

# STATS 3U03

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## Course Outline

- Textbook: Introduction to stochastic processes
- Requirement: 5 assignments, 2 tests, and 1 final
- Test 1: Friday, February 10th
- Test 2: Friday, March 17th

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# 1 Introduction

## 1.1 Review

**Definition 1.1** (Independent random variables).  *$X$  and  $Y$  are independent iff for any  $a, b \in \mathbb{R}$ ,  $P(X \leq a, Y \leq b) = P(X \leq a)P(Y \leq b)$*

## 1.2 Stochastic processes

**Definition 1.2** (Stochastic process). *Let  $T$  be a subset of  $[0, +\infty]$ . For each  $t \in T$ , let  $X_t$  be a random variable. Then, the collection of  $\{X_t : t \in T\}$  is called a stochastic process. Simply put, a stochastic process is just a family of random variables.*

**Example 1.2.1.** Let  $T = \{0\}$ . Then,  $\{X_0\}$  is a stochastic process.

**Example 1.2.2.** Let  $T = \{1, 2, 3, \dots, m\}$  be a set of finite natural numbers. Then,  $\{X_1, X_2, X_3, \dots, X_m\}$  is a stochastic process.

**Example 1.2.3.** Let  $T = \{0, 1, 2, \dots\}$  be a set of all non-negative integers. Then,  $\{X_1, X_2, X_3, \dots\}$  is a stochastic process.

**Example 1.2.4.** Let  $T = [0, +\infty)$  be a set of all non-negative real numbers. Then,  $\{X_t : t \geq 0\}$  is a stochastic process.

**Definition 1.3** (Time index). *Let  $T$  be time index. If  $T = \{0, 1, 2, \dots\}$ , then the time is discrete. If  $T = [0, \infty)$ , then time is continuous.*

**Definition 1.4** (State Space). *State space,  $S$ , is the space where the random variable takes the values.*

Given a sample space,  $S$ , and time index  $t \in T$ , we can define  $X_t(w) \in S$ , to describe a stochastic process. Here,  $\{X_t : t \in T\}$  describes the dependence relation.

We can further categorize a stochastic process by considering the following two cases: countable and uncountable state space. Time index can also be categorized as follows: discrete and continuous time. Note that each stochastic process must belong to one of the four categories.

*Remark.* Every stochastic process can be described by the following three factors:

1. Time index
2. State space
3. Dependence relation

**Example 1.2.5.** Let  $S = \{0, 1\}$  and  $T = \{0, 1, 2, \dots\}$ . Given,

$$X_n = \begin{cases} 1 & \text{with probability of } 1/2 \\ 0 & \text{with probability of } 1/2 \end{cases}$$

$\{X_0, X_1, X_2, \dots\}$  is a stochastic process and is often noted as Bernoulli trials.

## 2 Markov chains (Discrete time Markov chains)

We will only be dealing with discrete time Markov chains in chapter 1 and 2. In other words,  $T = \{0, 1, 2, \dots\}$ . It follows that the state space,  $S$ , will be at most countable. Finally, Markov describes the dependence relation:  $X_0, X_1, X_2, \dots$ .

In example 1.2.5, every trial of the Bernoulli trials was independent. On the other hand, in a Markov model,  $X_{n+1}$  depends on  $X_n$  but not on any past states,  $X_1, X_2, \dots, X_{n-1}$ .

### 2.1 Markov property

**Definition 2.1.** Markov property can be expressed as follows:

$$\begin{aligned} P(X_{n+1} = x_{n+1} | X_0 = x_0, \dots, X_{n-1} = x_{n-1}, X_n = x_n) \\ = P(X_{n+1} = x_{n+1} | X_n = x_n) \end{aligned}$$

$P(X_{n+1} = y | X_n = x)$  is noted as the transition probability and it describes the one step transition from  $x$  to  $y$  starting at time  $n$ . If

$$P(X_{n+1} = y | X_n = x) = P(X_1 = y | X_0 = x),$$

then the Markov chain is called to have stationary transition, or homogeneous.

**Definition 2.2.** Let  $\{X_n : n = 0, 1, 2, \dots\}$  be a homogeneous Markov chain. Then,

$$P_{xy} = P(X_1 = y | X_0 = x) = P(X_{n+1} = y | X_n = x),$$

is the one-step transition probability.

