Econometricks: Short guides to econometrics

Trick 07: The Generalized Method of Moments

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- 6. Asymptotic efficiency

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Minimize the quadratic form

The overidentified GMM estimator $\hat{\theta}_{GMM}(W_n)$ for K parameters in θ identified by L > K moment conditions is a function of the weighting matrix W_n for a sample of i = 1, ..., n observations:

$$\hat{ heta}_{GMM}(W_n) = \min_{ heta} q_n(heta),$$

where the quadratic form $q_n(\theta)$ is the criterion function and is given as a function of the sample moments $\bar{m}_n(\theta)$

$$q_n(\theta) = \bar{m}_n(\theta)' W \bar{m}_n(\theta).$$

The sample moments are a function

$$\bar{m}_n(\theta) = 1/n \sum_{i=1}^N m(X_i, Z_i, \theta_0)$$

of the model variables X_i , the instruments Z_i , and the true parameters θ_0 .

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What are the properties of the quadratic form

$$q_{n}(\theta) = \bar{m}_{n}(\theta)' \underset{l \times l}{W} \bar{m}_{n}(\theta).$$

Quadratic form criterion function $q_n(\theta) \geq 0$ is a scalar!

Weighting matrix W is symmetric (and positive definite that is x'Wx > 0 for all non-zero x)!

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Get an approximate deviation from the true θ_0

First order Taylor expansion of sample moments $\bar{m}_n(\hat{\theta}_{GMM})$ around $\bar{m}_n(\theta_0)$ at true parameters gives:

$$ar{m}_n(\hat{ heta}_{GMM}) pprox ar{m}_n(heta_0) + ar{G}_n(ar{ heta})(\hat{ heta}_{GMM} - heta_0),$$

where $\bar{G}_n(\bar{\theta}) = \frac{\partial \bar{m}_n(\bar{\theta})}{\partial \bar{\theta}'}$ and $\bar{\theta}$ is a point between $\hat{\theta}_{GMM}$ and θ_0 .

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Check the dimensions

First order Taylor expansion of sample moments $\bar{m}_n(\hat{\theta}_{GMM})$ around $\bar{m}_n(\theta_0)$ at true parameters gives:

$$ar{m}_n(\hat{ heta}_{GMM}) pprox ar{m}_n(heta_0) + ar{G}_n(ar{ heta})(\hat{ heta}_{GMM} - heta_0),$$

where $\bar{G}_n(\bar{\theta})=\frac{\frac{\partial \bar{m}_n(\bar{\theta})}{L\times 1}}{\frac{L\times 1}{\partial \bar{\theta}^i}}$ and $\bar{\theta}$ is a point between $\hat{\theta}_{GMM}$ and θ_0 , because of the

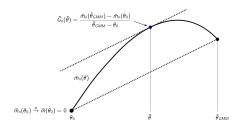
Mean value theorem...

Approximation introduced $\bar{ heta}$

...where $\bar{G}_n(\bar{\theta}) = \frac{\partial \bar{m}_n(\bar{\theta})}{\partial \bar{\theta}'}$ and $\bar{\theta}$ is a point between $\hat{\theta}_{GMM}$ and θ_0 .

Mean value theorem:

$$ar{G}_n(ar{ heta}) = rac{ar{m}_n(\hat{ heta}_{GMM}) - ar{m}_n(heta_0)}{\hat{ heta}_{GMM} - heta_0} ext{ for } heta_0 < ar{ heta} < \hat{ heta}_{GMM}.$$



Do the minimization

To minimize the quadratic form criterion, we take the first derivative of

$$q_n(\theta) = \bar{m}_n(\theta)' W \bar{m}_n(\theta)$$

$$\frac{\partial q_n(\hat{\theta}_{GMM})}{\partial \hat{\theta}_{GMM}} = 2\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{m}_n(\hat{\theta}_{GMM}) = 0.$$

Express as much as possible asymptotically

$$\frac{\partial q_n(\hat{\theta}_{GMM})}{\partial \hat{\theta}_{GMM}} = 2\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{m}_n(\hat{\theta}_{GMM}) = 0,$$

Plug in the approximation from before

$$ar{m}_n(\hat{ heta}_{GMM})pproxar{m}_n(heta_0)+ar{G}_n(ar{ heta})(\hat{ heta}_{GMM}- heta_0)$$

to obtain

$$\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{m}_n(\theta_0) + \bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{G}_n(\bar{\theta})(\hat{\theta}_{GMM} - \theta_0) \approx 0$$

which we rearrange to get the very useful

$$\hat{\theta}_{GMM} \approx \theta_0 - (\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{G}_n(\bar{\theta}))^{-1}\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{m}_n(\theta_0).$$

So the estimate $\hat{\theta}_{GMM}$ is approximately the true parameter θ_0 plus an sampling error that depends on the sample moment $\bar{m}_n(\theta_0)$.

