

Multinomial RegressionSummary of the last lecture

- Logistic regression models for multinomial response
- Ungrouped binary/Grouped binary

Key terms of this lecture

- Logistic regression (cont'd)
 - Other link function
 - Confounding and interaction
 - Different study designs

Reading

- McCullagh and Nelder (1989) Chapter 5
 - Dobson and Barnett (2008) Chapter 8
-

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- For the nominal logistic regression, we have two R functions to do this: one is to use `multinom()` in the package "nnet". The other is to use `vglm()` in the package "VGAM" with `family=multinomial`. `multinom()` is preferred, since it uses the first category as the reference and its result is consistent with the model form. However, `vglm()` can do both the nominal and ordinal logistic regression by changing the argument in "family=". Note: `multinom()` uses long form data and `vglm()` uses wide form data. [See Table8_2-VGAM.R]

```
## logistic regression for nominal response;
library(VGAM)
car.nl<-vglm(cbind(B_Imp, C_VeImp, A_NoImp) ~ c_age + c_sex,
             family=multinomial, data=car.wide)
```

- In fact, for the proportional odds logistic regression, we have two R functions to do this: one is to use `polr()` in the package "MASS". The other is to use `vglm()` in the package "VGAM". `vglm()` is preferred, since its result is consistent with the model form and the result from SAS, while `polr()` has a different model form and the signs of the coefficients are opposite from those from `vglm()`. Note: `polr()` uses long form data and `vglm()` uses wide form data. [See Table8_4-VGAM.R]

```
## Cumulative logit for ordinal response;
library(VGAM)
car.cl1 <-vglm(cbind(A_NoImp, B_Imp, C_VeImp) ~ c_age + c_sex,
               family=cumulative(parallel=TRUE), data=car.wide)
```

More Examples for Multi-category Logit Models: Political Ideology and Party Affiliation

In Agresti's 2007 book, Table 6.7, from a General Social Survey, relates political ideology to political party affiliation. Political ideology has a five-point ordinal scale, ranging from very liberal to very conservative.

Gender	Party	Ideology				
		VLib	SLib	Mod	SCon	VCon
Female	Dem	44	47	118	23	32
	Rep	18	28	86	39	48
Male	Dem	36	34	53	18	23
	Rep	12	18	62	45	51

Y = political ideology (very liberal, slightly liberal, moderate, slightly conservative, very conservative)

$x_1 = \text{gender}(1=\text{M}, 0=\text{F})$; $x_2 = \text{political party}(1 = \text{Rep}, 0 = \text{Dem})$

Cumulative Logit Model:

$$\text{logit}[P(Y \leq j)] = \alpha_j + \beta_1 x_1 + \beta_2 x_2, \quad j = 1, 2, 3, 4.$$

In R, we will fit these models using the VGAM package. Another R package ordinal also does the work. Install R package icda to get the data. See R code ideology.R. Note: VGAM treats the last category as the reference.

```
install.packages("icda", repos="http://www.stat.ufl.edu/~priesnell/R", type="source")
> library(VGAM)
> library(icda)
> data(ideology)
> head(ideology)
  Party Gender Ideology Freq
1  Dem Female  VLib    44
2  Rep Female  VLib    18
3  Dem  Male   VLib    36
4  Rep  Male   VLib    12
5  Dem Female  SLib    47
6  Rep Female  SLib    28

> library(reshape2)
> ideov <- dcast(ideology, Gender + Party ~ Ideology,
```

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```
      value.var="Freq")
> ideov
  Gender Party VLib SLib Mod SCon VCon
1 Female  Dem   44  47 118  23  32
2 Female  Rep   18  28  86  39  48
3  Male   Dem   36  34  53  18  23
4  Male   Rep   12  18  62  45  51
> ideov.cll <-
  vglm(cbind(VLib, SLib, Mod, SCon, VCon) ~ Gender + Party,
       family=cumulative(parallel=TRUE), data=ideov)
> summary(ideov.cll)