**Definition 2.3.** Following the definition 2.2, we can now define one-step transition matrix:

$$\mathbb{P} = (P_{xy})_{x,y \in S}$$

*Remark.* Given,  $X_0$ ,  $\pi_0(x) = P(X_0 = x)$  is called the *initial distribution*.

Given a Markov chain, we wish to answer the following fundamental questions:

1. Distribution of  $X_n$  for any  $n \geq 1$ .
2. Joint distribution of  $X_{n_1}, \dots, X_{n_k}$  for any  $1 \leq n_1 < n_2 < \dots < n_k$   $k, \geq 2$ .
3. Long time behaviour, i.e.

$$\lim_{n \rightarrow \infty} P(X_n = x)$$

**Example 2.1.1.** We have the following Markov chain:  $\{X_n : n = 0, 1, 2, \dots\}$  where  $S = \{0, 1\}$ . For this model, its initial distribution can be described as follows:

$$\begin{cases} \pi_0(0) = P(X_0 = 0) = a \\ \pi_0(1) = 1 - a \end{cases}$$

Transition probabilities can be written in a similar fashion:

$$\begin{aligned} P(X_1 = 1|X_0 = 0) &= p, & P(X_1 = 0|X_0 = 0) &= 1 - p \\ P(X_1 = 0|X_0 = 1) &= q, & P(X_1 = 1|X_0 = 1) &= 1 - q \end{aligned}$$

where  $0 \leq p, q \leq 1$ . For this Markov chain, we can consider the following three cases:

**Case 1.**  $p = q = 0$ .

This case is trivial.

**Case 2.**  $p = q = 1$ .

This case is also trivial.

**Case 3.**  $0 \leq p + q \leq 2$ .

$$\begin{aligned} P(X_{n+1} = 0) &= P(X_{n+1} = 0 \cap X_n = 0) + P(X_{n+1} = 0 \cap X_n = 1) \\ &= P(X_n = 0)P(X_{n+1} = 0|X_n = 0) + P(X_n = 1)P(X_{n+1} = 0|X_n = 1) \\ &= P(X_n = 0)(1 - p) + P(X_n = 1)q \\ &= P(X_n = 0)(1 - p) + (1 - P(X_n = 0))q \\ &= (1 - p - q)P(X_n = 0) + q \end{aligned}$$

We can further expand this as follows:

$$\begin{aligned} P(X_{n+1} = 0) &= (1 - p - q)P(X_n = 0) + q \\ &= (1 - p - q)[(1 - p - q)P(X_{n-1} = 0) + q] + q \\ &= (1 - p - q)^n P(X_0 = 0) + q \sum_{j=0}^{n-1} (1 - p - q)^j \end{aligned}$$

Note that

$$\sum_{j=0}^{n-1} (1 - p - q)^j = \frac{(1 - p - q)^n - 1}{(1 - p - q) - 1}$$

Therefore, we have

$$\begin{aligned}
P(X_{n+1} = 0) &= (1 - p - q)^n a + q \frac{(1 - p - q)^n - 1}{(1 - p - q) - 1} \\
&= (1 - p - q)^n a - \frac{q}{p + q} ((1 - p - q)^n - 1)
\end{aligned}$$

For this Markov chain, we find that

$$\lim_{n \rightarrow \infty} P(X_n = 0) = \frac{q}{p + q}$$

## 2.2 Transition function and initial distribution

**Example 2.2.1.**

$$\begin{aligned}
P_{xy} &= P(X_{n+1} = y | X_n = x) \\
&= P(X_1 = y | X_0 = x)
\end{aligned}$$

**Definition 2.4.** Transition function,  $P(x, y) : S \times X \rightarrow [0, 1]$ , satisfies the following conditions:

1.  $p(x, y) \geq 0$ .
2.  $\sum_{y \in S} p(x, y) = 1$  for all  $x \in S$ .

