

# deliverable2\_final

## Data Importing

```
library(tidyverse)
```

```
## -- Attaching packages ----- tidy
## v ggplot2 3.1.0      v purrr  0.2.5
## v tibble  2.0.1      v dplyr  0.7.8
## v tidyr   0.8.2      v stringr 1.4.0
## v readr   1.3.1      v forcats 0.4.0

## -- Conflicts -----
## x dplyr::filter() masks stats::filter()
## x dplyr::lag()     masks stats::lag()
```

```
library(car)
```

```
## Loading required package: carData
##
## Attaching package: 'car'
##
## The following object is masked from 'package:dplyr':
##
##     recode
##
## The following object is masked from 'package:purrr':
##
##     some
```

```
library(Hmisc)
```

```
## Loading required package: lattice
## Loading required package: survival
## Loading required package: Formula
##
## Attaching package: 'Hmisc'
##
## The following objects are masked from 'package:dplyr':
##
##     src, summarize
##
## The following objects are masked from 'package:base':
##
##     format.pval, units
```

```
#library(plotly)
(grad <- read_csv("data/grad.csv",
  col_types = cols_only(
    uni_name=col_character(),
    major=col_character(),
    degree=col_character(),
    season=col_character(),
```

```

decision=col_character(),
decision_date=col_character(),
decision_timestamp=col_double(),
ugrad_gpa=col_double(),
gre_verbal=col_double(),
gre_quant=col_double(),
gre_writing=col_double(),
is_new_gre=col_logical(),
status=col_character(),
comments=col_character()))))

## # A tibble: 345,303 x 14
##   uni_name major degree season decision decision_date decision_timest~
##   <chr>      <chr> <chr> <chr> <chr>      <chr>              <dbl>
## 1 Univers~ Ms. ~ MS      S16    Accepted (5, 11, 2015)      1446699600
## 2 Vanderb~ Educ~ MS      F16    Other    <NA>              NA
## 3 Univers~ Publ~ MS      F16    Accepted (16, 11, 201~    1447650000
## 4 Tufts U~ Comp~ PhD     S16    Accepted (16, 11, 201~    1447650000
## 5 Univers~ Theo~ MS      F16    Accepted (16, 11, 201~    1447650000
## 6 Univers~ Mast~ MS      S16    Rejected (14, 11, 201~    1447477200
## 7 Univers~ Publ~ MS      F16    Accepted (12, 11, 201~    1447304400
## 8 Tufts U~ MALD  MS      S16    Accepted (7, 11, 2015)    1446872400
## 9 New Yor~ Fina~ MS      S16    Accepted (15, 11, 201~    1447563600
## 10 Appalac~ Comm~ MS      S16    Accepted (13, 11, 201~    1447390800
## # ... with 345,293 more rows, and 7 more variables: ugrad_gpa <dbl>,
## #   gre_verbal <dbl>, gre_quant <dbl>, gre_writing <dbl>,
## #   is_new_gre <lgl>, status <chr>, comments <chr>

```

```
problems(grad)
```

```

## [1] row      col      expected actual
## <0 rows> (or 0-length row.names)

```

## Data Cleaning

```

grad <- grad %>% filter(str_detect(major, "Computer")|str_detect(major, "computer"))

# ,degree=="PhD") %>% mutate(research=(str_detect(comments, "research")|str_detect(comments, "Research"))
grad %>% group_by(uni_name) %>% count(uni_name) %>% arrange(desc(n))

## # A tibble: 416 x 2
## # Groups:   uni_name [416]
##   uni_name          n
##   <chr>          <int>
## 1 Carnegie Mellon University (CMU)      1414
## 2 Georgia Institute Of Technology (GTech)    985
## 3 University Of California, San Diego (UCSD)   954
## 4 University Of Illinois, Urbana-Champaign (UIUC) 954
## 5 Stanford University                    914
## 6 University Of California, Berkeley (UCB)     844
## 7 Purdue University                      745
## 8 University Of Washington, Seattle (UW)       713
## 9 University Of Texas, Austin (UT Austin)      706

```

```
## 10 Cornell University                                674
## # ... with 406 more rows

grad %>% group_by(uni_name,major) %>% count(uni_name, major) %>% arrange(desc(n))

## # A tibble: 2,749 x 3
## # Groups:   uni_name, major [2,749]
##   uni_name                major          n
##   <chr>                  <chr>        <int>
## 1 Carnegie Mellon University (CMU)    Computer Science    776
## 2 Stanford University                Computer Science    765
## 3 Georgia Institute Of Technology (GTech) Computer Science    610
## 4 Columbia University                Computer Science    585
## 5 University Of Illinois, Urbana-Champaign (UIUC) Computer Science    582
## 6 University Of Washington, Seattle (UW) Computer Science    581
## 7 University Of California, San Diego (UCSD) Computer Science    530
## 8 University Of California, Berkeley (UCB) Computer Science    521
## 9 University Of California, Los Angeles (UCLA) Computer Science    498
## 10 Cornell University                Computer Science    478
## # ... with 2,739 more rows
```

## Top 10 Dataset

```
grad1 <- grad %>% group_by(uni_name) %>% filter(decision == "Accepted") %>% count(uni_name) %>% arrange(desc(n))
grad2 <- grad %>% group_by(uni_name) %>% count(uni_name) %>% arrange(desc(n))
colnames(grad1)[2] = "accepted"
(top10 <- merge(grad1,grad2,by =("uni_name")) %>% mutate(rate = accepted/n) %>% filter(n>100) %>% arrange(desc(rate)))

##           uni_name accepted    n    rate
## 1 Carnegie Mellon University (CMU)    523 1414 0.3698727
## 2 Georgia Institute Of Technology (GTech)    413  985 0.4192893
## 3 University Of California, San Diego (UCSD)    349  954 0.3658281
## 4 University Of Illinois, Urbana-Champaign (UIUC)    367  954 0.3846960
## 5 Stanford University    245  914 0.2680525
## 6 University Of California, Berkeley (UCB)    155  844 0.1836493
## 7 Purdue University    335  745 0.4496644
## 8 University Of Washington, Seattle (UW)    167  713 0.2342216
## 9 University Of Texas, Austin (UT Austin)    282  706 0.3994334
## 10 Cornell University    253  674 0.3753709

grad3 <- grad %>% group_by(uni_name,major) %>% filter(decision == "Accepted") %>% count(uni_name,major)
grad4 <- grad %>% group_by(uni_name,major) %>% count(uni_name,major) %>% arrange(desc(n))
colnames(grad3)[3] = "accepted"
merge(grad3,grad4,by=c("uni_name","major")) %>% mutate(rate = accepted/n) %>% arrange(desc(n)) %>% head(10)

##           uni_name                major
## 1 Carnegie Mellon University (CMU)    Computer Science
## 2 Stanford University                Computer Science
## 3 Georgia Institute Of Technology (GTech) Computer Science
## 4 Columbia University                Computer Science
## 5 University Of Illinois, Urbana-Champaign (UIUC) Computer Science
## 6 University Of Washington, Seattle (UW) Computer Science
## 7 University Of California, San Diego (UCSD) Computer Science
## 8 University Of California, Berkeley (UCB) Computer Science
```

```

## 9      University Of California, Los Angeles (UCLA) Computer Science
## 10      Cornell University Computer Science
##      accepted    n      rate
## 1      238 776 0.3067010
## 2      198 765 0.2588235
## 3      247 610 0.4049180
## 4      216 585 0.3692308
## 5      238 582 0.4089347
## 6      132 581 0.2271945
## 7      192 530 0.3622642
## 8       83 521 0.1593090
## 9      194 498 0.3895582
## 10     164 478 0.3430962

merge(grad1,grad2,by =("uni_name")) %>% mutate(rate = accepted/n) %>% filter(uni_name == "Boston University")

##      uni_name accepted    n      rate
## 1 Boston University (BU)      79 184 0.4293478

merge(grad3,grad4,by=c("uni_name","major")) %>% mutate(rate = accepted/n) %>% filter(uni_name == "Boston University")

##      uni_name      major
## 1 Boston University (BU) (ECE) Electrical And Computer Engineering
## 2 Boston University (BU)      Computer Science
## 3 Boston University (BU)      Electrical And Computer Engineering
## 4 Boston University (BU)      ECE (Electrical & Computer Engineering)
## 5 Boston University (BU) ECE (Electrical And Computer Engineering)
## 6 Boston University (BU)      ECE(Electrical And Computer Engineering)
## 7 Boston University (BU) Electrical And Computer Engineering (ECE)
## 8 Boston University (BU)      Computer Engineering
## 9 Boston University (BU) Electrical and Computer Engineering (ECE)
## 10 Boston University (BU)      Electrical & Computer Engineering
##      accepted    n      rate
## 1         2     9 0.2222222
## 2        38   101 0.3762376
## 3         6    13 0.4615385
## 4         5    10 0.5000000
## 5         2     4 0.5000000
## 6         1     2 0.5000000
## 7         1     2 0.5000000
## 8        10    18 0.5555556
## 9         5     9 0.5555556
## 10        4     6 0.6666667

#filter for top 10 schools by ranking
top10 <- head(top10,10)$uni_name
grad <- subset(grad, uni_name %in% top10)
#filter
grad <- grad[complete.cases(grad), ]

