Time Series Analysis and Forecasting Vinod / Ashesh / Chris Workshop 4: Stationarity and ARIMA

(a) Dickey-Fuller Test

(i) Plotting correlogram

Correlogram is plotted to check if the time series is stationary. The autocorrelation graph is as shown,

> Sample: 1960M01 2018M12 Included observations: 708

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
		1	0.992	0.992	699.84	0.000
	<u> </u>	2	0.980		1384.0	0.000
	1	3	0.968	0.008	2051.6	0.000
	ינוףי	4	0.955	0.030	2703.3	0.000
	' '	5	0.944	0.004	3339.9	0.000
	I[I	6	0.931	-0.037	3961.0	0.000
	ינוןי	7	0.919	0.027	4567.3	0.000
	יוף ו	8	0.909	0.068	5160.5	0.000
	'P	9	0.901	0.111	5743.7	0.000
	' '	10		-0.004	6318.5	0.000
	q ·	11		-0.055	6884.1	0.000
	"	12		-0.113	7438.2	0.000
	100	13	0.866	0.016	7980.0	0.000
	ווי ו	14	0.856	0.069	8510.8	0.000
	1	15		-0.008	9031.3	0.000
	1111	16	0.838	0.019	9542.0	0.000
	1	17	0.830	0.021	10043.	0.000
	1 1	18	0.822	0.006	10535.	0.000
	III	19	0.814	0.010	11018.	0.000
	1	20	0.807	0.006	11494.	0.000
	(l)	21	0.800	-0.029	11963.	0.000
	(t)	22	0.792	-0.045	12423.	0.000
	ψ	23	0.783	-0.017	12873.	0.000
	1 1	24	0.774	0.002	13314.	0.000
	')	25	0.766	0.034	13745.	0.000
	l iĝi	26	0.758	0.031	14168.	0.000
	l ip	27	0.751	0.090	14585.	0.000
1	III	28	0.746	0.012	14997.	0.000
	III	29	0.741	0.024	15404.	0.000
1	1 1	30	0.738	0.000	15807.	0.000
		31	0.734	-0.002	16207.	0.000
1	·)	32	0.732	0.061	16605.	0.000

From the plot, we realize that the autocorrelation coefficients remain non-zero for many lags and do not die out quickly; hence it resembles characteristics similar to that of non-stationary series.

(ii) Dickey-Fuller test

Title:

Augmented Dickey-Fuller Test

Test Results: PARAMETER:

Lag Order: 19

STATISTIC:

Dickey-Fuller: -0.0962

P VALUE: 0.5857

Null Hypothesis: COPPER has a unit root Exogenous: None
Lag Length: 1 (Automatic - based on SIC, maxlag=19)

		t-Statistic	Prob.*
Augmented Dickey-Fu	ller test statistic	-0.525351	0.4888
Test critical values:	1% level	-2.570670	
	5% level	-1.941606	
	10% level	-1.616176	

*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation Dependent Variable: D(COPPER)
Method: Least Squares
Date: 02/14/19 Time: 22:26 Sample (adjusted): 1985M03 2018M12 Included observations: 406 after adjustments

Variable	Coefficient	Std. Error	t-Statistic	Prob.
COPPER(-1) D(COPPER(-1))	-0.001635 0.339270	0.003112 0.046923	-0.525351 7.230366	0.5996 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood Durbin-Watson stat	0.113315 0.111120 288.1898 33553558 -2874.516 1.975388	Mean dependent var S.D. dependent var Akaike info criterion Schwarz criterion Hannan-Quinn criter.		11.54389 305.6730 14.17003 14.18977 14.17784

The results from the Dickey-Fuller test is as shown above. The null hypothesis is to test whether COPPER series has unit root, meaning it is not stationary and imply random walk type behaviour. To check if the exchange rate series can be rejected, the t-statistic should be compared with Dickey-fuller distribution for unit root distribution. If t < DF or p-value is less than 0.05, then reject null hypotheses. In this case, since the result is opposite, we do not reject null hypothesis suggesting that series is indeed non-stationary.

(iii) Adding constant and trend in basic DF test

Dickey-Fuller test with constant and trend added is tested and the result is as shown,

Title: Augmented Dickey-Fuller Test

Test Results:
PARAMETER:
Lag Order: 19
STATISTIC:

Dickey-Fuller: -2.4675

P VALUE: 0.3804

Null Hypothesis: COPPER has a unit root Exogenous: Constant, Linear Trend Lag Length: 1 (Automatic - based on SIC, maxlag=19)

		t-Statistic	Prob.*
Augmented Dickey-Fu	ller test statistic	-2.790857	0.2015
Test critical values:	1% level	-3.980823	
	5% level	-3.420930	
	10% level	-3.133194	

*MacKinnon (1996) one-sided p-values.

