Homework 3

Questions: https://github.com/SUSTech-ML-Course/Machine-Learning-Course/blob/05b4e0bd406 3b9d486f34ce557720b519947cbe3/lesson-homework/hw3/HW3.md

Question 1

Compute the derivative with respect to $oldsymbol{w}$

$$abla E_D(oldsymbol{w}) = \sum_{n=1}^N r_n [t_n - oldsymbol{w}^T \phi(oldsymbol{x}_n)] \phi(oldsymbol{x}_n)^T$$

Set the derivative to 0, we have

$$\sum_{n=1}^N r_n t_n \phi(oldsymbol{x}_n)^T = oldsymbol{w}^T \sum_{n=1}^N r_n \phi(oldsymbol{x}_n) \phi(oldsymbol{x}_n)^T$$

If we denote $\phi'(\boldsymbol{x_n})=\sqrt{r_n}\phi(\boldsymbol{x_n})$ and $t'_n=\sqrt{r_n}t_n$, we can get the same equation (3.14) in the PRML textbook

$$\sum_{n=1}^N t_n' \phi'(oldsymbol{x}_n)^T = oldsymbol{w}^T \sum_{n=1}^N \phi'(oldsymbol{x}_n) \phi'(oldsymbol{x}_n)^T$$

Then we can get the solution

$$oldsymbol{w}^* = (\Phi^T \Phi)^{-1} \Phi^T \mathbf{t}$$

where

- $ullet \ oldsymbol{t} = [\sqrt{r_1}t_1, \sqrt{r_2}t_2, \ldots, \sqrt{r_n}t_n]^T$
- ullet Φ is an $N{ imes}M$ design matrix, whose elements are given by $\Phi_{ij}=\sqrt{r_i}\phi_j(m{x_i})$

$$\Phi = egin{bmatrix} \sqrt{r_1}\phi_0(oldsymbol{x_1}) & \sqrt{r_1}\phi_1(oldsymbol{x_1}) & \cdots & \sqrt{r_1}\phi_{M-1}(oldsymbol{x_1}) \ \sqrt{r_2}\phi_0(oldsymbol{x_2}) & \sqrt{r_2}\phi_1(oldsymbol{x_2}) & \cdots & \sqrt{r_2}\phi_{M-1}(oldsymbol{x_2}) \ dots & dots & dots & dots \ \sqrt{r_N}\phi_0(oldsymbol{x_N}) & \sqrt{r_N}\phi_1(oldsymbol{x_N}) & \cdots & \sqrt{r_N}\phi_{M-1}(oldsymbol{x_N}) \end{bmatrix}$$

The two alternative interpretation:

- (i) If we substitute β^{-1} by $r_n\beta^{-1}$ in the summation term, the equation in the question stem will become the expression above
- (ii) r_n can also be viewed as the effective number of observation of $(oldsymbol{x_n},t_n)$

Question 2

$$p(\boldsymbol{w}, \boldsymbol{\beta}) = \mathcal{N}(\boldsymbol{w}|\boldsymbol{m}_0, \boldsymbol{\beta}^{-1}\mathbf{S}_0) \operatorname{Gam}(\boldsymbol{\beta}|\boldsymbol{a}_0, \boldsymbol{b}_0)$$

$$\propto (\frac{\beta}{|\mathbf{S}_0|})^2 exp(-\frac{1}{2}(\boldsymbol{w} - \boldsymbol{m}_0)^T \boldsymbol{\beta} \mathbf{S}_0^{-1}(\boldsymbol{w} - \boldsymbol{m}_0)) b_0^{a_0} \boldsymbol{\beta}^{a_0 - 1} exp(-b_0 \boldsymbol{\beta})$$
 (i)

$$egin{align} p(\mathbf{t}|\mathbf{X},oldsymbol{w},oldsymbol{eta}) &= \prod_{n=1}^N \mathcal{N}(t_n|oldsymbol{w}^Toldsymbol{\phi}(oldsymbol{x}_n),oldsymbol{eta}^{-1}) \ &\propto \prod_{n=1}^N eta^{1/2} expig[-rac{eta}{2}(t_n-oldsymbol{w}^Toldsymbol{\phi}(oldsymbol{x}_n))^2ig] \end{align}$$

