A Simple FPTAS for Counting Edge Covers

Abstract

An edge cover of a graph is a set of edges such that every vertex has at least an adjacent edge in it. We design a very simple deterministic fully polynomial-time approximation scheme (FPTAS) for counting the number of edge covers for any graph. Previously, approximation algorithm is only known for 3 regular graphs and it is randomized. Our main technique is correlation decay, which is a powerful tool to design FPTAS for counting problems. In order to get FPTAS for general graphs without degree bound, we make use of a stronger notion called computationally efficient correlation decay, which is introduced in [Li, Lu, Yin SODA 2012].

1 Introduction

An edge cover of a graph is a set of edges such that every vertex has at least an adjacent edge in it. For any graph without isolated vertices, there is at least one edge cover: the set of all edges. So the decision problem is trivial. There is also a polynomial time algorithm based on maximum machining to compute a edge cover with minimum cardinality. In this paper, we study the counting version: For a given input graph, we count the number of edge covers for that graph. Unlike the decision or optimization problem, counting edge covers is a #Pcomplete problem even when we restrict the input to 3 regular graphs. In this paper, we study the approximation version. For any given parameter $\epsilon > 0$, the algorithm output a number \hat{N} such that $(1-\epsilon)N < \hat{N} < (1+\epsilon)N$, where N is the accurate number of edge covers of the input graph. We also require that the running time of the algorithm is bounded by $poly(n, 1/\epsilon)$, where n is the number of vertices of the given graph. This is called a fully polynomial-time approximation scheme (FPTAS). Our main result of this paper is a FPTAS for counting edge covers for any graph. Previously, approximation algorithm was only known for 3 regular graphs and the algorithm is randomized [BR09]. The randomized version of FPTAS is called fully polynomial-time randomized approximation scheme (FPRAS), which uses random bits in the algorithm and we require that the final output is within the range $[(1-\epsilon)N, (1+\epsilon)N]$ with high probability.

Edge cover is related to many other graph problems such as (perfect) matching, k-factor problems and so on. All of them are talking about a set of edges which satisfies some local constraints defined on each vertex. For edge cover, it says that at least one incident edge should be chosen; while for matching, it is at most one edge. For generic constraints, it is the Holant framework [CLX09, CLX11], which is well studied in terms of exactly counting [CLX11, HL12, CGW13], and recently in approximate counting [YZ13, McQ13, LWZ]. For counting matchings, there is a FPRAS based on Markov Chain Monte Carlo (MCMC) for any graph [JS89]. Deterministic FPTAS is only known for graphs with bounded degree [BGK⁺07]. For counting perfect matchings, it is a long standing open question if there is an FPRAS (or FPTAS) for it. For bipartite graphs, there is a FPRAS for counting perfect matchings. The

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weighted version can be viewed as computing permanent of a non-negative matrix [JSV04]. This is one great achievement of approximate counting. It is still widely open if there exists a FPTAS for it or not. The current best deterministic algorithm can only approximate the permanent with an exponential large factor [LSW98, GK10]. There are many other counting problems, for which there is a FPRAS and we do not know if there is a FPTAS or not. In this paper, we give a complete FPTAS for a problem, for which even FPRAS was only known for very special family of instances.

Another view point of the edge cover problem is read twice monotone CNF formula (Rtw-Mon-CNF): Each edge is viewed as a Boolean variable and it is connected with two vertices (read twice); the constraint on each vertex is exactly a monotone OR constrain as at least one edge variable is assigned to be True. Counting number of solutions for a Boolean formula is another set of interesting problems studied both in exact counting and approximate counting. One famous example is the FPRAS for counting the solutions for a DNF formula [KL83, KLM89]. It is an important open question to derandomize it [Tre04, GMR12]. Our FPTAS for counting edge covers can also be viewed as a FPTAS for counting solutions for a Rtw-Mon-CNF formula. If we do not restrict that each variable appears in at most two constraints, there is no FPTAS or FPRAS unless NP is equal to P or RP [DGJ10].

The common overall approach for designing approximate counting algorithms is to relate counting with probability distribution. This is usually referred as "counting vs sampling" paradigm when one mainly focuses on randomized counting. If we can compute (estimate) the marginal probability, which in our problem is the probability of a given edge is chosen when we sample a edge cover uniformly at random, we can in turn to approximate count. In randomized FPRAS, one estimate the marginal probability by sampling, and the most successful approach is sampling by Markov chain [JS97]. In deterministic FPTAS, one calculate the marginal probability directly, and the most successful approach is correlation decay as introduce in [BG08] and [Wei06]. We elaborate a bit on the ideas. The marginal probability is estimated using only a local neighborhood around the edge. To justify the precision of the estimation, we show that far-away edges have little influence on the marginal probability. One most successful example is in anti-ferromagnetic two-spin systems [LLY12, SST12, LLY13], including counting independent sets [Wei06]. The correlation decay based FPTAS is beyond the best known MCMC based FPRAS and achieves the boundary of approximability [SS12, GSV12]. To the best of our knowledge, that was the only example for which the best tractable range for correlation decay based FPTAS exceeds the sampling based FPRAS. This paper provides another such example. FPRAS was the solution concept for approximate counting [DGGJ00]. The recently development of correlation decay based FPTAS is changing the picture. It is interesting to investigate the deep relation between these two approaches.

