRUE, TRUE, FALSE, TRUE, TRUE, TRUE
- x == 4 | x != 3: TRUE, TRUE, FALSE, TRUE, TRUE, TRUE
- x == 4 & x != 3: FALSE, FALSE, FALSE, TRUE, FALSE, FALSE
- as.numeric(x == 4 | x != 3): 1, 1, 0, 1, 1, 1
- Arithmetically, TRUE is considered 1 and FALSE 0: sum(x == 4 | x != 3): 5

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Logical operators and subsetting

Subsetting will extract the values of the components that satisfy the given logical condition.

Let x < -c(1,2,3,4,5,6)

- $x[x \le 0]$: numeric(0) (empty set)
- $x[x \le 0 \mid x > 3]: 4, 5, 6$
- $x[x \le 0 \& x > 3]$: numeric(0) (empty set)
- $sum(x[x \le 0 \& x>3]): 0$

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Selection

$$y \leftarrow c(1, -2, 4, 6, 9, 2, 1)$$

- which(y <= 4): 1, 2, 3, 6, 7
 selects which entries satisfy the condition. Indexing of vectors starts from 1.
- y[which(y <= 4)]: 1, -2, 4, 2, 1 returns the values of the entries satisfying the condition.
- which(y == max(y)): 5 entry of the vector with the maximum value
- y[which(y == max(y))]: 9
 maximum value of the vector

Matrices

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A matrix A is specified by the number of its rows and columns, $m \times n$.

The **order** of rows and columns is **important**.

Matrix can be created by means of the matrix() function:

```
x <- 1:8
A <- matrix(x, nrow = 4, ncol = 2)
B <- matrix(x, nrow = 2, ncol = 4)
dim(A) # rows = 4; columns = 2</pre>
```

```
dim(B) # rows = 2; columns = 4
```

[1] 2 4

[1] 4 2

Matrices

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```
## [,1] [,2]
## [1,] 1 5
## [2,] 2 6
## [3,] 3 7
## [4,] 4 8
```

Matrices: Subsetting

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Timo corios

• Select the entry of the matrix A in the first row, second column:

A[1,2]

[1] 5

• Select the third column of the matrix B

B[,3]

[1] 5 6

Select the third row of the matrix A

A[3,]

[1] 3 7

Matrices: Subsetting

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Select the first and fouth column of the matrix B

```
## [,1] [,2]
## [1,] 1 7
## [2,] 2 8
```

Binding vectors: by columns

```
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```
u <- 1:4
v = rev(u)
w <- rep(1,4)
C <- cbind(u,v,w)
rownames(C) <- c("1st", "2nd", "3rd", "4th")
C</pre>
```

Binding vectors: by rows

```
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```
u <- 1:4
v = rev(u)
w <- rep(1,4)
D <- rbind(u,v,w)
colnames(D) <- c("1st", "2nd", "3rd", "4th")
D</pre>
```

```
## 1st 2nd 3rd 4th
## u 1 2 3 4
## v 4 3 2 1
## w 1 1 1 1
```

Matrix operations

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- The sum of two compatible matrices is A + B
- The subtraction of two compatible matrices is A − B
- The product of two compatible matrices is A *** B
- The transpose of a matrix A is t(A)
- The inverse of a square matrix, A, if it exists, is solve(A)

Inverse matrix

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```
x <- c(1,2,3,5, 4, 1, 2,2,1)
A <- matrix(x, nrow = 3, ncol = 3)
A1 <- solve(A)
A1 %*% A</pre>
```

The result is the identity matrix, up to round-off errors.

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 We can use the function all.equal() to compare the results of solve(A) and A1 %*% A with the identity matrix:

```
all.equal(A1 %*% A, diag(3))
```

```
## [1] TRUE
```

 The function diag() can also be used to extract the diagonal elements of a matrix

```
diag(A)
```

```
## [1] 1 4 1
```

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$diag((1:5)^{(.5)})$

```
[,1] [,2] [,3] [,4] [,5]
##
   [1,]
            1 0.00 0.00
                            0 0.00
   [2,]
            0 1.41 0.00
                            0 0.00
##
   [3.]
           0 0.00 1.73
                            0 0.00
   [4,]
           0 0.00 0.00
                            2 0.00
##
   [5.]
            0 0.00 0.00
                            0 2.24
##
```

```
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```

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```
diag(log(1):log(5))

## [,1] [,2]
## [1,] 0 0
## [2,] 0 1
diag(log(1:5))
```

```
[,2] [,3] [,4] [,5]
##
        [,1]
   [1,]
             0.000
                    0.0 0.00 0.00
##
##
   [2,]
           0 0.693
                    0.0 0.00 0.00
   [3,]
            0.000
                    1.1 0.00 0.00
   [4,]
           0.000
                    0.0 1.39 0.00
##
  [5,]
           0.000
                    0.0 0.00 1.61
```

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R has some built-in functions to deal with matrices. Consider ${\tt A}$

- rowSums(A) = 8, 8, 5
- colSums(A) = 6, 10, 5
- rowMeans(A) = 2.667, 2.667, 1.667

More matrices

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The operator **%*%**

The operator **%*%** works differently on vectors and matrices.