Call:
vglm(formula = cbind(VLib, SLib, Mod, SCon, VCon) ~ Gender +
      Party, family = cumulative(parallel = TRUE), data = ideov)
```

Coefficients:

	Estimate	Std. Error	z value	Pr(> z)
(Intercept):1	-1.4518	0.1228	-11.818	< 2e-16 ***
(Intercept):2	-0.4583	0.1068	-4.333	1.47e-05 ***
(Intercept):3	1.2550	0.1145	10.955	< 2e-16 ***
(Intercept):4	2.0890	0.1292	16.174	< 2e-16 ***
GenderMale	-0.1169	0.1268	-0.921	0.357
PartyRep	-0.9636	0.1294	-7.449	9.39e-14 ***

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

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Number of linear predictors: 4

Names of linear predictors:

logit(P[Y<=1]), logit(P[Y<=2]), logit(P[Y<=3]), logit(P[Y<=4])

Residual deviance: 15.0556 on 10 degrees of freedom

Log-likelihood: -47.415 on 10 degrees of freedom

```
> deviance(ideo.c11)
[1] 15.056
> df.residual(ideo.c11)
[1] 10
> pchisq(deviance(ideo.c11), df.residual(ideo.c11),
        lower.tail = FALSE)
[1] 0.13005
```

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Political Ideology and Party Affiliation (cont'd)

- Cumulative logit model fit: for $j = 1, 2, 3, 4$. gender party

$$\text{logit}[\hat{P}(Y \leq j)] = \alpha_j + (-0.117)X_1 + (-0.964)X_2$$

- $\alpha_1 = -1.452$, $\alpha_2 = -0.458$, $\alpha_3 = 1.255$, $\alpha_4 = 2.089$
- Controlling for gender, estimated odds that a Rep's response is in liberal direction ($Y \leq j$) rather than conservative ($Y > j$) are $\frac{\exp(-0.964)}{\exp(0.964)} = 0.38$ times estimated odds for a Dem. (0.38136)
- Equivalently: controlling for gender, estimated odds that a Dem's response is in liberal direction ($Y \leq j$) rather than conservative ($Y > j$) are

$\exp(0.964) = 2.62$ times estimated odds for a Rep.

- 95% CI for true odds ratio is $\exp(\hat{\beta} \pm SE(\hat{\beta}) \times 1.96) = (0.30, 0.49)$.
 $= (\exp(-1.21684), \exp(-0.7116)) = (0.296, 0.491)$
- Contingency table data. No evidence of lack of fit: deviance = 15.1, df = 10, p-value = 0.13 (0.30, 0.49)

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- Test for party effect (controlling for gender), i.e., $H_0: \beta_2 = 0$

$$\text{Wald: } z = \frac{\hat{\beta}_2 - 0}{\text{se}(\hat{\beta}_2)} \quad (z^2 = 55.49) = \frac{-0.964}{0.129} = -7.4729$$

$$\text{LR: } D_0 - D_1 = \frac{71.902 - 15.057}{56.846}, \quad df = 11 - 10 = 1$$

$p\text{-value} < 0.0001$, (either test)

Strong evidence that Republicans tend to be less liberal (more conservative than Democrats (for each gender).

- No evidence of gender effect (controlling for party). ($p\text{-value} \approx 0.36$ using either Wald or LR test).

```
> ideo.cl2 <-
  vglm(cbind(VLib,SLib,Mod,SCon,VCon) ~ Gender,
       family=cumulative(parallel=TRUE), data=ideow)
> deviance(ideo.cl2)
[1] 71.902
> df.residual(ideo.cl2)
[1] 11
```

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```
> deviance(ideo.cl2) - deviance(ideo.cl1)
[1] 56.847
> pchisq(deviance(ideo.cl2) - deviance(ideo.cl1),
         df.residual(ideo.cl2) - df.residual(ideo.cl1),
         lower.tail=FALSE)
[1] 4.711e-14
```

- The ordinal logistic regression model can also be fitted by `polr()` in the R package "MASS". Note the opposite signs in coefficients, but the intercept are the same.

```
##Method II: Using polr()
library(MASS)
##Use the long form data
## Variables: Party Gender Ideology Freq
ideo.polr<-polr(factor(Ideology)~factor(Gender)
               +factor(Party), weights=Freq, data=ideology)
```

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```
print(summary(ideo.polr))
```

Call:

```
polr(formula = factor(Ideology) ~ factor(Gender)
      + factor(Party), data = ideology, weights = Freq)
```

Coefficients:

		Value	Std. Error	t value
Compare this to that from vglm(), the signs are opposite.	factor(Gender)Male	0.1169	0.1273	0.9177
	factor(Party)Rep	0.9636	0.1297	7.4311

Intercepts:

		Value	Std. Error	t value
Compare this to that from vglm(), they are the same!	VLib SLib	-1.4518	0.1226	-11.8373
	SLib Mod	-0.4583	0.1048	-4.3746
	Mod SCon	1.2550	0.1142	10.9874
	SCon VCon	2.0890	0.1293	16.1588

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Residual Deviance: 2474.142

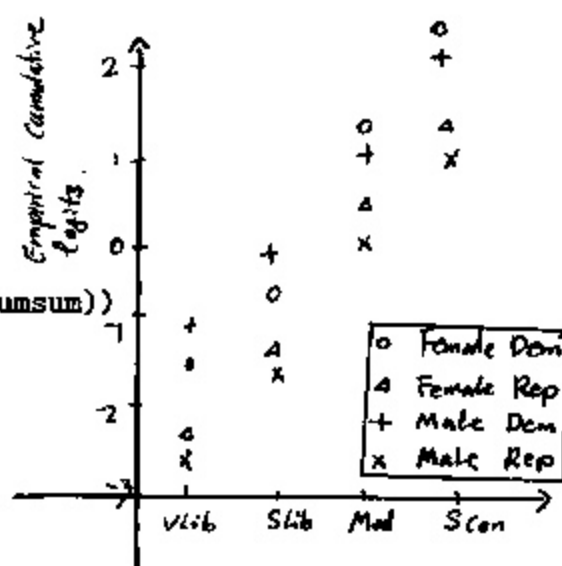
AIC: 2486.142

Political Ideology and Party Affiliation (cont'd): Party-Gender interaction?

```
> ideow
  Gender Party VLib SLib Mod SCon VCon
1 Female  Dem  44  47 118   23   32
2 Female  Rep  18  28  86   39   48
3 Male    Dem  36  34  53   18   23
4 Male    Rep  12  18  62   45   51
```

```
> ideo.csum <- t(apply(ideow[, -(1:2)], 1, cumsum))
```

```
> ideo.csum
      VLib SLib Mod SCon VCon
[1,]  44  91 209 232 264
[2,]  18  46 132 171 219
[3,]  36  70 123 141 164
[4,]  12  30  92 137 188
```



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```
> ideo.cprop <- ideo.csum[,1:4]/ideo.csum[,5]
> ideo.ecl <- qlogis(ideo.cprop) # empirical cumul. logits
> ideo.cl3 <-
+   vglm(cbind(VLib,SLib,Mod,SCon,VCon) ~ Gender*Party,
+       family=cumulative(parallel=TRUE), data=ideow)
> coef(summary(ideo.cl3))
```

	Estimate	Std. Error	z value	Pr(> z)
(Intercept):1	-1.5520853	0.1335312	-11.6233940	3.134323e-31
(Intercept):2	-0.5549908	0.1170310	-4.7422535	2.113539e-06
(Intercept):3	1.1646526	0.1233734	9.4400633	3.725543e-21
(Intercept):4	2.0012144	0.1368228	14.6263266	1.908149e-48
GenderMale	0.1430828	0.1793647	0.7977197	4.250331e-01
PartyRep	-0.7562072	0.1669103	-4.5306197	5.881093e-06
GenderMale:PartyRep	-0.5091332	0.2540825	-2.0038111	4.509030e-02

```
> deviance(ideo.cl3)
[1] 11.06338
> df.residual(ideo.cl3)
```