A set of tools was developed for establishing correlation decay property. These are something like coupling argument, canonic path and so on for establish to the rapid mixing property for Markov Chains. They are self-avoid walk tree, computation tree, potential function, dangling instance, bounded variables and so on. Armed with these powerful tools, there are recently many FPTAS designed for many counting problems [LLY12,SST12,LLY13,YZ13,LWZ]. Many of these techniques are also used in this paper for designing and analyzing the FPTAS for counting edge covers.

Usually, the correlation decay property only implies FPTAS for system with bounded degree. The reason is that we need to explore a local neighborhood with radius of order $\log n$, then the total running time is $n^{\log n}$, which is not a polynomial, if there is no degree bound. To overcome this, we make use of stronger notion called computationally efficient correlation decay as introduced in [LLY12]. The observation is that when we go through a vertex with superconstant degree, the error is also decreased by a super-constant rate. Thus we do not need to explore a depth of $\log n$ if the degrees are large. The tradeoff relation between degree and decay rate defined by computationally efficient correlation decay can support FPTAS with unbounded

degree systems. Previously, this notation was only used in anti-ferromagnetic two-spin systems. In this paper, we prove that the distribution defined by edge covers also satisfies this stronger version of correlation decay and thus we give FPTAS for counting edge covers for any graph.

2 Preliminaries

2.1 Definitions

An edge cover of a graph is a set of edges such that every vertex has at least an adjacent edge in it.

Given a graph G = (V, E) with $e \in E$, we use EC(G) to denote the set of all edge covers of graph G, and P(G, e) to denote the marginal probability over EC(G) that edge e is not chosen, or formally, with $X \sim EC(G)$ uniformly,

$$P(G, e) \triangleq \mathbb{P} (\text{edge } e \text{ is not chosen in } X)$$
 (1)

In this paper, we deal with an extended notion of undirected graphs where dangling edges and free edges may be allowed.

Definition 1. A dangling edge $e = (u, \underline{\ })$ of a graph is such singleton edge with only one end-point vertex u, as shown in the Figure 1a.

A free edge $e = (_, _)$ of a graph is such edge with no end-point vertex. Note that a free edge is not a dangling edge.

So we use graph to refer graphs with or without dangling edges or free edges. Edges in the usual sense (i.e. not dangling and not free), will be referred to as normal edges.

We remark that an alternative view to these combinatorial definitions is from Rtw-Mon-CNF. A dangling edge is just a variable which only appears at one clause, and a free edge is a variable that does not appear at all, whereas normal edge just corresponds to variables appearing twice.

For a graph G = (V, E), an edge $e = (u, v) \in E$ and a vertex $u \in V$, define

$$G - e \triangleq (V, E - e)$$

 $e - u \triangleq (\underline{\hspace{0.5cm}}, v)$ (note that here v could be $\underline{\hspace{0.5cm}}$)

$$G-u\triangleq (V-u,\{e:e\in E,e\text{ is not incident with }u\}\cup \{e-u:e\in E,e\text{ is incident with }u\})$$

Note that here in edge set E, duplicates are allowed. We may have multiple dangling edges $(v, _)$, many free edges $(_, _)$ and even parallel edges between two vertices (u, v).

For example, given a degree-3 vertex u with dangling edge e shown in Figure 1a , the result of G-e-u is shown in Figure 1b.

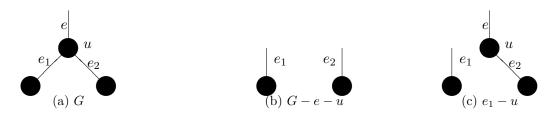


Figure 1: Dangling edges examples.

We use 0 to denote scalar value 0, and $\mathbf{0}$ to denote vector value 0, and $\{e_i\}_{i=1}^d$ denote the d-dimensional vector with i-th coordinate being e_i , so $\{e_i\} = \mathbf{0}$ means $\forall i, e_i = 0$. We also use the convention that when $d = 0, \prod_i^d p_i \triangleq 1$.

In general we use n to refer to the number of vertices in a given graph, and m to refer to the number of edges.

3 The Recursion

First we show the recursion for computing the marginal probability P(G, e).

3.1 e is free edge

Proposition 2.

$$P(G,e) = \frac{1}{2}$$

Proof. If e is a free edge, then any edge cover with e chosen is in one-to-one correspondence to any edge cover with e not chosen. Hence exactly half of the edge covers in EC(G) doesn't choose e, so $P(G, e) = \frac{1}{2}$.