- On vectors it computes the dot product
- On matrices, the matrix multiplication (matrix multiplication is a form of ordered, vectorized dot product)

```
a \leftarrow c(1,2,3)
```

$$b \leftarrow c(4,5,6)$$

The dot product can also be implemented as

```
sum(a*b)
```

More matrices

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Vector can be coerced to work as matrices; in this case, columns matrices:

```
as.matrix(a)
       [,1]
##
## [1,]
## [2,]
## [3,]
          3
t(as.matrix(b)) %*% as.matrix(a) # dot product
       [,1]
##
## [1,]
         32
```

More matrices

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```
as.matrix(a) %*% t(as.matrix(b))
```

```
## [,1] [,2] [,3]
## [1,] 4 5 6
## [2,] 8 10 12
## [3,] 12 15 18
```

Dataframes

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Timo sorios

- Statistical analysis is done using datasets. A dataset contains a certain number of variables and observations.
- It is a good practice to have each variable set as a column vector and each observation as a row vector.
- A dataframe, in R, is the data structure of an observed dataset.
- A dataframe can be thought of as a matrix in which different columns may have different data types.

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When loading a dataset for analysis, there are some aspects to consider.

- Understand the data: what the dataset is about; what are its variables; how many obervation the dataset contains.
- Oetermine whether there are missing observations; some functions will not work properly otherwise.
- Visualize some of its variables.

Import the dataset with RStudio

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Information about the dataset

- The dataset is taken from the UCI, Machine Learning Repository website.
- The dataset can be found here.
 - Download the file, change its name to cleve.csv and move it to the project (working directory).
 - The names of the variables will be assigned later.
- The dataset will be called cleve hereinafter in the presentation.
- The dataset format is csv (comma separated variables).

Import the dataset with RStudio

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- Open an R script and save it in the project.
- Write in the script cleve <- read.csv("cleve.csv", header = FALSE) and hit CTRL + Enter (Windows); CMD + Enter (Mac OSX) to execute.
 - a. If you get an error in the console, you may be not in the project (working directory)
- Type head(cleve) in the script and execute the instruction. It will show in the console the first six observations.

Working with the dataset

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 Set the names of the variables. Copy the instructions and paste it in the script:

- The documentation of the dataset is found here
- The gender variable sex: 0 for females, 1 for males.

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 To work with only one variable from the dataset, e.g., age, we can extract it by means of \$.

age <- cleve\$age

 On the other hand, if we work with several variables in a dataset, it is better to use attach(namedataset) at the beginning; once done, use detach(namedataset).

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The dataset is a dataframe

```
class(cleve)
```

```
## [1] "data.frame"
```

with

```
dim(cleve)
```

```
## [1] 303 14
```

```
nrow(cleve) = 303 observations and ncol(cleve) = 14 variables.
```

Properties of the dataset

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The names of the variables can also be obtained by colnames(cleve).

The first six observations can be displayed by

head(cleve)

(Output omitted because it does not fit the slide.)

Slicing

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Display the first 3 observations of the first, third and seventh through ninth variables:

```
cleve[1:3,c(1,3,7:9)]
```

```
## age cp restcg thalac exang
## 1 63 1 2 150 0
## 2 67 4 2 108 1
## 3 67 4 2 129 1
```

or with (omitted for space)

Slicing

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Time series

Examples

• Compute the number of individuals whose age is greater than, or equal to, 50 years:

```
length(cleve$age[cleve$age >= 50])
```

```
## [1] 216
```

2 Compute the number of individuals without a diagnosis of heart disease:

```
length(cleve$sex[cleve$diagnostic == 0])
```

```
## [1] 164
```

Remark The variable diagnostic is strictly positive if there is indication of a heart condition.

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Examples

Ompute the number of individuals with a diagnosis of heart disease and with fasting blood sugar > 120 mg/dl

```
## [1] 22
```

Compute the number of individuals with a diagnosis of heart disease or with fasting blood sugar > 120 mg/dl

```
## [1] 162
```

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```
Consider, e.g., the variable age. Compute:
```

• the mean (average)

mean(cleve\$age)

[1] 54.4

4 the median

median(cleve\$age)

[1] 56

the interquartile range

IQR(cleve\$age) # Q3 - Q1

[1] 13

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• the summary of the principal statistics

```
summary(cleve$age)
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 29.0 48.0 56.0 54.4 61.0 77.0
```

Remark The function summary() does not return neither the variance nor the standard deviation.

Quantile distribution

```
quantile(cleve$age, c(.1, .25, .40, .60, .80))
```

```
## 10% 25% 40% 60% 80%
## 42 48 53 58 62
```

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It is possible to apply the summary() function to a full dataset, or to several variables of it

```
summary(cleve[,c(1,5,8)])
```

```
##
                       chol
                                    thalac
        age
                  Min .126
##
   Min.
          .29 0
                                Min.
                                       • 71
##
   1st Qu.:48.0
                  1st Qu.:211
                                1st Qu.:134
##
   Median:56.0
                  Median:241
                                Median: 153
##
   Mean :54.4
                  Mean : 247
                                Mean
                                       :150
##
   3rd Qu.:61.0
                  3rd Qu.:275
                                3rd Qu.:166
##
   Max. :77.0
                  Max
                         :564
                                Max
                                       :202
```

(Only three variables are selected for space.)

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Variance and standard deviation

- The variance (resp., standard deviation) is computed by var() (resp., sd()).
- If the variable has NA (missing values), then var(), sd() return NA:

```
var(cleve$thal)
## [1] NA
sd(cleve$thal)
```

```
## [1] NA
```

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Variance

We can apply the function var() to a single variable or several of them; in the latter case, we obtain the variance-covariance matrix of the selected variables:

```
var(cleve[, c(1, 3, 5, 6)])
## age cp
```

```
## age cp chol fbs

## age 81.697 0.9037 97.787 0.3816

## cp 0.904 0.9218 3.595 -0.0137

## chol 97.787 3.5951 2680.849 0.1815

## fbs 0.382 -0.0137 0.181 0.1269
```

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Standard deviation

The standard deviation sd(), on the other hand, can only be applied to a single variable:

```
sd(cleve$age)
```

```
## [1] 9.04
```

sd(cleve\$chol)

```
## [1] 51.8
```

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Standard deviation of a set of variables

We have to vectorize the function sd() by means of apply():

```
apply(cleve[,11:13], 2, sd)
```

- The value 2 in the second parameter of the function apply() computes the standard deviation, sd, of each variable (column).
- The value 1 would compute the standard deviation of each row (observation).

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Standard deviation of a set of variables

```
apply(cleve[,11:13], 2, sd)
```

slope ca thal

0.616 NA NA

There are missing values in the dataset.