```

## Modeling

```

grad <- grad[complete.cases(grad[, -14]), ] %>% filter(is_new_gre == TRUE, ugrad_gpa <=4, status!="Other")
# models

```

```
full_mod_int <- glm(decision1 ~ (ugrad_gpa+GRE_Total+gre_writing)*status -1, data = grad, family = binomial)
(gradmodel_int <- step(full_mod_int))
```

```
## Start: AIC=1459.38
## decision1 ~ (ugrad_gpa + GRE_Total + gre_writing) * status -
## 1
##
##           Df Deviance   AIC
## - ugrad_gpa:status    2   1436.5 1456.5
## - GRE_Total:status    2   1437.5 1457.5
## - gre_writing:status  2   1438.4 1458.4
## <none>                 1435.4 1459.4
##
## Step: AIC=1456.54
## decision1 ~ ugrad_gpa + GRE_Total + gre_writing + status + GRE_Total:status +
## gre_writing:status - 1
##
##           Df Deviance   AIC
## - GRE_Total:status    2   1438.4 1454.4
## - gre_writing:status  2   1439.3 1455.3
## <none>                 1436.5 1456.5
## - ugrad_gpa           1   1454.0 1472.0
##
## Step: AIC=1454.44
## decision1 ~ ugrad_gpa + GRE_Total + gre_writing + status + gre_writing:status -
## 1
##
##           Df Deviance   AIC
## <none>                 1438.4 1454.4
## - gre_writing:status  2   1444.7 1456.7
## - GRE_Total           1   1449.8 1463.8
## - ugrad_gpa           1   1455.9 1469.9
##
## Call: glm(formula = decision1 ~ ugrad_gpa + GRE_Total + gre_writing +
## status + gre_writing:status - 1, family = binomial, data = grad)
##
## Coefficients:
##                ugrad_gpa
##                1.14664
##                GRE_Total
##                0.03128
##                gre_writing
##               -0.26817
##               statusAmerican
##              -13.40324
##               statusInternational
##              -12.78241
## statusInternational with US Degree
##              -15.54470
## gre_writing:statusInternational
##              -0.25273
## gre_writing:statusInternational with US Degree
##               0.54078
```

```
##
## Degrees of Freedom: 1091 Total (i.e. Null); 1083 Residual
## Null Deviance: 1512
## Residual Deviance: 1438 AIC: 1454
summary(gradmodel_int)

##
## Call:
## glm(formula = decision1 ~ ugrad_gpa + GRE_Total + gre_writing +
##      status + gre_writing:status - 1, family = binomial, data = grad)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.5368  -1.0575  -0.8203   1.2008   1.9261
##
## Coefficients:
##                                Estimate Std. Error
## ugrad_gpa                    1.146643    0.283057
## GRE_Total                    0.031278    0.009388
## gre_writing                  -0.268174    0.189752
## statusAmerican              -13.403241    2.973278
## statusInternational         -12.782405    2.852552
## statusInternational with US Degree -15.544697    3.081698
## gre_writing:statusInternational -0.252731    0.225536
## gre_writing:statusInternational with US Degree 0.540781    0.358197
##                                z value Pr(>|z|)
## ugrad_gpa                    4.051 5.10e-05 ***
## GRE_Total                    3.332 0.000863 ***
## gre_writing                  -1.413 0.157571
## statusAmerican              -4.508 6.55e-06 ***
## statusInternational         -4.481 7.43e-06 ***
## statusInternational with US Degree -5.044 4.55e-07 ***
## gre_writing:statusInternational -1.121 0.262468
## gre_writing:statusInternational with US Degree 1.510 0.131112
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 1512.4  on 1091  degrees of freedom
## Residual deviance: 1438.4  on 1083  degrees of freedom
## AIC: 1454.4
##
## Number of Fisher Scoring iterations: 4
full_mod <- glm(decision1 ~ ugrad_gpa+GRE_Total+gre_writing+status-1, data = grad, family = binomial)
(gradmodel <- step(full_mod))

## Start: AIC=1456.66
## decision1 ~ ugrad_gpa + GRE_Total + gre_writing + status - 1
##
##              Df Deviance    AIC
## <none>          1444.7 1456.7
## - gre_writing  1  1455.5 1465.5
```

```
## - GRE_Total      1    1455.7 1465.7
## - ugrad_gpa      1    1462.8 1472.8
## - status         3    1477.0 1483.0

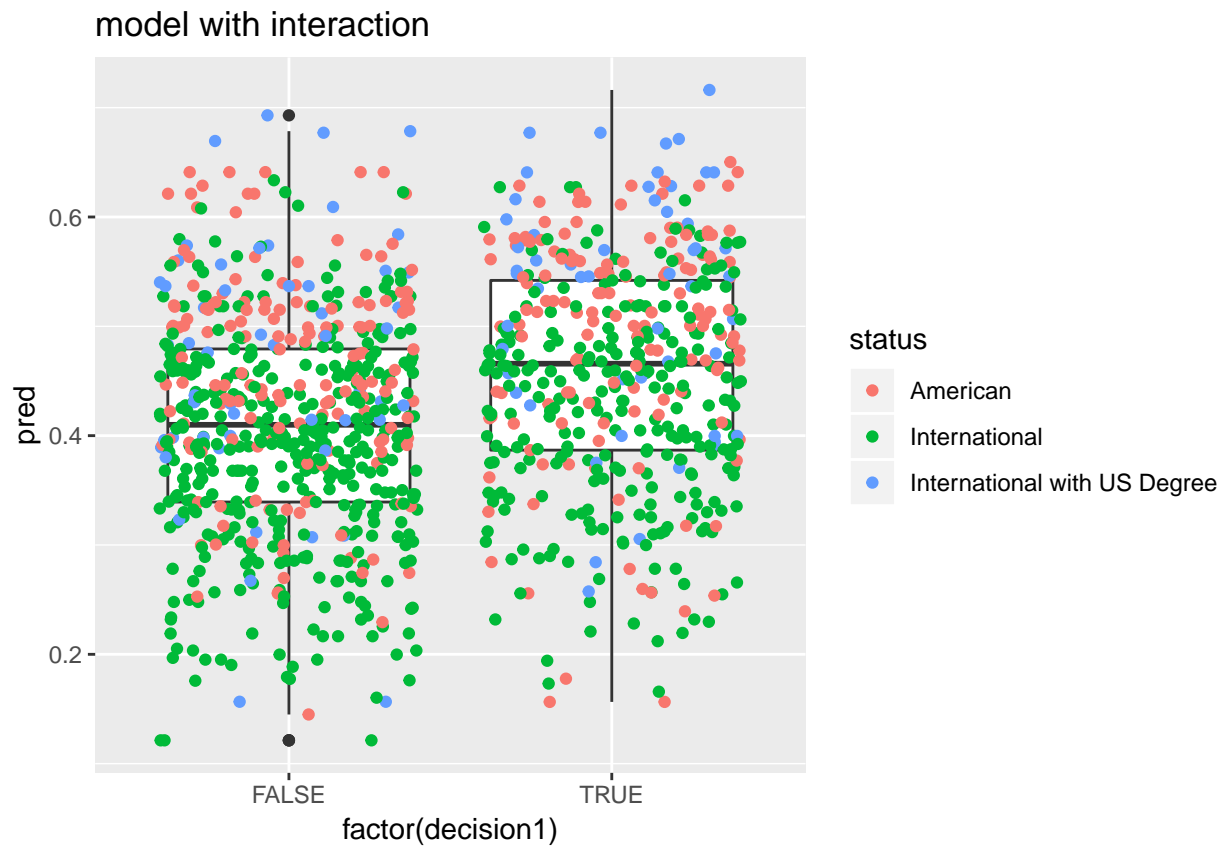
##
## Call:  glm(formula = decision1 ~ ugrad_gpa + GRE_Total + gre_writing +
##          status - 1, family = binomial, data = grad)
##
## Coefficients:
##                ugrad_gpa                GRE_Total
##                1.16848                0.03074
##                gre_writing                statusAmerican
##                -0.35978                -12.89275
##                statusInternational statusInternational with US Degree
##                -13.30241                -12.98166
##
## Degrees of Freedom: 1091 Total (i.e. Null);  1085 Residual
## Null Deviance:      1512
## Residual Deviance: 1445  AIC: 1457
```

```
summary(gradmodel)
```

```
##
## Call:
## glm(formula = decision1 ~ ugrad_gpa + GRE_Total + gre_writing +
##          status - 1, family = binomial, data = grad)
##
## Deviance Residuals:
##      Min       1Q   Median       3Q      Max
## -1.4334  -1.0585  -0.8327   1.2071   1.9526
##
## Coefficients:
##                Estimate Std. Error z value Pr(>|z|)
## ugrad_gpa          1.168482    0.283037   4.128 3.65e-05
## GRE_Total          0.030744    0.009342   3.291 0.000998
## gre_writing        -0.359779    0.110128  -3.267 0.001087
## statusAmerican     -12.892745    2.846590  -4.529 5.92e-06
## statusInternational -13.302409    2.834117  -4.694 2.68e-06
## statusInternational with US Degree -12.981663    2.846232  -4.561 5.09e-06
##
## ugrad_gpa          ***
## GRE_Total          ***
## gre_writing        **
## statusAmerican     ***
## statusInternational ***
## statusInternational with US Degree ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for binomial family taken to be 1)
##
##      Null deviance: 1512.4  on 1091  degrees of freedom
## Residual deviance: 1444.7  on 1085  degrees of freedom
## AIC: 1456.7
##
```