Quickly check dimensions

Useful approximation

$$\hat{\theta}_{\underset{K\times 1}{GMM}} \approx \theta_0 - (\bar{G}_n(\hat{\theta}_{\underset{K\times L}{GMM}})' \underset{\underset{L\times L}{W_n}}{W_n} \bar{G}_n(\bar{\theta}))^{-1} \bar{G}_n(\hat{\theta}_{\underset{K\times L}{GMM}})' \underset{\underset{L\times L}{W_n}}{W_n} \bar{m}_n(\theta_0).$$

So the estimate $\hat{\theta}_{GMM}$ is approximately the true parameter θ_0 plus an sampling error that depends on the sample moment $\bar{m}_n(\theta_0)$.

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Three assumptions: moment conditions

Definition

GMM1: Moment Conditions and Identification.

$$\bar{m}(\theta_a) \neq \bar{m}(\theta_0) = E[m(X_i, Z_i, \theta_0)] = 0.$$

Identification implies that the probability limit of the GMM criterion function is uniquely minimized at the true parameters.

Three assumptions: law of large numbers

Definition

GMM2: Law of Large Numbers Applies.

$$\bar{m}_n(\theta) = 1/n \sum_{i=1}^N m(X_i, Z_i, \theta_0) \stackrel{p}{\rightarrow} E[m(X_i, Z_i, \theta_0)].$$

The data meets the conditions for a law of large numbers to apply, so that we may assume that the empirical moments converge in probability to their expectation. Three assumptions: central limit theorem

Definition

GMM3: Central Limit Theorem Applies.

$$\sqrt{n}\bar{m}_n(\theta) = \sqrt{n}/n\sum_{i=1}^N m(X_i, Z_i, \theta_0) \stackrel{d}{\rightarrow} N[0, \Phi].$$

The empirical moments obey a central limit theorem. This assumes that the moments have a finite asymptotic covariance matrix.

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Consistency

Recall the useful approximation of the estimator:

$$\hat{\theta}_{GMM} \approx \theta_0 - (\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{G}_n(\bar{\theta}))^{-1}\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{m}_n(\theta_0).$$
 Assumption GMM2 implies that

$$\bar{m}_n(\theta_0) = 1/n \sum_{i=1}^N m(X_i, Z_i, \theta_0) \stackrel{p}{\rightarrow} E[m(X_i, Z_i, \theta_0)] = \bar{m}(\theta_0).$$

That is, the sample moment equals the population moment in probability. Assumption GMM1 implies that

$$\bar{m}(\theta_0)=0.$$

Then

$$\bar{m}_n(\theta_0) \stackrel{p}{\to} \bar{m}(\theta_0) = 0,$$

such that

$$\hat{ heta}_{\textit{GMM}} \overset{\textit{p}}{ o} heta_0 \text{ for } extstyle N o \infty$$

That is, by GMM1 and GMM2 the GMM estimator is consistent.

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Asymptotic normality

Recall the useful approximation of the estimator:

$$\hat{\theta}_{GMM} \approx \theta_0 - (\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{G}_n(\bar{\theta}))^{-1}\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{m}_n(\theta_0).$$

Rewrite to obtain

$$\sqrt{n}(\hat{\theta}_{GMM} - \theta_0) \approx -(\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{G}_n(\bar{\theta}))^{-1}\bar{G}_n(\hat{\theta}_{GMM})'W_n\sqrt{n}\bar{m}_n(\theta_0),$$

The right hand side has several parts for which we made assumptions on what happens when $N \to \infty$. Under the central limit theorem (GMM3)

$$\sqrt{n}\bar{m}_n(\theta_0)\stackrel{d}{ o} N[0,\Phi]$$

$$plimW_n = W$$

$$plim\bar{G}_n(\hat{\theta}_{GMM}) = plim\bar{G}_n(\bar{\theta}) = plim\frac{\partial m(X_i, Z_i, \theta_0)}{\partial \theta_0'} = E\left[\frac{\partial \bar{m}(\theta_0)}{\partial \theta_0'}\right] = \Gamma(\theta_0)$$