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Count and remove NA's

```
Let's count them:
```

```
sum(is.na(cleve))
```

[1] 6

We can eliminate them (less than 2% of the observed values)

```
cleve <- na.omit(cleve)
sum(is.na(cleve))</pre>
```

[1] 0

Remark

R works in memory (RAM), so only the dataset **cleve** loaded in memory is changed. The file **cleve.csv** with the original observations is not changed (reproducibility in research).

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Standard deviation of a set of variables

```
apply(cleve[,11:13], 2, sd)
```

```
## slope ca thal
## 0.618 0.939 1.939
```

The variable slope has a slightly larger standard deviation now.

Contingency Tables

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Write a contingency table of gender and diagnosis of heart disease:

```
table(cleve$sex, cleve$diagnostic)
```

Using the with() function:

```
with(cleve, table(age, diagnostic))
```

(Same result; output omitted for space.)

Contingency Tables

```
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```

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```
Alternative to table
Using the xtabs() function:
with(cleve, xtabs(~ sex + diagnostic))
## diagnostic
## sex 0 1 2 3 4
## 0 71 9 7 7 2
## 1 89 45 28 28 11
```

Estimated Frequencies

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Redefine the variable diagnostic as dummy, with value 1 if some heart problem is observed:

```
cleve$diagnostic[cleve$diagnostic >0] <- 1
```

Table of estimated frequencies:

```
with(cleve, xtabs(~ sex + diagnostic)/nrow(cleve))
```

```
## diagnostic
## sex 0 1
## 0 0.2391 0.0842
## 1 0.2997 0.3771
```

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```
barplot(table(cleve$diagnostic),
    main = "Diagnostic (observed)")
```

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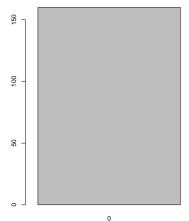
Statistica Inference

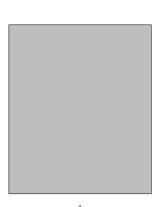
regression models

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Diagnostic (observed)





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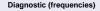
Barplot with observed frequencies

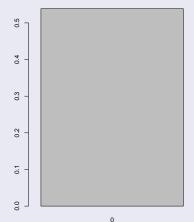
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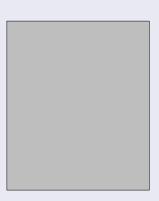
Lino AA

Preliminary analysis

Barplot with observed frequencies







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Barplot with a contingency table

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Grouped barplots

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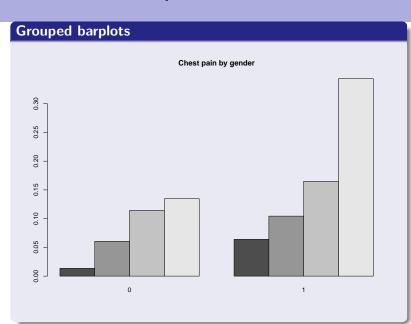
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Grouped barplots with colors

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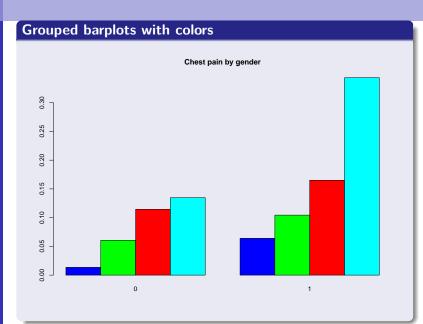
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Variable "age"

boxplot(cleve\$age)

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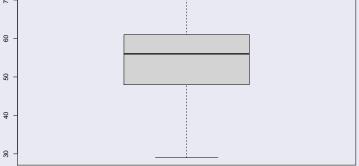
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Age as a function of gender

boxplot(age ~ sex, data = cleve)

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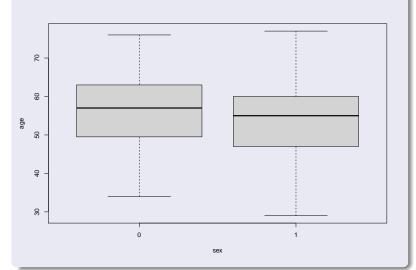
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Age as a function of gender, with positive diagnosis

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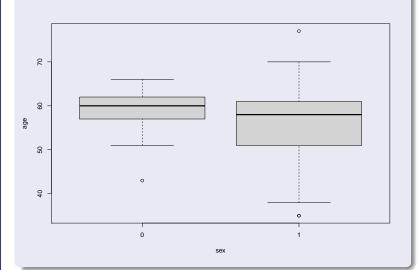
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Age as a function of gender, with positive diagnosis



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Timo carios

Age as a function of gender

Justification

See, e.g., John M. Chalmers, William S. Cleveland, Beat Kleiner, Paul A. Tukey, "Graphical Methods for Data Analysis", Wadsworth International Group, Duxbury Press, 1983, pp. 60-63.

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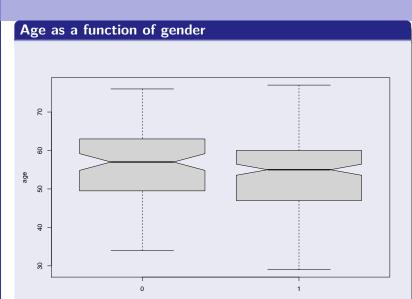
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sex

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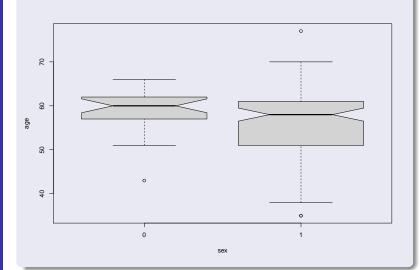
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Two graphs side by side

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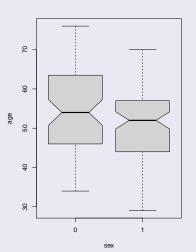
Statistica Inference

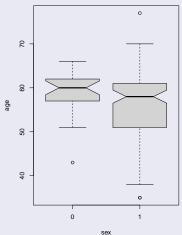
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Two graphs side by side





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The variable cp has four levels.

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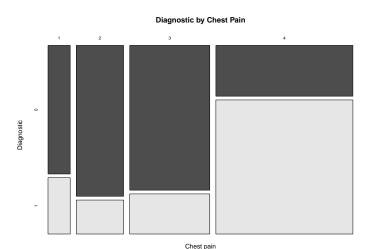
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The variable exang (exercise induced angina): 1 = yes; 0 = no

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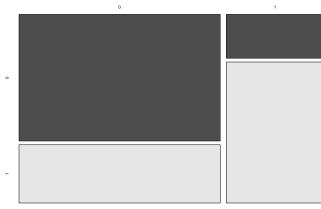
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Diagnostic by Angina



Exercise-induced angina

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Consider the variable age.

plot(cleve\$age)

gives the scatterplot of the variable.

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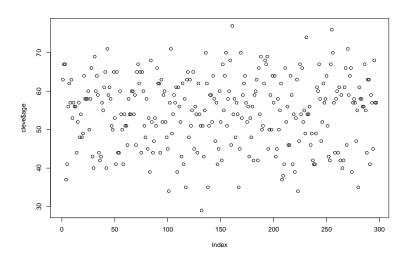
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```
plot(cleve$age, cleve$chol)
```