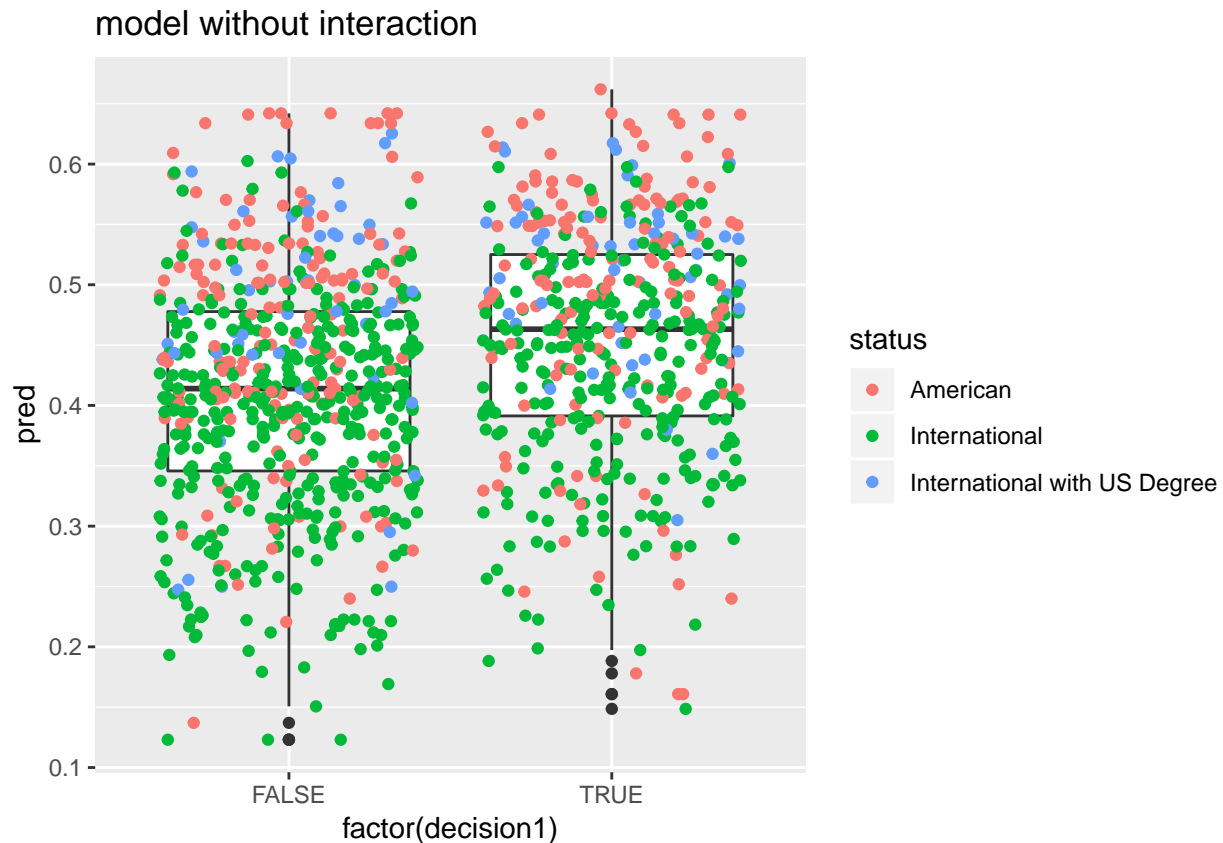
```
## Number of Fisher Scoring iterations: 4
null_mod <- glm(decision1 ~ 1, data = grad, family=binomial)

grad %>%
  mutate(pred = predict(gradmodel_int,
    type = "response")) %>%
  ggplot(aes(factor(decision1), pred)) +
    geom_boxplot() +
    geom_point(aes(color = status),
      position = "jitter") +
    labs(title = "model with interaction")
```



```
grad %>%
  mutate(pred = predict(gradmodel,
    type = "response")) %>%
  ggplot(aes(factor(decision1), pred)) +
    geom_boxplot() +
    geom_point(aes(color = status),
      position = "jitter") +
    labs(title = "model without interaction")
```





```
gg_int <- grad %>%
  mutate(pred = predict(gradmodel_int,
    type = "response")) #>% select(decision, pred)
gg <- grad %>%
  mutate(pred = predict(gradmodel,
    type = "response")) #>% select(decision, pred)
```

## Distribution of the Respond

In model without intarnation distribution, the mean of predicted probabilities of rejected students is around 0.41, while the mean of predicted probabilities of accepted students is around 0.43, which is slightly higher than the mean of predicted probabilities of rejected students. Since their interval are overlapped, it means that the prediction may not be significant enough to explain the success of a student being accepted. In addition, the plots are fairly scattered, meaning that there does no exist a certain pattern to explain the trend.

In model with interaction, the mean of predicted probabilities of rejected students is around 0.4, while the mean of predicted probabilities of accepted students is around 0.45, which is slightly higher than the mean of predicted probabilities of rejected students. Since their interval are overlapped, it means that the prediction may not be significant enough to explain the success of a student being accepted. However, the plots are more densely concentrated than the one without interaction.

## Coefficient Interpretation

For the model that includes interaction: The regression coefficient for `ugrad_gpa` is  $\beta(\hat{ugradgpa}) = 1.146643$  meaning that for a one-unit increase in undergraduate gpa the logit-transformed probability of getting accepted to the program will increase by 1.15. Predictor `GRE_Total` has a coefficient  $\beta(\hat{GREtotal}) = 0.031106$ , showing that for a one-unit increase in GRE total scores the log odds will increase by 0.03. We also include categorical variable `status` representing applicant's identity. The corresponding coefficient  $\beta(\hat{American}) = -13.403241$  shows that if the applicant is an American student, the log odds will decrease by 13.4, holding all other independent variables constant,  $\beta(\hat{International}) = -12.782405$  shows the change in log odds given the student is an international student, and  $\beta(\hat{USdegree}) = -15.544697$  shows the change in log odds given the student is an international student with a US degree.

