Name: George Loo and Daniel Yuan

# In-Class Workshop 5

# Edit Expectations for Values Using Simple 12 Month Average

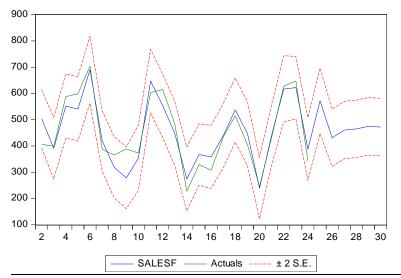
obs	sales	prom	adv	index
1	504.72	15.6	30	100
2	406.59	22.2	36	102
3	398.55	0	45	104
4	587.76	0	57	104
5	598.92	0	39	104
6	703.62	31.8	21	100
7	387.24	21.3	12	98
8	365.67	3.9	6	96
9	388.71	0	6	98
10	372.96	8.4	30	103
11	603.3	45.3	30	105
12	614.73	50.1	33	107
13	484.38	39.6	6	107
14	227.76	4.2	33	107
15	329.13	0	6	108
16	308.25	0	3	105
17	433.86	0	45	103
18	514.98	13.8	48	108
19	404.7	17.7	0	110
20	245.43	0	15	112
21	433.2	17.4	9	113
22	627.24	37.8	54	112
23	647.61	42.3	36	113
24	342.81	11.4	39	114
25		15	25	109
26		13	26	110
27		14	26	110
28		15	27	110
29		16	29	110
30		18	28	111

# A colleague's model

Dependent Variable: SALES
Method: Least Squares
Date: 02/16/19 Time: 11:50
Sample (adjusted): 2 24
Included observations: 23 after adjustments

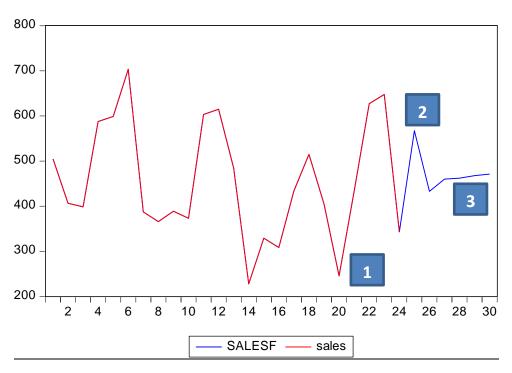
Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	341.2918	28.95320	11.78770	0.0000
ADV(-1)	3.238252	0.875865	3.697203	0.0016
PROM	5.818424	1.002906	5.801565	0.0000
PROM(-1)	-3.621209	0.969695	-3.734381	0.0015
DLOG(INDEX)	-1197.232	682.6878	-1.753704	0.0965
R-squared	0.782850	Mean dependent var		453.3652
Adjusted R-squared	0.734594	S.D. dependent var		134.3372
S.E. of regression	69.20722	Akaike info criterion		11.50175
Sum squared resid	86213.50	Schwarz criterion		11.74859
Log likelihood	-127.2701	Hannan-Quinn criter.		11.56383
F-statistic	16.22300	Durbin-Wats	on stat	1.717798
Prob(F-statistic)	0.000009			

Plot salesf and actuals



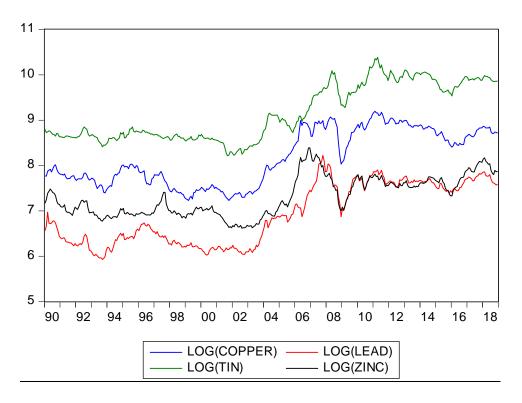
Forecast: SALESF Actual: SALES Forecast sample: 1 30 Adjusted sample: 2 30 Included observations: 29 Root Mean Squared Error 45.91298 Mean Absolute Error 37.43406 Mean Abs. Percent Error 9.133037 Theil Inequality Coef. 0.048750 Bias Proportion 0.000000 Variance Proportion 0.032550 Covariance Proportion 0.967450 Theil U2 Coefficient 0.290669 Symmetric MAPE 9.092548

### Salesf against sales



Forecasts (2) seem to mirror the same pattern observed 12 months ago (month 14 onwards) (1) before tapering off in month 27-31 (3). This is possibly because I used a L12M average for the variables for prom adv and index.

