MSiA 400 Lab Basic Statistics with R

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Data Set

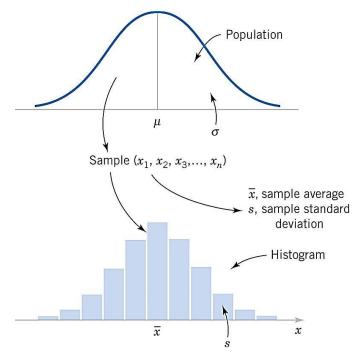
Average heights of men and women of 71 countries (collected from Wikipedia)

	A	В	С	D	
1		Male	Female		
2	Argentina	1.735	1.608		
3	Australia	1.784	1.645		
4	Austria	1.792	1.676		
5	Azerbaijan	1.718	1.654		
6	Bahrain	1.651	1.542		
7	Belgium	1.786	1.681		
8	Bolivia	1.6	1.422		
9	Brazil	1.707	1.588		
10	Bulgaria	1.752	1.632		
11	Cameroon	1.706	1.613		
12	Canada	1.76	1.633		
13	Chile	1.71	1.591		
14	China, People's Republic of	1.663	1.57		
15	Colombia	1.706	1.587		
16	Côte d'Ivoire	1.701	1.591		
17	Croatia	1.805	1.663		
18	Cuba	1.68	1.56		
19	Czech Republic	1.803	1.672		
20	Denmark	1.826	1.687		
21	Dinaric Alps	1.856	1.711		
วว	Fount	1 703	1 590		

- Load height.txt from the last lab
 - > height = read.table("z:\\ msia400lab1 \\height.txt", header=T)

Numerical Summary of Data

- Definitions
 - Data: the numeric observations of a phenomenon of interest
 - Population: the complete set of objects of interest
 - Sample: A subset of the population, or a portion used for analysis
- Description of population or sampled data
 - o center: measured by the mean
 - spread: measured by the variance



^{*} Source: Applied Statistics and Probability for Engineers, Montgomery and Runger

Mean

• Sample mean: n observations x_1, x_2, \dots, x_n in a random sample

$$\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum_{i=1}^{n} x_i}{n}$$

■ Population mean: N observations x_1, x_2, \dots, x_N in a population

$$\mu = \frac{x_1 + x_2 + \dots + x_N}{N} = \frac{\sum_{i=1}^{N} x_i}{N}$$

R: Calculate the sample mean of male and female

```
> M = height$Male;
> F = height$Female;
> n = length(F);
> mean(M);  # equivalent to mean(height[,1]) or mean(height[, "Male"])
> mean(F);
```

Variance

Sample variance

$$s^2 = \frac{\sum_{i=1}^{n} (x_i - \bar{x})^2}{n-1}$$

Population variance

$$\sigma^2 = \frac{\sum_{i=1}^N (x_i - \mu)^2}{N}$$

R: Calculate the variance of male and female

```
> var(M);  # sample variance
> var(F);  # sample variance
> (n-1)/n*var(M);  # population variance if we assume data is from population
> (n-1)/n*var(F);  # population variance if we assume data is from population
> sd(F);  # standard deviation
```

Covariance

- Covariance of two variables x and y
 - measures how the two are linearly related
- Sample covariance of x and y

$$s_{xy} = \frac{1}{n-1} \sum_{i=1}^{n} (x_i - \bar{x})(y_i - \bar{y})$$

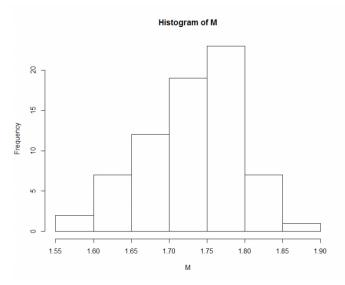
• Population covariance of x and y

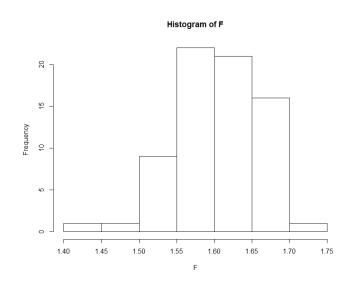
$$\sigma_{xy} = \frac{1}{N} \sum_{i=1}^{N} (x_i - \mu_x)(y_i - \mu_y)$$