$\beta(\hat{GREwriting}) = -0.267686$  is the regression coefficients for GRE writing score, and  $\beta(\hat{GREwriting:International}) = -0.252731$  and  $\beta(\hat{GREwriting:USdegree}) = 0.540781$  are the coefficients of GRE writing scores with respect to student's status. However, the hypothesis tests for coefficient indicates that those terms would not significantly impact the prediction of our model.

```
# prediction of model with interaction term
(mod_coef <- coef(gradmodel_int))
```

```
##                ugrad_gpa
##                1.14664296
##                GRE_Total
##                0.03127823
##                gre_writing
##                -0.26817442
##                statusAmerican
##                -13.40324104
##                statusInternational
##                -12.78240546
##                statusInternational with US Degree
##                -15.54469677
##                gre_writing:statusInternational
##                -0.25273087
##                gre_writing:statusInternational with US Degree
##                0.54078123
```

```
prediction_american <- mod_coef[1]*mean(grad$ugrad_gpa)+mod_coef[2]*mean(grad$GRE_Total)+mod_coef[3]*mean(
exp(prediction_american) / (1 + exp(prediction_american)))
```

```
## ugrad_gpa
## 0.4785392
```

```
prediction_inter <- mod_coef[1]*mean(grad$ugrad_gpa)+mod_coef[2]*mean(grad$GRE_Total)+mod_coef[3]*mean(
exp(prediction_inter) / (1 + exp(prediction_inter)))
```

```
## ugrad_gpa
## 0.5700877
```

```
prediction_inter_us <- mod_coef[1]*mean(grad$ugrad_gpa)+mod_coef[2]*mean(grad$GRE_Total)+mod_coef[3]*mean(
exp(prediction_inter_us) / (1 + exp(prediction_inter_us)))
```

```
## ugrad_gpa
## 0.1562274
```

Using our model that includes the interaction between student's status and GRE writing score, we use mean GPA, GRE total score and writing score to compute the probability of a student getting accepted. There's

47.9% chance that the student will be admitted to the program if the student is an American student, and 57% and 15.6% respectively if the student is an international student or international student with a US degree.

For the model that does not include interaction terms: The regression coefficient for `ugrad_gpa` is  $\beta(\hat{ugradgpa}) = 1.168482$ , which indicates that for a one-unit increase in undergraduate gpa the logit-transformed probability of getting accepted to the program will increase by 1.15.  $\beta(\hat{GREtotal}) = 0.030744$  is the coefficient for predictor `GRE_Total` showing that for a one-unit increase in GRE total scores the log odds will increase by 0.03.  $\beta(\hat{GREwriting}) = -0.359779$  shows that GRE writing score is negatively related with probability of getting admitted, and for every one unit increase in writing score leads to a 0.36 drop in log odds. If the applicant is an American students, our model predicts a drop equals to  $\beta(\hat{American}) = -12.892745$  in the log odds, holding all other independent variables constant. If the applicant is a international student, log odds decreases by  $\beta(\hat{International}) = -13.302409$ , and if the student has earned a US degree, log odds drops by  $\beta(\hat{USdegree}) = -12.981663$ .

```
# prediction of model without interaction term
(mod_coef_n <- coef(gradmodel))

##                ugrad_gpa                GRE_Total
##                1.1684823                0.0307436
##                gre_writing                statusAmerican
##                -0.3597787                -12.8927451
##                statusInternational statusInternational with US Degree
##                -13.3024091                -12.9816626

prediction_american_n <- mod_coef_n[1]*mean(grad$ugrad_gpa)+mod_coef_n[2]*mean(grad$GRE_Total)+mod_coef_n[3]*
exp(prediction_american_n) / (1 + exp(prediction_american_n))

## ugrad_gpa
## 0.4909546

prediction_inter_n <- mod_coef_n[1]*mean(grad$ugrad_gpa)+mod_coef_n[2]*mean(grad$GRE_Total)+mod_coef_n[3]*
exp(prediction_inter_n) / (1 + exp(prediction_inter_n))

## ugrad_gpa
## 0.390348

prediction_inter_us_n <- mod_coef_n[1]*mean(grad$ugrad_gpa)+mod_coef_n[2]*mean(grad$GRE_Total)+mod_coef_n[3]*
exp(prediction_inter_us_n) / (1 + exp(prediction_inter_us_n))

## ugrad_gpa
## 0.468765
```

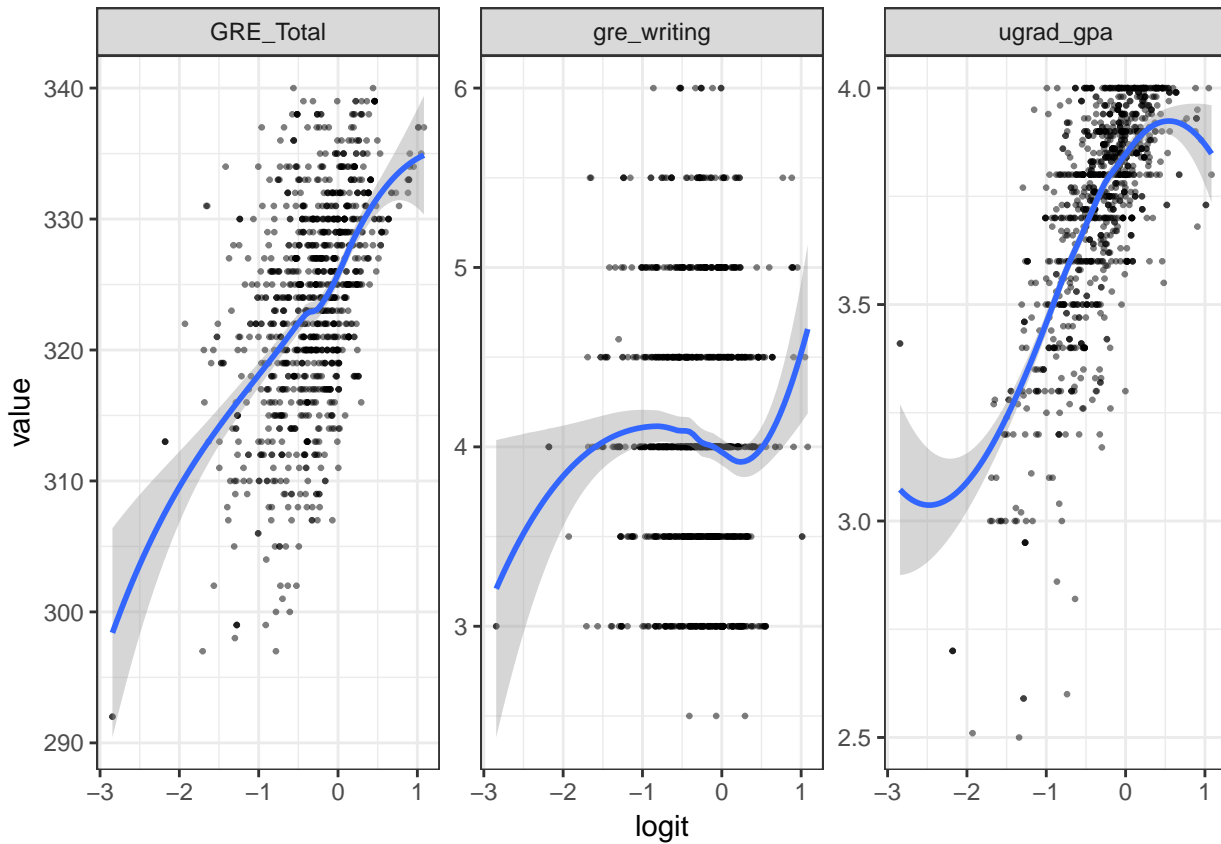
Using same mean level GPA, GRE total score and writing score, our simple logistic model predicts that the probability of an American student getting accepted to the program is 49.1% and the probability for international student without a US degree and those with a US degree is 39% and 46.9% respectively.

```
# check assumptions
# 1. outcome is binary
# 2. linear relationship between the logit of the outcome and each predictor variables
# 3. no influential values
# 4. no high intercorrelations
library(broom)
p_int <- predict(full_mod_int, type = "response")
grad_mod_int <- grad %>%
  select_if(is.numeric) %>% select(-1, -gre_quant, -gre_verbal)
predictors_int <- colnames(grad_mod_int)
grad_mod_int <- (grad_mod_int %>%
```

```

mutate(logit = log(p_int/(1-p_int))) %>%
gather(key = "predictors_int", value = "value", -logit))
# check linearity between x and logit of the outcome
ggplot(grad_mod_int, aes(logit, value))+
  geom_point(size = 0.5, alpha = 0.5) +
  geom_smooth(method = "loess") +
  theme_bw() +
  facet_wrap(~predictors_int, scales = "free_y")

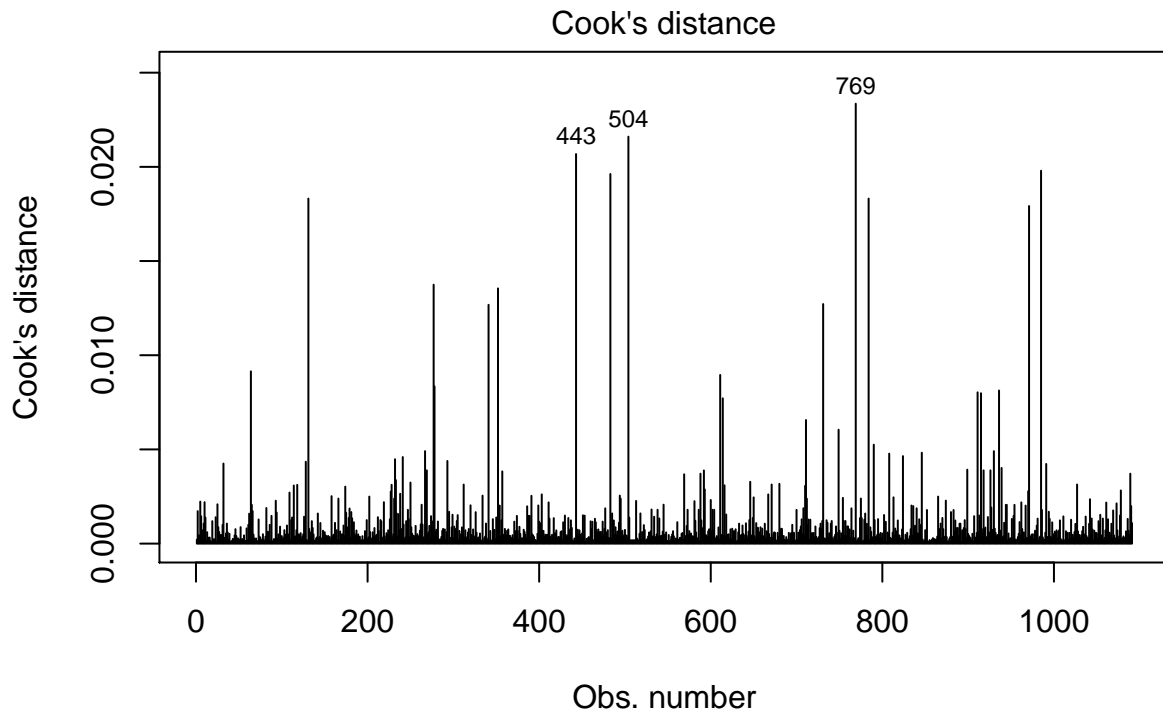
```



```

# check influential values
# top3 largest values
plot(full_mod_int, which = 4, id.n = 3)