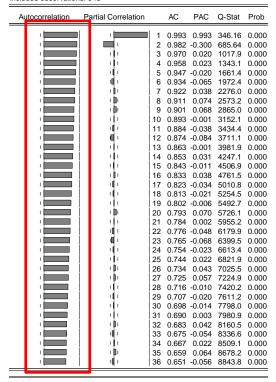
Metal Prices (Monthly Log(Prices) from 1990-2018)



### To check for stationarity we used the correlogram

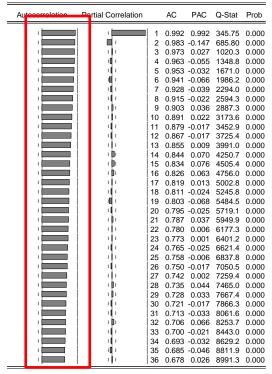
### Correlogram of LOG Copper

Date: 02/16/19 Time: 15:29 Sample: 1990M01 2018M12 Included observations: 348



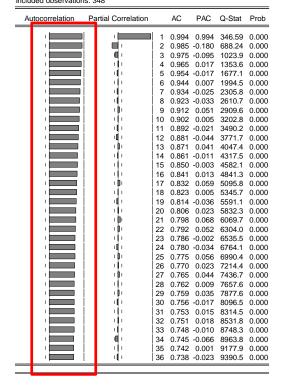
Correlogram of LOG Lead

Date: 02/16/19 Time: 15:30 Sample: 1990M01 2018M12 Included observations: 348



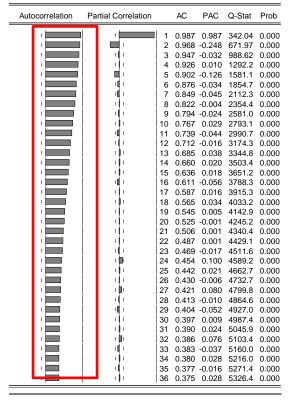
#### Correlogram of LOG Tin

Date: 02/16/19 Time: 15:32 Sample: 1990M01 2018M12 Included observations: 348



Correlogram of Log Zinc

Date: 02/16/19 Time: 15:33 Sample: 1990M01 2018M12 Included observations: 348



A check against the above correlograms indicates that they are non stationary.

Johansen-Cointegration Test

Date: 02/16/19 Time: 15:35
Sample (adjusted): 1990M03 2018M12
Included observations: 346 after adjustments
Trend-assumption: No deterministic trend (restricted constant)
Series: LOG(COPPER) LOG(LEAD) LOG(TIN) LOG(ZINC)
Lags interval (in first differences): 1 to 1

Unrestricted Cointegration Rank Test (Trace)							
Hypothesized No. of CE(s)	Eigenvalue	Trace Statistic	0.05 Critical Value	Prob.**			
None *	0.078118	54.58767	54.07904	0.0450			
None * At most 1	0.078118 0.044796	54.58767 26.44455	54.07904 35.19275	0.0450			
				0.0.00			

Trace test indicates 1 cointegrating eqn(s) at the 0.05 level

Unrestricted Cointegration Rank Test (Maximum Eigenvalue)