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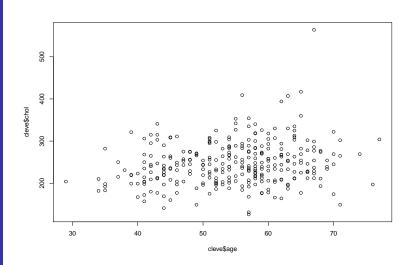
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With regression line:

```
plot(cleve$age, cleve$chol)
abline(lm(chol ~ age, data = cleve), col = "red")
```

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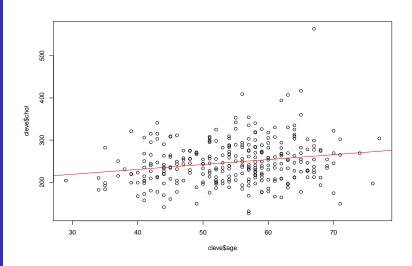
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Classification by K-means clustering

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See the script classification.R

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We want to determine whether there is a difference in age between individuals with no heart condition and those with an indication of a heart condition.

```
age.pos <- cleve$age[cleve$diagnostic == 1]
age.neg <- cleve$age[cleve$diagnostic == 0]</pre>
```

Formally

 H_0 : age.pos = age.neg

 H_1 : age.pos \neq age.neg

(significance level: $\alpha = .05$)

```
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```

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95% Confidence interval

```
t.test(age.pos, age.neg, mu = 0)
##
##
   Welch Two Sample t-test
##
## data: age.pos and age.neg
## t = 4, df = 295, p-value = 6e-05
## alternative hypothesis: true difference in means i
## 95 percent confidence interval:
## 2.12 6.11
## sample estimates:
## mean of x mean of y
       56.8
                 52.6
##
```

```
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```

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99% Confidence interval

```
t.test(age.pos, age.neg, mu = 0, conf.level = .99)
##
##
   Welch Two Sample t-test
##
## data: age.pos and age.neg
## t = 4, df = 295, p-value = 6e-05
## alternative hypothesis: true difference in means i
## 99 percent confidence interval:
## 1.49 6.74
## sample estimates:
## mean of x mean of y
       56.8 52.6
##
```

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- The function t.test() uses by default the Welch t.test, with Welch-Satterthwaite correction for degrees of freedom.
- When the variances are equal, we can use

```
t.test(age.pos, age.neg, mu = 0,
     var.equal = TRUE)
```

• The test for equal variances is

```
var.test(age.pos, age.neg, mu = 0, ratio = 1)
```

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Logit model:

Timo carios

Power calculation

Having rejected the null hypothesis, we compute the power of the test under the alternative

$$H_1$$
: $age.pos - age.neg = 2.5$

A power calculation needs four parameters, α , sd, n (sample size), δ (true difference in mean). It must also specify whether the test is one sided or two sided.

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Power calculation

```
power.t.test(n = nrow(cleve), sd =
  sqrt(var(age.pos)+var(age.neg)),
  sig.level = .05, delta = 2.5,
  alternative = "two.sided"
)
```

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Power calculation

```
##
##
        Two-sample t test power calculation
##
##
                  n = 297
##
             delta = 2.5
##
                 sd = 12.4
##
         sig.level = 0.05
##
             power = 0.689
       alternative = two.sided
##
##
## NOTE: n is number in *each* group
```

Hypothesis testing

```
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```

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Power calculation

Compute the sample size needed so that the power of the test is .90, when delta = 2.5; sd = 12.395, with $\alpha = .05$ in a two sided alternative:

```
##
##
        Two-sample t test power calculation
##
##
                  n = 518
##
             delta = 2.5
##
                 sd = 12.4
##
         sig.level = 0.05
             power = 0.9
##
##
       alternative = two.sided
##
## NOTE: n is number in *each* group
```

Power curve

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```
curve(power.t.test(n=50,delta = x,
    sd = sqrt(var(age.pos)+var(age.neg)),
        type="two.sample",
        alternative="two.sided")$power,
        from=.1, to=10, xlab="delta",
        ylab="power")
```