3.2 e is dangling

Proposition 3. For graph G = (V, E) with a dangling edge $e = (u, _)$, denote the d edges incident with u except e as e_1, e_2, \ldots, e_d , let $G_i = G - e - u - \sum_{k=1}^{i-1} e_k$ (specifically, $G_1 = G - e - u$), we have

$$P(G,e) = \frac{1 - \prod_{i=1}^{d} P(G_i, e_i)}{2 - \prod_{i=1}^{d} P(G_i, e_i)}$$
(2)

Proof. For $\alpha \in \{0,1\}^d$, let $EC_{\alpha}(G-e-u)$ be the set of edge coverings in G-e-u such that its restriction onto $\{e_i\}_{i=1}^d$ is consistent with α , denote $Z_{\alpha} = ||EC_{\alpha}(G-e-u)||$, and $Z = \sum_{\alpha \in \{0,1\}^d} Z_{\alpha} = ||EC(G-e-u)||$.

Also note that as long as $\alpha \neq 0$, counting edge coverings with restriction α is the same in either G, G-e, or G-e-u, so it's enough to work with G-e-u. Note that in G-e-u, for every i, e_i is either dangling or free, but not normal.

$$\begin{split} P(G,e) = & \frac{\|EC(G-e)\|}{\|EC(G)\|} \\ = & \frac{\sum_{\alpha \in \{0,1\}^d, \alpha \neq \mathbf{0}} Z_{\alpha}}{Z_{\mathbf{0}} + 2\sum_{\alpha \in \{0,1\}^d, \alpha \neq \mathbf{0}} Z_{\alpha}} \\ = & \frac{1 - \frac{Z_{\mathbf{0}}}{Z}}{2 - \frac{Z_{\mathbf{0}}}{Z}}. \end{split}$$

Now consider the term $\frac{Z_0}{Z}$, it says the probability that a uniformly random edge cover drawn from EC(G-e-u) picked none of $\{e_i\}_{i=1}^d$, so

$$\frac{Z_{\mathbf{0}}}{Z} = \mathbb{P}\left(\{e_i\} = \mathbf{0}\right) = \mathbb{P}(e_1 = 0) \prod_{i=2}^{d} \mathbb{P}\left(e_i = 0 \mid \{e_j\}_{j=1}^{i-1} = \mathbf{0}\right) = \prod_{i=1}^{d} P(G_i, e_i).$$

Hence concludes the proof.

3.3 e is normal edge

By definition we have

$$P(G, e) = \frac{\|EC(G - e)\|}{\|EC(G - e)\| + \|EC(G - e - u - v)\|}.$$
 (3)

For e = (u, v) as a normal edge, let $\{e_i\}$ be the set of edges incident with vertex u except e, and $\{f_i\}$ is the set of edges incident with vertex v except e, and $d_1 = \|\{e_i\}\|$, $d_2 = \|\{f_i\}\|$, now for $\alpha \in \{0, 1\}^{d_1}$, $\beta \in \{0, 1\}^{d_2}$, we use $E_{\alpha, \beta}^G$ to denote the set of edge coverings in G such that its restriction to $\{e_i\}_{i=1}^{d_1}$ is consistent with α , and restriction to $\{f_i\}_{i=1}^{d_2}$ is consistent with β .

restriction to $\{e_i\}_{i=1}^{d_1}$ is consistent with α , and restriction to $\{f_i\}_{i=1}^{d_2}$ is consistent with β . Denote $Z_{\alpha,\beta}^G \triangleq \|EC_{\alpha,\beta}(G)\|$, $G_1 \triangleq G - e$, $G_2 \triangleq G - e - u - v$, now for $\alpha \neq \mathbf{0}$, $\beta \neq \mathbf{0}$, we also have working with G_1 and working with G_2 is the same with restriction to α and β , or formally,

$$||EC(G-e)|| = \sum_{\alpha \neq \mathbf{0}, \beta \neq \mathbf{0}} Z_{\alpha,\beta}^{G_1} = \sum_{\alpha \neq \mathbf{0}, \beta \neq \mathbf{0}} Z_{\alpha,\beta}^{G_2}$$

Since we are only working with G_2 , we may simply denote $Z_{\alpha,\beta} \triangleq Z_{\alpha,\beta}^{G_2}$, now let $Z = \sum_{\alpha,\beta} Z_{\alpha,\beta}$, and $\mathbb{P}(\alpha = 0, \beta = 0) \triangleq \frac{Z_{\mathbf{0},\mathbf{0}}}{Z}, \mathbb{P}(\alpha = \mathbf{0}) \triangleq \frac{\sum_{\beta} Z_{\mathbf{0},\beta}}{Z}, \mathbb{P}(\beta = \mathbf{0}) \triangleq \frac{\sum_{\alpha} Z_{\alpha,\mathbf{0}}}{Z}$ we have,

Proposition 4.

$$P(G, e) = 1 - \frac{1}{2 + \mathbb{P}(\alpha = 0, \beta = 0) - \mathbb{P}(\alpha = 0) - \mathbb{P}(\beta = 0)}$$

Proof.

