

Homework 3

Questions: <https://github.com/SUSTech-ML-Course/Machine-Learning-Course/blob/05b4e0bd4063b9d486f34ce557720b519947cbe3/lesson-homework/hw3/HW3.md>

Question 1

Compute the derivative with respect to \mathbf{w}

$$\nabla E_D(\mathbf{w}) = \sum_{n=1}^N r_n [t_n - \mathbf{w}^T \phi(\mathbf{x}_n)] \phi(\mathbf{x}_n)^T$$

Set the derivative to 0, we have

$$\sum_{n=1}^N r_n t_n \phi(\mathbf{x}_n)^T = \mathbf{w}^T \sum_{n=1}^N r_n \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T$$

If we denote $\phi'(\mathbf{x}_n) = \sqrt{r_n} \phi(\mathbf{x}_n)$ and $t'_n = \sqrt{r_n} t_n$, we can get the same equation (3.14) in the PRML textbook

$$\sum_{n=1}^N t'_n \phi'(\mathbf{x}_n)^T = \mathbf{w}^T \sum_{n=1}^N \phi'(\mathbf{x}_n) \phi'(\mathbf{x}_n)^T$$

Then we can get the solution

$$\mathbf{w}^* = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{t}$$

where

- $\mathbf{t} = [\sqrt{r_1} t_1, \sqrt{r_2} t_2, \dots, \sqrt{r_n} t_n]^T$
- Φ is an $N \times M$ **design matrix**, whose elements are given by $\Phi_{ij} = \sqrt{r_i} \phi_j(\mathbf{x}_i)$

$$\Phi = \begin{bmatrix} \sqrt{r_1} \phi_0(\mathbf{x}_1) & \sqrt{r_1} \phi_1(\mathbf{x}_1) & \cdots & \sqrt{r_1} \phi_{M-1}(\mathbf{x}_1) \\ \sqrt{r_2} \phi_0(\mathbf{x}_2) & \sqrt{r_2} \phi_1(\mathbf{x}_2) & \cdots & \sqrt{r_2} \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \sqrt{r_N} \phi_0(\mathbf{x}_N) & \sqrt{r_N} \phi_1(\mathbf{x}_N) & \cdots & \sqrt{r_N} \phi_{M-1}(\mathbf{x}_N) \end{bmatrix}$$

The two alternative interpretation:

- If we substitute β^{-1} by $r_n \beta^{-1}$ in the summation term, the equation in the question stem will become the expression above
- r_n can also be viewed as the effective number of observation of (\mathbf{x}_n, t_n)

Question 2

$$p(\mathbf{w}, \beta) = \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \beta^{-1} \mathbf{S}_0) \text{Gam}(\beta | a_0, b_0) \\ \propto \left(\frac{\beta}{|\mathbf{S}_0|} \right)^2 \exp\left(-\frac{1}{2}(\mathbf{w} - \mathbf{m}_0)^T \beta \mathbf{S}_0^{-1} (\mathbf{w} - \mathbf{m}_0)\right) b_0^{a_0} \beta^{a_0-1} \exp(-b_0 \beta) \quad (i)$$

$$p(\mathbf{t} | \mathbf{X}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n | \mathbf{w}^T \phi(\mathbf{x}_n), \beta^{-1}) \\ \propto \prod_{n=1}^N \beta^{1/2} \exp\left[-\frac{\beta}{2}(t_n - \mathbf{w}^T \phi(\mathbf{x}_n))^2\right] \quad (ii)$$

According to Bayesian Inference, we have $p(\mathbf{w}, \beta | \mathbf{t}) \propto p(\mathbf{t} | \mathbf{X}, \mathbf{w}, \beta) p(\mathbf{w}, \beta)$

- The **quadratic term** in the exponent is

$$\text{quadratic term} = -\frac{\beta}{2} \mathbf{w}^T \mathbf{S}_0^{-1} \mathbf{w} + \sum_{n=1}^N -\frac{\beta}{2} \mathbf{w}^T \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T \mathbf{w} \\ = -\frac{\beta}{2} \mathbf{w}^T [\mathbf{S}_0^{-1} + \sum_{n=1}^N \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T] \mathbf{w}$$

where the first term is from (i) and the second from (ii), then we can get

$$\mathbf{S}_N = \left[\mathbf{S}_0^{-1} + \sum_{n=1}^N \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T \right]^{-1} \quad (1)$$

- The **linear term** in the exponent is

$$\text{linear term} = \beta \mathbf{m}_0^T \mathbf{S}_0^{-1} \mathbf{w} + \sum_{n=1}^N \beta t_n \phi(\mathbf{x}_n)^T \mathbf{w} \\ = \beta \left[\mathbf{m}_0^T \mathbf{S}_0^{-1} + \sum_{n=1}^N t_n \phi(\mathbf{x}_n)^T \right] \mathbf{w}$$

Similarly, the first term is from (i) and the second from (ii), then we can also get

$$\mathbf{m}_N^T \mathbf{S}_N^{-1} = \mathbf{m}_0^T \mathbf{S}_0^{-1} + \sum_{n=1}^N t_n \phi(\mathbf{x}_n)^T \\ \Rightarrow \mathbf{m}_N = \mathbf{S}_N \left[\mathbf{S}_0^{-1} \mathbf{m}_0 + \sum_{n=1}^N t_n \phi(\mathbf{x}_n) \right] \quad (2)$$

- The **constant term** in the exponent is

$$\text{constant term} = \left(-\frac{\beta}{2} \mathbf{m}_0^T \mathbf{S}_0^{-1} \mathbf{m}_0 - b_0 \beta\right) - \frac{\beta}{2} \sum_{n=1}^N t_n^2 \\ = -\beta \left[\frac{1}{2} \mathbf{m}_0^T \mathbf{S}_0^{-1} \mathbf{m}_0 + b_0 + \frac{1}{2} \sum_{n=1}^N t_n^2 \right]$$

