计量经济学:作业一

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1.

$$\hat{\beta}_{1} = \frac{\sum_{i=1}^{n} (x_{i} - \bar{x}) y_{i}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}} \\
= \frac{\sum_{i=1}^{n} (x_{i} - \bar{x}) y_{i}}{SST_{x}} \\
= \frac{\sum_{i=1}^{n} (x_{i} - \bar{x}) (\beta_{0} + \beta_{1} x_{i} + u_{i})}{SST_{x}} \\
= \frac{0 + \beta_{1} SST_{x} + \sum_{i=1}^{n} u_{i} (x_{i} - \bar{x})}{SST_{x}} \\
= \beta_{1} + \frac{\sum_{i=1}^{n} u_{i} (x_{i} - \bar{x})}{SST_{x}} \\
= \beta_{1} + \frac{\sum_{i=1}^{n} d_{i} u_{i}}{SST_{x}} \quad (d_{i} = x_{i} - \bar{x})$$

$$Var(\hat{\beta}_{1}) = (\frac{1}{SST_{x}})^{2} Var(\sum_{i=1}^{n} d_{i} u_{i}) \\
= (\frac{1}{SST_{x}})^{2} \sum_{i=1}^{n} d_{i}^{2} Var(u_{i}) \\
= (\frac{\sigma}{SST_{x}})^{2} \sum_{i=1}^{n} d_{i}^{2} \\
= (\frac{\sigma}{SST_{x}})^{2} \\
= \frac{\sigma^{2}}{\sum_{i=1}^{n} (x_{i} - \bar{x})^{2}}$$

$$(1)$$

$$E[(\hat{\beta}_{1} - \beta_{1})\bar{u})] = E[\bar{u}\frac{\sum_{i=1}^{n} d_{i}u_{i}}{SST_{x}}]$$

$$= \frac{1}{SST_{x}} \sum_{i=1}^{n} d_{i}\bar{u}E(u_{i})$$

$$= \frac{1}{SST_{x}} \sum_{i=1}^{n} (\bar{u})^{2}$$

$$= \frac{\bar{u}^{2}}{SST_{x}} \sum_{i=1}^{n} d_{i}$$

$$= 0$$
(3)

$$Cov(\hat{\beta}_1, \bar{u}) = E[\hat{\beta}_1 \bar{u}] - E[\hat{\beta}_1] E[\bar{u}]$$

$$= 0 - 0$$

$$= 0$$
(4)

$$Var(\hat{\beta}_{0}) = Var[\beta_{0} + \bar{u} + (\beta_{1} - \hat{\beta}_{1})\bar{x}]$$

$$= Var[\bar{u} - \hat{\beta}_{1}\bar{x}]$$

$$= Var[\bar{u}] + Var[\hat{\beta}_{1}]\bar{x}^{2} \quad (Cov(\hat{\beta}_{1}, \bar{u}) = 0)$$

$$= Var[\bar{u}] + \bar{x}^{2} \frac{\sigma^{2}}{SST_{x}}$$

$$= \frac{\sigma^{2}}{n} + \frac{\bar{x}^{2}\sigma^{2}}{SST_{x}}$$

$$= \frac{\sigma^{2}}{nSST_{x}} (SST_{x} + n\bar{x}^{2})$$

$$= \frac{\sigma^{2} \sum_{i=1}^{n} x_{i}^{2}}{n \sum_{i=1}^{n} (x_{i} - \bar{x})^{2}}$$
(5)

2. 不正确。

对于给定抽样 x,虽然有 E[u|x]=0,但是因为 u_i 仅有有限个样本,不能从统计意义上推断 $\sum_{i=1}^n u_i=0$.

此外, $E[\hat{\beta}_1] = \beta_1$ 的无偏性也不能说明 $\hat{\beta}_1 = \beta_1$.

此题混淆了统计意义和抽样意义的区别。

3. 代码如下:

```
cd "D:\大三上\计量\作业\" //注意修改数据路径
use "EduIncome.dta", clear

summarize gender birthyear marriage empjob_twage schooling_yr

generate female = 1 if gender == 1
replace female = 0 if female==.
tab female

reg empjob_twage schooling_yr female
```

```
generate pred_twage = 1984.774*schooling_yr - 14048.56 * female + 46554.54

generate pred_u = empjob_twage - pred_twage

summarize empjob_twage pred_twage pred_u

drop pred_twage pred_u

drop pred_twage pred_u
```

(a) 除 id 之外所有变量的均值、标准差、最小值和最大值:

. summarize g	ender birthyear	marriage	empjob_twage	schooling_	yr
Variable	Obs	Mean	Std. Dev.	Min	Max
gender birthyear marriage empjob_twage schooling_yr	2,539 2,539 2,539 2,539 2,539	1.688854 1974.824 .6025994 57321.29 7.627019	.4630536 11.30764 .4894565 40688.22 2.942299	1 1914 0 2113.569	2 1997 1 608707.9

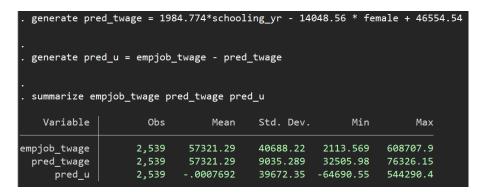
(b) 创建新变量 female 并统计女性比例 女性占比 31.11%

```
generate female = 1 if gender == 1
(1,749 missing values generated)
. replace female = 0 if female==.
(1,749 real changes made)
. tab female
     female
                   Freq.
                              Percent
                                              Cum.
          0
                   1,749
                                68.89
                                             68.89
                                31.11
          1
                     790
                                            100.00
      Total
                   2,539
                               100.00
```