$$P(G, e) = \frac{\sum_{\alpha \neq \mathbf{0}, \beta \neq \mathbf{0}} Z_{\alpha, \beta}}{Z + \sum_{\alpha \neq \mathbf{0}, \beta \neq \mathbf{0}} Z_{\alpha, \beta}}$$

$$= \frac{Z - \sum_{\alpha} Z_{\alpha, \mathbf{0}} - \sum_{\beta} Z_{\mathbf{0}, \beta} + Z_{\mathbf{0}, \mathbf{0}}}{2Z - \sum_{\alpha} Z_{\alpha, \mathbf{0}} - \sum_{\beta} Z_{\mathbf{0}, \beta} + Z_{\mathbf{0}, \mathbf{0}}}$$

$$= 1 - \frac{1}{2 + \mathbb{P}(\alpha = 0, \beta = 0) - \mathbb{P}(\alpha = 0) - \mathbb{P}(\beta = 0)}$$

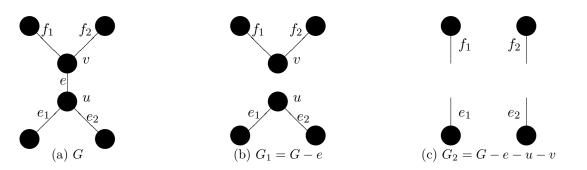


Figure 2: Normal edge examples.

Let
$$G_i^1 \triangleq G - e - u - v - \sum_{k=1}^{i-1} e_k$$
, $G_i^2 \triangleq G - e - u - v - \sum_{k=1}^{d_1} e_k - \sum_{k=1}^{i-1} f_k$, $G_i^3 \triangleq G - e - u - v - \sum_{k=1}^{i-1} f_k$, so we have

Proposition 5.

$$\mathbb{P}(\alpha = 0) = \prod_{i=1}^{d_1} P(G_i^1, e_i)$$

$$\mathbb{P}(\beta = 0) = \prod_{i=1}^{d_2} P(G_i^3, f_i)$$

$$\mathbb{P}(\alpha = 0, \beta = 0) = \prod_{i=1}^{d_1} P(G_i^1, e_i) \cdot \prod_{i=1}^{d_2} P(G_i^2, f_i)$$

Proof.

$$\mathbb{P}(\alpha = 0) = \mathbb{P}(\{e_i\} = \mathbf{0}) = \prod_{i=1}^{d_1} P(G_i^1, e_i)$$

$$\mathbb{P}(\beta = 0) = \mathbb{P}(\{f_i\} = \mathbf{0}) = \prod_{i=1}^{d_2} P(G_i^3, f_i)$$

$$\mathbb{P}(\alpha = 0, \beta = 0) = \mathbb{P}(\alpha = 0) \cdot \mathbb{P}(\beta = 0 \mid \alpha = 0)$$

$$= \mathbb{P}(\{e_i\} = \mathbf{0}) \cdot \mathbb{P}(\{f_i\} = \mathbf{0} \mid \{e_i\} = \mathbf{0})$$

$$= \prod_{i=1}^{d_1} \mathbb{P}(e_i = 0 \mid \{e_j\}_{j=1}^{i-1} = \mathbf{0}) \cdot \prod_{i=1}^{d_2} \mathbb{P}(f_i = 0 \mid \{e_j\}_{j=1}^{d_1} = \mathbf{0}, \{f_j\}_{j=1}^{i-1} = \mathbf{0})$$

$$= \prod_{i=1}^{d_1} P(G_i^1, e_i) \cdot \prod_{i=1}^{d_2} P(G_i^2, f_i)$$

Corollary 6.

$$P(G, e) = 1 - \frac{1}{2 + \prod_{i=1}^{d_1} P(G_i^1, e_i) \cdot \prod_{i=1}^{d_2} P(G_i^2, f_i) - \prod_{i=1}^{d_1} P(G_i^1, e_i) - \prod_{i=1}^{d_2} P(G_i^3, f_i)}$$

Note that for every i, e_i is dangling or free in G_i^1 , f_i is dangling or free in G_i^3 , and in G_i^2 , neither e_i nor f_i is normal.

4 Computation Tree for Marginal Probability

Algorithm 1: Estimate P(G, e) up to depth L

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\begin{array}{l} \textit{function } P(G,e,L): \\ \textbf{input} : \mathsf{Graph} \ G; \ \mathsf{edge} \ e; \ \mathsf{Recursion} \ \mathsf{depth} \ L; \\ \textbf{output}: \ \mathsf{Estimate} \ \mathsf{of} \ P(G,e) \ \mathsf{up} \ \mathsf{to} \ \mathsf{depth} \ L \ . \\ \textbf{begin} \\ & \quad \quad | \mathbf{f} \ L \leq 0 \ \mathbf{then} \\ & \quad \quad | \mathbf{return} \ \frac{1}{2} \\ & \quad \quad | \mathbf{else} \ \mathsf{if} \ e \ \mathsf{is} \ \mathsf{free} \ \mathbf{then} \\ & \quad \quad | \mathbf{return} \ \frac{1}{2}; \\ & \quad \quad | \mathbf{else} \ \mathsf{if} \ e \ \mathsf{is} \ \mathsf{dangling} \ \mathbf{then} \\ & \quad \quad | L' \leftarrow L - \lceil \log_6 \left(d+1\right) \rceil; \\ & \quad \quad | \mathbf{return} \ \frac{1-\prod_{i=1}^d P(G_i,e_i,L')}{2-\prod_{i=1}^d P(G_i,e_i,L')}; \\ & \quad \quad | \mathbf{else} \ // \ e \ \mathsf{is} \ \mathsf{normal} \\ & \quad | X \leftarrow \prod_{i=1}^{d_1} P(G_i^1,e_i,L); \\ & \quad \quad | Y \leftarrow \prod_{i=1}^{d_2} P(G_i^2,f_i,L); \\ & \quad \quad | Z \leftarrow \prod_{i=1}^{d_2} P(G_i^3,f_i,L); \\ & \quad \quad | \mathbf{return} \ 1 - \frac{1}{2+X \cdot Y - X - Z}; \\ \end{array}
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We may compute the marginal probability P(G, e) with the previous recursion, but that could take recursion depth of O(n) which results in exponential computation time.

