

CS 589 Fall 2020

Latent semantic indexing

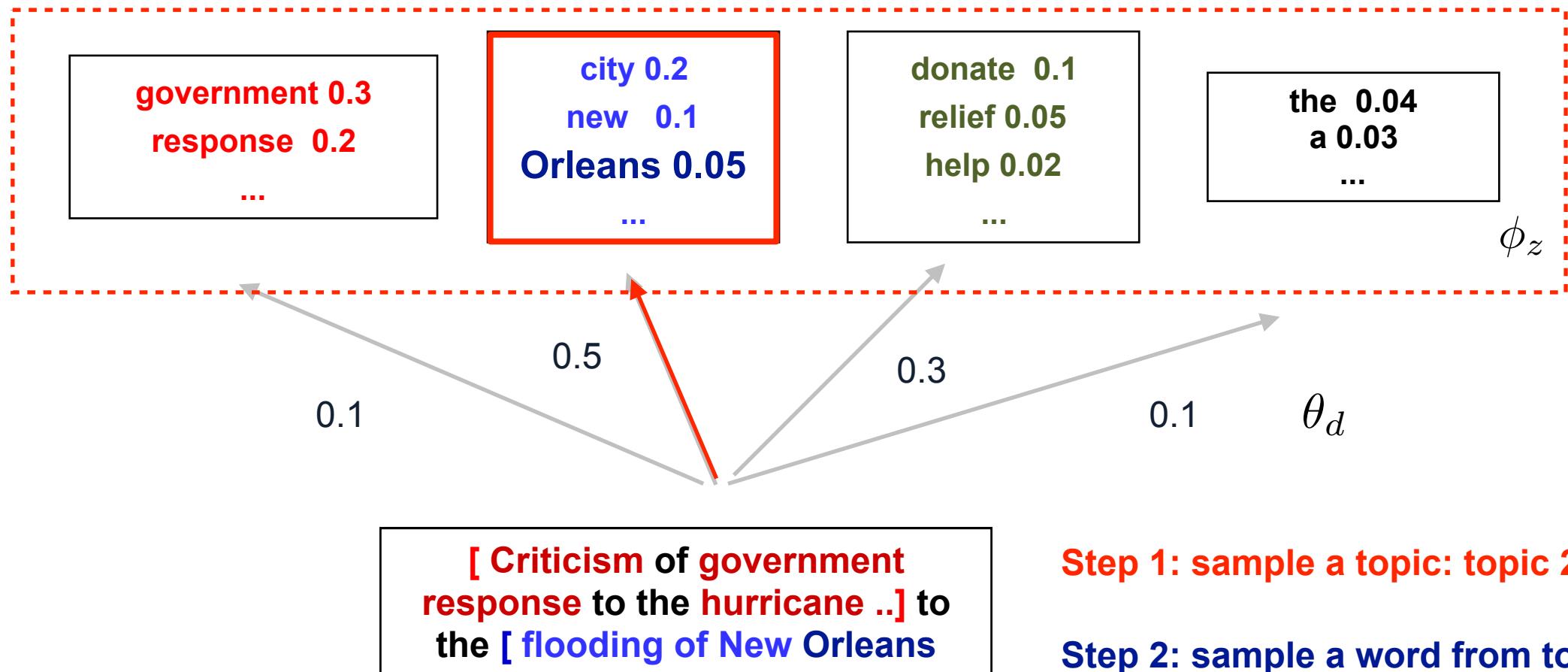
Latent Dirichlet allocation

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Review of PLSA

Phi $T \times V$, V : vocabulary size ~50,000, T : #topics, $T=20$
theta $D \times T$, D : #documents: ~10,000, T : #topics, $T=20$



Today's lecture

- Latent semantic indexing
- Continue on topic model: Latent Dirichlet allocation
 - Bayesian inference of topic model
 - Variational inference for LDA
 - Gibbs sampling, Markov chain Monte-Carlo

Vocabulary gap problem with vector space model

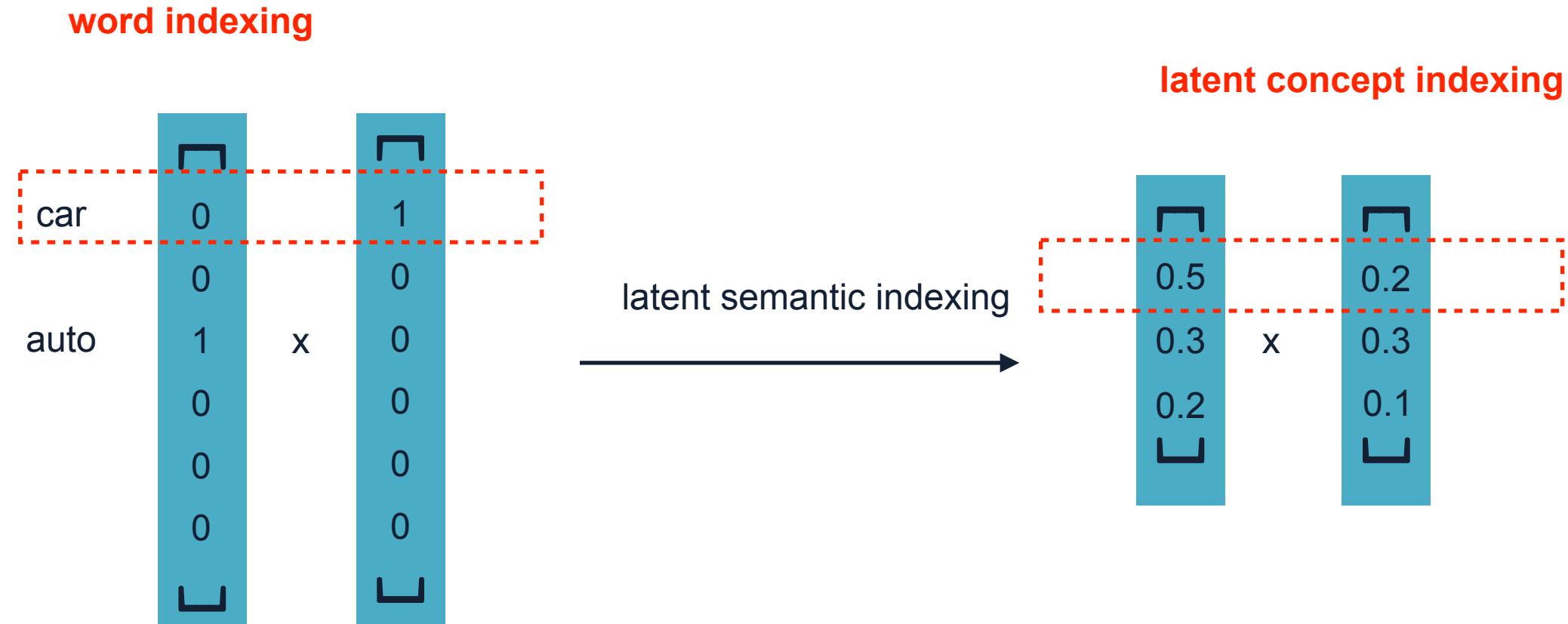
- Vector space model:

$$score(q, d) = \frac{q \cdot d}{\|q\| \cdot \|d\|}$$

- Challenges matching synonyms
 - e.g., auto vs. car
- Challenges matching polysemy
 - e.g., apple (fruit vs. company)

	doc1	doc2
car	0	1
auto	1	0
	x	
	0	0
	0	0
	0	0
	0	0

Low-dimensional, dense vector representation [Deerwester et al. 1998]

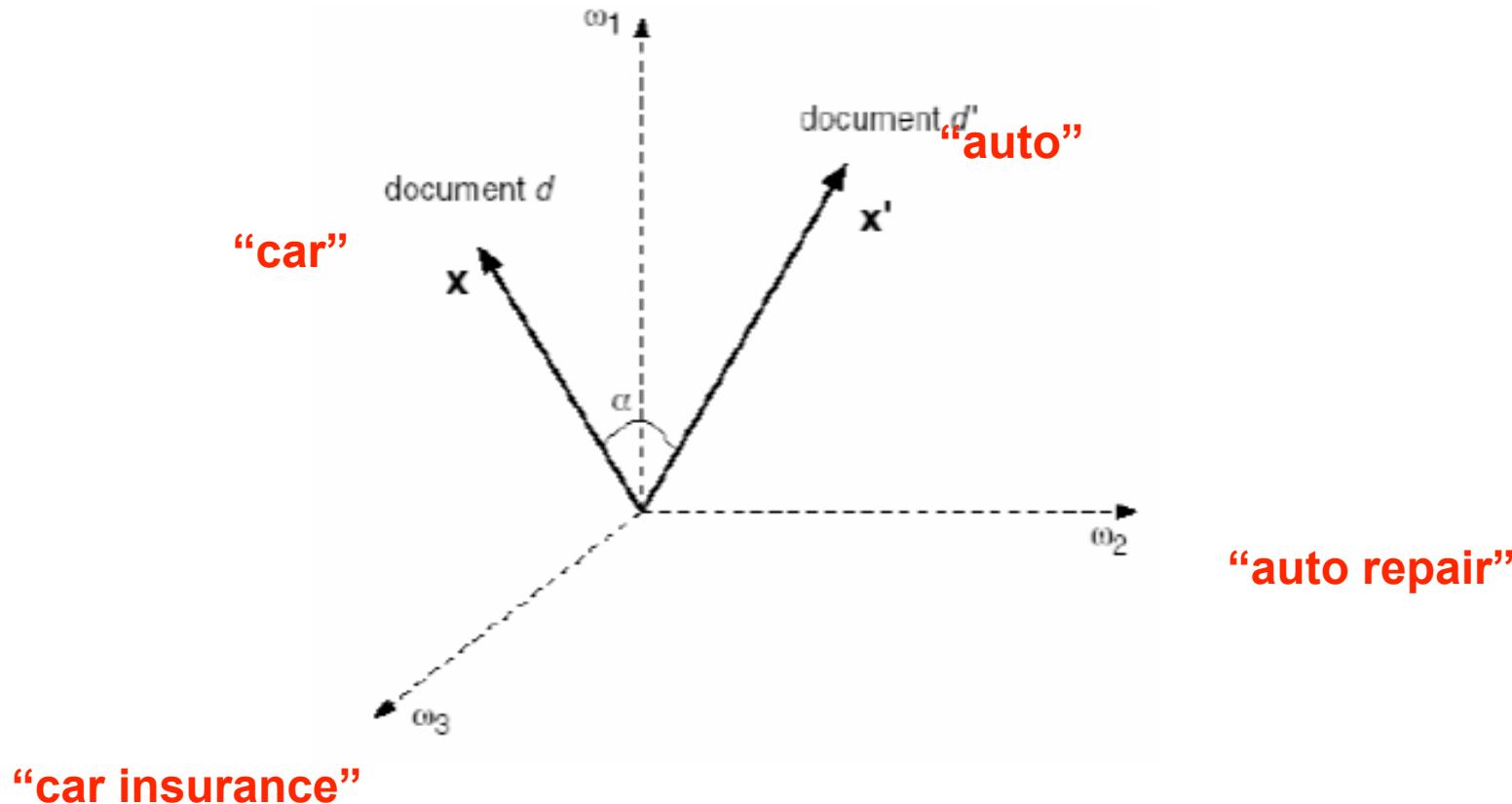


Latent semantic indexing

- Uses statistically derived **conceptual indices** instead of individual word indexing for retrieval
- Assumes that there is some underlying or latent structure in word usage that is obscured by variability in word choice
- Key idea: instead of representing documents and queries as vectors in a **low dimensional space** of terms

Low dimensional vector representation of words

- Axes are concepts, also called principal components (PCA)



Singular value decomposition (matrix factorization)