```

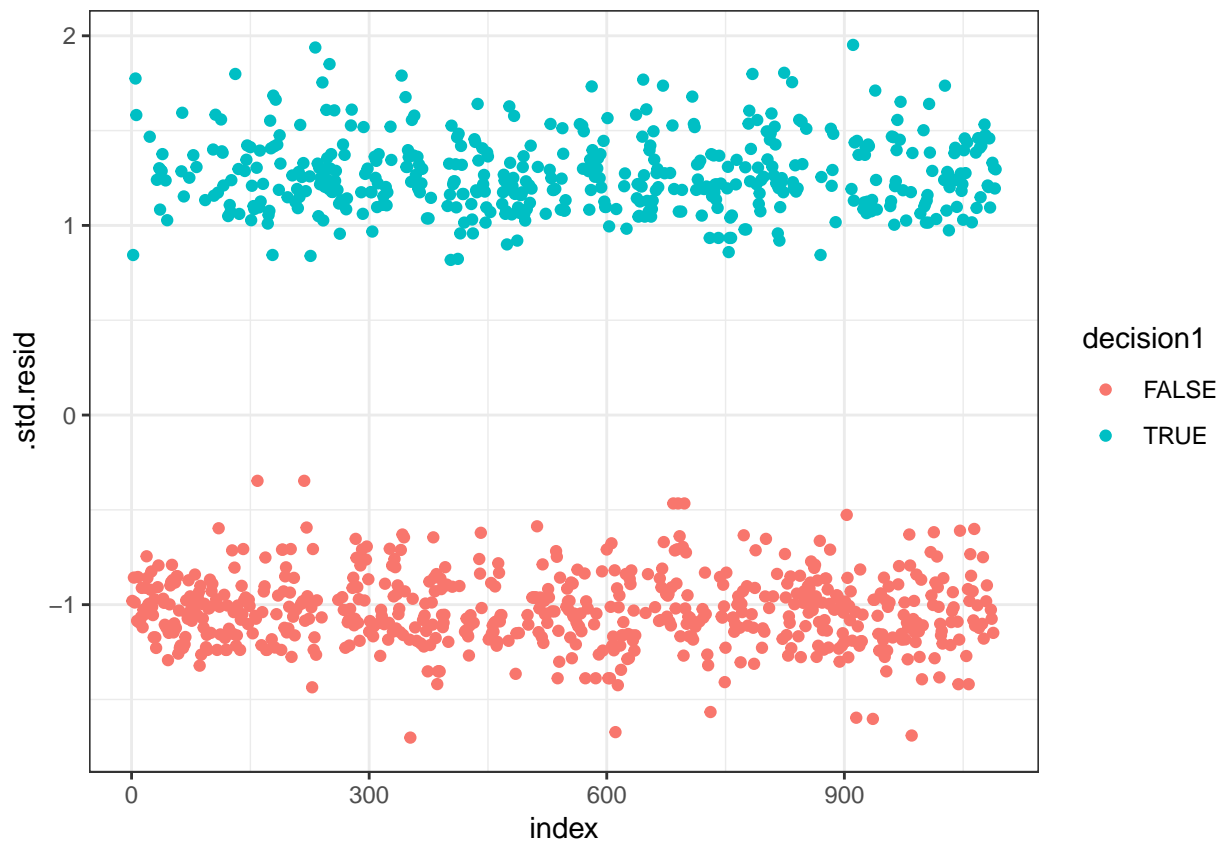


`glm(decision1 ~ (ugrad_gpa + GRE_Total + gre_writing) * status - 1)`

```
# plot the standardized residual
data_int <- augment(full_mod_int) %>%
  mutate(index = 1:n())
data_int %>% top_n(3, .cooksds)
```

```
## # A tibble: 3 x 13
##   decision1 ugrad_gpa GRE_Total gre_writing status .fitted .se.fit .resid
##   <lg1>      <dbl>    <dbl>    <dbl> <chr>      <dbl>    <dbl> <dbl>
## 1 TRUE      3.17      324      3.5 Inter~ -0.299    0.749  1.31
## 2 TRUE      3.2       324      3   Inter~ -0.325    0.755  1.32
## 3 FALSE     3.3       325     5.5 Inter~ -0.00218  0.863 -1.18
## # ... with 5 more variables: .hat <dbl>, .sigma <dbl>, .cooksds <dbl>,
## #   .std.resid <dbl>, index <int>
```

```
ggplot(data_int, aes(index, .std.resid)) +
  geom_point(aes(color = decision1)) +
  theme_bw()
```



```
# if standardized residual is greater than 3 -> Influential
data_int %>%
  filter(abs(.std.resid) > 3)
```

```
## # A tibble: 0 x 13
## #   ... with 13 variables: decision1 <lgl>, ugrad_gpa <dbl>,
## #   GRE_Total <dbl>, gre_writing <dbl>, status <chr>, .fitted <dbl>,
## #   .se.fit <dbl>, .resid <dbl>, .hat <dbl>, .sigma <dbl>, .cooksd <dbl>,
## #   .std.resid <dbl>, index <int>
```

```
#Correlation matrix
```

```
library(Hmisc)
```

```
grad_noNA = grad %>% filter(is.na(ugrad_gpa) == FALSE, is.na(gre_verbal) ==FALSE, is.na(gre_quant) ==F
(grad_noNA = grad_noNA %>% mutate(gre_total = gre_verbal + gre_quant))
```

```
## # A tibble: 1,091 x 18
##   uni_name major degree season decision decision_date decision_timest~
##   <chr>    <chr> <chr> <chr> <chr>    <chr>                <dbl>
## 1 Purdue ~ (Com~ MS    S16    Rejected (2, 11, 2015)    1446440400
## 2 Univers~ Comp~ MS    S16    Accepted (28, 9, 2015)     1443412800
## 3 Univers~ (Com~ MS    F15    Rejected (24, 5, 2015)    1432440000
## 4 Carnegi~ ( EC~ MS    F15    Other    (27, 6, 2015)    1435377600
## 5 Carnegi~ Elec~ MS    F15    Accepted (2, 6, 2015)     1433217600
## 6 Univers~ Elec~ MS    F15    Accepted (14, 4, 2015)    1428984000
## 7 Univers~ Comp~ PhD    F15    Other    (20, 4, 2015)    1429502400
## 8 Cornell~ Comp~ MS    F15    Rejected (7, 4, 2015)    1428379200
## 9 Univers~ Comp~ PhD    F15    Other    (16, 4, 2015)    1429156800
## 10 Univers~ Comp~ PhD    F15    Other    (16, 4, 2015)    1429156800
```

```