So here we use a truncated computation tree for an estimate of P(G, e).

We remark that the recursion depth used here is actually the so-called M-based depth introduced in [LLY12] with M=6. In fact M could take any value as long as $M \geq 6$.

Note that the normal case is invoked only once, so the algorithm keeps exploring in the third cases, until it hits the first 2 cases. We remark that an alternative view of the recursion depth is, we replace every node with degree greater than 6 with a 6-ary branching subtree. Now with this alternative view, it's easy to see that the nodes involved in the branching tree up to depth L is at most 6^L , so the algorithm P(G, e, L) has running time $O(n \cdot 6^L)$.

5 Correlation Decay Property

In the last section we show an algorithm P(G, e, L) for estimating the marginal probability P(G, e), so here we establish the exponential correlation decay property, in the stronger sense with the M-based depth, of the estimation error in P(G, e, L).

Theorem 7. Given graph G, edge e and depth L,

$$|P(G, e, L) - P(G, e)| \le 3 \cdot (\frac{1}{2})^{L+1}$$

Such phenomenon is usually referred to as exponential correlation decay. Before we prove the main theorem, we will introduce a few useful propositions and lemmas.

Proposition 8.

$$P(G, e) \le \frac{1}{2}$$

Proof. Although one may examine this case by case algebraically, this propositions can be seen quite obvious combinatorially in that, for any edge cover $X \in EC(G)$ s.t. $e \notin X$, X + e is also

an edge cover in G, and $\forall X, Y \in EC(G)$ s.t. $X \neq Y, e \notin X, e \notin Y$, we have $X + e \neq Y + e$. So the edge covers with e chosen is at least as many as the edge covers with e not chosen, hence the proposition follows.

We remark that our algorithm also guarantees that P(G, e, L), since $\frac{1-\prod_i x_i}{2-\prod_i x_i} = \frac{1}{2} - \frac{\prod_i x_i}{2(2-\prod_i x_i)}$, and $X \cdot Y - X - Z \leq 0$.

For notational convenience, given a d-dimensional vector $\mathbf{x} \in [0, \frac{1}{2}]^d$, we denote

$$f(\mathbf{x}) \triangleq \frac{1 - \prod_i x_i}{2 - \prod_i x_i}$$

Given a d_1 -dimensional vector $\mathbf{x} \in [0, \frac{1}{2}]^{d_1}$ and two d_2 -dimensional vectors $\mathbf{y}, \mathbf{z} \in [0, \frac{1}{2}]^{d_2}$, let

$$g(\mathbf{x}, \mathbf{y}, \mathbf{z}) \triangleq 1 - \frac{1}{2 + \prod_{i} x_{i} \cdot \prod_{i} y_{i} - \prod_{i} x_{i} - \prod_{i} z_{i}}$$

Lemma 9. For d-variate function $f(\mathbf{x})$, and a d-dimensional vector $\mathbf{x} \in [0, \frac{1}{2}]^d$,

$$\left| \sum_{i}^{d} \left| \frac{\partial f(\mathbf{x})}{\partial x_{i}} \right| \leq \min \left\{ \frac{1}{2}, d \left(\frac{1}{2} \right)^{d-1} \right\} \right|$$

Proof. Denote i^* be one of the indices of smallest x_i , since $x_i \leq \frac{1}{2}$, we have

$$\sum_{i}^{d} \left| \frac{\partial f(\mathbf{x})}{\partial x_{i}} \right| = \sum_{i}^{d} \frac{\prod_{k \neq i}^{d} x_{k}}{(2 - \prod_{i} x_{i})^{2}}$$

$$\leq d \prod_{k \neq i^{*}}^{d} x_{k}$$

$$\leq d \left(\frac{1}{2}\right)^{d-1}$$

So for $d \ge 4$ we have $\sum_i \left| \frac{\partial f(\mathbf{x})}{\partial x_i} \right| \le \frac{1}{2}$.

For
$$d = 0, \sum_{i} \left| \frac{\partial f(\mathbf{x})}{\partial x_i} \right| = 0.$$

Now consider
$$d = 1$$
, $\sum_{i} \left| \frac{\partial f(\mathbf{x})}{\partial x_{i}} \right| = \frac{1}{(2-x_{1})^{2}} \leq \frac{4}{9}$.