- For a matrix $A \in \mathbb{R}^{m \times n}$ of rank r, there exists a factorization (SVD) as follows:

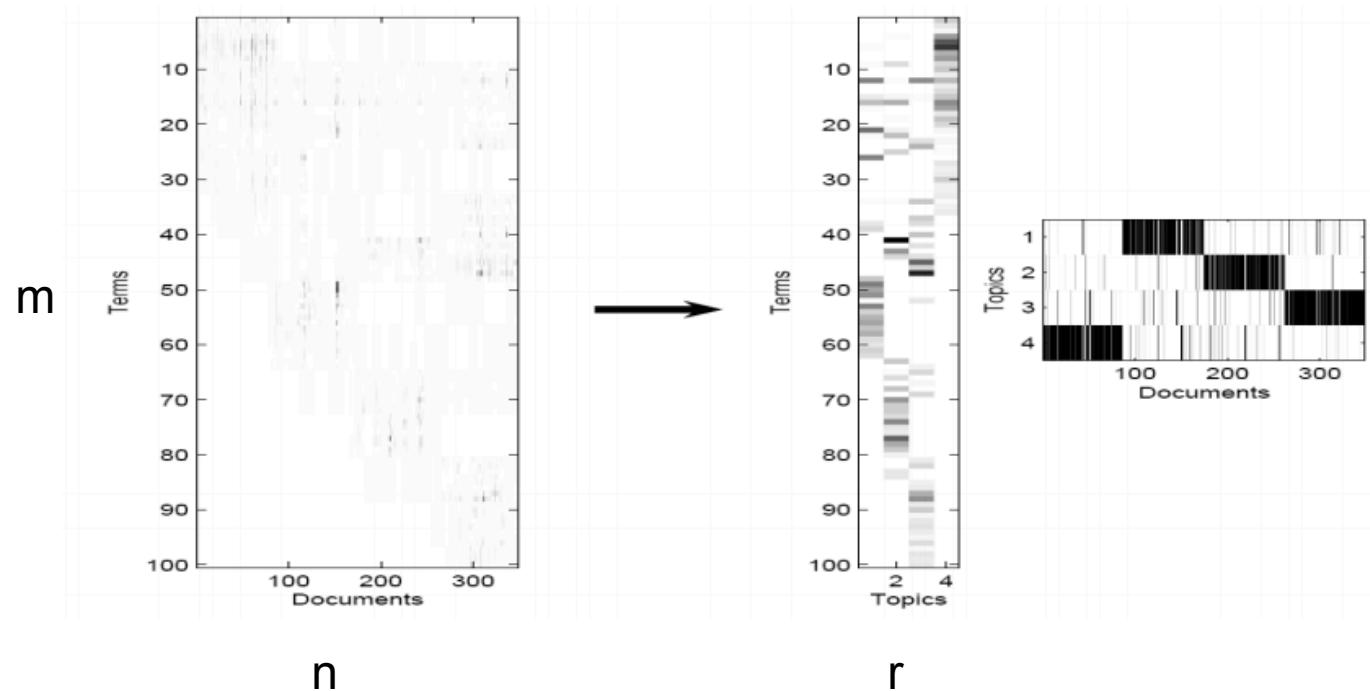
$$A = U\Sigma V^T$$

- U: orthogonal eigen vectors of AA^T
- V: orthogonal eigen vectors of A^TA
- Sigma: eigen values

$$\Sigma = \text{diag}(\sigma_1 \dots \sigma_r)$$

Dimension reduction

- Map documents and queries to a low dimensional space
- Retrieval in this space may be superior to retrieval in the original space

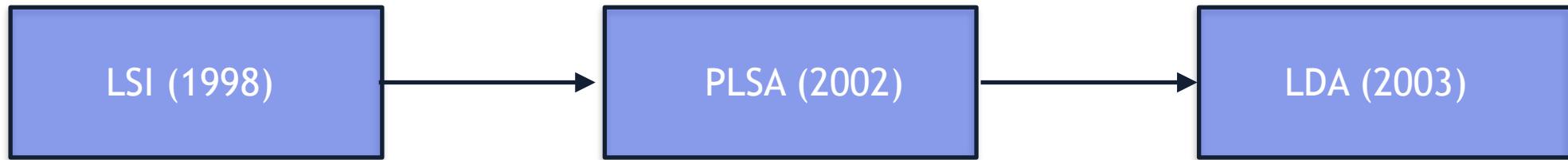


What LSI can do

- LSI effectively does
 - Dimensionality reduction
 - Noise reduction
 - Exploitation of redundant data (linearity)
 - Correlation analysis and query expansion (with related words)
- Some of the individual effects can be achieved with simpler techniques (e.g. thesaurus construction). LSI does them together
- LSI handles **synonymy** well, not so much **polysemy** (vs word embedding)
- Challenge: SVD is complex to compute ($O(n^3)$) – Needs to be updated as new documents are found/updated

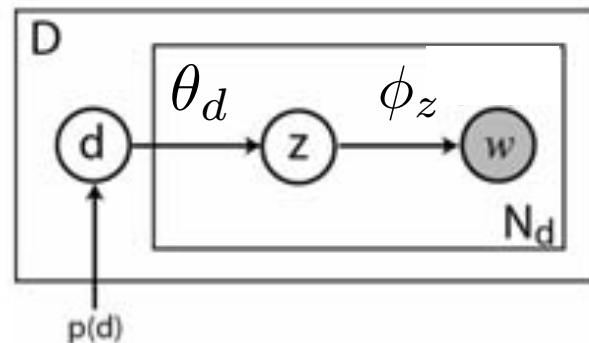
Latent Dirichlet allocation [Blei et al. 03]

- A Bayesian topic model by considering the prior distribution of topics
 - That is, assuming the topic portions are randomly sampled from another distribution