**Definition 2.5.** Given a transition function,  $p(x, y)$ , a transition matrix is defined as follows:

$$\mathbb{P} = (p(x, y))_{x, y \in S}$$

**Example 2.2.2.**

$$\begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$

**Example 2.2.3.**

$$\begin{pmatrix} 1/4 & 1/2 & 1/4 \\ 1/8 & 1/4 & 5/8 \\ 0 & 1/4 & 3/4 \end{pmatrix}$$

**Definition 2.6.** Initial distribution is a probability mass function (pmf) that is defined as follows

$$\pi_0(x) = p(X_0 = x).$$

Note that it must satisfy the following conditions:

1.  $\pi_0(x) \geq 0$
2.  $\sum_{x \in S} \pi_0(x) = 1$

**Theorem 2.1.** Let  $\{x_n : n = 0, 1, 2, \dots\}$  be a Markov chain with initial distribution  $\pi_0(x)$ , and one-step transition matrix  $\mathbb{P} = (p(x, y))_{x, y \in S}$ . Then, the distribution of  $X_n$  is

$$\begin{aligned} P(X_n = x_n) &= \sum_{x_0 \in S} \sum_{x_1 \in S} \cdots \sum_{x_{n-1} \in S} \pi_0(x_0) p(x_0, x_1) \cdots p(x_{n-1}, x_n) \\ &= \pi_0 \underbrace{\mathbb{P} \mathbb{P} \cdots \mathbb{P}}_n \end{aligned}$$

*Proof.* For any  $n \geq 1$ ,  $x_n \in S$

$$\begin{aligned} P(X_n = x_n) &= P(X_n = x_n, x_0 \in S, X_1 \in S, \dots, X_{n-1} \in S) \\ &= \sum_{x_0 \in S} \sum_{x_1 \in S} \cdots \sum_{x_{n-1} \in S} P(X_n = x_n, X_0 = x_0, \dots, X_{n-1} = x_{n-1}) \end{aligned}$$

Note that

$$\begin{aligned} &P(X_n = x_n, X_0 = x_0, \dots, X_{n-1} = x_{n-1}) \\ &= P(X_0 = x_0) P(X_1 = x_1 | X_0 = x_0) P(X_2 = x_2 | X_0 = x_0, X_1 = x_1) \\ &\cdots P(X_n = x_n | X_0 = x_0, \dots, X_{n-1} = x_{n-1}) \end{aligned}$$

Using the Markov property, it is evident that the equation above is equivalent to  $P(X_0 = x_0) P(X_1 = x_1 | X_0 = x_0) \cdots P(X_n = x_n | X_{n-1} = x_{n-1})$ .  $\square$

**Example 2.2.4.** Simple random walk is a Markov chain:

$$\begin{aligned} X_0 &= 0 \\ X_1 &= \begin{cases} 1 & p \\ -1 & q \end{cases} \\ X_2 &= \begin{cases} X_1 + 1 & p \\ X_1 - 1 & q \end{cases} \\ X_n &= \begin{cases} X_{n-1} + 1 & p \\ X_{n-1} - 1 & q \end{cases} \end{aligned}$$

where  $S = \{0, \pm 1, \pm 2, \dots\}$ .

**Example 2.2.5** (Ehrenfest chain). Suppose that we have a box and an invisible bar that divides the box into region I and II.  $d$  balls are placed in a box. Initially,  $n$  balls are distributed in region I and  $d - n$  balls are distributed in region II. You pick a ball at random. If it's from region I, you put it back in region II. If it's from region II, you put it back in region I.

First, note that this Markov chain has a state space of  $S = \{0, 1, 2, \dots, d\}$ . We observe that

$$P(0, y) = \begin{cases} 0 & y > 1 \\ 1 & y = 1 \end{cases}$$

$$P(1, y) = \begin{cases} 0 & y \neq 0, 2 \\ \frac{1}{d} & y = 0 \\ 1 - \frac{1}{d} & y = 2 \end{cases}$$

In general, we have

$$P(x, y) = \begin{cases} 0 & y \neq x \pm 1 \\ 1 - \frac{x}{d} & y = x + 1 \\ \frac{x}{d} & y = x - 1 \end{cases}$$

Combining these results, we have the following transition matrix:

$$\begin{pmatrix} 0 & 1 & 0 & \dots & \\ \frac{1}{d} & 0 & 1 - \frac{1}{d} & & \\ & & & \ddots & \\ & & & & \ddots \end{pmatrix}_{(d+1) \times (d+1)}$$