Next consider
$$d = 2$$
, $\sum_{i} \left| \frac{\partial f(\mathbf{x})}{\partial x_i} \right| = \frac{x_1 + x_2}{(2 - x_1 x_2)^2} \le \frac{16}{49}$.

Finally for
$$d = 3$$
, $\sum_{i} \left| \frac{\partial f(\mathbf{x})}{\partial x_{i}} \right| = \frac{x_{1} + x_{2} + x_{3}}{(2 - x_{1} x_{2} x_{3})^{2}} \le \frac{32}{75}$.

Lemma 10. Given a d_1 -dimensional vector $\mathbf{x} \in [0, \frac{1}{2}]^{d_1}$ and two d_2 -dimensional vectors $\mathbf{y}, \mathbf{z} \in [0, \frac{1}{2}]^{d_2}$,

$$\sum_{i} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial x_{i}} \right| \leq 1$$

$$\sum_{i} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial y_{i}} \right| \leq 1$$

$$\sum_{i} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial z_{i}} \right| \leq 1$$

Proof.

$$\sum_{k}^{d_{1}} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial x_{k}} \right| = \sum_{k}^{d_{1}} \frac{\prod_{i \neq k}^{d_{1}} x_{i} \cdot \left(1 - \prod_{i}^{d_{2}} y_{i}\right)}{(2 + \prod_{i} x_{i} \cdot \prod_{i} y_{i} - \prod_{i} x_{i} - \prod_{i} z_{i})^{2}} \\
\leq d_{1} \frac{1}{2^{d_{1} - 1}} \leq 1$$

$$\sum_{k}^{d_{2}} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial y_{k}} \right| = \sum_{k}^{d_{2}} \frac{\prod_{i \neq k}^{d_{2}} y_{i} \cdot \prod_{i}^{d_{1}} x_{i}}{(2 + \prod_{i} x_{i} \cdot \prod_{i} y_{i} - \prod_{i} x_{i} - \prod_{i} z_{i})^{2}} \\
\leq d_{2} \frac{1}{2^{d_{1} + d_{2} - 1}} \leq 1$$

$$\sum_{k}^{d_{2}} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial z_{k}} \right| = \sum_{k}^{d_{2}} \frac{\prod_{i \neq k}^{d_{2}} z_{i}}{(2 + \prod_{i} x_{i} \cdot \prod_{i} y_{i} - \prod_{i} x_{i} - \prod_{i} z_{i})^{2}} \\
\leq d_{2} \frac{1}{2^{d_{2} - 1}} \leq 1$$

Now we are ready for the main theorem.

Proof. First we note that by Mean Value Theorem, given estimated $\hat{\mathbf{x}}$, let \mathbf{x} be the true value, let $\epsilon = \max_i |x_i - \hat{x_i}|$, we have for d-variate function f,

$$|f(\hat{\mathbf{x}}) - f(\mathbf{x})| \le \sum_{i} \left| \frac{\partial f(\mathbf{x})}{\partial x_{i}} \right| \cdot \epsilon \le d \cdot \left(\frac{1}{2}\right)^{d-1} \cdot \epsilon \tag{4}$$

Given estimated $\hat{\mathbf{x}}, \hat{\mathbf{y}}, \hat{\mathbf{z}}$, let $\epsilon = \max\{|x_i - \hat{x}_i|, |y_i - \hat{y}_i|, |z_i - \hat{z}_i|\}.$

$$|g(\hat{\mathbf{x}}, \hat{\mathbf{y}}, \hat{\mathbf{z}}) - g(\mathbf{x}, \mathbf{y}, \mathbf{z})| \le \left(\sum_{i} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial x_{i}} \right| + \sum_{i} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial y_{i}} \right| + \sum_{i} \left| \frac{\partial g(\mathbf{x}, \mathbf{y}, \mathbf{z})}{\partial z_{i}} \right| \right) \cdot \epsilon \le 3\epsilon$$
(5)

Also note that the recursion for normal edge case is applied only once, so it's sufficient to show for free or dangling edge e,

$$|P(G, e, L) - P(G, e)| \le (\frac{1}{2})^{L+1}$$
, for free or dangling edge e

then the case of normal edge automatically follows from 5 that

$$|P(G, e, L) - P(G, e)| \le 3 \cdot (\frac{1}{2})^{L+1}$$
, for normal edge e

Now we prove by induction with induction hypothesis:

$$|P(G, e, L) - P(G, e)| \le (\frac{1}{2})^{L+1}$$
, for free or dangling edge e

For base case $L = 0, |P(G, e, L) - P(G, e)| \le \frac{1}{2}$ holds when e is free or dangling.

Now suppose for L < k we have the induction hypothesis true, now we try to show it's true for L = k.

Case 1, e is free edge, then |P(G, e, L) - P(G, e)| = 0.