Power curve

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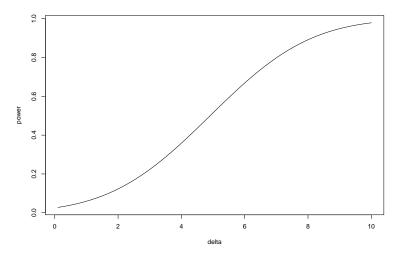
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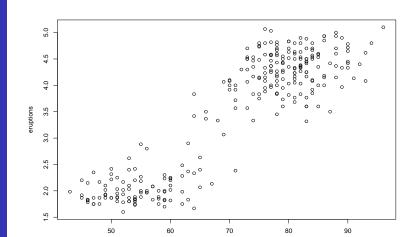
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attach(faithful)
plot(waiting, eruptions)



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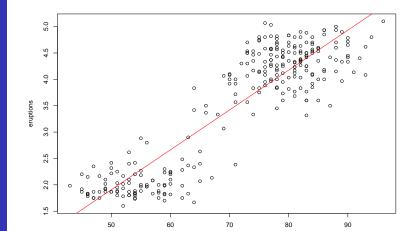
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```
plot(waiting, eruptions)
abline(lm(eruptions ~ waiting), col = "red")
```



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Let us consider the model

$$eruptions = \beta_0 + \beta_1 waiting + u$$

slrm <- lm(eruptions ~ waiting, data = faithful)</pre>

```
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```

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```
slrm.summary <- summary(slrm)
slrm.summary$coefficients</pre>
```

```
## Estimate Std. Error t value Pr(>|t|)
## (Intercept) -1.8740 0.16014 -11.7 7.36e-26
## waiting 0.0756 0.00222 34.1 8.13e-100
```

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Estimation

```
Estimate eruptions, when waiting = 90.
```

```
eruptions.fit.coef <- coefficients(slrm)
c <- c(1,90)
eruptions.fit.coef %*% c
## [,1]
## [1,] 4.93</pre>
```

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Confidence interval

```
Find a 99% confidence interval for eruptions, when waiting = 90.
```

```
## fit lwr upr
## 1 4.93 4.8 5.07
```

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Prediction interval

Find a 99% confidence interval for a prediction of *eruptions*, when waiting = 90.

```
## fit lwr upr
## 1 4.93 3.64 6.23
```

Remark

After using a dataset, detach it:

```
detach(faithful)
```

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Donwload and then load the library openintro

install.packages("openintro") # donwload the library
library(openintro)

load the library in the current session

Load the dataset hsb2:

data(hsb2)
attach(hsb2)

The description of the variables is obtained typing

?hsb2

or can be found here.

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Example 1

Estimate the standardized math score by race:

$$math = \beta_0 + \beta_1 race + u$$
.

The variable *race* is **categorical**.

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Example 1

Model estimation

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Example 1

.....

Model estimation: Extraction of the estimated coefficients

summary.m1\$coefficients

##	Estimate	Sta.	Error	t value	Pr	(>
## (Intercept)	54.40		1.32	41.082	2.	93
## factor(race)black	-1.34		1.87	-0.716	4.	75
## factor(race)hisp	-2.98		1.87	-1.591	1.	13
<pre>## factor(race)white</pre>	-2.70		1.87	-1.442	1.	51

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Example 1

Use of vectorization

```
summary.m1\$coefficients[2,]
```

Pr(>|t|) Estimate Std. Error t value ## ##

1.873 -0.716-1.3400.475

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Example 2

Factor interaction

Estimate the standardized math score by gender, race.

R code

```
m2 <- lm(math ~ female*factor(race))
summary.m2 <- summary(m2)</pre>
```

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Example 2

##

Estimated model

summary(m2)\$coefficients

female:factor(race)black

female:factor(race)hisp

female:factor(race)white

ππ		PPCTINGCE	blu.	ETIOI	U V C	1
##	(Intercept)	55.91		1.99	28.	0
##	female	-2.69		2.66	-1	0
##	factor(race)black	-3.21		2.78	-1	1
##	factor(race)hisp	-2.39		2.78	-0.	8
##	factor(race)white	-6.13		2.78	-2.	2

Estimate Std

3.37

-1.20

6.25

0.8

1.6

3.75

3.75

3.75

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Example 3

Estimate the effect of the scores of reading skills, *read*; social studies, *socst*, y science, *science* on the score of mathematics, *math*.

Model

$$math = \beta_0 + \beta_1 read + \beta_2 socst + \beta_3 science + u$$
.

R code

```
m3 <- lm(math ~ read + socst + science)
m3.estim <- summary(m3)</pre>
```

m3.estim\$coefficients

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Example 3

Estimated model

##		Estimate	Std. Error	t	value	Pr(> t)
##	(Intercept)	11.887	2.7937		4.25	3.24e-05
##	read	0.313	0.0654		4.79	3.24e-06
##	socst	0.154	0.0548		2.82	5.37e-03
##	science	0.315	0.0599		5.25	3.96e-07

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We randomly split the data and use one set to estimate the model and the rest to estimate the goodness of fit. To ensure reproducibility, we set the RNG seed to set.seed(13).

We shall validate firstly the second model, m2, with interaction of factors.

To validate other models, the only change needed is to change below the code where the actual model is specified,

linmodel.Train.

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Example 2

```
set.seed(13)
TrainRows <- sample(1:nrow(hsb2),
      .8*nrow(hsb2), replace = FALSE)
Train <- hsb2[TrainRows,]</pre>
Valid <- hsb2[-TrainRows. ]
linmodel.Train <-
  lm(math ~ female*factor(race), data = Train)
mathPredic <-
  predict(linmodel.Train, Valid)
actual.predic <-
  data.frame(cbind(actual = Valid$math,
                    predicted = mathPredic))
```

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Example 2

Correlation matrix