(c) 回归方程,给出结果

. reg empjob_t	twage schoolin	g_yr femal	e				
Source	SS	df	MS	Numbe	er of obs		2,539
				F(2,	2536)		65.77
Model	2.0719e+11	2	1.0360e+11	Prob	> F		0.0000
Residual	3.9945e+12	2,536	1.5751e+09	R-sq	uared	=	0.0493
				Adj I	R-squared		0.0486
Total	4.2017e+12	2,538	1.6555e+09	Root	MSE		39688
empjob_twage	Coef.	Std. Err.	t F	۰> t	[95% Con	f.	Interval]
schooling_yr	1984.774	268.381	7.40	0.000	1458.506		2511.042
female	-14048.56	1705.326	-8.24	0.000	-17392.54		-10704.59
_cons	46554.54	2289.326	20.34 6	0.000	42065.4		51043.68

- (d) 受教育年限一样时,平均来讲男性年收入比女性年收入高 14048.56 元。
- (e) 教育年限每增加一年,收入平均增加 1984.774 元
- (f) 计算均值



- (g) $R^2 = 0.0493$,所以有 4.93% 的比例的收入波动被教育年限和性别解释。(修正的 $R^2 = 0.0486$)
- 4. (a) 不对。满足 $x_1 + x_2 = 1$, 并不是完全线性关系 (如 $x_1 = \alpha x_2$)。两个变量存在相关性,但并不是完全线性。所以依然满足 MLR.3,可以用 OLS。

(b)

$$E[ux_1] = E[E[(ux_1|x_1, x_2)]] = E[x_1E[u|x_1, x_2]] = E[x_10] = 0$$
(6)

$$E[ux_2] = E[E[(ux_2|x_1, x_2)]] = E[x_2E[u|x_1, x_2]] = E[x_20] = 0$$
(7)

(c) 由 $E[ux_1] = 0$ 和 $E[ux_2] = 0$ 得到

$$\frac{1}{n} \sum_{i=1}^{n} x_{1_i} (y_i - \hat{\beta}_1 x_{1_i} - \hat{\beta}_2 x_{2_i}) = 0$$

$$\frac{1}{n} \sum_{i=1}^{n} x_{2_i} (y_i - \hat{\beta}_1 x_{1_i} - \hat{\beta}_2 x_{2_i}) = 0$$
(8)

(d) 目标函数

$$H = \sum_{i=1}^{n} (y_i - b_1 x_{1_i} - b_2 x_{2_i})^2$$
(9)

$$(\hat{\beta}_1, \hat{\beta}_2) = \arg_{(b_1, b_2)} \min H \tag{10}$$

(e) 一阶条件:

$$\frac{\partial H}{\partial b_1} = -2\sum_{i=1}^n x_{1_i} (y_i - b_1 x_{i_1} - b_2 x_{i_2}) = 0$$

$$\frac{\partial H}{\partial b_2} = -2\sum_{i=1}^n x_{2_i} (y_i - b_1 x_{i_1} - b_2 x_{i_2}) = 0$$
(11)

- (f) 参数估计是一样的。从数学上看,两者的约束方程是等价的。因此得到的解是一致的。
- (g) 解得

$$\hat{\beta}_{1} = \frac{\sum_{i=1}^{n} x_{2_{i}}^{2} \sum_{i=1}^{n} x_{1_{i}} y_{i} - \sum_{i=1}^{n} x_{2_{i}} y_{i} \sum_{i=1}^{n} x_{1_{i}} x_{2_{i}}}{\sum_{i=1}^{n} x_{1_{i}}^{2} x_{2_{i}}^{2} - (\sum_{i=1}^{n} x_{1_{i}} x_{2_{i}})^{2}}$$

$$\hat{\beta}_{2} = \frac{\sum_{i=1}^{n} x_{1_{i}}^{2} \sum_{i=1}^{n} x_{2_{i}} y_{i} - \sum_{i=1}^{n} x_{1_{i}} y_{i} \sum_{i=1}^{n} x_{1_{i}} x_{2_{i}}}{\sum_{i=1}^{n} x_{1_{i}}^{2} x_{2_{i}}^{2} - (\sum_{i=1}^{n} x_{1_{i}} x_{2_{i}})^{2}}$$

$$(12)$$

由数据计算的 $\hat{\beta}_1$ 和 $\hat{\beta}_2$ 为

$$\hat{\beta}_1 = 5.0 \hat{\beta}_2 = 6.0$$
 (13)

附上述计算的 Python 程序:

```
import numpy as np
y = np.array([5, 9, 7, 3, 1, 6, 4, 8, 2, 10])
x1 = np.array([1, 1, 1, 1, 1, 0, 0, 0, 0])
x2 = np.array([0, 0, 0, 0, 1, 1, 1, 1])
x1_x2 = x1 * x2

beta1 = (x2.dot(x2) * x1.dot(y) - x2.dot(y) * x1.dot(x2)) / (x1.dot(x1) * x2.dot(x2) - (x1_x2.sum())**2)
beta2 = (x1.dot(x1) * x2.dot(y) - x1.dot(y) * x1.dot(x2)) / (x1.dot(x1) * x2.dot(x2) - (x1_x2.sum())**2)
print(beta1, beta2)
```