Case 2, $e=(u,\underline{\ })$ is dangling with deg(u)=d+1, then by induction hypothesis we have $\epsilon \leq \frac{1}{2}^{L-\lceil \log_6 d+1 \rceil}$. First we need to show that for $d \leq 4$,

$$\frac{1}{2^{1+L-\lceil \log_6\left(d+1\right)\rceil}} \le \frac{1}{2^L}$$

which is obvious because $\lceil \log_6 (d+1) \rceil \le 1$.

Next we show for $d \geq 5$,

$$d \cdot \left(\frac{1}{2}\right)^{d-1 + L - \lceil \log_6\left(d+1\right) \rceil} \leq \left(\frac{1}{2}\right)^L$$

Namely for $d \geq 5$,

$$\log_2 d + \lceil \log_6 (d+1) \rceil \le d - 1$$

For d = 5, 6, one can directly examine that as $\log_2 d < 3$ and $\log_6 6 = 1, \log_6 7 < 2$.

For $d \geq 7$, note that the function $f(x) = d - 2 - \log_2 d - \log_6 (d+1)$ is monotonically increasing, and f(7) > 0, so we have

$$\log_2 d + \log_6 (d+1) + 1 \le d - 1$$

Therefore, the hypothesis for L = k is verified.

$$|P(G, e, L) - P(G, e)| \le (\frac{1}{2})^{L+1}$$
, for free or dangling edge e

To sum up, the case of free or dangling edge and the case of normal edge together conclude the proof for our main theorem.

6 Counting Edge Covers

Finally, we present the procedures for approximately counting edge covers given good estimations of the marginal probability P(G,e), hence an FPTAS for the approximate counting of edge covers problem.

Proposition 11. Let $Z(G) \triangleq ||EC(G)||$, order the set of edges in G in any order as $\{e_i\}$, define $G_1 \triangleq G, G_i \triangleq G_{i-1} - e_{i-1} - u_{i-1} - v_{i-1}, 1 < i \le m.$

$$Z(G) = \frac{1}{\prod_{i=1}^{m} (1 - P(G_i, e_i))}$$

Proof. Note that $EC(G) \neq \emptyset$, since the set of all edges E is an edge cover.

Now with $X \sim EC(G)$ uniformly, $\mathbb{P}(X = E)$ has two equivalent expressions,

$$\mathbb{P}(X = E) = \frac{1}{Z(G)}$$

$$\mathbb{P}(X = E) = \prod_{i} \mathbb{P}\left(e_i = 1 \mid \{e_j\}_{j=1}^{i-1} = \mathbf{1}\right)$$

$$= \prod_{i} (1 - P(G_i, e_i))$$

Therefore we have

$$Z(G) = \frac{1}{\prod_{i=1}^{m} (1 - P(G_i, e_i))}$$

We now show the main theorem of this section. Define $Z(G, L) \triangleq \frac{1}{\prod_{i=1}^{m} (1 - P(G_i, e_i, L))}$ as the estimated number of edge covers given estimated $P(G_i, e_i, L)$

Theorem 12. For $0 < \epsilon < 1$, take $L = \log m + \log(8/\epsilon)$,

$$1 - \epsilon \le \frac{Z(G)}{Z(G, L)} \le 1 + \epsilon$$

Proof.

$$\frac{Z(G)}{Z(G,L)} = \prod_{i=1}^{m} \frac{1 - P(G_i, e_i, L)}{1 - P(G_i, e_i)}$$

By Theorem 7,

$$|P(G_i, e_i, L) - P(G_i, e_i)| \le \frac{\epsilon}{4m}$$

Since $1 - P(G_i, e_i) \ge \frac{1}{2}$,

$$\frac{|P(G_i, e_i, L) - P(G_i, e_i)|}{1 - P(G_i, e_i)} \le \frac{\epsilon}{2m}$$

Namely $\forall i$,

$$\left(1 - \frac{\epsilon}{2m}\right) \le \frac{1 - P(G_i, e_i, L)}{1 - P(G_i, e_i)} \le \left(1 + \frac{\epsilon}{2m}\right)$$

So we have

$$\left(1 - \frac{\epsilon}{2m}\right)^m \le \prod_{i=1}^m \frac{1 - P(G_i, e_i, L)}{1 - P(G_i, e_i)} \le \left(1 + \frac{\epsilon}{2m}\right)^m$$
$$1 - \epsilon \le \frac{Z(G)}{Z(G, L)} \le 1 + \epsilon$$

To sum up, run Z(G, L) with $L = \log m + \log(8/\epsilon)$, is the FPTAS for counting edge covers, the total running time is $O(m \cdot n \cdot (m \cdot \frac{1}{\epsilon})^{\log_2 6})$.