```
corr.accuracy <- cor(actual.predic)
corr.accuracy</pre>
```

```
## actual predicted
## actual 1.0000 0.0959
## predicted 0.0959 1.0000
```

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Accuracy: Min-Max

The Min-Max method compute the mean of the minimum over the mean of the maximum.

Values very close to 1 (Min-Max > .90) denotes excellent accuracy.

R code

[1] 0.854

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Accuracy: Mean Absolute Error (MAE)

If Ai are the actual (observed) values of the response and F_i are the forecast ones, then

$$MAE = \frac{1}{T} \sum_{i=1}^{I} |A_i - F_i|;$$

T is the sample size.

R code

mae

[1] 8.41

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Accuracy: Mean Absolute Percentage Error (MAPE)

The mean absolute percentage error (MAPE) is defined as

$$MAPE = \frac{1}{T} \sum_{i=1}^{I} \left| \frac{A_i - F_i}{A_i} \right|.$$

- MAPE can be interpreted as the average percentage error.
- Sometimes, MAPE can be very large, even though the forecast is reasonably good. If, e.g., $A_i \approx 10^{-3}$ and $|F_i A_i| \approx 10^{-1}$, entonces

$$\left|\frac{A_i - F_i}{A_i}\right| \approx 10^2$$

• If the forecast is exact, then MAPE = 0.

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MAPE: R code

```
mape <- mean(
   abs((actual.predic$predicted -
     actual.predic$actual))/actual.predic$actual)
mape</pre>
```

[1] 0.169

On the average, the error is 16.9%.

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RMSE

$$RMSE = \sqrt{\frac{1}{T} \sum_{i=1}^{T} (y_i - \widehat{y}_i)^2}$$

- RMSE is the square root of the variance of the residuals.
- It can be interpreted as the standard deviation of the unexplained variance of the model.
- It has the same units as the response variable.

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The library lmtest permits to run tests to determine

- normality of errors;
- heteroskedasticity;
- serial autocorrelation, and
- correct specification of the model (RESET)

The library sandwich permits to estimate HC, HAC¹ standard errors.

We shall run the diagnostics on the model m3 (Example 3).

¹HC: Heteroskedastic Consistent standard errors; HAC: Heteroskedastic and Autocorrelation Consistent errors.

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Normality

There are several normality test available.

- The Shapiro-Wilk test is available in the base library (loaded by default), but it cannot be applied to vectors with more than 5,000 observations. For this reason, we also present the Jarque-Bera test.
- The normality of errors is less of a concern when the sample size is sufficiently large.

Hyopthesis test

 H_0 : errors are normal

 H_1 : errors are not normal

$$(\alpha = .05)$$

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Normality: Shapiro-Wilk test

```
shapiro.test(m3$residuals)
##
## Shapiro-Wilk normality test
##
## data: m3$residuals
## W = 1, p-value = 0.5
```

Conclusion

There is no evidence of non-normality of errors.

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Jarque-Bera test

The Jarque-Bera test is available in the library tseries:

```
library(tseries)
jarque.bera.test(m3$residuals)

##

## Jarque Bera Test

##

## data: m3$residuals

## X-squared = 3, df = 2, p-value = 0.2
```

Conclusion

There is no evidence of non-normality of errors.

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QQ plot

The QQ plot compares the observed quantiles of the errors to the theoretical quantiles of the standard normal distribution.

```
qqnorm(m3$residuals)
qqline(m3$residuals, col = "red")
```

We can see that the distribution of the errors follows pretty closely the standard normal distribution.

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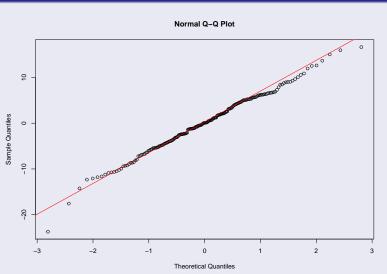
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Heteroskedasticity

The underlying hypothesis test is

 H_0 : the model is homoskedastic

 H_1 : the model is heteroskedastic

Function bptest()

Apply the function bptest() to the fitted model.

library(lmtest)
bptest(m3)

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Heteroskedasticity

```
bptest(m3)
##
## studentized Breusch-Pagan test
##
## data: m3
## BP = 9, df = 3, p-value = 0.03
```

Conclusion

- There is evidence of heteroskedasticity.
- If there is evidence of autocorrelation of errors, then neither OLS nor HC standard errors are correct.

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Serial autocorrelation

The purpose of the test is to determine whether there is any linear dependence among terms of the innovations.

We can apply the Durbin-Watson test, dwtest(), and the Breusch-Godfrey test, bgtest().

```
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Serial autocorrelation

```
##
## Durbin-Watson test
##
## data: m3
## DW = 2, p-value = 0.3
## alternative hypothesis: true autocorrelation is no
```

Result

There is no evidence that $corr(u_t, u_{t-1}) \neq 0$.

dwtest(m3, alternative = "two.sided")

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Serial autocorrelation

```
bgtest(m3, order = 10)
##
## Breusch-Godfrey test for serial correlation of or
##
## data: m3
## LM test = 12, df = 10, p-value = 0.3
```

Result

There is no evidence of linear dependence among the first 10 terms of the errors.

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Heteroskedasticity

We have detected heteroskedasticity, but not autocorrelation of the errors, so HC standard errors are appropriate.

HC standard errors are obtained using the function coeftest() passing the parameter vcov. = vcovHC^a

```
library(sandwich)
coeftest(m3, vcov. = vcovHC)
```

All controls are still significant.

^aThere are several variants of HC standard errors.

```
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```

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-*....

Heteroskedasticity

```
##
## t test of coefficients:
##
## Estimate Std. Error t value Pr(>|t|)
```

(Intercept) 11.8865 2.4585 4.83 2.7e-06 * ## read 0.3134 0.0600 5.22 4.5e-07 *

read 0.3134 0.0600 5.22 4.5e-07 *
socst 0.1542 0.0572 2.70 0.0076 *

science 0.3145 0.0567 5.54 9.5e-08 *

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05

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RESET: Hypothesis test

 H_0 : the model does not need nonlinear combination of controls

 H_1 : the model needs nonlinear combination of controls

- Typically, cuadratic and cubic powers are considered.
- The test statistics may need heteroskedastic (HC; HAC) corrections.

R code

```
resettest(m3, power = 2:3, vcov = vcovHC)
```

```
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RESET

```
##
## RESET test
##
## data: m3
```

RESET = 1, df1 = 2, df2 = 196, p-value = 0.3

Conclusion

There is no need of nonlinear combinations of regressors.

