Exercise 14: This problem focuses on the collinearity problem.

Import notebook funcs

```
from notebookfuncs import *
```

Import user funcs

```
from userfuncs import *
```

Import libraries

```
from sympy import symbols, poly
import numpy as np
import seaborn as sns
import pandas as pd
import statsmodels.formula.api as smf
```

(a) Perform the following commands in Python:

```
rng = np.random.default_rng (10)
x1 = rng.uniform (0, 1, size =100)
x2 = 0.5 * x1 + rng.normal(size =100) / 10
y = 2 + 2 * x1 + 0.3 * x2 + rng.normal(size =100)
```

The last line corresponds to creating a linear model in which y is a function of x_1 and x_2 . Write out the form of the linear model. What are the regression coefficients?

```
x1, x2, y = symbols("x_1 x_2 y")
beta_0, beta_1, beta_2 = symbols(r"\beta_0 \beta_1 \beta_2")
equation = beta_0 + beta_1 * x1 + beta_2 * x2
display(equation)
equation = equation.subs([(beta_0,2), (beta_1,2),(beta_2, 0.3)])
equation = poly(equation)
```

```
\begin{split} \beta_0 + \beta_1 x_1 + \beta_2 x_2 \\ \text{Poly} \left( 2.0 x_1 + 0.3 x_2 + 2.0, x_1, x_2, domain = \mathbb{R} \right) \end{split}
```

```
rng = np.random.default_rng (10)
x1 = rng.uniform (0, 1, size =100)
x2 = 0.5 * x1 + rng.normal(size =100) / 10
y = 2 + 2 * x1 + 0.3 * x2 + rng.normal(size =100);
```

(b) What is the correlation between x_1 and x_2 ? Create a scatterplot displaying the relationship between the variables.

Correlation between x1 and x2

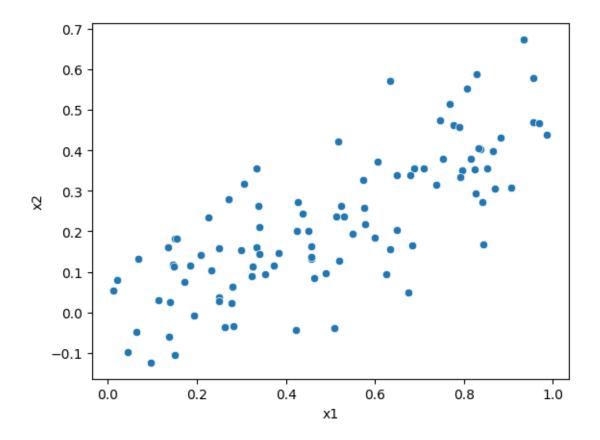
```
np.corrcoef(x1,x2)[0][1]
```

0.772324497691354

Display scatterplot of x1 against x2

```
def construct_df(x1, x2,y):
    df = pd.DataFrame({"x1": x1,"x2": x2, "y": y})
    return df

df = construct_df(x1,x2,y)
sns.scatterplot(df, x="x1", y="x2");
```



(c) Using this data, fit a least squares regression to predict y using x_1 and x_2 . Describe the results obtained. What are β_0 , β_1 , and β_2 ? How do these relate to the true β_0 , β_1 , and β_2 ? Can you reject the null hypothesis $H_0:\beta_1=0$? How about the null hypothesis $H_0:\beta_2=0$?

Fit a least squares regression

```
# Fit combined regression
def fit_combined(df):
    formula = "y ~ x1 + x2"
    model = smf.ols(f"{formula}", df)
    results = model.fit()
    print(results.summary())
    return results

results = fit_combined(df);
```