Asymptotic normality

With $plimW_n = W$ and

$$plim\bar{G}_n(\hat{\theta}_{GMM}) = plim\bar{G}_n(\bar{\theta}) = \Gamma(\theta_0)$$

the expression

$$\sqrt{n}(\hat{\theta}_{GMM}-\theta_0)\approx -(\bar{G}_n(\hat{\theta}_{GMM})'W_n\bar{G}_n(\bar{\theta}))^{-1}\bar{G}_n(\hat{\theta}_{GMM})'W_n\sqrt{n}\bar{m}_n(\theta_0)$$

becomes

$$\sqrt{n}(\hat{\theta}_{GMM} - \theta_0) \approx -(\Gamma(\theta_0)'W\Gamma(\theta_0))^{-1}\Gamma(\theta_0)'W\sqrt{n}\bar{m}_n(\theta_0)$$

from which we get the variance V. So

$$\sqrt{n}(\hat{\theta}_{GMM} - \theta_0) \stackrel{d}{\rightarrow} N[0, V]$$

with

$${\displaystyle \bigvee_{\kappa imes \kappa}} = 1/n [\Gamma(heta_0)'W\Gamma(heta_0)]^{-1} [\Gamma(heta_0)'W\Phi W'\Gamma(heta_0)] [\Gamma(heta_0)'W\Gamma(heta_0)]^{-1}$$

That is by GMM1, GMM2, and GMM3 the GMM estimator is asymptotic normal.

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Asymptotic efficiency

Which weighting matrix W gives us the smallest possible asymptotic variance of the GMM estimator $\hat{\theta}_{GMM}$.

The variance of the GMM estimator V depends on the choice of W

$$V = 1/n[\Gamma(\theta_0)'W\Gamma(\theta_0)]^{-1}[\Gamma(\theta_0)'W\Phi W'\Gamma(\theta_0)][\Gamma(\theta_0)'W\Gamma(\theta_0)]^{-1}$$

So let us minimize V to get the optimal weight matrix. Try from GMM3

$$\underset{n\to\infty}{plim}W_n=W=\Phi^{-1}$$

$$V_{GMM,optimal} = 1/n[\Gamma(\theta_0)'\Phi^{-1}\Gamma(\theta_0)]^{-1}[\Gamma(\theta_0)'\Phi^{-1}\Phi\Phi^{-1'}\Gamma(\theta_0)][\Gamma(\theta_0)'\Phi^{-1}\Gamma(\theta_0)]^{-1}$$

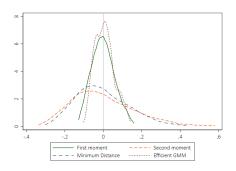
Which can be simplified to

$$V_{GMM,optimal} = 1/n[\Gamma(\theta_0)'\Phi^{-1}\Gamma(\theta_0)]^{-1}$$

Asymptotic efficiency

$$V_{GMM,optimal} = 1/n[\Gamma(\theta_0)'\Phi^{-1}\Gamma(\theta_0)]^{-1}$$

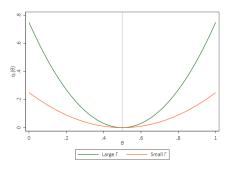
If Φ is small, there is little variation of this specific sample moment around zero and the moment condition is very informative about θ_0 . So it is best to assign a high weight to it.



Asymptotic efficiency

$$V_{GMM,optimal} = 1/n[\Gamma(\theta_0)'\Phi^{-1}\Gamma(\theta_0)]^{-1}$$

If Γ is large, there is a large penalty from violating the moment condition by evaluating at $\theta \neq \theta_0$. Then the moment condition is very informative about θ_0 . V is inversely related to Γ .



Estimate the variance in practice

$$\hat{V}_{GMM,optimal} = 1/n[\Gamma(\theta_0)'\Phi_n^{-1}\Gamma(\theta_0)]^{-1}$$

Consistent estimator

$$\Phi_n = NV(\bar{m}_n(\theta))$$

$$\bar{G}_n(\bar{\theta}) = \frac{\partial m(X_i, Z_i, \hat{\theta})}{\partial \hat{\theta}'}$$

References I

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