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Plots

We can also apply plot() to a lm() object to obtain these tests.

It is an interactive plot and it is convenient to do it in the console.

plot(m3)

Plots: documentation

It can be found here (check "Details").

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Plots: Leverage and influential observations

The concepts of leverage and influential observations can be used as a proxy for finding whether a given observation is an outlier.

- Leverage is a measure of how far the independent variables of an observation are from those of other observations.
- An influential observation is one whose deletion will affect greatly the estimate.

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Plots: Leverage and influential observations

- The leverage in a linear regression can be computed using the projection matrix (F. Hayashi, 2000, *Econometrics*, Princeton University Press, pp. 21-23).
- The Cook's distance defined in the next slide can be often used as a measure of observaions with high leverage;
 Cook's distance can also be used to detect highly influential observations.
- In R, the <u>residual-leverage</u> plot draws a red, dashed line identifying observations with Cook's distance greater than .5.
- In R, observations with leverage greater than 1 are omitted with a warning; check Residual-Leverage plot in "Details" here,

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Cook's distance

The Cook's distance is a tool that is used to check whether a single observation has a large influence on the estimate of the linear regression:

$$D_{i} = \frac{\sum_{j=1}^{T} \left(\widehat{y}_{j} - \widehat{y}_{j(i)}\right)^{2}}{(k+1)TSS}$$

where $\hat{y}_{j(i)}$ is the fitted response value when excluding i; TSS is the total sum of square (mean square error):

$$TSS = (1/T) \sum_{i=1}^{T} (y_i - \hat{y}_i)^2$$
.

detach(hsb2)

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Estimate the probability of a positive diagnosis of heart disease, controlling for cholesterol levels and age:

$$logit(diagnostic) = \beta_0 + \beta_1 chol + \beta_2 age,$$

where

$$logit(diagnostic) = log\left(\frac{diagnostic}{1 - diagnostic}\right)$$

is the log-odds of a positive diagnosis.

Thus, e.g., β_1 is the change of the log-odds of a positive diagnosis with respect to a unit increase of *chol*.

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Estimation

Call:

Signif. codes:

##

##

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Lino AA

```
Notarantonio
(lino@tec.mx)
          ## glm(formula = diagnostic ~ chol + age, family =
          ##
          ## Deviance Residuals:
Data Types
          ##
                Min
                          1Q
                              Median
                                          30
                                                  Max
          ## -1.593 -1.085 -0.815
                                       1.158
                                                1.729
          ##
          ## Coefficients:
                          Estimate Std. Error z value Pr(>|z|)
          ##
          ## (Intercept) -3.32713
                                      0.89145
                                                 -3.73 0.00019 *
                                                 0.62 0.53411
          ## chol
                          0.00147
                                      0.00237
                           0.05129
                                      0.01405
                                                  3.65 0.00026 *
          ## age
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```

0.001

'**' 0.01

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Goodness of fit

A measure of goodness of fit is a comparison of the estimated logit model to the null modelo

$$logit(diagnostic) = \beta_0$$

The comparison is done by the hypothesis test

 H_0 : The null model fits better the data

 H_1 : The logit model fits better the data

$$(\alpha = .05.)$$

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Goodness of fit

The test statistics is the difference of deviations, which has χ^2 distribution, with df. null - df.residual degrees of freedom:

[1] 0.000322

Conclusion

The logit model better fit the data.

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Diagnosis of heart disease (diagnostic = 1), if the estimated value of diagnostic in the logit model is greater than mean(diagnostic).

Validation

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Off-diagonal terms are false positives and false negatives.

Validation

```
## probs.predicted2
## 0 1
## 0 0.323 0.215
## 1 0.148 0.313
```

There is a 21.5% of false positives and a 14.8% of false negatives.

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##

0 0.374 0.165

1 0.226 0.236

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Diagnosis of heart disease (diagnostic = 1), if the estimated value of diagnostic in the logit model is greater than .5.

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Needed libraries are forecast, tseries.

Download them,

install.packages("forecast", "tseries")

and then load then in your session

library(forecast)
library(tseries)

Time series

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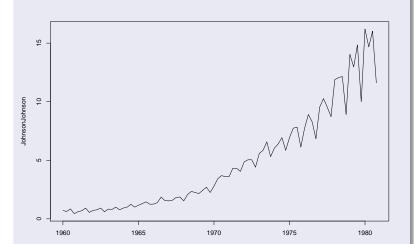
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Plots (Multiplicative model)

plot.ts(JohnsonJohnson)



Time series

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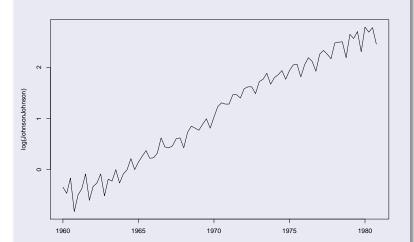
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Plots (Additive model)

plot.ts(log(JohnsonJohnson))



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Let y_t be a stochastic process.

The idea of exponential smoothing is to compute the *one-step* ahead forecast, $\hat{y}_{T+1|T}$, as a weighted mean of the previous observed terms:

$$\widehat{y}_{T+1|T} = \alpha y_T + \alpha (1-\alpha) y_{T-1} + \cdots$$
$$= \alpha y_T + (1-\alpha) \widehat{y}_{T|T-1}, \qquad 0 \le \alpha < 1;$$

rearranging terms,

$$\widehat{y}_{t+1|t} = \alpha y_t + (1 - \alpha)\widehat{y}_{t|t-1},$$

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which gives

$$\widehat{y}_{t+1|t} = \alpha y_t + (1 - \alpha)\widehat{y}_{t|t-1} = \widehat{y}_{t|t-1} + \alpha (y_t - \widehat{y}_{t|t-1})
= \widehat{y}_{t|t-1} + \alpha e_t,$$

where $e_t = y_t - \hat{y}_{t|t-1}$ is the forecast error.

The value α is estimated optimizing the errors squared.

The work by Holt & Winters allowed the inclusion of seasonal, s_t , and trending, b_t , terms, beside the level term, ℓ_t .

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Holt-Winters Modeling (Additive model)

$$y_{t+h|t} = \ell_t + hb_t + s_{t-m+h_m^+}$$

$$\ell_t = \alpha(y_t - s_{t-m}) + (1 - \alpha)(\ell_{t-1} - b_{t-1})$$

$$b_t = \beta^*(\ell_t - \ell_{t-1}) + (1 - \beta^*)b_{t-1}$$

$$s_t = \gamma(y_t - \ell_{t-1} - b_{t-1}) + (1 - \gamma)s_{t-m}$$

where m denotes the seasonality period (per year) The symbol

$$h_m^+ = \lfloor (h-1) \mod(m) \rfloor + 1$$

makes sure that the estimation of the seasonality is the last year of the sample.

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Estimation