OLS Regression Results

=========	======		=====	=====			
Dep. Variable	:		У	R-sqı	ared:		0.291
Model:			OLS	Adj.	R-squared:		0.276
Method:		Least Squa	ares	F-sta	atistic:		19.89
Date:		Wed, 16 Oct 2	2024	Prob	(F-statistic)	:	5.76e-08
Time:		18:03	3:33	Log-I	Likelihood:		-130.62
No. Observation	ons:		100	AIC:			267.2
Df Residuals:			97	BIC:			275.1
Df Model:			2				
Covariance Ty	pe:	nonrol	oust				
=========	======	=========		======			=======
	coef				P> t	[0.025	0.975]
Intercept	1.9579				0.000	1.581	2.334
x1	1.6154	0.527	;	3.065	0.003	0.569	2.661
x2	0.9428	0.831		1.134	0.259	-0.707	2.592
	======		===== .051	Durb	n-Watson:	.=======	1.964
Prob(Omnibus)	:		.975		ne-Bera (JB):		0.041
Skew:	•		.036	-			0.979
Kurtosis:			. 931	Cond			11.9
		_			-		

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Describe the results

- The regression tests whether the coefficients β_0 , β_1 and β_2 are 0. This is the null hypothesis
- From the p-values, we deduce that the intercept and β_1 are significant and hence we do not accept the null hypothesis for them.
- β_2 , however, is not significant and thus its null hypothesis is accepted.
- The adjusted R^2 is 0.276 i.e., 27.6% of the variance of the response (y) is explianed by the regressors x1 and x2.

What are β_0 , β_1 and β_2 ?

```
params = results.params.to_frame().transpose()
params["Index"] = ["Estimate"]
params.set_index("Index")
```

	Intercept	x1	x2
Index			
Estimate	1.957909	1.615368	0.942777

How do these relate to the true β_0 , β_1 and β_2 ?

Index	Intercept	x1	x2
Estimate		1.615368	0.942777
Original		2.00000000000000000	0.300000000000000000

Influential points

```
get_influence_points(results)
```

```
(
   dfb_Intercept
                dfb_x1
                          dfb_x2 cooks_d hat_diag student_resid \
99
        0.628518 -0.530616 0.253443 0.133708
                                                      2.934959
                                          0.04782
     dffits student_resid_pvalue hat_influence cooks_d_pvalue
                      0.002087
99 0.657732
                                   0.140351
                                                0.939756
{'n': 100.0,
 'p': 3,
 'dfbetas_cutoff': 0.3,
 'dffits_cutoff': 0.34641016151377546,
 'studentized_residuals_cutoff': 3.0,
 'studentized_residuals_pvalue_cutoff': 0.01,
 'cooks_d_cutoff': 1.0,
 'cooks_d_pvalue_cutoff': 0.05})
```

(d) Now fit a least squares regression to predict y using only x_1 . Comment on your results. Can you reject the null hypothesis $H_0:\beta_1=0$?

```
def fit_x1(df):
   formula = "y ~ x1"
   model = smf.ols(f"{formula}", df)
   results = model.fit()
   print(results.summary())
   return results

results = fit_x1(df);
```

OLS Regression Results

Dep. Variable:	у	R-squared:	0.281
Model:	OLS	Adj. R-squared:	0.274
Method:	Least Squares	F-statistic:	38.39
Date:	Wed, 16 Oct 2024	Prob (F-statistic):	1.37e-08
Time:	18:03:33	Log-Likelihood:	-131.28
No. Observations:	100	AIC:	266.6
Df Residuals:	98	BIC:	271.8
Df Model:	1		
Covariance Type:	nonrobust		

	coef	std err	t	P> t	[0.025	0.975]
Intercept x1	1.9371 2.0771	0.189 0.335	10.242 6.196	0.000	1.562 1.412	2.312 2.742
Omnibus: Prob(Omnibus Skew: Kurtosis:):	0	.903 Jaro .046 Prob	in-Watson: que-Bera (JB) (JB): . No.):	1.931 0.042 0.979 4.65
=========	=======	========	=======	========		========

Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Can you reject the null hypothesis $H_0: \beta_1 = 0$?

• Yes, we can reject the null hypothesis $H_0: \beta_1 = 0$ since the p-value for the coefficient of x_1 is significant.