Hypothesized No. of CE(s)	Eigenvalue	Max-Eigen Statistic	0.05 Critical Value	Prob.**	= _ 2 Cointegrating E	Equation(s):	Log likelihood	2200.512	
None	0.078118	28.14312	28.58808	0.0569	N 1 - 1 -		:t- (-tll		\
At most 1	0.044796	15.85728	22.29962	0.3085			eients (standard er		
At most 2	0.026528	9.302613	15.89210	0.4020	LOG(COPPER)		LOG(TIN)	LOG(ZINC)	С
At most 3	0.003706	1.284661	9.164546	0.9101	1.000000	0.000000	-1.045597	0.196117	-0.001223
					=		(0.17043)	(0.24422)	(1.03506)
		o cointegration a			0.000000	1.000000	-1.060957	0.011848	2.730669
**MacKinnon-Ha		hesis at the 0.05 999) p-values	ievei				(0.10856)	(0.15556)	(0.65931)
							d error in parenth	eses)	
Unrestricted Col	integrating Coe	mcients (normaliz	zed by b'*S11*b=l	):	D(LOG(COP	-0.037830	0.051048		
LOG(COPPER)	LOG(LEAD)	LOG(TIN)	LOG(ZINC)	С	_	(0.01632)	(0.02104)		
-2.664133	6.755345	-4.381520	-0.442444	18.44987	D(LOG(LEAD))	0.032134	-0.038308		
4.599743	-1.138564	-3.601509	0.888597	-3.114668		(0.01957)	(0.02522)		
3.798540	1.556980	-2.014232	-5.279514	15.14970	D(LOG(TIN))	0.004214	0.058059		
-0.254843	0.728597	0.615509	0.310015	-11.42763		(0.01477)	(0.01904)		
					D(LOG(ZINC))	-0.006969	0.011657		
						(0.01671)	(0.02154)		
Unrestricted Adj	ustment Coeffic	cients (alpha):			=				
D(LOG(COP	0.006838	-0.004264	-0.001509	-0.002849	3 Cointegrating E	Equation(s):	Log likelihood	2205.164	
D(LOG(LEAD))	-0.004979	0.004102	-0.001740	-0.003780		1	- 3		
D(LOG(TIN))	0.009695	0.006532	0.000283	-0.001506	Normalized coint	egrating coeffic	ients (standard er	ror in parenthes	es)
D(LOG(ZINC))	0.001629	-0.000571	0.005745	-0.002782	LOG(COPPER)		LOG(TIN)	LOG(ZINC)	C
					1.000000	0.000000	0.000000	-1.554440	3.157166
4 Caintannetina I		Las Olabara	0400 504					(0.22633)	(1.65523)
1 Cointegrating I	equation(s):	Log likelihood	2192.584		- 0.000000	1.000000	0.000000	-1.764425	5.935456
Normalized coint	earating coeffic	iante (etandard a	rror in parenthese	ae)				(0.25745)	(1.88281)
LOG(COPPER)		LOG(TIN)	LOG(ZINC)	C C	0.000000	0.000000	1.000000	-1.674218	3.020657
1.000000	-2.535663	1.644632	0.166074	-6.925281	0.00000	0.00000		(0.27570)	(2.01632)
1.000000	(0.45645)	(0.40814)	(0.34076)	(1.93706)				(0.21010)	(2.01002)
		, ,		/	Adjustment coeff	icients (standar	d error in parenth	eses)	
Adjustment coeff		d error in parenth	neses)		D(LOG(COP	-0.043561	0.048699	-0.011566	
D(LOG(COP	-0.018218					(0.02006)	(0.02157)	(0.01848)	
	(0.00820)				D(LOG(LEAD))	0.025526	-0.041016	0.010547	
D(LOG(LEAD))	0.013266				,,,	(0.02404)	(0.02585)	(0.02215)	
D/I 00/Th:::	(0.00982)				D(LOG(TIN))	0.005290	0.058500	-0.066575	
D(LOG(TIN))	-0.025830				_(LOO(1114))	(0.01816)	(0.01953)	(0.01673)	
D(LOG(ZINC))	(0.00746) -0.004341				D(LOG(ZINC))	0.014855	0.020603	-0.016654	
D(LOG(ZINC))	(0.00838)				D(200(Zii40))	(0.02044)	(0.02198)	(0.01883)	
	(0.00000)								

From the above, we can see that we can reject the null hypothesis that there are no cointegrated variables at a 95% confidence interval under the trace test and 94% confidence interval under the maximum eigen value test.

**Vector Error Correction Estimates** 

<sup>\*</sup> denotes rejection of the hypothesis at the 0.05 level \*\*MacKinnon-Haug-Michelis (1999) p-values

Vector Error Correction Estimates Date: 02/16/19 Time: 15:47

Sample (adjusted): 1990M03 2018M12 Included observations: 346 after adjustments Standard errors in ( ) & t-statistics in [ ]

Cointegrating Eq:	CointEq1	
LOG(COPPER(-1))	1.000000	
LOG(LEAD(-1))	-2.535663 (0.45645) [-5.55520]	
LOG(TIN(-1))	1.644632 (0.40814) [4.02953]	
LOG(ZINC(-1))	0.166074 (0.34076) [ 0.48736]	,
С	-6.925281 (1.93706) [-3.57515]	

### The equation indicates that:

LOG(COPPER(-1)) = 5.5 LOG(LEAD(-1)) - 1.65 LOG(TIN(-1)) + 0.17 LOG(ZINC(-1)) - 6.93

There is a positive correlation between the price of copper and lead, while there is a negative correlation between the price of copper and the price of tin.

We note from the t statistics that LOG ZINC is not statistically significant, meaning it is not entering into the long term equation with copper, lead, and tin.