From which we can obtain

$$\begin{aligned}\frac{1}{2}\mathbf{m}_N^T \mathbf{S}_N^{-1} \mathbf{m}_N + b_N &= \frac{1}{2}\mathbf{m}_0^T \mathbf{S}_0^{-1} \mathbf{m}_0 + b_0 + \frac{1}{2} \sum_{n=1}^N t_n^2 \\ \Rightarrow b_N &= \frac{1}{2}\mathbf{m}_0^T \mathbf{S}_0^{-1} \mathbf{m}_0 + b_0 + \frac{1}{2} \sum_{n=1}^N t_n^2 - \frac{1}{2}\mathbf{m}_N^T \mathbf{S}_N^{-1} \mathbf{m}_N\end{aligned}\quad (3)$$

- The **exponential term** based on β is

$$\begin{aligned}\text{exponent term} &= (2 + a_0 - 1) + \frac{N}{2} \\ &= 2 + a_N - 1 \\ \Rightarrow a_N &= a_0 + \frac{N}{2}\end{aligned}\quad (4)$$

Question 3

According to the standard form of a multivariate normal distribution, we know that

$$\begin{aligned}\int \frac{1}{(2\pi)^{M/2}} \frac{1}{|\mathbf{A}|^{1/2}} \exp\left[-\frac{1}{2}(\mathbf{w} - \mathbf{m}_N)^T \mathbf{A}(\mathbf{w} - \mathbf{m}_N)\right] d\mathbf{w} &= 1 \\ \Rightarrow \int \exp\left[-\frac{1}{2}(\mathbf{w} - \mathbf{m}_N)^T \mathbf{A}(\mathbf{w} - \mathbf{m}_N)\right] d\mathbf{w} &= (2\pi)^{M/2} |\mathbf{A}|^{1/2}\end{aligned}$$

Since $E(\mathbf{m}_N)$ doesn't depend on \mathbf{w} , we can get the equation

$$\int \exp[-E(\mathbf{w})] d\mathbf{w} = \exp[-E(\mathbf{m}_N)] (2\pi)^{M/2} |\mathbf{A}|^{-1/2}$$

Then we substitute it into the equation (3.78) in PRML textbook

$$p(\mathbf{t}|\alpha, \beta) = \left(\frac{\beta}{2\pi}\right)^{N/2} \left(\frac{\alpha}{2\pi}\right)^{M/2} \int \exp[-E(\mathbf{w})] d\mathbf{w} \quad (3.78)$$

Finally we can get

$$\ln p(\mathbf{t}|\alpha, \beta) = \frac{M}{2} \ln \alpha + \frac{N}{2} \ln \beta - E(\mathbf{m}_N) - \frac{1}{2} \ln |\mathbf{A}| - \frac{N}{2} \ln(2\pi)$$

Question 4

$$\begin{aligned}F(a) &= \frac{1}{2} \sum_i (Y_i - aX_i)^2 \\ \frac{\partial F(a)}{\partial a} &= \sum_i \hat{a} X_i^2 - X_i Y_i = 0 \\ \hat{a} &= \frac{\sum_i X_i Y_i}{\sum_i X_i^2}\end{aligned}$$

Question 5

$$\begin{aligned}
\text{Likelihood } L(\theta|y_1, y_2, \dots, y_n) &= P(X = x_1|\theta)P(X = x_2|\theta) \cdots P(X = x_n|\theta) \\
&= e^{-\theta} \frac{\theta^{y_1}}{y_1!} \cdots e^{-\theta} \frac{\theta^{y_n}}{y_n!} \\
&= e^{-n\theta} \frac{\theta^{y_1+y_2+\dots+y_n}}{y_1!y_2! \cdots y_n!} \\
&= e^{-n\theta} \frac{\theta^{\sum_{i=1}^n y_i}}{\prod_{i=1}^n y_i!}
\end{aligned}$$

$$\text{Log-likelihood } \ln L(\theta|y_1, y_2, \dots, y_n) = -n\theta + \left(\sum_{i=1}^n y_i \right) \ln \theta - \ln \left(\prod_{i=1}^n y_i! \right)$$

Then we can get the MLE of θ

$$\begin{aligned}
\frac{\partial \ln L(\theta|y_1, y_2, \dots, y_n)}{\partial \theta} &= -n + \frac{\sum_{i=1}^n y_i}{\hat{\theta}} = 0 \\
\hat{\theta} &= \frac{\sum_{i=1}^n y_i}{n}
\end{aligned}$$

Question 6

$$\begin{aligned}
\text{Likelihood } L(\mathbf{X}|\alpha, \lambda) &= \prod_{i=1}^n f(X_i|\alpha, \lambda) \\
&= \left(\frac{\lambda^\alpha}{\Gamma(\alpha)} \right)^n \prod_{i=1}^n X_i^{\alpha-1} e^{-\lambda X_i}
\end{aligned}$$

$$\text{Log-likelihood } \ln L(\mathbf{X}|\alpha, \lambda) = (\alpha - 1) \sum_{i=1}^n \ln X_i - \lambda \sum_{i=1}^n X_i + n\alpha \ln \lambda - n \ln \Gamma(\alpha)$$

$$\begin{aligned}
\frac{\partial \ln L(\mathbf{X}|\alpha, \lambda)}{\partial \lambda} &= - \sum_{i=1}^n X_i + \frac{n\alpha}{\hat{\lambda}} = 0 \\
\Rightarrow \hat{\lambda} &= \frac{n\alpha}{\sum_{i=1}^n X_i} = \frac{\alpha}{\bar{X}}
\end{aligned}$$