

# 可靠性数据与生存分析作业

计 91 刘程华 20180116897

清华大学计算机系

日期: 2021 年 3 月 7 日

## 1 第一次作业

**1.1 Prove: If  $\Lambda(t)$  is cumulative hazard function of  $T$ , then  $\Lambda(T) \sim EXP(1)$ , the unit exponential distribution.**

We have known that Cumulative hazard function  $\Lambda(t)$

$$\Lambda(t) = \int_0^t \lambda(u) du = -\log\{S(t)\}$$

Where  $S(t) = P[T > t] = 1 - F(t)$  is survival function. Since  $S(0) = 1$ ,

$$S(t) = \exp\{-\Lambda(t)\} = \exp\left\{-\int_0^t \lambda(u) du\right\}$$

So we have

$$P(\Lambda(T) \geq t) = P\left(T \geq \Lambda^{-1}(t)\right) = S\left(\Lambda^{-1}(t)\right) = \exp\left\{-\Lambda\left(\Lambda^{-1}(t)\right)\right\} = e^{-t}$$

Q.E.D.

**1.2 Prove: If  $T_1$  and  $T_2$  are two independent survival times with hazard function  $\lambda_1(t)$  and  $\lambda_2(t)$ , respectively, then  $T = \min(T_1, T_2)$  has a hazard function  $\lambda_T(t) = \lambda_1(t) + \lambda_2(t)$**

We have survival function

$$S_T(t) = P(\min(T_1, T_2) \geq t) = P(T_1 \geq t) P(T_2 \geq t) = S_1(t) S_2(t)$$

Then we have

$$\lambda_T(t) = -\frac{d \log \{S_T(t)\}}{dt} = -\frac{d \log \{S_1(t)\}}{dt} - \frac{d \log \{S_2(t)\}}{dt} = \lambda_1(t) + \lambda_2(t)$$

Q.E.D.

**1.3 Prove the formula from the lecture notes**

$$mrl(t_0) = E[T - t_0 | T \geq t_0] = \frac{\int_{t_0}^{\infty} S(t) dt}{S(t_0)}$$

We remark the nonnegative random variable  $T$  with distribution function  $F$ . The mean residual life is defined as

$$mrl(t_0) = E[T - t_0 | T \geq t_0] = \frac{E[(T - t_0)I_{\{T > t_0\}}]}{P\{T > t_0\}} = \frac{1}{1 - F(t_0)} \int_{t_0}^{\infty} (t - t_0) dF(t)$$

for  $t_0 > 0$ , according to the Fubini's theorem

$$\int_{t_0}^{\infty} (t - t_0) dF(t) = \int_{t_0}^{\infty} \left( \int_{t_0}^t du \right) dF(t) = \int_{t_0}^{\infty} \left( \int_u^{\infty} dF(t) \right) du = \int_{t_0}^{\infty} P\{X > u\} du = \int_{t_0}^{\infty} (1 - F(u)) du$$

Therefore,

$$mrl(t_0) = \frac{1}{1 - F(t_0)} \int_{t_0}^{\infty} (1 - F(u)) du = \frac{\int_{t_0}^{\infty} S(t) dt}{S(t_0)}$$

Q.E.D.

**1.4 The time in days to development of a tumor for rats exposed to a carcinogen follows a Weibull distribution with  $\alpha = 2$  and  $\lambda = 0.01$**

$$f(t) = \alpha \lambda t^{\alpha-1} e^{-\lambda t^\alpha}$$

**(a) Find the probabilities that a (random) rat will be tumor free at 50 days. (b) What is the average time to tumor development? (Hint  $\Gamma(0.5) = \sqrt{\pi}$ ) (c) Find the hazard rate of time to tumor development at 50 days. (d) Find the median time to tumor development.**

(a) according to  $f(t) = \alpha \lambda t^{\alpha-1} e^{-\lambda t^\alpha}$ , we have known that  $S(t) = e^{-\lambda t^\alpha}$ . So,

$$P(T \geq 50) = S(50) = e^{-0.01 \times (50)^2} = 0.082085$$

(b) the average time to tumor development is

$$E(T) = \int_0^{\infty} S(t) dt = \int_0^{\infty} e^{-0.01 \times (t)^2} dt = 5\sqrt{\pi} = 8.86227$$

(c) the hazard rate of time to tumor development is

$$\lambda(t) = -\frac{d \log \{S(t)\}}{dt} = \alpha \lambda t^{\alpha-1}$$

the hazard rate of time to tumor development at 50 days is

$$\lambda(50) = 2 \times 0.01 \times 50 = 1$$

(d) the median time to tumor development is

$$t_{0.5} = \left[ \frac{\log 2}{\lambda} \right]^{\frac{1}{\alpha}} = 8.325546$$

**1.5 Suppose we have a small data set with different kinds of censoring: 2+, 3, 4, 5-, 6, 7+, [8, 10], where +(-) means right (left) censored observations and [a, b] means an interval censored observation. Suppose the distribution of the underlying survival time is an exponential distribution with a constant hazard  $\lambda$ . Write down the likelihood function of  $\lambda$  for this given data set.**

Obviously  $\lambda > 0$ . And the likelihood function is

$$\begin{aligned} L(\lambda) &= S(2) \cdot f(3) \cdot f(4) \cdot F(5) \cdot f(6) \cdot S(7) \cdot [F(10) - F(8)] \\ &= \lambda^3 e^{-30\lambda} (1 - e^{-5\lambda}) (1 - e^{-2\lambda}) \end{aligned}$$

**1.6 Derive the following properties for a) exponential distribution and b) Weibull distribution. (1) Survival function (2) Hazard function (3) Cumulative hazard function (4) Mean survival time**

Weibull distribution with the parametrization  $W(\lambda, p)$ ,  $\lambda > 0$  and  $p > 0$  :

(1)  $S(t) = e^{-(\lambda t)^p}$

(2)  $\lambda(t) = p\lambda(\lambda t)^{p-1}$

(3)  $\Lambda(t) = (\lambda t)^p$

(4)  $E(t) = \frac{\Gamma(1+1/p)}{\lambda}$

in the case of exponential distribution, just set  $p = 1$  in the above formula.