## # ... with 1,081 more rows, and 11 more variables: ugrad_gpa <dbl>,
## #   gre_verbal <dbl>, gre_quant <dbl>, gre_writing <dbl>,
## #   is_new_gre <lgl>, status <chr>, comments <chr>, season1 <chr>,
## #   decision1 <lgl>, GRE_Total <dbl>, gre_total <dbl>

my_data1 <- grad_noNA[, c(8,11,14)]
my_data2 <- grad_noNA[, c(8,9,10,11)]
#(rcorr(as.matrix(my_data)))
#This is the correlation matrix for ugrad_gpa, gre_verbal, gre_quant, gre_writing
(rcorr(as.matrix(my_data2)))

##           ugrad_gpa gre_verbal gre_quant gre_writing
## ugrad_gpa      1.00      0.19      0.18      0.13
## gre_verbal      0.19      1.00     -0.08      0.54
## gre_quant       0.18     -0.08      1.00      0.03
## gre_writing      0.13      0.54      0.03      1.00
##
## n= 1091
##
##
## P
##           ugrad_gpa gre_verbal gre_quant gre_writing
## ugrad_gpa      0.0000      0.0000      0.0000
## gre_verbal 0.0000      0.0069      0.0000
## gre_quant 0.0000      0.0069      0.3114
## gre_writing 0.0000      0.0000      0.3114
```

## Assumption

First, since we set the accepted decision as dependent variables and the decision is binary, either 1, accepted or 0, rejected. Therefore, the predicted probability is bind within the interval between 0 and 1. It meets the first assumption of dependent variable to be binary.

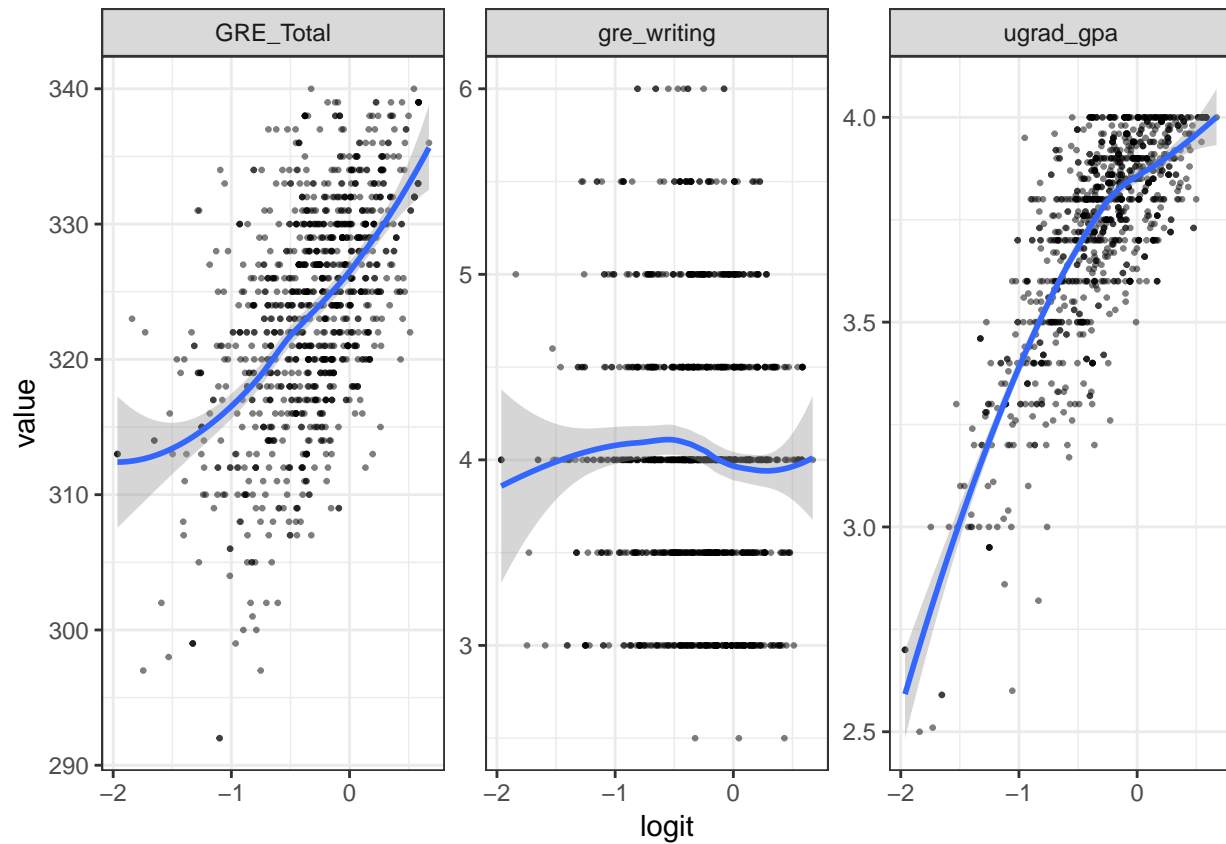
Second, logistic regression also assumes the linearity of independent variables. As shown in “The linearity of independent variables”, the logit of GRE is quite linear to the accepted probability in logit scale. Even though there exists an U-shaped trend at the end of the parabala, the majority of gpa points associated linearly to the logit outcome of undergraduate gpa. However, the scatter plots of gre\_writing shows non\_linearity, similar to a cubic term.

Third, some outliers may be influential enough to alter the quality of the logistic regression model. Therefore, we calculated the Cook’s distance for each points; the higher the leverage and residuals of that point, the higher its Cook’s distance. As demonstrated in Cook’s distance graph, there exist couple of spikes in the graph. To further investigate this issue, the deviance residuals plots has ben constructed. Since it does not have any observations whose cook’s value is large than 3, we conclude that the dataset does not have any influential outliers.

Last but not least, since the variables are intercorrelated, we take this into consideration and use iteration terms to overcome this issue.