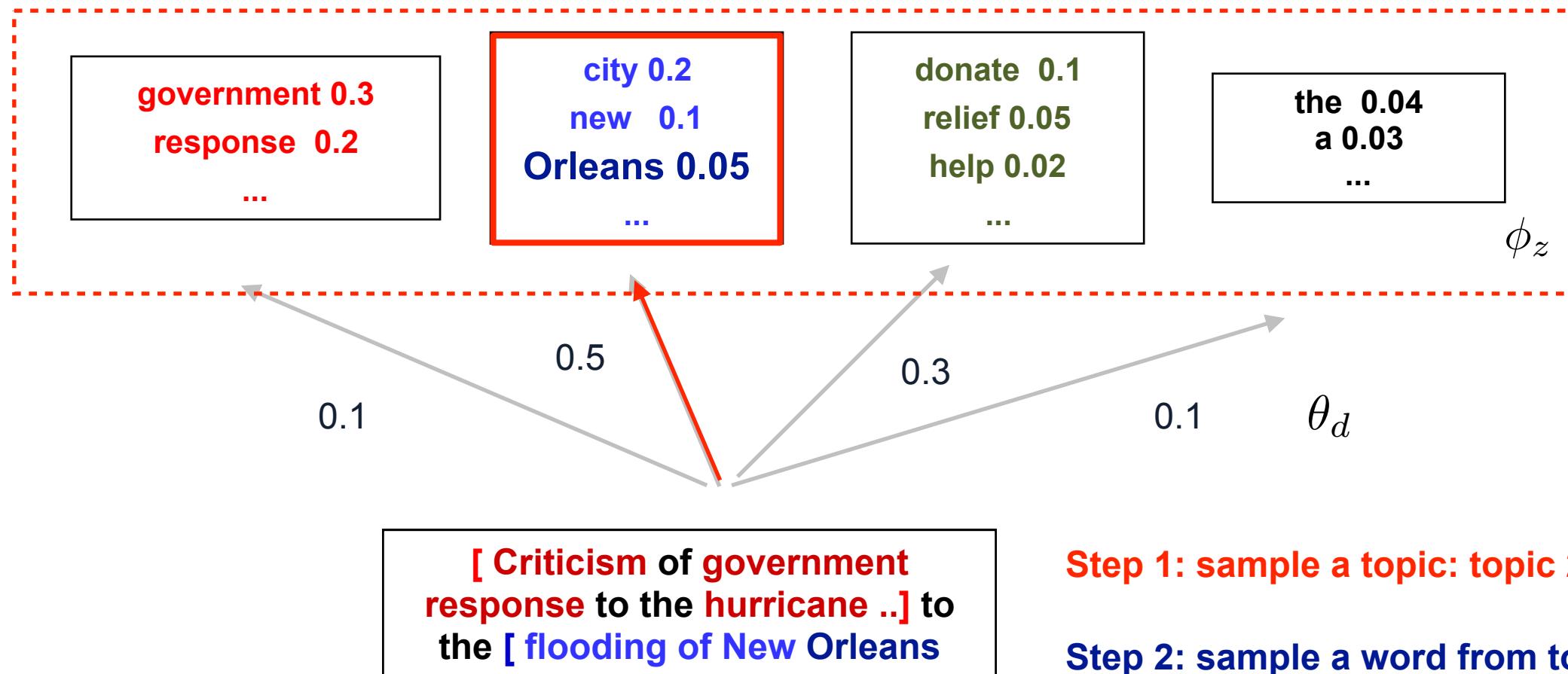
Review of PLSA

- Documents are generated by sampling words from k latent topics
- For each document d:
 - For each token position i
 - Choose a topic $z \sim \text{Multinomial}(\theta_d)$ *(θ_d is a fixed distribution)*
 - Choose a term $w \sim \text{Multinomial}(\phi_z)$



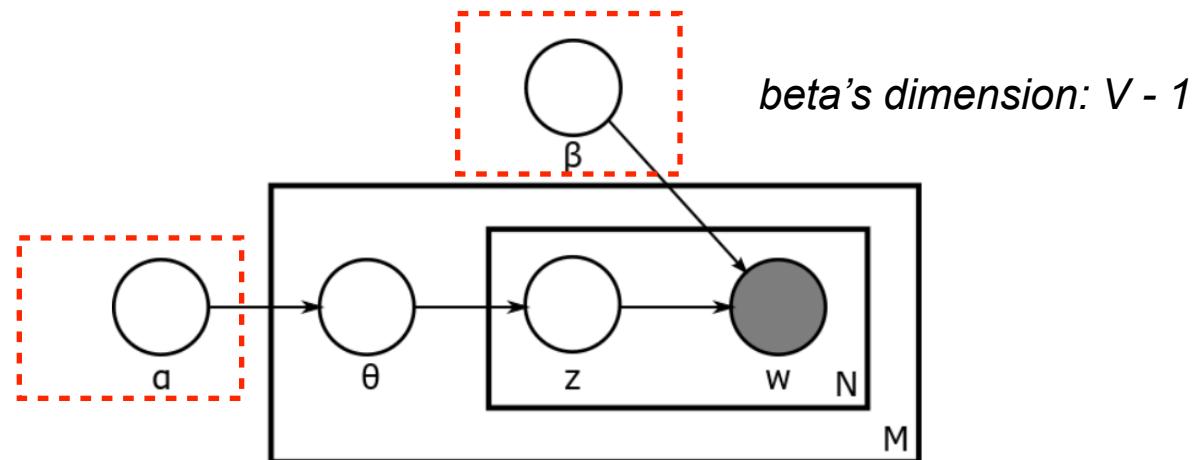
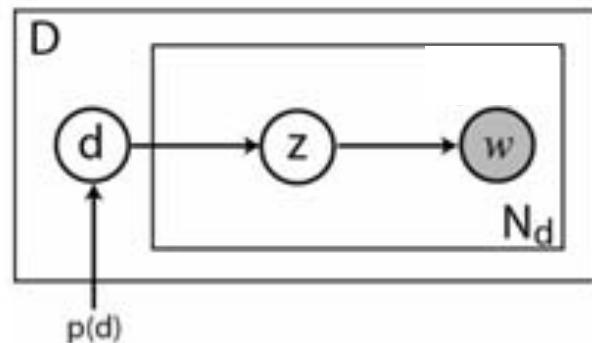
Probabilistic Latent Semantic Analysis. Thomas Hoffman. 2001.