- R: Calculate covariance of male and female heights
 - > cov(M,F);

Histogram

- Histogram (Frequency Distribution)
 - presents the counts of observations grouped within pre-specified classes or groups





R: Generate histograms

```
> hist(M);
> hist(F);
> hist(F, breaks=4); # Try
> hist(F, breaks=seq(1.40,1.75,by=0.03); # Try
```

Point estimation of population mean

■ Sample mean: n observations x_1, x_2, \dots, x_n in a random sample

$$\bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n} = \frac{\sum_{i=1}^{n} x_i}{n}$$

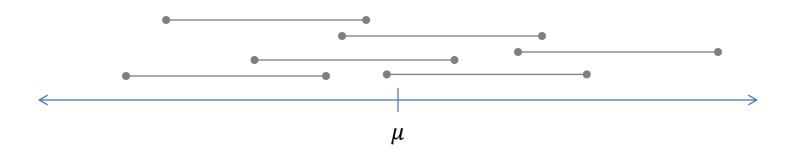
R: Find a point estimation of women's height

```
> mean(F);
```

Confidence Interval

• Confidence interval (CI) on μ with known population variance (σ^2) $(\bar{x}-z_{\alpha/2}\frac{\sigma}{\sqrt{n}}$, $\bar{x}+z_{\alpha/2}\frac{\sigma}{\sqrt{n}})$

- Interpretation of CI
 - Note: The confidence interval is a random interval!!
 - Yes: The observed interval includes the true value of μ with confidence $100(1-\alpha)$
 - \circ No: The true value of μ lies in the observed interval with confidence 100(1-lpha)



Confidence Interval of population mean

• Known population variance (σ^2)

$$\bar{x} \pm z_{\alpha/2} \frac{\sigma}{\sqrt{n}}$$

 \circ R: Confidence interval of men's height with known σ =0.0036 and $\alpha=0.05$

```
> sigma = 0.0036; Mbar = mean(M);
> E = qnorm(0.975)*sigma/sqrt(n); # margin of error
```

 $> CI_M = Mbar + c(-E,E);$

• Unknown population variance (s^2)

$$\bar{x} \pm t_{\alpha/2} \frac{s}{\sqrt{n}}$$

 \circ R: interval estimation of men's height with unknown σ and $\alpha=0.05$

```
> S = sd(M);
> E = qt(0.975, df=n-1)*S/sqrt(n); # margin of error
> CI_M = Mbar + c(-E,E);
```

Hypothesis testing

Statistical hypotheses

Two-tailed alternative hypothesis

$$H_0$$
: $\mu_m = 1.728$ H_1 : $\mu_m \neq 1.728$

Lower-tailed alternative hypothesis

$$H_0$$
: $\mu_m = 1.728$
 H_1 : $\mu_m < 1.728$

- Hypothesis testing
 - Obtains information in a random sample from the population
 - If the information is consistent with the hypothesis, we conclude the hypothesis is true, otherwise, false.

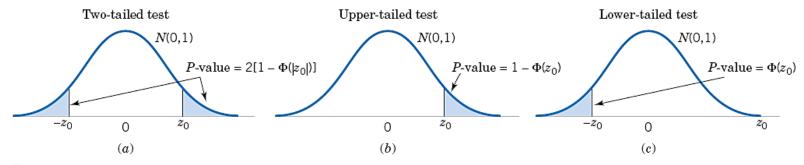


Figure 9-7 The *P*-value for a *z*-test. (a) The two-sided alternative $H_1: \mu \neq \mu_0$. (b) The one-sided alternative $H_1: \mu > \mu_0$. (c) The one-sided alternative $H_1: \mu < \mu_0$.