According to Bayesian Inference, we have $p({m w}, eta | {f t}) \propto p({f t} | {f X}, {m w}, eta) p({m w}, eta)$

• The quadratic term in the exponent is

$$\begin{aligned} \text{quadratic term} &= -\frac{\beta}{2} \boldsymbol{w}^T \mathbf{S_0}^{-1} \boldsymbol{w} + \sum_{n=1}^N -\frac{\beta}{2} \boldsymbol{w}^T \boldsymbol{\phi}(\boldsymbol{x_n}) \boldsymbol{\phi}(\boldsymbol{x_n})^T \boldsymbol{w} \\ &= -\frac{\beta}{2} \boldsymbol{w}^T [\boldsymbol{S_0}^{-1} + \sum_{n=1}^N \boldsymbol{\phi}(\boldsymbol{x_n}) \boldsymbol{\phi}(\boldsymbol{x_n})^T] \boldsymbol{w} \end{aligned}$$

where the first term is from (i) and the second from (ii), then we can get

$$oldsymbol{S_N} = \left[oldsymbol{S_0}^{-1} + \sum_{n=1}^N \phi(oldsymbol{x_n}) \phi(oldsymbol{x_n})^T
ight]^{-1}$$
 (1)

• The **linear term** in the exponent is

$$\begin{aligned} \text{linear term} &= \beta \boldsymbol{m_0^T} \boldsymbol{S_0^{-1}} \boldsymbol{w} + \sum_{n=1}^N \beta t_n \phi(\boldsymbol{x_n})^T \boldsymbol{w} \\ &= \beta \left[\boldsymbol{m_0^T} \mathbf{S_0}^{-1} + \sum_{n=1}^N t_n \phi(\boldsymbol{x_n})^T \right] \boldsymbol{w} \end{aligned}$$

Similarly, the first term is from (i) and the second from (ii), then we can also get

$$egin{aligned} oldsymbol{m}_N{}^T oldsymbol{S}_N{}^{-1} &= oldsymbol{m}_0{}^T oldsymbol{S}_0{}^{-1} + \sum_{n=1}^N t_n \phi(oldsymbol{x}_n)^T \ &\Rightarrow oldsymbol{m}_N &= oldsymbol{S}_N \left[oldsymbol{S}_0{}^{-1} oldsymbol{m}_0 + \sum_{n=1}^N t_n \phi(oldsymbol{x}_n)
ight] \end{aligned}$$

• The constant term in the exponent is

$$\begin{aligned} \text{constant term} &= (-\frac{\beta}{2} \boldsymbol{m}_{\boldsymbol{0}}^T \boldsymbol{S}_{\boldsymbol{0}}^{-1} \boldsymbol{m}_{\boldsymbol{0}} - b_0 \beta) - \frac{\beta}{2} \sum_{n=1}^N t_n^2 \\ &= -\beta \left[\frac{1}{2} \boldsymbol{m}_{\boldsymbol{0}}^T \boldsymbol{S}_{\boldsymbol{0}}^{-1} \boldsymbol{m}_{\boldsymbol{0}} + b_0 + \frac{1}{2} \sum_{n=1}^N t_n^2 \right] \end{aligned}$$

From which we can obtain

$$\frac{1}{2}\boldsymbol{m}_{N}{}^{T}\boldsymbol{S}_{N}{}^{-1}\boldsymbol{m}_{N} + b_{N} = \frac{1}{2}\boldsymbol{m}_{0}{}^{T}\boldsymbol{S}_{0}{}^{-1}\boldsymbol{m}_{0} + b_{0} + \frac{1}{2}\sum_{n=1}^{N}t_{n}^{2}$$

$$\Rightarrow b_{N} = \frac{1}{2}\boldsymbol{m}_{0}^{T}\boldsymbol{S}_{0}{}^{-1}\boldsymbol{m}_{0} + b_{0} + \frac{1}{2}\sum_{n=1}^{N}t_{n}^{2} - \frac{1}{2}\boldsymbol{m}_{N}^{T}\boldsymbol{S}_{N}{}^{-1}\boldsymbol{m}_{N} \qquad (3)$$