Review of PLSA



Latent Dirichlet allocation [Blei et al. 2003]

- A generative statistical topic model
- Adding a prior hyper parameter alpha to PLSA model
 - The document-topic probability is **randomly sampled** from the same prior distribution



alpha's dimension: $k - 1$

Notation table for LDA

- d (or n): index of document, $d=1, \dots, D$
- i : word inside document d , $i = 1, 2, \dots, N_d$
- z : topic index, $z = 1, 2, \dots, k$
- w or v : word index
- $w_{\{d,i\}}$: the i -th word in document d
- $\theta_{\{d,z\}}$: probability of the z -th topic in document d
- $\phi_{\{z, w\}}$: probability of the w -th word in topic z
- α : prior distribution of θ , dimension = $k - 1$
- β : prior distribution of ϕ , dimension = $V - 1$

Bayes' rules

Chain rule: joint distribution

$$P(A, B) = P(A \cap B) = P(A|B)P(B) = P(B|A)P(A)$$

Bayes' rule:

posterior likelihood prior

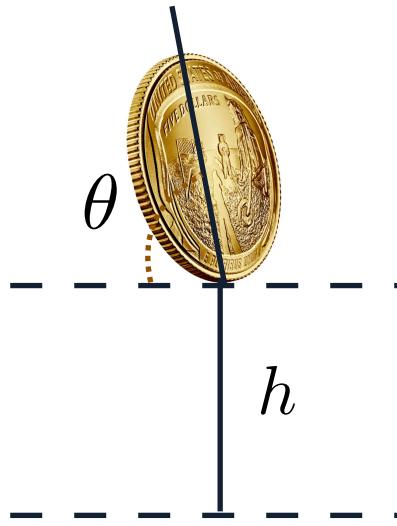
$$P(A|B) = \frac{P(B|A)P(A)}{P(B)} = \left[\frac{P(B|A)}{\sum_{X \in \{A, \bar{A}\}} P(B|X)P(X)} \right] P(A)$$

$$\boxed{P(A|B) \propto P(B|A)P(A)}$$
$$\sum_A P(A|B) = 1$$

skipping estimating $P(B)$

trick for estimating the posterior

Frequentist vs. Bayesian inference



- Frequentist treats parameters as **fixed**

sequence = 0, 1, 0, 0,
1, 1, 0, 1, 0, 1, 0, 0

$$\mu = \frac{\#ups}{\#ups + \#downs}$$

$$x_i \sim \mu$$

- Bayesians treats parameters as **random**
 - Consider the force that causes the coin to be biased
 - The forces are explained a prior distribution:

$$x_i \sim \mu$$

$$\mu \sim Beta(\alpha, \beta)$$

Maximum a Posterior (MAP) estimation

- Maximum likelihood estimation:

$$\theta_{MLE} = \arg \max_{\theta} p(x_1, \dots, x_n | \theta)$$

- Choose the parameter with the maximum posterior probability

$$\begin{aligned}\theta_{MAP} &= \arg \max_{\theta} p(\theta | x_1, \dots, x_n) \\ &= \arg \max_{\theta} p(x_1, \dots, x_n | \theta) p(\theta)\end{aligned}$$

How many parameters in PLSA vs. LDA? PLSA: $d^*(k-1) + k^*(V-1)$, LDA:
 $d^*(k-1)+k^*(V-1) + (k - 1) + (V - 1)$

PLSA -> LDA

- PLSA:

$$p_d(w \mid \{\theta_d\}, \{\phi_z\}) = \sum_{z=1}^k \theta_{d,z} \phi_{z,w}$$

$$\log p(D \mid \{\phi_z\}, \{\theta_d\}) = \sum_{d \in D} \sum_{w \in V} c(w, d) \log \left[\sum_{z=1}^k \theta_{d,z} \phi_{z,w} \right]$$

- LDA:

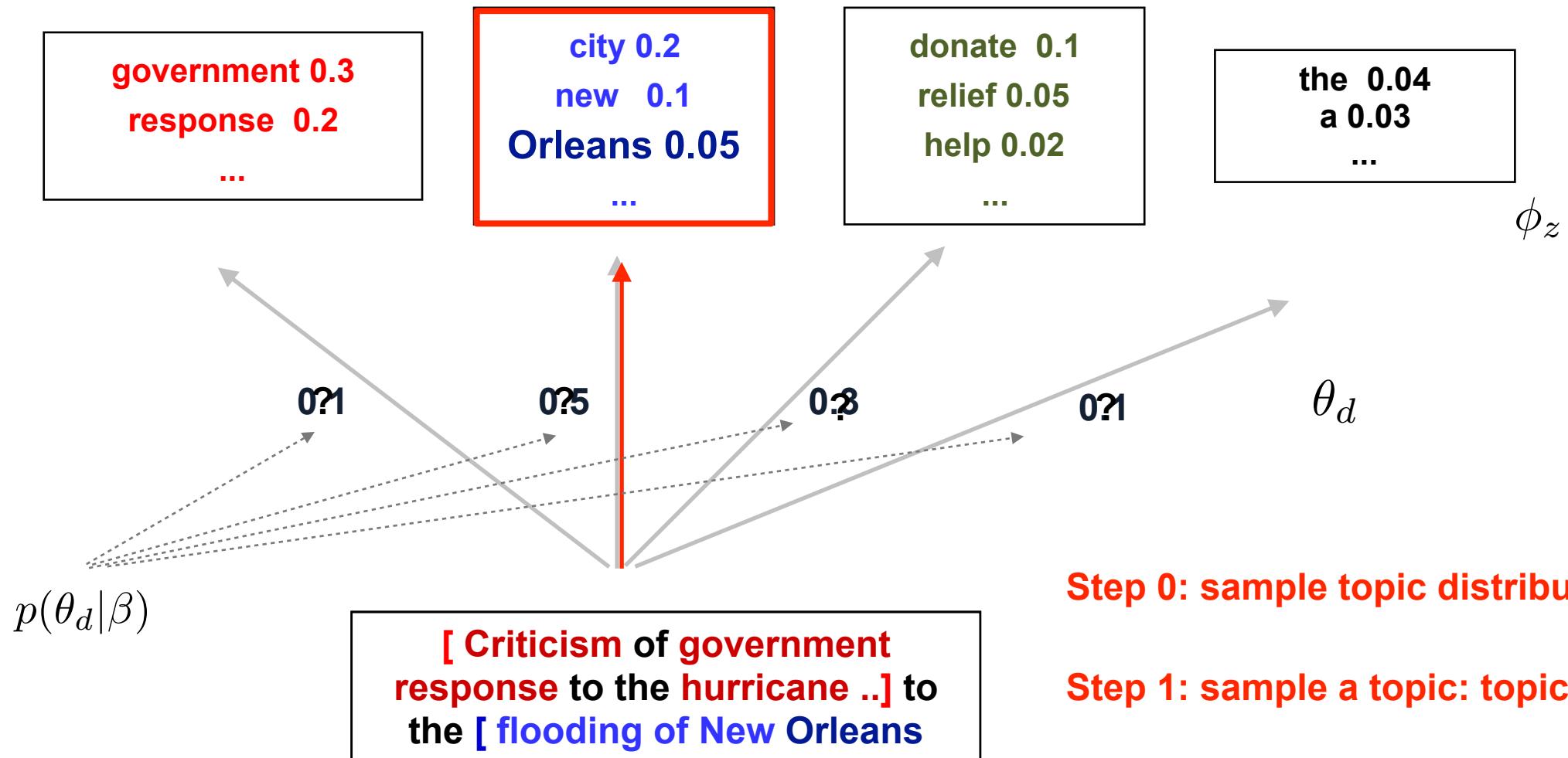
$$p_d(w \mid \{\theta_d\}, \{\phi_z\}) = \sum_{z=1}^k \theta_{d,z} \phi_{z,w}$$

$$\log p(d \mid \alpha, \{\phi_z\}) = \int \sum_{w \in V} c(w, d) \log \left[\sum_{z=1}^k \theta_{d,z} \phi_{z,w} \right] p(\theta_d \mid \alpha) d\theta_d$$

marginalized probability:

$$\log p(D \mid \alpha, \beta) = \int \sum_{d \in D} \log p(d \mid \alpha, \{\phi_z\}) \prod_{z=1}^k p(\phi_z \mid \beta) d\phi_1 \dots d\phi_k$$

PLSA \rightarrow LDA



Solving Maximum a Posteriori inference for LDA

- Maximum likelihood estimation:

$$\begin{aligned}\mathcal{L} = \log p(\mathcal{W} | \mathbf{R}, \Phi, \Theta) &= \sum_d^D \sum_{d_i}^{N_d} \sum_z^T R_{(w_{d_i}, z)} (\log \phi_{(z, w_{d_i})} + \log \theta_{(d, z)}) \\ &+ \left[\sum_{d=1}^D \lambda_d \left(1 - \sum_{z=1}^T \theta_{(d, z)} \right) + \sum_{z=1}^T \sigma_k \left(1 - \sum_{w=1}^V \phi_{z, w} \right) \right]\end{aligned}$$

M step:

$$\theta_{d,z} \propto \sum_{\substack{v=1 \\ v=D}}^V R_{d,v,z} \text{count}(v, d)$$

$$\phi_{z,v} \propto \sum_{d=1}^D [z_{d,v} == z] \text{count}(v, d)$$

E step:

$$R_{(w_{d_i}, z)} \propto \phi_{z,v} \theta_{d,z}$$

Solving Maximum a Posteriori inference for LDA

- Exact inference is intractable:

$$p(Z, \Phi, \Theta | D, \alpha, \beta) = \frac{p(D, Z, \Phi, \Theta | \alpha, \beta)}{p(D | \alpha, \beta)}$$

$$\log p(D | \alpha, \beta) = \int \sum_{d \in D} \log p(d | \alpha, \{\phi_z\}) \prod_{z=1}^k p(\phi_z | \beta) d\phi_1 \dots d\phi_k$$

- Equation (1) is computationally intractable due to the coupling of beta and phi in the denominator
- Question: **why do we need to compute the denominator?**

Variational inference

- Key idea: use a **surrogate distribution** to approximate the posterior distribution of latent variables
 - Surrogate distribution is simpler to estimate than the true posterior
- Goal: finding the “best” surrogate distribution from a certain parametric family by **minimizing** the KL-divergence between the surrogate function (Q) to the true posterior (P)
- Typical surrogate distributions
 - Mean-field approximation [Blei et al. 03]
 - Expectation propagation [Minka et al. 02]
 - Collapsed variational Bayes [Teh et al. 07]

Evidence lower bound (ELBO)

- Given that $p(Z|D) = \frac{p(D, Z)}{p(D)}$, we want to minimize the KL divergence between $Q(Z)$ and $p(Z|D)$:

$$\begin{aligned} D_{\text{KL}}(q\|p) &= \int_Z q(Z) \left[\log \frac{q(Z)}{p(Z, D)} + \log p(D) \right] \\ &= \int_Z q(Z) [\log q(Z) - \log p(Z, D)] + \int_Z q(Z) [\log p(D)] \\ &= \int_Z q(Z) [\log q(Z) - \log p(Z, D)] + \log p(D) \end{aligned}$$

$$\Rightarrow \log p(D) = D_{\text{KL}}(q\|p) - \mathbb{E}_q[\log q(Z) - \log p(Z, D)] = D_{\text{KL}}(q\|p) + \mathcal{L}(q)$$

Evidence lower bound (ELBO)