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```

[1] 9
> deviance(ideo.cl1) - deviance(ideo.cl3)
[1] 3.992192
> pchisq(deviance(ideo.cl1) - deviance(ideo.cl3),
+         df.residual(ideo.cl1) - df.residual(ideo.cl3),
+         lower.tail=FALSE)
[1] 0.04571157

```

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Political Ideology and Party Affiliation (with interaction)

Plot of empirical logits suggests interaction between party and gender. The model is

$$\text{logit}[\hat{P}(Y \leq j)] = \alpha_j + \beta_1 x_1 + \beta_2 x_2 + \beta_3 x_1 x_2, \quad j = 1, 2, 3, 4.$$

- ML fit:

$$\text{logit}[\hat{P}(Y \leq j)] = \alpha_j + \underset{\substack{\text{Gender} \\ \downarrow}}{0.143} x_1 + \underset{\substack{\text{Party} \\ \downarrow}}{1.0756} x_2 + \underset{\substack{\text{Gender} \times \text{Party} \\ \downarrow}}{(-0.509)} x_1 * x_2$$

- Test for party \times gender interaction ($H_0: \beta_3 = 0$):

$$\text{LR: } D_0 - D_1 = \boxed{15.056 - 11.063} = \boxed{3.993}$$

$$df = \boxed{1 (10 - 9)}, \quad p\text{-value} = \boxed{0.0457}$$

- Some evidence (significant at 0.05 level) that effect of Party varies with Gender (and vice versa).

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Political Ideology and Party Affiliation (with interaction)

- Estimated odds ratio for party effect (x_2) is

$$e^{-0.756} = 0.47 \quad \text{when } x_1 = 0 \text{ (F)}$$

$$e^{-0.756-0.509} = e^{-1.265} = 0.28 \quad \text{when } x_1 = 1 \text{ (M)}$$

Implying: Among males and females, Reps are more conservative than Dems.

- Estimated odds ratio for gender effect (x_1) is

$$e^{0.143} = 1.15 \quad \text{when } x_2 = 0 \text{ (Dem)}$$

$$e^{0.143-0.509} = e^{-0.336} = 0.69 \quad \text{when } x_2 = 1 \text{ (Rep)}$$

- Among Dems, males tend to be more liberal than females. Among Reps, males tend to be more conservative than females.

$$OR < 1$$

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Political Ideology and Party Affiliation (with interaction) (cont'd)

- Probability of very liberal for male and female Republicans, that is $\hat{P}(Y = 1)$ when $x_1 = 1$ or 0 and $x_2 = 1$.

- Notice $\hat{P}(Y \leq j) = \frac{\exp(\hat{\alpha}_j + 0.143x_1 - 0.756x_2 - 0.509x_1x_2)}{1 + \exp(\hat{\alpha}_j + 0.143x_1 - 0.756x_2 - 0.509x_1x_2)}$. For $j = 1$, $\hat{\alpha}_1 = -1.55$.

- Male Republicans ($x_1 = 1, x_2 = 1$):

$$\hat{P}(Y = 1) = \frac{\exp(-1.55 + 0.143 - 0.756 - 0.509)}{1 + \exp(-1.55 + 0.143 - 0.756 - 0.509)} = \frac{\exp(-2.67)}{1 + \exp(-2.67)} = 0.065.$$

- Female Republicans ($x_1 = 0, x_2 = 1$):

$$\hat{P}(Y = 1) = \frac{\exp(-1.55 - 0.756)}{1 + \exp(-1.55 - 0.756)} = \frac{\exp(-2.31)}{1 + \exp(-2.31)} = 0.090.$$

- Similarly, $\hat{P}(Y = 2) = \hat{P}(Y \leq 2) - \hat{P}(Y \leq 1)$, etc. Note:

$$\hat{P}(Y = 5) = \hat{P}(Y \leq 5) - \hat{P}(Y \leq 4) = 1 - \hat{P}(Y \leq 4).$$

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Remarks

- Reversing order of response categories changes signs of "slope" estimates (cumulative odds ratio \rightarrow 1/cumulative odds ratio). To show this, let $Z = J - Y + 1$, $k = J - j$, when $j = 1, \dots, J - 1$, $k = J - 1, \dots, 1$, then, for example,

$$\text{logit}[P(Z \leq k)] = -\text{logit}[P(Y \leq j)] = -\alpha_{J-k} - \beta_1 x_1 - \beta_2 x_2.$$

- For ordinal response, only two sensible orderings, either $1, 2, \dots, J$ or $J, J - 1, \dots, 1$