* Source: Applied Statistics and Probability for Engineers, Montgomery and Runger

Two tailed test with known variance

Two-tailed test with known variance

$$H_0$$
: $\mu_m = 1.729$
 H_1 : $\mu_m \neq 1.729$

Test statistics

$$z_0 = \frac{\bar{x} - \mu_0}{\frac{\sigma}{\sqrt{n}}}$$

* Fail to reject H_0 if $-z_{\alpha/2} \le z_0 \le z_{\alpha/2}$

• R: Test the hypothesis with $\alpha = 0.05$

- > z = (Mbar 1.729)/(sigma/sqrt(n));
- > z.half.alpha = qnorm(1-0.05/2);
- > c(-z.half.alpha, z.half.alpha); # check if z is inside the interval

Lower-tailed test with known variance

Two-tailed test with known variance

$$H_0$$
: $\mu_m = 1.729$
 H_1 : $\mu_m < 1.729$

Test statistics

$$z_0 = \frac{\bar{x} - \mu_0}{\frac{\sigma}{\sqrt{n}}}$$

* Fail to reject H_0 if $z_0 \leq -z_\alpha$

• R: Test the hypothesis with $\alpha = 0.05$

```
> z = (Mbar - 1.729)/(sigma/sqrt(n));
```

> z.alpha = -qnorm(1-0.05);

Two tailed test with unknown variance

Two-tailed test with unknown variance

$$H_0$$
: $\mu_m = 1.729$
 H_1 : $\mu_m \neq 1.729$

Test statistics

$$t_0 = \frac{\bar{x} - \mu_0}{\frac{S}{\sqrt{n}}}$$

* Fail to reject H_0 if $-t_{\alpha/2} \le t_0 \le t_{\alpha/2}$

• R: Test the hypothesis with $\alpha = 0.05$

```
> t = (Mbar – 1.729)/(sd(M)/sqrt(n));
> t.half.alpha = qt(1-0.05/2, df=n-1);
> c(-t.half.alpha, t.half.alpha); # check if t is inside the interval
```

Example: Tensile Strength

Tensile Strength

A manufacturer of paper used for making grocery bags is interested in improving the tensile strength of the product. Product engineering thinks that tensile strength is a function of the hardwood concentration in the pulp and that the range of hardwood concentrations of practical interest is between 5 and 20%. A team of engineers decided to investigate four levels of hardwood concentration: 5%, 10%, 15%, and 20%. They decided to make up six test specimens at each concentration level, using a pilot plant. All 24 specimens are tested on a laboratory tensile tester, in random order.

Tensile Strength of Paper (psi)

Hardwood	Observations						Totals	A
concentration (%)	1	2	3	4	5	6	— Totals	Averages
5	7	8	15	11	9	10	60	10.00
10	12	17	13	18	19	15	94	15.67
15	14	18	19	17	16	18	102	17.00
20	19	25	22	23	18	20	127	21.17

^{*} Source: Applied Statistics and Probability for Engineers, Montgomery and Runger

ANOVA (1)

- Goal
 Check if the tensile strength for 4 different concentrations are equal
- Hypothesis

$$H_0$$
: $\mu_1 = \mu_2 = \mu_3 = \mu_4 = 0$
 H_1 : $\mu_i \neq 0$ for at least one i

ANOVA table

Source of Variation	Sum of Squares	Degree of Freedom	Mean Square	F_0
Treatments	$SS_{Treatments}$	k-1	$MS_{Treatments}$	$\frac{MS_{Treatments}}{MS_E}$
Error	SS_E	k(n-1)	MS_E	
Total	SS_T	kn-1		

ANOVA (2)

- Import data
 - > tensile = read.table("z:\\ msia400lab1 \\tensile.txt", header=T)
- Constructing ANOVA table

Step 1. Turn the data into a single vector

```
> resp = c(t(as.matrix(Tensile)));
```

Step 2. Setup values

- > treats = c("HC5","HC10","HC15","HC20");
- > k = 4;

> n = 6;

- Step 3. Create a vector of treatment factors that corresponds to each element of resp
 > tm = gl(k,1,n*k,factor(treats));
- Step 4. Apply the function aov and print summary
 - > myANOVA = aov(resp ~ tm);
 - > summary(myANOVA);

Note

gl(): generate levels (# levels, # replications, length, labels)