<https://stats.stackexchange.com/questions/205506/why-do-we-use-the-mean-field-approximation-for-variational-bayes>
<https://www.cs.colorado.edu/~jbg/>

- $p(D)$ does not rely on the latent variable Z :

$$\Rightarrow \log p(D) = D_{\text{KL}}(q||p) - \mathbb{E}_q[\log q(Z) - \log p(Z, D)] = D_{\text{KL}}(q||p) + \mathcal{L}(q)$$

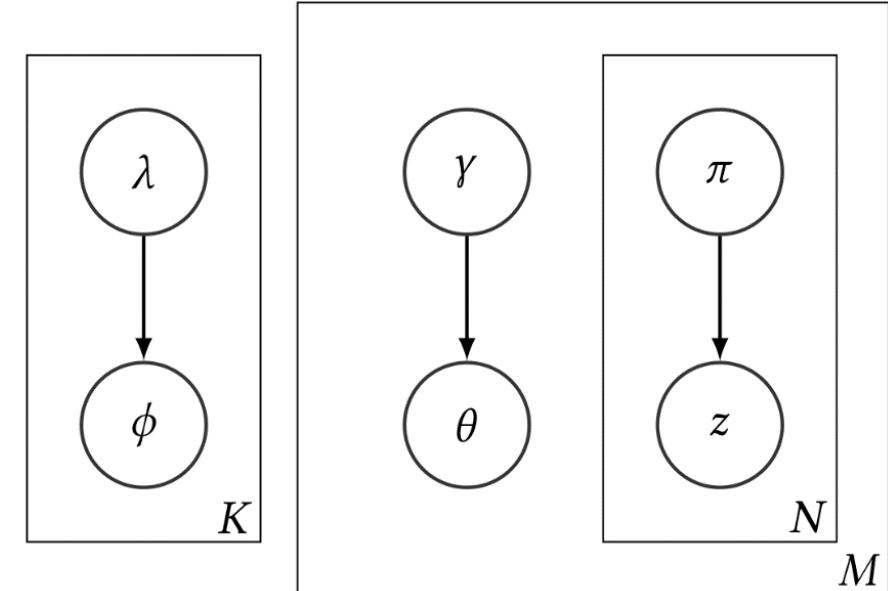
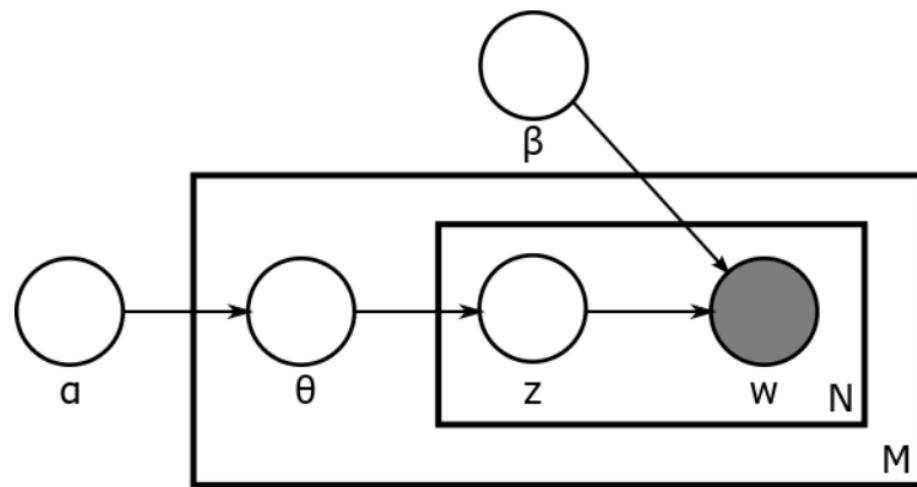
- Therefore minimizing KL divergence = maximizing $\mathcal{L}(q)$, which we call evidence lower bound, as KL divergence is non-negative
- The lower bound is more tractable to optimize than the evidence:

$$\mathcal{L}(q) = \int_Z q(Z)[\log q(Z) - \log p(Z, D)]$$

LDA variational inference

- Assumption: q belongs to a family of distribution which can be factorized

$$q(\theta, \mathbf{z} \mid \gamma, \phi) = q(\theta \mid \gamma) \prod_{n=1}^N q(z_n \mid \phi_n)$$



LDA variational inference

$$L(q) = \int_Z q(Z)[\log q(Z) - \log p(Z, D)]$$

$$p(Z, \Phi, \Theta | D, \alpha, \beta) = \frac{p(D, Z, \Phi, \Theta | \alpha, \beta)}{p(D | \alpha, \beta)}$$

- ELBO of LDA under the factorization assumption:

$$\begin{aligned} L(\gamma, \phi; \alpha, \beta) &= \mathbb{E}_q[\log p(\theta | \alpha)] + \mathbb{E}_q[\log p(\mathbf{z} | \theta)] + \mathbb{E}_q[\log p(\mathbf{w} | \mathbf{z}, \beta)] \\ &\quad - \mathbb{E}_q[\log q(\theta)] - \mathbb{E}_q[\log q(\mathbf{z})] \end{aligned}$$

$$\begin{aligned} L(\gamma, \phi; \alpha, \beta) &= \log \Gamma \left(\sum_{j=1}^k \alpha_j \right) - \sum_{i=1}^k \log \Gamma (\alpha_i) + \sum_{i=1}^k (\alpha_i - 1) \left(\Psi (\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right) \right) \\ &\quad + \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} \left(\Psi (\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right) \right) \\ &\quad + \sum_{n=1}^N \sum_{i=1}^k \sum_{j=1}^V \phi_{ni} w_n^j \log \beta_{ij} \\ &\quad - \log \Gamma \left(\sum_{j=1}^k \gamma_j \right) + \sum_{i=1}^k \log \Gamma (\gamma_i) - \sum_{i=1}^k (\gamma_i - 1) \left(\Psi (\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right) \right) \\ &\quad - \sum_{n=1}^N \sum_{i=1}^k \phi_{ni} \log \phi_{ni} \end{aligned}$$