```
logJJ.forecast <- HoltWinters(
  log(JohnsonJohnson), beta = TRUE, gamma = TRUE)
logJJ.forecast</pre>
```

```
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```

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Estimation

```
logJJ.forecast$coefficients[1]
```

a ## 2.61

logJJ.forecast\$SSE # measure of estimate error

[1] 0.661

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Plots

plot(logJJ.forecast)

The original series is plotted in black and the forecast is in red.

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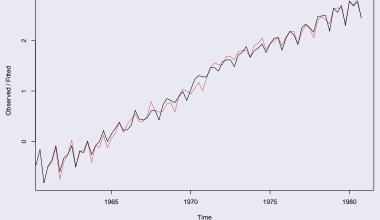
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Estimate with an initial value

```
logJJ.forecast2 <- HoltWinters(
  log(JohnsonJohnson), beta = TRUE, gamma =
   TRUE, 1.start = .91) # arbitrary initial value
logJJ.forecast2</pre>
```

```
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```

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Estimate with an initial value

```
logJJ.forecast2$coefficients[1]
##
     а
## 2.68
logJJ.forecast$coefficients[1]
##
    а
## 2.61
logJJ.forecast$SSE # measure of estimate error
## [1] 0.661
logJJ.forecast2$SSE
## [1] 6.01
```

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Comparing plots

```
plot(logJJ.forecast2)
lines(logJJ.forecast$fitted[,1], col = "green")
```

The object is a matrix whose columns are, respectively, the fitted time series; its level part; its trend part and the seasonality. We select above the fitted part.

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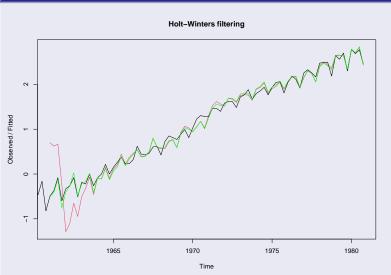
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Forecast

Forecast can be performed using the library forecast

library(forecast)

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Forecast

logJJ.forecast8095 <-</pre>

forecast.HoltWinters(logJJ.forecast, h=8)

- We use the first estimate, logJJ.forecast, as it is proven to be the better.
- The parameter h=8 will forecast the estimate eight periods in the future (two years).
- The forecast also plots a confidence interval (80%; 95% is the default).

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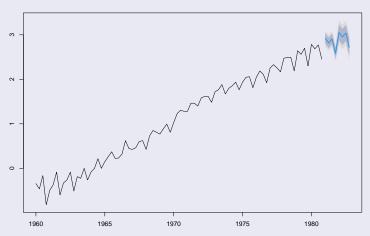
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Forecast: Plot





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Forecast: Plot

```
logJJ.forecast8599 <- forecast(
  logJJ.forecast, h=8, level = c(85,99))
plot(logJJ.forecast8599)</pre>
```

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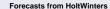
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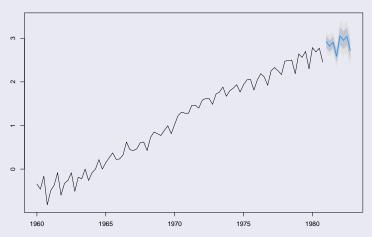
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Diagnostics

```
library(tseries)
ts8095.fitted <- as.ts(logJJ.forecast8095$fitted)
ts8095.fitted <- na.omit(ts8095.fitted)
ts8095.residuals <-
   as.ts(logJJ.forecast8095$residuals)
ts8095.residuals <-
   na.omit(as.ts(logJJ.forecast8095$residuals))</pre>
```

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Diagnostics: ACF

acf(ts8095.residuals)

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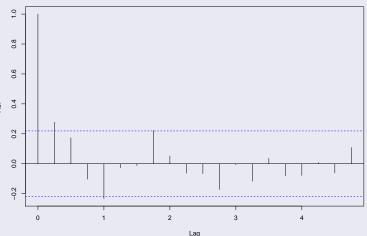
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Diagnostics: Ljung-Box Test

```
Box.test(ts8095.residuals, lag = 20)
##
## Box-Pierce test
##
## data: ts8095.residuals
## X-squared = 25, df = 20, p-value = 0.2
```

Conclusion

We do not reject the null hypothesis of no autocorrelation (for the first 20 lags).

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Diagnostics: Augmented Dickey-Fuller Test

```
adf.test(ts8095.fitted)
##
## Augmented Dickey-Fuller Test
##
## data: ts8095.fitted
## Dickey-Fuller = -1, Lag order = 4, p-value = 0.8
## alternative hypothesis: stationary
```

Conclusion

There is evidence that the errors are not white noise.

Box-Jenkins approach

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See the script BoxJenkins.R