Augmented Dickey-Fuller Test Equation Dependent Variable: D(COPPER) Method: Least Squares Date: 02/14/19 Time: 22:30 Sample (adjusted): 1985M03 2018M12 Included observations: 406 after adjustments

Variable	Coefficient	Std. Error t-Statist		Prob.
COPPER(-1)	-0.025480	0.009130 -2.790857		0.0055
D(COPPER(-1))	0.349569	0.046755 7.476613		0.0000
C	-91.23446	73.00562 -1.249691		0.2121
@TREND("1960M01")	0.393199	0.188703 2.083698		0.0378
R-squared	0.130186	Mean dependent var		11.54389
Adjusted R-squared	0.123695	S.D. dependent var		305.6730
S.E. of regression	286.1440	Akaike info criterion		14.16067
Sum squared resid	32915119	Schwarz criterion		14.20014
Log likelihood F-statistic Prob(F-statistic)	-2870.616 20.05600 0.000000	Hannan-Quin Durbin-Watso	14.17629 1.985789	

From the result, we can see that the p-value is still high and t-statistic is not less than DF, therefore it does not give enough evidence to reject the null hypotheses, indicating that it's non-stationary data.

Since both above cases indicate Non Stationary in the dataset, we introduce stationary in the dataset by using the DIFF(dataset) function and then checking the Dickey Fuller test to verify if Non-Stationary is removed.

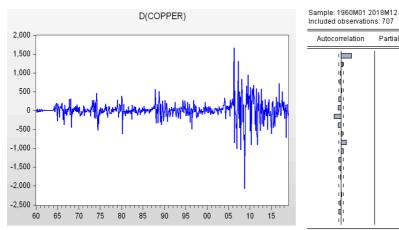
(iv) Augmented DF test including lagged values of dependent variables

We check the autocorrelation coefficients of the residuals to see if series is stationary.

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
1/10	1 1	1 0.01	4 0.014	0.1372	0.711
ı(i	l (fi	1	0 -0.031	0.7918	0.673
ılı .	l di	3 -0.02	1 -0.020	1.0975	0.778
ı d ı	l di	4 -0.02	8 -0.028	1.6539	0.799
ı j ı	<u> </u>	5 0.07	0.070	5.1854	0.394
d i	(d)	6 -0.05	2 -0.056	7.0883	0.313
I I	ļ iļi	7 0.00	2 0.007	7.0918	0.419
	<u> </u>	8 -0.16	9 -0.173	27.567	0.001
ı q ı	ψ.	9 -0.03	0 -0.022	28.218	0.001
- III		10 0.01	4 -0.006	28.349	0.002
' !		11 0.18		51.674	0.000
ų p	'P	12 0.07		55.268	0.000
q٠	4'	13 -0.10		63.049	0.000
1 1	Ψ.	1	3 -0.016	63.054	0.000
ų.	1	15 0.01		63.169	0.000
111	'¶'		5 -0.047	63.185	0.000
' p	ן יי	17 0.05		65.147	0.000
<u>"</u> "	' '	18 -0.02		65.470	0.000
q.	'['	19 -0.05		67.624	0.000
Ψ.	'l!'	20 -0.00		67.632	0.000
<u>'P</u>	! ! !	21 0.07		72.074	0.000
<u>'l'</u>	"!	22 0.05		74.594	0.000
7.	1 12 T	23 0.05		76.734	0.000
91	"1."		0 -0.031	78.556	0.000
11.	l ' <u>'</u> '.	25 -0.00		78.621	0.000
31.	l 5'.	26 -0.03		79.653	0.000
11.	1 7.	27 -0.02 28 -0.02	4 -0.019	80.064	0.000
31	J. J.	29 -0.04		80.699	0.000
1.	1 3	30 0.03		81.912	0.000
7.	""	31 -0.08		82.863 87.749	0.000
3.	1 1	32 0.01		87.946	0.000
36	" ;	33 0.02		88.244	0.000
ili.	ا أا	34 -0.01		88.404	0.000
ili.	l 16	35 0.04		89.927	0.000
Ψ.	1 P	100 0.04	0.040	00.021	5.566

From the residual ACF plot, we are able to see that the probability is high till lag 7 and then the probability drops below 0.05, indicating statistically significant values. This shows that there is autocorrelation and that the series is not stationary.