Influential points

get_influence_points(results)

```
n = 100.0, p = 2
Average Hat Leverage: 0.020000000000000004
Hat Leverage Cutoff = 2 * Average Hat Leverage = 0.04000000000000001
DFBetas Cutoff = 3 / sqrt(n) = 0.3
DFFITS Cutoff = 2 * sqrt(p/n) = 0.282842712474619
Cooks Distance Cutoff = 1.0
Cooks Distance p-value Cutoff = 0.05
Studentized Residuals Cutoff = 3.0
Studentized Residuals p-value Cutoff = 0.01
(
     dfb_Intercept dfb_x1
                             cooks_d hat_diag student_resid
                                                                  dffits \
                                                      3.027795 0.623819
 99
         0.623723 -0.541833 0.179606
                                        0.04072
     student_resid_pvalue hat_influence cooks_d_pvalue
                                               0.835874 ,
 99
                0.001578
                               0.123292
 {'n': 100.0,
```

(e) Now fit a least squares regression to predict y using only x_2 . Comment on your results. Can you reject the null hypothesis $H_0:\beta_1=0$?

```
def fit_x2(df):
   formula = "y ~ x2"
   model = smf.ols(f"{formula}", df)
   results = model.fit()
   print(results.summary())
   return results

results = fit_x2(df);
```

OLS Regression Results

==========						
Dep. Variable	e:		y R-sq	uared:		0.222
Model:		(OLS Adj.	R-squared:		0.214
Method:		Least Squa	res F-st	atistic:		27.99
Date:	W	ed, 16 Oct 20	024 Prob	(F-statistic	:):	7.43e-07
Time:		18:03	:33 Log-	Likelihood:		-135.24
No. Observat:	ions:		100 AIC:			274.5
Df Residuals	:		98 BIC:			279.7
Df Model:			1			
Covariance T	ype:	nonrob	ust			
========	coef	std err	t	P> t	[0.025	0.975]
Intercept	2.3239	0.154	15.124	0.000	2.019	2.629
x2	2.9103	0.550	5.291	0.000	1.819	4.002
Omnibus:	======	0.1	======= 191 Durb	in-Watson:		1.943

```
      Prob(Omnibus):
      0.909
      Jarque-Bera (JB):
      0.373

      Skew:
      -0.034
      Prob(JB):
      0.830

      Kurtosis:
      2.709
      Cond. No.
      6.11
```

Notes

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

Can you reject the null hypothesis $H_0: \beta_1 = 0$?

'cooks_d_pvalue_cutoff': 0.05})

• Yes, we can reject the null hypothesis $H_0: \beta_1 = 0$ since the p-value for the coefficient of x_2 is significant.

Influential points

```
get_influence_points(results)
n = 100.0, p = 2
Average Hat Leverage: 0.0199999999999993
DFBetas Cutoff = 3 / sqrt(n) = 0.3
DFFITS Cutoff = 2 * sqrt(p/n) = 0.282842712474619
Cooks Distance Cutoff = 1.0
Cooks Distance p-value Cutoff = 0.05
Studentized Residuals Cutoff = 3.0
Studentized Residuals p-value Cutoff = 0.01
(Empty DataFrame
Columns: [dfb_Intercept, dfb_x2, cooks_d, hat_diag, student_resid, dffits, student_resid_pv
Index: [],
 {'n': 100.0,
  'p': 2,
  'average_hat': 0.0199999999999999,
  'dfbetas_cutoff': 0.3,
  'dffits_cutoff': 0.282842712474619,
  'studentized_residuals_cutoff': 3.0,
  'studentized_residuals_pvalue_cutoff': 0.01,
  'cooks_d_cutoff': 1.0,
```

- (f) Do the results obtained in (c)–(e) contradict each other? Explain your answer.
 - No, the results do not contradict each other since the two variables are collinear and contain the same information.
 - Thus, they can be interchanged for each other without much loss of information in the regression model.
- (g) Suppose we obtain one additional observation, which was unfortunately mismeasured. We use the function np.concatenate() to add this additional observation to each of x_1, x_2 and y.

```
x1 = np.concatenate ([x1 , [0.1]])
x2 = np.concatenate ([x2 , [0.8]])
y = np.concatenate ([y, [6]])
```

Re-fit the linear models from (c) to (e) using this new data. What effect does this new observation have on the each of the models? In each model, is this observation an outlier? A high-leverage point? Both? Explain your answers.