References

- [BG08] Antar Bandyopadhyay and David Gamarnik, Counting without sampling: Asymptotics of the log-partition function for certain statistical physics models, Random Structures & Algorithms 33 (2008), no. 4, 452–479.
- [BGK⁺07] Mohsen Bayati, David Gamarnik, Dimitriy Katz, Chandra Nair, and Prasad Tetali, Simple deterministic approximation algorithms for counting matchings, Proceedings of STOC, 2007, pp. 122–127.
- [BR09] Ivona Bezáková and William A Rummler, Sampling edge covers in 3-regular graphs, Mathematical Foundations of Computer Science 2009, Springer, 2009, pp. 137–148.
- [CGW13] Jin-Yi Cai, Heng Guo, and Tyson Williams, A complete dichotomy rises from the capture of vanishing signatures: extended abstract, STOC, 2013, pp. 635–644.
- [CLX09] Jin-Yi Cai, Pinyan Lu, and Mingji Xia, *Holant problems and counting CSP*, Proceedings of the 41st annual ACM symposium on Theory of computing (New York, NY, USA), STOC '09, ACM, 2009, pp. 715–724.

- [CLX11] Jin-Yi Cai, Pinyan Lu, and Mingji Xia, Computational complexity of holant problems, SIAM J. Comput. 40 (2011), no. 4, 1101–1132.
- [DGGJ00] Martin E. Dyer, Leslie Ann Goldberg, Catherine S. Greenhill, and Mark Jerrum, On the relative complexity of approximate counting problems, The 13th International Workshop on Approximation Algorithms for Combinatorial Optimization Problems (Klaus Jansen and Samir Khuller, eds.), APPROX '00, vol. 1913, Springer, 2000, pp. 108–119.
- [DGJ10] Martin E. Dyer, Leslie Ann Goldberg, and Mark Jerrum, *An approximation tri*chotomy for boolean #csp, Journal of Computer and System Sciences **76** (2010), no. 3-4, 267–277.
- [GK10] David Gamarnik and Dmitriy Katz, A deterministic approximation algorithm for computing the permanent of a 0, 1 matrix, Journal of Computer and System Sciences 76 (2010), no. 8, 879–883.
- [GMR12] Parikshit Gopalan, Raghu Meka, and Omer Reingold, *Dnf sparsification and a faster deterministic counting algorithm*, Computational Complexity (CCC), 2012 IEEE 27th Annual Conference on, IEEE, 2012, pp. 126–135.
- [GSV12] A. Galanis, D. Stefankovic, and E. Vigoda, Inapproximability of the partition function for the antiferromagnetic ising and hard-core models, Arxiv preprint arXiv:1203.2226 (2012).
- [HL12] Sangxia Huang and Pinyan Lu, A dichotomy for real weighted holant problems, IEEE Conference on Computational Complexity, 2012, pp. 96–106.
- [JS89] Mark Jerrum and Alistair Sinclair, Approximating the permanent, SIAM journal on computing 18 (1989), no. 6, 1149–1178.
- [JS97] _____, The markov chain monte carlo method: an approach to approximate counting and integration, pp. 482–520, PWS Publishing Co., Boston, MA, USA, 1997.
- [JSV04] Mark Jerrum, Alistair Sinclair, and Eric Vigoda, A polynomial-time approximation algorithm for the permanent of a matrix with nonnegative entries, Journal of the ACM 51 (2004), 671–697.
- [KL83] Richard M. Karp and Michael Luby, Monte-carlo algorithms for enumeration and reliability problems, FOCS, 1983, pp. 56–64.
- [KLM89] Richard M. Karp, Michael Luby, and Neal Madras, *Monte-carlo approximation algorithms for enumeration problems*, J. Algorithms **10** (1989), no. 3, 429–448.
- [LLY12] Liang Li, Pinyan Lu, and Yitong Yin, Approximate counting via correlation decay in spin systems, Proceedings of SODA, 2012, pp. 922–940.
- [LLY13] _____, Correlation decay up to uniqueness in spin systems, Proceedings of SODA, 2013, pp. 67–84.
- [LSW98] Nathan Linial, Alex Samorodnitsky, and Avi Wigderson, A deterministic strongly polynomial algorithm for matrix scaling and approximate permanents, Proceedings of the thirtieth annual ACM symposium on Theory of computing, ACM, 1998, pp. 644–652.
- [LWZ] Pinyan Lu, Menghui Wang, and Chihao Zhang, FPTAS for weighted fibonacci gates and its applications, Submitted to SODA 2014.

- [McQ13] Colin McQuillan, Approximating holant problems by winding, CoRR abs/1301.2880 (2013).
- [SS12] Allan Sly and Nike Sun, The computational hardness of counting in two-spin models on d-regular graphs, Proceedings of FOCS, 2012, pp. 361–369.
- [SST12] Alistair Sinclair, Piyush Srivastava, and Marc Thurley, Approximation algorithms for two-state anti-ferromagnetic spin systems on bounded degree graphs, Proceedings of SODA, 2012, pp. 941–953.
- [Tre04] Luca Trevisan, A note on approximate counting for k-dnf, APPROX-RANDOM, 2004, pp. 417–426.
- [Wei06] Dror Weitz, Counting independent sets up to the tree threshold, Proceedings of STOC, 2006, pp. 140–149.
- [YZ13] Yitong Yin and Chihao Zhang, Approximate counting via correlation decay on planar graphs, Proceedings of SODA, 2013, pp. 47–66.