• The **exponential term** based on β is

Question 3

According to the standard form of a multivariate normal distribution, we know that

$$\int rac{1}{(2\pi)^{M/2}} rac{1}{|\mathbf{A}|^{1/2}} exp\left[-rac{1}{2}(oldsymbol{w} - oldsymbol{m}_N)^T \mathbf{A}(oldsymbol{w} - oldsymbol{m}_N)
ight] doldsymbol{w} = 1 \ \Rightarrow \int exp\left[-rac{1}{2}(oldsymbol{w} - oldsymbol{m}_N)^T \mathbf{A}(oldsymbol{w} - oldsymbol{m}_N)
ight] doldsymbol{w} = (2\pi)^{M/2} |\mathbf{A}|^{1/2}$$

Since $E(oldsymbol{m_N})$ doesn't depend on $oldsymbol{w}$, we can get the equation

$$\int \exp[-E(\mathbf{w})]\mathrm{d}\mathbf{w} = \exp[-E(\mathbf{m}_N)](2\pi)^{M/2}|\mathbf{A}|^{-1/2}$$

Then we substitute it into the equation (3.78) in PRML textbook

$$p(\mathbf{t}|\alpha, \beta) = \left(\frac{\beta}{2\pi}\right)^{N/2} \left(\frac{\alpha}{2\pi}\right)^{M/2} \int \exp\left[-E(\mathbf{w})\right] d\mathbf{w}$$
 (3.78)

Finally we can get

$$\ln p(\mathbf{t}|lpha,eta) = rac{M}{2} \ln lpha + rac{N}{2} \ln eta - E(\mathbf{m}_N) - rac{1}{2} \ln |\mathbf{A}| - rac{N}{2} \ln (2\pi)$$

Question 4

$$F(a) = rac{1}{2} \sum_i (Y_i - aX_i)^2 \ rac{\partial F(a)}{\partial a} = \sum_i \hat{a}X_i^2 - X_iY_i = 0 \ \hat{a} = rac{\sum_i X_iY_i}{\sum_i X_i^2}$$

Question 5

Likelihood
$$L(\theta|y_1, y_2, \dots, y_n) = P(X = x_1|\theta)P(X = x_2|\theta) \cdots P(X = x_n|\theta)$$

$$= e^{-\theta} \frac{\theta^{y_1}}{y_1!} \cdots e^{-\theta} \frac{\theta^{y_n}}{y_n!}$$

$$= e^{-n\theta} \frac{\theta^{y_1+y_2+\dots+y_n}}{y_1!y_2! \cdots y_n!}$$

$$= e^{-n\theta} \frac{\theta^{\sum_{i=1}^n y_i}}{\prod_{i=1}^n y_i!}$$

$$ext{Log-likelihood} \quad lnL(heta|y_1,y_2,\ldots,y_n) = -n heta + \left(\sum_{i=1}^n y_i
ight) \ln heta - \ln(\prod_{i=1}^n y_i!)$$

Then we can get the MLE of heta

$$\frac{\partial lnL(\theta|y_1, y_2, \dots, y_n)}{\partial \theta} = -n + \frac{\sum_{i=1}^n y_i}{\hat{\theta}} = 0$$

$$\hat{\theta} = \frac{\sum_{i=1}^n y_i}{n}$$

Question 6

$$\begin{array}{ll} \text{Likelihood} & L(\boldsymbol{X}|\alpha,\lambda) = \prod_{i=1}^n f(X_i|\alpha,\lambda) \\ & = \left(\frac{\lambda^\alpha}{\Gamma(\alpha)}\right)^n \prod_{i=1}^n X_i^{\alpha-1} e^{-\lambda X_i} \end{array}$$

$$\text{Log-likelihood} \quad lnL(\boldsymbol{X}|\alpha,\lambda) = (\alpha-1)\sum_{i=1}^n lnX_i - \lambda \sum_{i=1}^n X_i + n\alpha ln\lambda - nln\Gamma(\alpha)$$

$$egin{aligned} rac{\partial lnL(m{X}|lpha,\lambda)}{\partial \lambda} &= -\sum_{i=1}^n X_i + rac{nlpha}{\hat{\lambda}} = 0 \ \\ &\Rightarrow \quad \hat{\lambda} &= rac{nlpha}{\sum_{i=1}^n X_i} = rac{lpha}{ar{X}} \end{aligned}$$