(v) Examining differences in series to check for stationarity



ncluded observation	s: 707					
Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
· =		1	0.333	0.333	78.954	0.000
ı þ	(d)	2	0.063	-0.055	81.743	0.000
10	(I)	3	-0.020	-0.027	82.020	0.000
4	10	4	-0.037	-0.021	82.987	0.000
ı j ı	i ju	5	0.009	0.033	83.043	0.000
d)	-	6	-0.081	-0.107	87.696	0.000
d)	(1)	- 7	-0.098	-0.045	94.588	0.000
<u></u> -	<u> </u>	8	-0.205	-0.176	124.70	0.000
d)	i ji	9	-0.090	0.037	130.57	0.000
i)n	ı)	10	0.025	0.041	131.02	0.000
· 🖿	· 🗖	11	0.166	0.160	150.89	0.000
ı j ı	(I)	12	0.074	-0.058	154.87	0.000
d)	-	13	-0.082	-0.104	159.68	0.000
ığı	1 1	14	-0.040	-0.004	160.86	0.000
ı(tı	1 1	15	-0.012	-0.004	160.96	0.000
1 1	(I)	16	-0.002	-0.046	160.96	0.000
ıþι	10	17	0.021	0.053	161.29	0.000
ı d ı	l di	18	-0.039	-0.037	162.37	0.000
d:	1 1	19	-0.066	0.001	165.51	0.000
ı ı	l li	20	-0.007	0.026	165.54	0.000

Title: Augmented Dickey-Fuller Test

Test Results:
PARAMETER:
Lag Order: 19
STATISTIC:

Dickey-Fuller: -6.1267

P VALUE: 0.01

Null Hypothesis: D(COPPER) has a unit root Exogenous: Constant, Linear Trend Lag Length: 0 (Automatic - based on SIC, maxlag=19)

		t-Statistic	Prob.*
Augmented Dickey-Fu Test critical values:	ller test statistic 1% level 5% level 10% level	-18.74024 -3.971104 -3.416195 -3.130392	0.0000

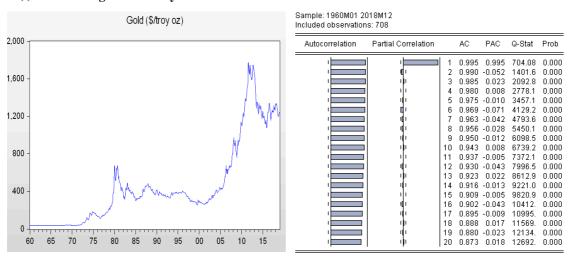
^{*}MacKinnon (1996) one-sided p-values.

Differencing is performed on the time series variable, copper; and its line plot, correlogram and unit root test is analyzed to check if its first differences are stationary.

One can observe from its correlogram that autocorrelation coefficients die down to zero after the first lag indicating stationarity. DF unit root test also proves the same as the t-statistic is lower than the DF critical values. The p-value is also lower than 0.05 in this case, hence signifying evidence against the null hypothesis of having a unit root(non-stationarity).

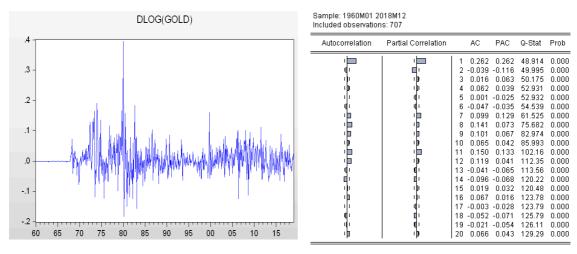
(b) ARIMA Modeling

(i) Establishing Stationarity



Correlogram of Gold time series is plotted and its autocorrelation coefficients are studied. Sine the coefficients decreases slowly without dying down quickly, the series is not stationary.

Differencing is done on the log return of the series and its correlogram is studied as shown,



From the correlogram, its apparent that the coefficient pattern resemble that of a stationary series. Since stationarity is established, Box-Jenkins methodology can be applied using Moving Average with lag 1.