Note

as.matrix() : convert table to matrix
t() : transpose

ANOVA (3)

ANOVA table

Source of Variation	Sum of Squares	Degree of Freedom	Mean Square	F_0	P-value
Treatments	382.8	3	127.60	19.61	3.59 E-6
Error	130.2	20	6.51		
Total	513.0	23			

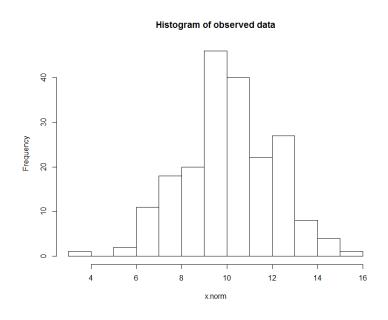
- Two approaches
 - o $F_{0.01,3,20} = 4.94 < 19.61 = F_0$, reject H_0
 - P-value =3.59 E-6 < $\alpha = 0.01$, reject H_0

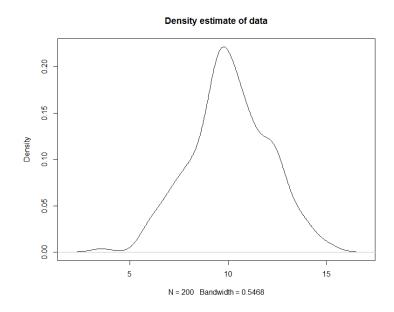
* Note > qf(1-0.01,df1=3,df2=20)

Conclusion: Hardwood concentration affects the strength of the paper

Distribution Fitting 1

- Samples from Normal distribution
 - > x.norm = rnorm(n=200,m=10,sd=2); # Generate 200 random samples from N(10,2)
 - hist(x.norm,breaks = 10,main="Histogram of observed data");
 - > plot(density(x.norm), main="Density estimate of data");

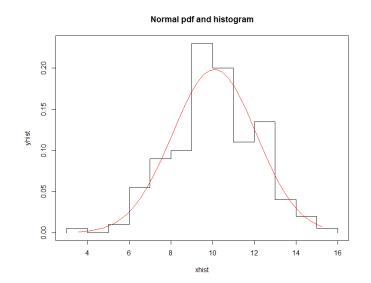


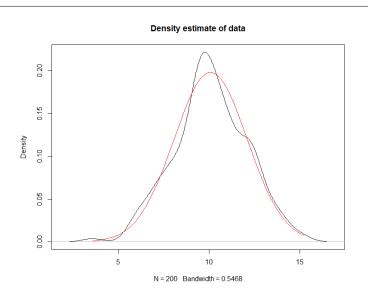


Distribution Fitting 2

Comparisons

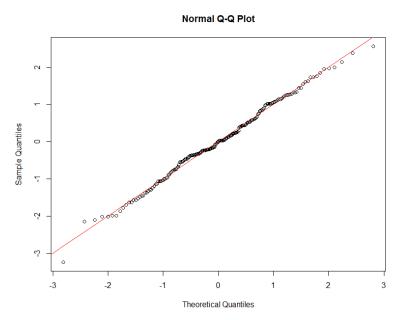
- > h = hist(x.norm,breaks=10);
- > x.hist = c(min(h\$breaks),h\$breaks); y.hist = c(0,h\$density,0);
- > x.fit = seq(min(x.norm),max(x.norm),length=40);
- > y.fit = dnorm(x.fit,mean=mean(x.norm),sd=sd(x.norm));
- > plot(x.hist,y.hist,type="s",ylim=c(0,max(y.hist,y.fit)), main="Normal pdf and histogram")
- > lines(x.fit,y.fit, col="red");
- > plot(density(x.norm),main="Density estimate of data")
- > lines(x.fit,y.fit, col="red")





Distribution Fitting 3

QQ plot
 A graphical method for comparing two probability distributions by plotting their quantiles against each other



R: QQ plot generation

```
> z.norm = (x.norm-mean(x.norm))/sd(x.norm);
```

- > qqnorm(z.norm);
- > abline(0,1, col ="red");