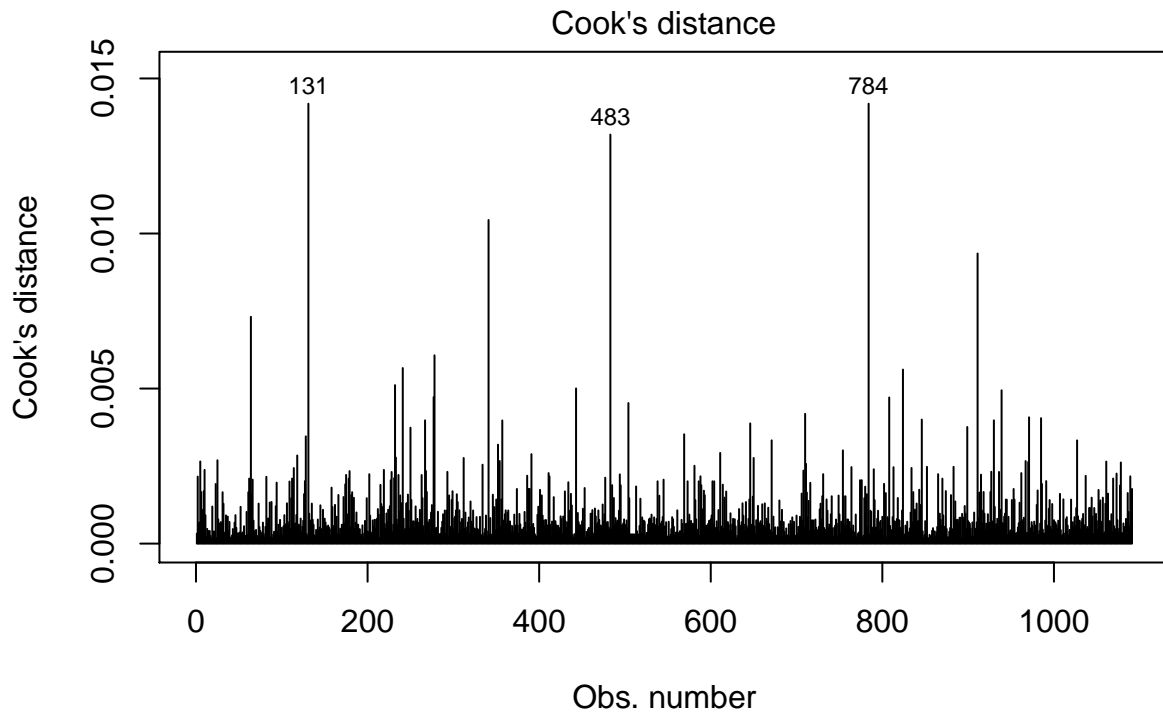
```
p <- predict(full_mod, type = "response")
grad_mod <- grad %>%
  select_if(is.numeric) %>% select(-1, -gre_quant, -gre_verbal)
predictors <- colnames(grad_mod)
grad_mod <- (grad_mod %>%
  mutate(logit = log(p/(1-p))) %>%
  gather(key = "predictors", value = "value", -logit))
```

```
# check linearity between x and logit of the outcome
ggplot(grad_mod, aes(logit, value))+
  geom_point(size = 0.5, alpha = 0.5) +
  geom_smooth(method = "loess") +
  theme_bw() +
  facet_wrap(~predictors, scales = "free_y")
```



```
# check influential values
# top3 largest values
plot(full_mod, which = 4, id.n = 3)
```

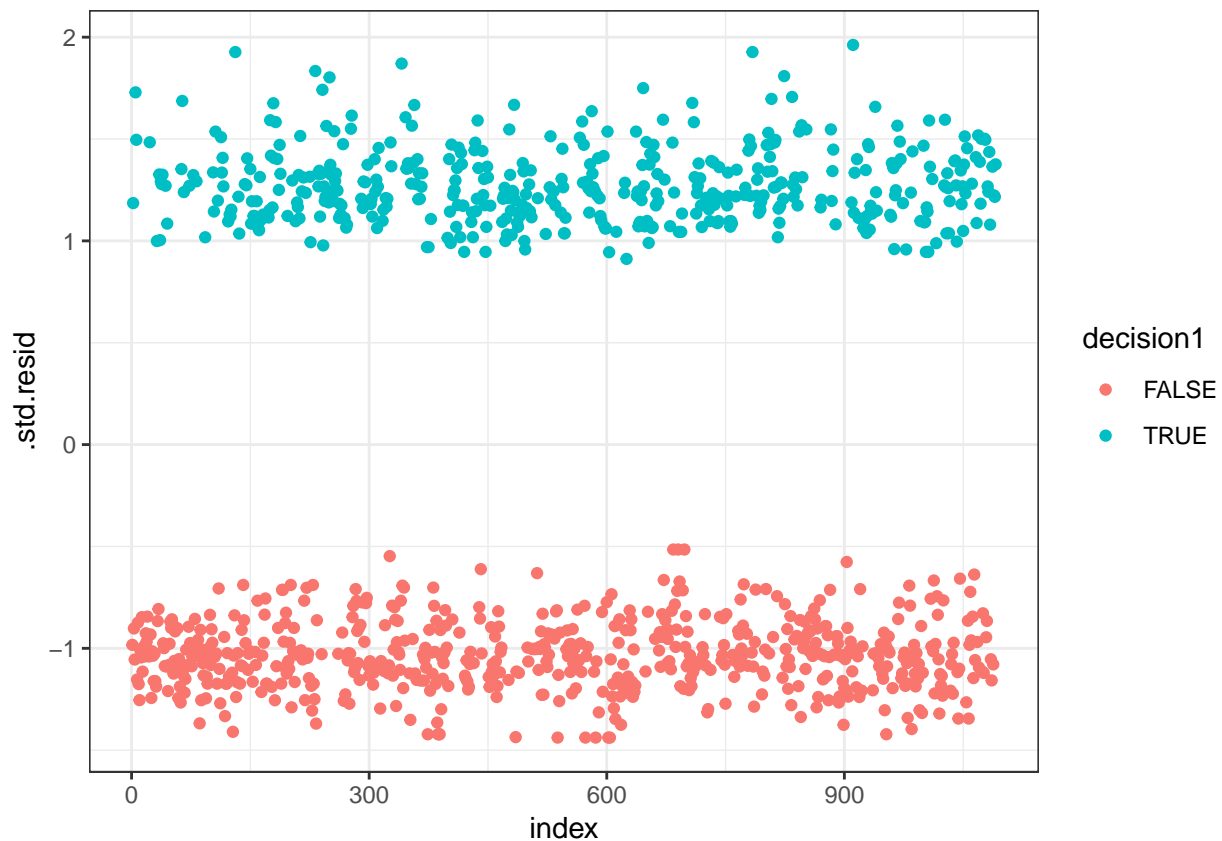




```
# plot the standardized residual
data <- augment(full_mod) %>%
  mutate(index = 1:n())
data %>% top_n(3, .cooksd)
```

```
## # A tibble: 3 x 13
##   decision1 ugrad_gpa GRE_Total gre_writing status .fitted .se.fit .resid
##   <lgl>      <dbl>    <dbl>      <dbl> <chr>    <dbl>  <dbl>  <dbl>
## 1 TRUE      2.59      314         4 Ameri~ -1.65   0.342  1.91
## 2 TRUE      2.6       333         4 Ameri~ -1.06   0.369  1.65
## 3 TRUE      2.59      314         4 Ameri~ -1.65   0.342  1.91
## # ... with 5 more variables: .hat <dbl>, .sigma <dbl>, .cooksd <dbl>,
## #   .std.resid <dbl>, index <int>
```

```
ggplot(data, aes(index, .std.resid)) +
  geom_point(aes(color = decision1)) +
  theme_bw()
```



```
# if standardized residual is greater than 3 -> Influential
data %>%
  filter(abs(.std.resid) > 3)
```

```
## # A tibble: 0 x 13
## #   ... with 13 variables: decision1 <lgl>, ugrad_gpa <dbl>,
## #   GRE_Total <dbl>, gre_writing <dbl>, status <chr>, .fitted <dbl>,
## #   .se.fit <dbl>, .resid <dbl>, .hat <dbl>, .sigma <dbl>, .cooksd <dbl>,
## #   .std.resid <dbl>, index <int>
```

```
# correlation covariance matrix
```

```
grad_noNA = grad %>% filter(is.na(ugrad_gpa) == FALSE, is.na(gre_verbal) ==FALSE, is.na(gre_quant) ==F
(grad_noNA = grad_noNA %>% mutate(gre_total = gre_verbal + gre_quant))
```

```
## # A tibble: 1,091 x 18
##   uni_name major degree season decision decision_date decision_timest~
##   <chr>      <chr> <chr>   <chr>   <chr>      <chr>              <dbl>
## 1 Purdue ~ (Com~ MS     S16     Rejected (2, 11, 2015)    1446440400
## 2 Univers~ Comp~ MS     S16     Accepted (28, 9, 2015)    1443412800
## 3 Univers~ (Com~ MS     F15     Rejected (24, 5, 2015)    1432440000
## 4 Carnegi~ ( EC~ MS     F15     Other    (27, 6, 2015)    1435377600
## 5 Carnegi~ Elec~ MS     F15     Accepted (2, 6, 2015)    1433217600
## 6 Univers~ Elec~ MS     F15     Accepted (14, 4, 2015)    1428984000
## 7 Univers~ Comp~ PhD    F15     Other    (20, 4, 2015)    1429502400
## 8 Cornell~ Comp~ MS     F15     Rejected (7, 4, 2015)    1428379200
## 9 Univers~ Comp~ PhD    F15     Other    (16, 4, 2015)    1429156800
## 10 Univers~ Comp~ PhD    F15     Other    (16, 4, 2015)    1429156800
## # ... with 1,081 more rows, and 11 more variables: ugrad_gpa <dbl>,
```

```
## # gre_verbal <dbl>, gre_quant <dbl>, gre_writing <dbl>,
## # is_new_gre <lgl>, status <chr>, comments <chr>, season1 <chr>,
## # decision1 <lgl>, GRE_Total <dbl>, gre_total <dbl>
```

```
(my_data1 <- grad_noNA[, c(8,11,17)])
```

```
## # A tibble: 1,091 x 3
##   ugrad_gpa gre_writing GRE_Total
##   <dbl>     <dbl>     <dbl>
## 1      3.5      3.5      325
## 2      3.68     4.5      335
## 3      3.96     5       318
## 4      3.93     5       332
## 5      3.3      4       314
## 6      3.76     5       325
## 7      4       5       337
## 8      3.25     3.5      325
## 9      3.7     3.5      322
## 10     3.7      3       322
## # ... with 1,081 more rows
```

```
my_data2 <- grad_noNA[, c(8,9,10,11)]
```

```
#This is the correlation matrix for ugrad_gpa, gre_total, gre_writing
(rcorr(as.matrix(my_data1)))
```

```
##           ugrad_gpa gre_writing GRE_Total
## ugrad_gpa      1.00      0.13      0.27
## gre_writing     0.13      1.00      0.48
## GRE_Total       0.27      0.48      1.00
```

```
##
## n= 1091
```

```
##
## P
##           ugrad_gpa gre_writing GRE_Total
## ugrad_gpa           0           0
## gre_writing 0           0
## GRE_Total    0           0
```

```
#This is the correlation matrix for ugrad_gpa, gre_verbal, gre_quant, gre_writing
(rcorr(as.matrix(my_data2)))
```

```
##           ugrad_gpa gre_verbal gre_quant gre_writing
## ugrad_gpa      1.00      0.19      0.18      0.13
## gre_verbal     0.19      1.00     -0.08      0.54
## gre_quant      0.18     -0.08      1.00      0.03
## gre_writing     0.13      0.54      0.03      1.00
```

```
##
## n= 1091
```

```
##
## P
##           ugrad_gpa gre_verbal gre_quant gre_writing
## ugrad_gpa           0.0000      0.0000      0.0000
## gre_verbal 0.0000           0.0069      0.0000
## gre_quant  0.0000      0.0069           0.3114
```

```
## gre_writing 0.0000    0.0000    0.3114
```