**System of Equations** 

Error Correction:	D(LOG(CO	D(LOG(LE	D(LOG(TIN))	D(LOG(ZINC))
CointEq1	-0.018358	0.014190	-0.025730	-0.004139
	(0.00838)	(0.01002)	(0.00762)	(0.00855)
	[-2.19159]	[1.41546]	[-3.37608]	[-0.48402]
D(LOG(COPPER(-1)))	0.420359	0.088349	0.091142	0.190775
	(0.06853)	(0.08201)	(0.06235)	(0.06996)
	[ 6.13430]	[ 1.07727]	[ 1.46181]	[ 2.72676]
D(LOG(LEAD(-1)))	-0.090019	0.227931	0.017069	-0.051980
	(0.05867)	(0.07022)	(0.05338)	(0.05990)
	[-1.53429]	[ 3.24603]	[ 0.31975]	[-0.86775]
D(LOG(TIN(-1)))	-0.022938	-0.155624	0.199468	-0.222176
	(0.06518)	(0.07801)	(0.05931)	(0.06655)
	[-0.35192]	[-1.99493]	[ 3.36338]	[-3.33850]
D(LOG(ZINC(-1)))	0.032855	0.024887	-0.016007	0.303320
	(0.06805)	(0.08144)	(0.06192)	(0.06948)
	[ 0.48281]	[ 0.30558]	[-0.25854]	[ 4.36574]
C	0.001830	0.002201	0.002413	0.001508
	(0.00309)	(0.00370)	(0.00281)	(0.00315)
	[ 0.59238]	[ 0.59525]	[ 0.85857]	[ 0.47818]

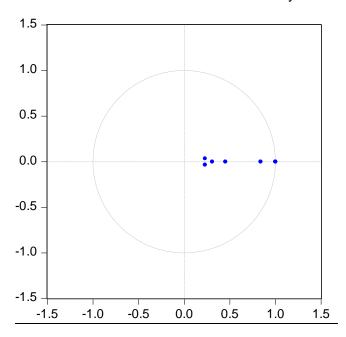
From the VECR lead and tin have a statistically significant long term relationship with copper, while zinc does not.

From the system of equations, tin(-1) is the most important variable in forecasting the prices of lead, tin, and zinc, but not copper.

I would improve the model by dropping the Zinc variable which is does not have a long term relationship with the other variables, and instead use past Zinc prices to forecast Zinc prices in the future.

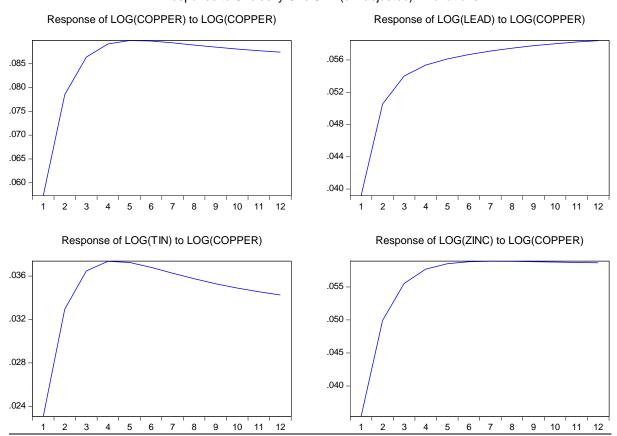
Inverse roots of AR characteristic polynomial

# Inverse Roots of AR Characteristic Polynomial



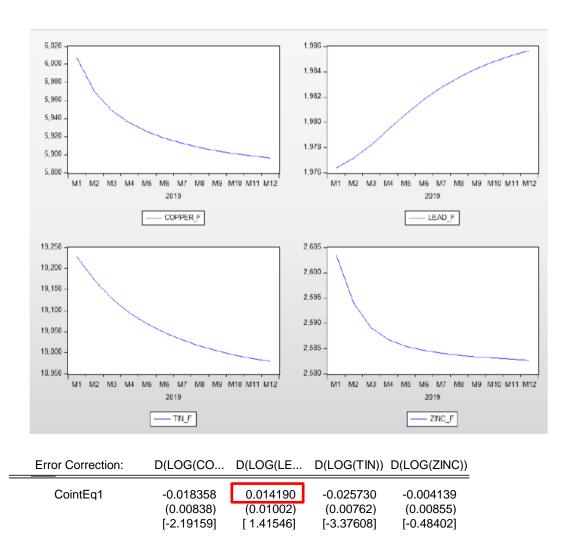
### **Impulse Response Graph**

Response to Cholesky One S.D. (d.f. adjusted) Innovations



Finally we test the model's forecasts.

### Out of Sample Forecasting



From the out of sample forecasting above, we can conclude:

In the next 12 months, prices of copper, tin, and zinc are expected to decrease, while prices of lead are expected to increase.

This is explained by the above circled long term change in price of lead which is expected to have a positive coefficient, while the other metals have a negative coefficient, forecasting a downward trend.