(ii)Identification

After removing the initial data, the correlogram is plotted again as shown,

Sample: 1985M01 2018M12

Included observation	is: 407					
Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
		1	0.136	0.136	7.6014	0.006
od (<u> </u>	2	-0.085	-0.106	10.580	0.005
1 1		3	-0.000	0.028	10.580	0.014
ւիլ		4	0.047	0.035	11.506	0.021
1)1	1 1 1	5	0.018	0.009	11.647	0.040
1 1	1 1 1	6	0.005	0.009	11.657	0.070
ı İ zi	<u> </u>	7	0.072	0.074	13.829	0.054
1)1	1 1	8	0.024	0.001	14.060	0.080
1 1	1 1	9	-0.013	-0.004	14.130	0.118
1 1	1 1 1	10	0.009	0.013	14.163	0.166
ı 🗖		11	0.154	0.148	24.152	0.012
ı j i	1 1 1	12	0.030	-0.017	24.534	0.017
1 1		13	-0.001	0.027	24.534	0.027
1) 11		14	0.035	0.027	25.056	0.034
ւիլ		15	0.045	0.027	25.923	0.039
ւիլ		16	0.042	0.034	26.664	0.045
1)1	1 1	17	0.019	0.015	26.820	0.061
ւիլ		18	0.057	0.038	28.209	0.059
1)1	1 1	19	0.021	0.002	28.390	0.076
ı İ zi	• b	20	0.081	0.089	31.181	0.053
1) 11	1 1	21	0.027	-0.002	31.496	0.066
1) 1	1 1	22	0.036	0.020	32.060	0.076
ւիլ		23	0.037	0.025	32.656	0.087
ւիլ		24	0.047	0.037	33.608	0.092
ı (h	1 1	25	0.009	-0.016	33.647	0.116

From the ACF and PACF, there are two significant spikes and then the coefficients die down with random spikes appearing in lag 7, 11 and 20. A number of ARMA processes could result in this pattern. The following reasonable candidate models are verified: ARMA(2,0), ARMA(0,2), ARMA(1,0), ARMA(0,1), and ARMA(1,1).

(iii) Estimation

ARMA(0,1) is tried out through linear regression to see if it's a reasonable candidate.

Sample: 1985M02 2018M12 Included observations: 407

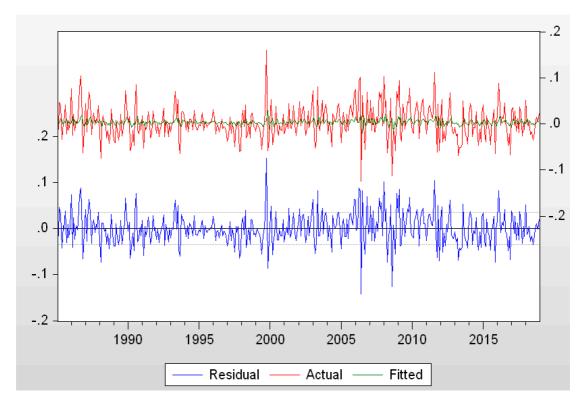
Convergence achieved after 16 iterations
Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	0.003484	0.002086	1.670293	0.0956
MA(1)	0.167650	0.040449	4.144712	0.0000
SIGMASQ	0.001196	6.34E-05	18.84932	0.0000
R-squared	0.022957	Mean depend	ont var	0.003484
•				
Adjusted R-squared	0.018120	S.D. depende		0.035029
S.E. of regression	0.034710	Akaike info cri	terion	-3.876179
Sum squared resid	0.486726	Schwarz criter	rion	-3.846630
Log likelihood	791.8023	Hannan-Quin	n criter.	-3.864485
F-statistic	4.746237	Durbin-Watso	n stat	2.026107
Prob(F-statistic)	0.009174			
Inverted MA Roots	17			

Both its regression coefficients are statistically significant hence the model looks reasonable.

(iv) Testing

Consider the characteristics residuals by plotting it out as shown,



The residuals seem to be random in nature, having same variance. To verify its nature further, correlogram for residuals is plotted and studied,