## Assumption\_w/o interaction

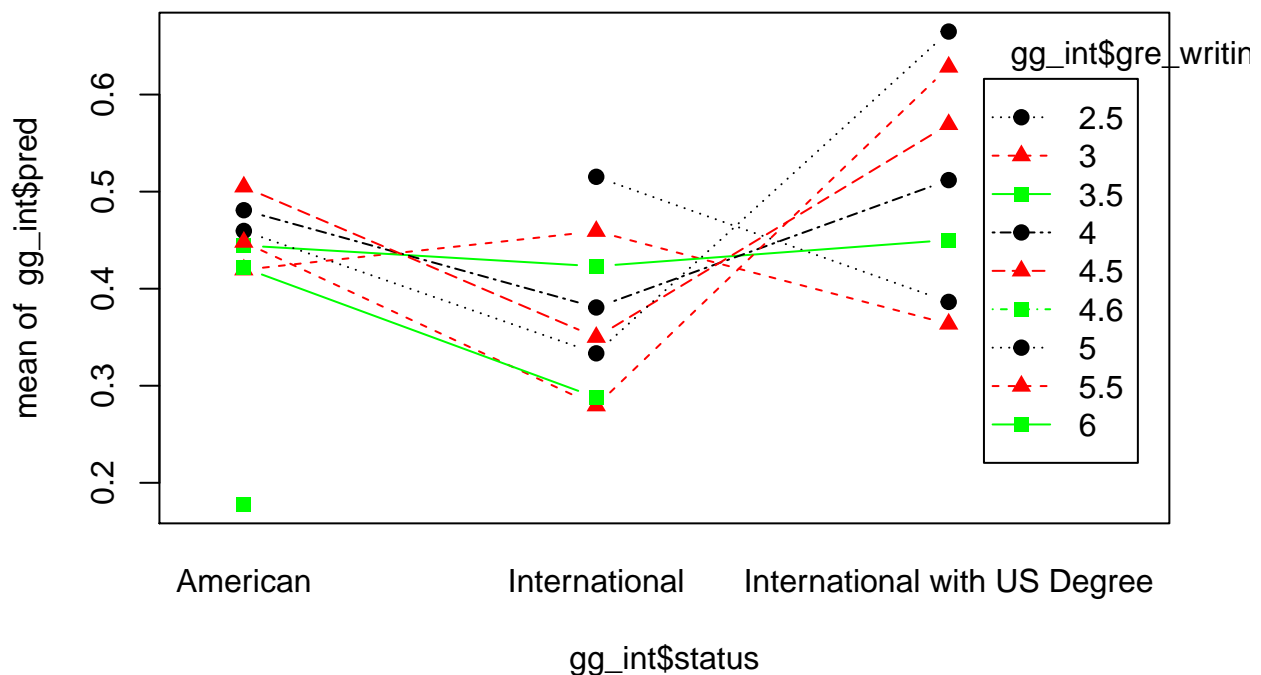
First, since we set the accepted decision as dependent variables and the decision is binary, either 1, accepted or 0, rejected. Therefore, the predicted probability is bind within the interval between 0 and 1. It meets the first assumption of dependent variable to be binary.

Second, logistic regression also assumes the linearity of independent variables. As shown in “The linearity of independent variables”, the logit of GRE and undergraduate gpa are fairly linear to the accepted probability in logit scale. However, the scatter plots of gre\_writing fits a parabola, instead of a linear line.

Third, some outliers may be influential enough to alter the quality of the logistic regression model. Therefore, we calculated the Cook’s distance for each points; the higher the leverage and residuals of that point, the higher its Cook’s distance. As demonstrated in Cook’s distance graph, there exist couple of spikes in the graph. To further investigate this issue, the deviance residuals plots has ben constructed. Since it does not have any observations whose cook’s value is large than 3, we conclude that the dataset does not have any influential outliers.

Last but not least, from the covariance matrix, we can tell that each term are correlated with each other since its p value is near 0. Therefore, we incorporate interaction terms in our further model to overcome this disadvantage.

```
interaction.plot(x.factor      = gg_int$status,
                 trace.factor  = gg_int$gre_writing,
                 response      = gg_int$pred,
                 fun = mean,
                 type="b",
                 col=c("black","red","green"),   ### Colors for levels of trace var.
                 pch=c(19, 17, 15),             ### Symbols for levels of trace var.
                 fixed=TRUE,                    ### Order by factor order in data
                 leg.bty = "o")
```



```
(anova( full_mod, full_mod_int, test = "Chisq"))

## Analysis of Deviance Table
##
## Model 1: decision1 ~ ugrad_gpa + GRE_Total + gre_writing + status - 1
## Model 2: decision1 ~ (ugrad_gpa + GRE_Total + gre_writing) * status -
##      1
##   Resid. Df Resid. Dev Df Deviance Pr(>Chi)
## 1      1085      1444.7
## 2      1079      1435.4  6   9.2779   0.1585

(Anova(full_mod_int, type = "II"))

## Analysis of Deviance Table (Type II tests)
##
## Response: decision1
##              LR Chisq Df Pr(>Chisq)
## ugrad_gpa      17.480  1  2.904e-05 ***
## GRE_Total      11.910  1  0.0005584 ***
## gre_writing     11.308  1  0.0007716 ***
## status         32.334  3  4.449e-07 ***
## ugrad_gpa:status    1.154  2  0.5616588
## GRE_Total:status    2.097  2  0.3504976
## gre_writing:status   3.014  2  0.2215837
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

## Test for the inclusion of a Categorical Variable

$H_0$ : full\_mod = full\_mod

$H_a$ : full\_mod = full\_mod\_int

**Significant Level: 0.05**

Pr(>Chi) for two models is 0.1581, which is bigger than significant level 0.05. Therefore, two models are not significantly different. Pr(>Chi) for ugrad\_gpa, GRE\_Total, gre\_writing and status are all smaller than significant level 0.05, while all the interaction effect is not significant. Therefore, the anova table indicates that the main effect are significant, and interaction effect is not significant.

## Discussion