Add an additional observation

```
x1 = np.concatenate ([x1 , [0.1]])
x2 = np.concatenate ([x2 , [0.8]])
y = np.concatenate ([y, [6]]);
```

```
x1[-1], x2[-1], y[-1]
```

```
(0.1, 0.8, 6.0)
```

```
df = construct_df(x1,x2,y)
df.tail(1)
```

	x1	x2	У
100	0.1	0.8	6.0

Combined regression

```
results = fit_combined(df);
```

OLS Regression Results

		ULS K	egression 	Kesuits 		
Dep. Variable	e:		y R-s	quared:		0.292
Model:	•		•	. R-squared:		0.277
Method:		Least Squ	-	tatistic:		20.17
Date:	,	Wed, 16 Oct		b (F-statistic):	4.60e-08
Time:		18:0	3:34 Log	-Likelihood:		-135.30
No. Observat:	ions:		101 AIC):		276.6
Df Residuals	:		98 BIC	: :		284.5
Df Model:			2			
Covariance T	ype:	nonro	bust			
=========	======		=======	========		
	coef	std err	t	P> t	[0.025	0.975]
Intercept	2.0618	0.192	10.720	0.000	1.680	2.443
x1	0.8575	0.466	1.838	0.069	-0.068	1.783
x2	2.2663	0.705	3.216	0.002	0.868	3.665
			=======			
Omnibus:		0	.139 Dur	bin-Watson:		1.894
Prob(Omnibus):	0	.933 Jar	que-Bera (JB):		0.320
Skew:		0	.013 Pro	b(JB):		0.852
Kurtosis:		2	.725 Cor	d. No.		9.68

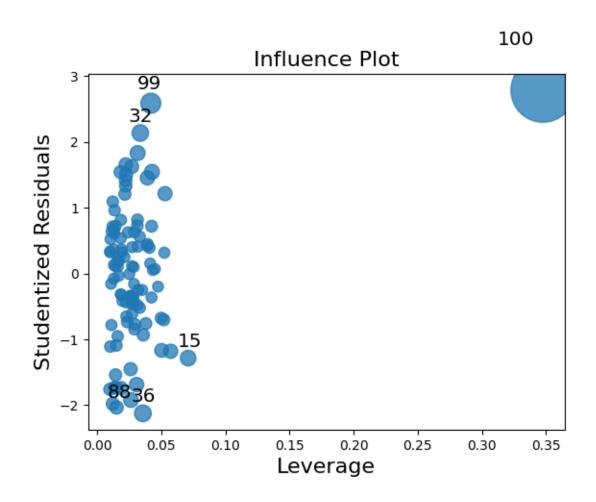
Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

- Here, we see the effect of the additional mismeasured data point.
- The effect on the combined regression is to switch the significance of the regressors x1 and x2.
- Now, the coefficient of x1 is not statistically significant with a p-value of 0.07.

Residuals, outliers, leverage and influence

display_cooks_distance_plot(results);



display_hat_leverage_plot(results)

Unable to display output for mime type(s): text/html

Unable to display output for mime type(s): application/vnd.plotly.v1+json, text/html

get_influence_points(results)

```
n = 101.0, p = 3
Average Hat Leverage: 0.0297029702972
Hat Leverage Cutoff = 2 * Average Hat Leverage = 0.05940594059405944
DFBetas Cutoff = 3 / sqrt(n) = 0.29851115706299675
DFFITS Cutoff = 2 * sqrt(p/n) = 0.3446909937728556
Cooks Distance Cutoff = 1.0
Cooks Distance p-value Cutoff = 0.05
Studentized Residuals Cutoff = 3.0
Studentized Residuals p-value Cutoff = 0.01
(
      dfb_Intercept
                                          cooks_d hat_diag student_resid \
                       dfb_x1
                                 dfb_x2
           0.522200 -0.421657
                               0.129666 0.092112 0.041905
                                                                   2.585431
 99
 100
           0.558412 -1.679554 1.941733 1.287988 0.347672
                                                                  2.783731
        dffits
                student_resid_pvalue
                                     hat_influence
                                                     cooks_d_pvalue
 99
      0.540708
                            0.005607
                                           0.108343
                                                           0.964231
 100 2.032257
                            0.003230
                                           0.967824
                                                           0.282850
 {'n': 101.0,
  'p': 3,
  'average hat': 0.02970297029702972,
  'hat_leverage_cutoff': 0.05940594059405944,
  'dfbetas_cutoff': 0.29851115706299675,
  'dffits_cutoff': 0.3446909937728556,
  'studentized_residuals_cutoff': 3.0,
  'studentized_residuals_pvalue_cutoff': 0.01,
  'cooks_d_cutoff': 1.0,
  'cooks_d_pvalue_cutoff': 0.05})
```