Sample: 1985M01 2018M12 Included observations: 407

Q-statistic probabilities adjusted for 1 ARMA term

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
П	l di	1	-0.014	-0.014	0.0774	
d :	d ·	2	-0.083	-0.084	2.9417	0.086
1 1	1 1	3	0.006	0.004	2.9572	0.228
ı j ir	I <mark>]</mark> II	4	0.044	0.037	3.7588	0.289
())	1 11	5	0.012	0.015	3.8229	0.431
ų i	1 1	6	-0.009	-0.002	3.8558	0.570
ı þi	יום י	7	0.071	0.073	5.9516	0.429
1)1	1 1	8	0.014	0.013	6.0305	0.536
ų.	1 1	9	-0.013	-0.002	6.0984	0.636
41	10	10	-0.014	-0.013	6.1854	0.721
' 		11	0.155	0.150	16.262	0.092
1 1	1 1	12	0.006	0.005	16.277	0.131
1 1	1 11	13	-0.007	0.020	16.296	0.178
1 11	I	14	0.030	0.028	16.689	0.214
ı j ir	']	15	0.034	0.025	17.181	0.247
i j ii		16	0.035	0.037	17.715	0.278
1 1	1 11	17	0.004	0.014	17.723	0.340
(þ)	ווןי	18	0.056	0.042	19.070	0.325
1 1	1 1	19	-0.001	-0.006	19.071	0.387
ı þ i	'Þ	20	0.079	0.087	21.741	0.297

The ACF plot shows random spikes at certain lags and are independent proving that the time series is stationary.

Q-statistic figures are all above 0.05, hence they are not statistically significant. Other measures like RSS, SBC, AIC and HQ are noted to understand the quality of the fit of model.

RSS	0.487
SBC	-3.847
AIC	-3.88
HQ	-3.86

From the figures, we realize SBC, AIC and HQ being negative indicating reasonable fitness of the model.

To find an alternative model, other reasonable candidate models are tried out and their results are validated as shown,

Figure 1: ARMA(2,0)

Sample: 1985M02 2018M12 Included observations: 407 Convergence achieved after 15 iterations Coefficient covariance computed using outer product of gradients

		•		
Variable	Coefficient	Std. Error	t-Statistic	Prob.
c	0.003473	0.001897	1.830899	0.0679
AR(1)	0.150628	0.039910	3.774163	0.0002
AR(2)	-0.105918	0.043571	-2.430929	0.0155
SIGMASQ	0.001188	6.36E-05	18.67461	0.0000
R-squared	0.029621	Mean dependent var		0.003484
Adjusted R-squared	0.022398	S.D. depende	ent var	0.035029
S.E. of regression	0.034634	Akaike info cr	iterion	-3.878078
Sum squared resid	0.483406	Schwarz crite	rion	-3.838679
Log likelihood	793.1888	Hannan-Quir	ın criter.	-3.862486
F-statistic	4.100590	Durbin-Watson stat		1.992539
Prob(F-statistic)	0.006942			
Inverted AR Roots	.08+.32i	.0832i		

Figure 2 : ARMA(1,1)

Sample: 1985M02 2018M12 Included observations: 407

Convergence achieved after 14 iterations

Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C AR(1) MA(1) SIGMASQ	0.003484 -0.330921 0.490279 0.001190	0.002013 0.215195 0.203645 6.33E-05	1.731338 -1.537772 2.407521 18.80114	0.0842 0.1249 0.0165 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.027818 0.020581 0.034666 0.484304 792.8132 3.843834 0.009822	Mean depend S.D. depende Akaike info cr Schwarz crite Hannan-Quir Durbin-Watso	ent var iterion rion in criter.	0.003484 0.035029 -3.876232 -3.836833 -3.860640 2.011501
Inverted AR Roots Inverted MA Roots	33 49	<u> </u>	<u> </u>	

Sample: 1985M02 2018M12 Included observations: 407

Convergence achieved after 11 iterations

Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
C AR(1) SIGMASQ	0.003486 0.136010 0.001201	0.002101 0.039860 6.43E-05	1.659647 3.412218 18.67665	0.0978 0.0007 0.0000
R-squared Adjusted R-squared S.E. of regression Sum squared resid Log likelihood F-statistic Prob(F-statistic)	0.018565 0.013706 0.034788 0.488914 790.8945 3.820989 0.022702	Mean depend S.D. depende Akaike info cr Schwarz crite Hannan-Quin Durbin-Watso	nt var iterion rion n criter.	0.003484 0.035029 -3.871717 -3.842168 -3.860024 1.969347
Inverted AR Roots	.14			

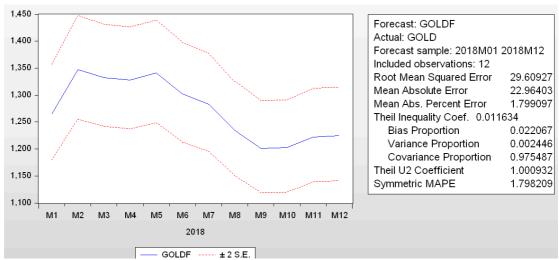
Sample: 1985M02 2018M12 Included observations: 407 Convergence achieved after 16 iterations Coefficient covariance computed using outer product of gradients

Variable	Coefficient	Std. Error	t-Statistic	Prob.
С	0.003479	0.001948	1.786143	0.0748
MA(1)	0.151899	0.040588	3.742491	0.0002
MA(2)	-0.079066	0.043340	-1.824306	0.0688
SIGMASQ	0.001188	6.36E-05	18.67115	0.0000
R-squared	0.029432	Mean dependent var		0.003484
Adjusted R-squared	0.022207	S.D. dependent var		0.035029
S.E. of regression	0.034637	Akaike info criterion		-3.877886
Sum squared resid	0.483500	Schwarz criterion		-3.838487
Log likelihood	793.1497	Hannan-Quinn criter.		-3.862294
F-statistic	4.073607	Durbin-Watson stat		1.996835
Prob(F-statistic)	0.007200			
Inverted MA Roots	.22	37		

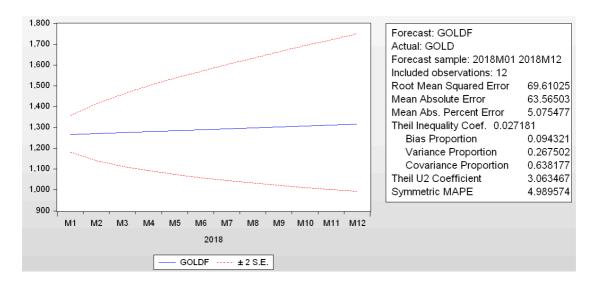
From the four different models, we are able to see that the ARMA(2,0) and ARMA(1,0) are more reasonable models comparable with ARMA(0,1) model since they have regression coefficients which are statistically significant.

(v) Forecasting

The last 12 observations are removed first before forecasting is done. After keying in the equation for the model ARMA(0,1), different forecasting methods are evaluated and the graphs are as shown,



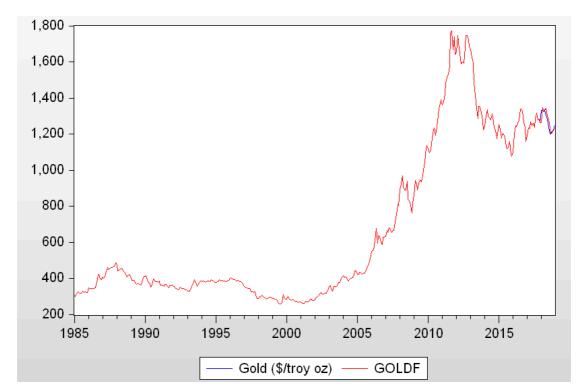
The graph above shows the forecast of gold using static method.



The graph above shows the forecast of gold using dynamic method.

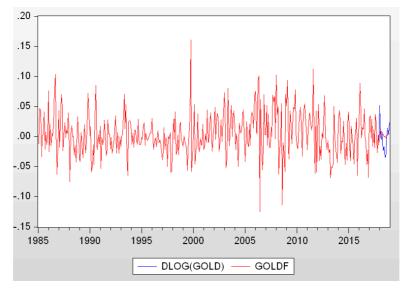
Dynamic produces a 1-step-ahead forecast, a 2-step-ahead forecast, a 3-step-ahead forecast, all the way till 12-step-ahead forecast for a 12 month forecast. By contrast, Static forecasting produces twelve 1 –step ahead forecasts.

Static forecast for gold are plotted to see the fitness of model.



One can observe that the forecast charts the actual time series values of gold, hence it's proves to be a good fit.

Static forecasting is also tried with DLOG(gold) and its fitness is validated.



The forecast for DLOG(gold), GOLDF shows a upward trend around 2018 time period, indicating prices to go up over that year.

Do you expect gold prices to go up or down in the next 12 months? How about risk? What is the prediction interval?

Prediction interval in this case is 12 months, from 2019 month1 to 2019 month12.

Forecasting for DLOG(gold) is done through dynamic forecasting and studied,



Prices of gold are expected to increase as seen from the forecasting model.

Risk of forecast in this case would be unpredictable decrease in gold price when it was expected to increase. An example would be buyinh GOLD commodity stocks assuming their level would appreciate over time, when in fact the stock prices decrease. This could cause them lose money invested.

Static forecasting makes use of actual values to make the next step forecast, which in this case in unavailable. In this scenario, forecast for the next 12 months are to be made based on past data. For this, dynamic forecasting is more appropriate as it previous forecasted values for prediction.