- From the above, we can see that there are two influential datapoints, 99 and 100.
- This is initially surprising until we compute the influential points without the freshly added mismeasured data point and discover that point 99 was influential in the earlier regression.

Regress on x1

```
OLS Regression Results

------
Dep. Variable:

y R-squared:

0.217
```

Model: Method: Date: Time: No. Observations:	OLS Least Squares Wed, 16 Oct 2024 18:03:35 101	Adj. R-squared: F-statistic: Prob (F-statistic) Log-Likelihood: AIC:	:	0.209 27.42 9.23e-07 -140.37 284.7
Df Residuals:	99	BIC:		290.0
Df Model:	1			
Covariance Type:	nonrobust			
coei	std err	t P> t	[0.025	0.975]
Intercept 2.0739	0.201 1	0.310 0.000	1.675	2.473
x1 1.8760	0.358	5.236 0.000	1.165	2.587
Omnibus: Prob(Omnibus): Skew: Kurtosis:	8.232 0.016 0.396 4.391	Durbin-Watson: Jarque-Bera (JB): Prob(JB): Cond. No.	=====	1.636 10.781 0.00456 4.61

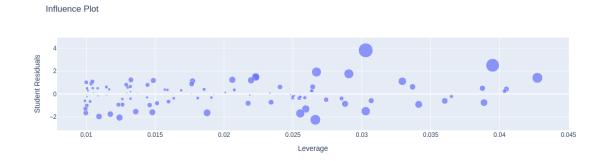
Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

display_cooks_distance_plot(results);

100 Influence Plot 4 99 Studentized Residuals 3 52 2 1 0 -1 -2 0.015 0.020 0.025 0.030 0.035 0.040 0.010 Leverage

display_hat_leverage_plot(results)



get_influence_points(results)

```
n = 101.0, p = 2
Average Hat Leverage: 0.0198019801980198
Hat Leverage Cutoff = 2 * Average Hat Leverage = 0.0396039603960396
DFBetas Cutoff = 3 / sqrt(n) = 0.29851115706299675
DFFITS Cutoff = 2 * sqrt(p/n) = 0.2814390178921167
Cooks Distance Cutoff = 1.0
Cooks Distance p-value Cutoff = 0.05
Studentized Residuals Cutoff = 3.0
Studentized Residuals p-value Cutoff = 0.01
(
      dfb_Intercept
                       dfb_x1
                                cooks_d hat_diag student_resid
                                                                    dffits \
           0.532991 -0.461445 0.134079 0.039495
 99
                                                        2.628842 0.533075
 100
           0.734700 -0.605898 0.233830 0.030283
                                                        4.179207 0.738540
      student_resid_pvalue hat_influence cooks_d_pvalue
 99
                  0.004974
                                 0.103827
                                                 0.874679
 100
                  0.000032
                                 0.126561
                                                 0.791933
 {'n': 101.0,
  'p': 2,
  'average_hat': 0.0198019801980198,
  'hat_leverage_cutoff': 0.0396039603960396,
  'dfbetas_cutoff': 0.29851115706299675,
  'dffits_cutoff': 0.2814390178921167,
  'studentized_residuals_cutoff': 3.0,
  'studentized_residuals_pvalue_cutoff': 0.01,
  'cooks_d_cutoff': 1.0,
  'cooks_d_pvalue_cutoff': 0.05})
```

• Similarly, in the regression of y on x1 only, we find points 99 and 100 to be influential.

Regress on x2

```
OLS Regression Results

------
Dep. Variable:

y R-squared:

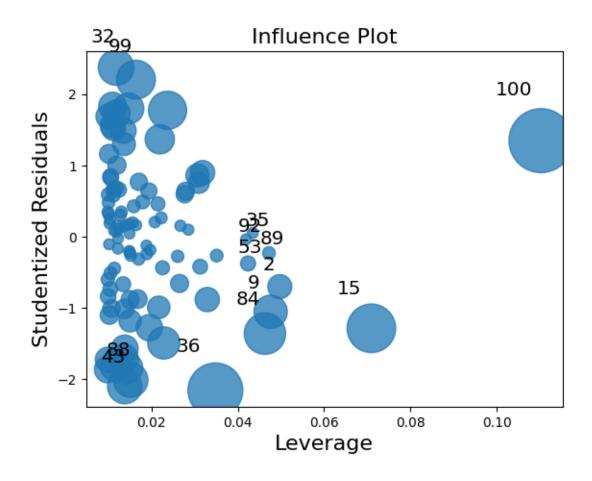
0.267
```

Model:			OLS	Adi.	R-squared:		0.260
Method:		Least Squ	ares	-	atistic:		36.10
Date:		Wed, 16 Oct	2024	Prob	(F-statistic)	:	3.13e-08
Time:		18:0	3:54	Log-	Likelihood:		-137.01
No. Observati	ons:		101	AIC:			278.0
Df Residuals:			99	BIC:			283.3
Df Model:			1				
Covariance Ty	pe:	nonro	bust				
========	======		=====		=========	:=======	
	coei	f std err		t	P> t	[0.025	0.975]
Intercept	2.2840	0.151	15	5.088	0.000	1.984	2.584
x2	3.1458	0.524	6	3.008	0.000	2.107	4.185
 Omnibus:	======		===== . 495	Durb	======== in-Watson:	-======	1.939
Prob(Omnibus)	:	_	.781		ue-Bera (JB):		0.631
Skew:		-0	.041	-	(JB):		0.729
Kurtosis:		2	.621	Cond	. No.		5.84

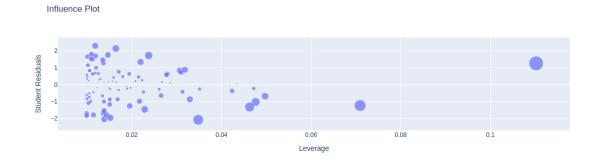
Notes:

[1] Standard Errors assume that the covariance matrix of the errors is correctly specified.

display_cooks_distance_plot(results);



display_hat_leverage_plot(results)



get_influence_points(results)

$$n = 101.0, p = 2$$

```
Average Hat Leverage: 0.019801980198019806
Hat Leverage Cutoff = 2 * Average Hat Leverage = 0.03960396039603961
DFBetas Cutoff = 3 / sqrt(n) = 0.29851115706299675
DFFITS Cutoff = 2 * sqrt(p/n) = 0.2814390178921167
Cooks Distance Cutoff = 1.0
Cooks Distance p-value Cutoff = 0.05
Studentized Residuals Cutoff = 3.0
Studentized Residuals p-value Cutoff = 0.01
(Empty DataFrame
   Columns: [dfb_Intercept, dfb_x2, cooks_d, hat_diag, student_resid, dffits, student_resid_pvalue columns cooks_d, hat_diag, student_resid, dffits, student_resid_pvalue columns cooks_d, hat_diag, student_resid, dffits, student_resid_pvalue cooks_d, hat_diag, student_resid, dffits, student_resid_pvalue cooks_d, hat_diag, studen
   Index: [],
   {'n': 101.0,
       'p': 2,
       'average_hat': 0.019801980198019806,
        'hat_leverage_cutoff': 0.03960396039603961,
        'dfbetas_cutoff': 0.29851115706299675,
        'dffits_cutoff': 0.2814390178921167,
        'studentized_residuals_cutoff': 3.0,
        'studentized_residuals_pvalue_cutoff': 0.01,
        'cooks_d_cutoff': 1.0,
        'cooks_d_pvalue_cutoff': 0.05})
```

• In the regression of y on x2, no data point is influential since neither the studentized residuals or their associated p-values cross the thresholds for these parameters.

allDone();

<IPython.lib.display.Audio object>