← no θ , why?

$$\mathbb{E}_q [\log (\theta_i) | \gamma] = \Psi (\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right)$$

LDA variational inference

- LDA variational inference algorithm: optimizing γ , φ , β
1. Randomly initialize variational parameters (can't be uniform)
 2. For each iteration:
 1. For each document, update γ and φ
 2. For corpus, update β
 3. Compute L for diagnostics
 3. Return **expectation of variational parameters** for solution to latent variables

$$\gamma_i = \alpha_i + \sum_n \phi_{ni}$$

$$\beta_{ij} \propto \sum_d \sum_n \phi_{dni} w_{dn}^j$$

$$\phi_{ni} \propto \beta_{iv} \exp \left(\psi(\gamma_i) - \Psi \left(\sum_j \gamma_j \right) \right)$$

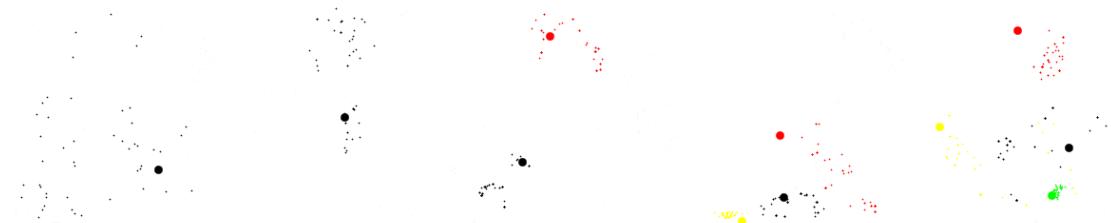
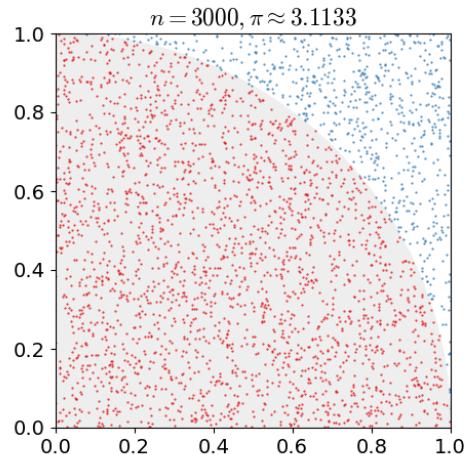
$$E_q [\log (\theta_i) | \gamma] = \Psi(\gamma_i) - \Psi \left(\sum_{j=1}^k \gamma_j \right)$$

Pros and Cons of Variational inference

- Pros
 - Deterministic algorithm - easy to tell when converged
 - Embarrassingly parallelizable
- Cons:
 - Speed: make many calls to transcendental functions (no close form solution)
 - Quality is questionable due to the factorization assumption
 - Memory usage: requires $O(MNK)$ to store the per-token variational distributions

Monte Carlo for LDA inference

- Monte Carlo
 - A broad class of computational algorithms that rely on repeated random sampling to obtain numerical results
- Markov chain Monte Carlo (MCMC)
 - A Monte Carlo method with the Markov assumption, i.e., the next sample only rely on the previous state



Conjugate prior distribution

- That is, the **posterior distributions** are in the same probability distribution family as the prior probability distribution
- The prior is called the **conjugate prior** of the likelihood probability
- Conjugate prior makes denominator easy to compute because you can integrate out theta
- e.g., Beta distribution is the conjugate prior of Bernoulli distribution

Bernoulli: $p(x | \theta) = \theta^x(1 - \theta)^{1-x}.$

Beta: $p(\theta | \alpha) = K(\alpha)\theta^{\alpha_1-1}(1 - \theta)^{\alpha_2-1}.$

Demominator:
$$\begin{aligned} K(\alpha) &= \left(\int \theta^{\alpha_1-1}(1 - \theta)^{\alpha_2-1} d\theta \right)^{-1} \\ &= \frac{\Gamma(\alpha_1 + \alpha_2)}{\Gamma(\alpha_1)\Gamma(\alpha_2)} \end{aligned}$$

Collapsed variational inference [Teh et al. 07]

- If your priors are conjugate, they can be integrated out!
- Let n_{dzv} be the number of times that word v in document d has topic z
- Let $n_{.zv}$ be the number of times word v in any document has topic z

$$p(\mathbf{z}, \mathbf{x} \mid \alpha, \beta) = \prod_d \frac{\Gamma(k\alpha)}{\Gamma(k\alpha + n_{d..})} \prod_z \frac{\Gamma(\alpha + n_{dz.})}{\Gamma(\alpha)} \prod_z \frac{\Gamma(V\beta)}{\Gamma(V\beta + n_{.z.})} \prod_v \frac{\Gamma(\beta + n_{.zv})}{\Gamma(\beta)}$$

No theta or phi in the above formula

Collapsed variational inference [Teh et al. 07]

$$p(\mathbf{z}, \mathbf{x} \mid \alpha, \beta) = \prod_d \frac{\Gamma(k\alpha)}{\Gamma(k\alpha + n_{d..})} \prod_z \frac{\Gamma(\alpha + n_{dz.})}{\Gamma(\alpha)} \prod_z \frac{\Gamma(V\beta)}{\Gamma(V\beta + n_{.z.})} \prod_v \frac{\Gamma(\beta + n_{.zv})}{\Gamma(\beta)}$$

- How do we get back phi and theta? MAP estimates from σ and δ

$$\theta_{dz} = \frac{\alpha + n_{dk.}}{K\alpha + n_{d..}}$$

$$\phi_{zv} = \frac{\beta + n_{.zv}}{W\beta + n_{.z.}}$$

- How do we get n_{dzv} ?
 - Gibbs sampling: sample values of Z, count from those examples

Collapsed Gibbs sampling [Griffths et al. 04; Teh et al. 07]

- **Key idea:** construct a well-behaved Markov chain such that
 1. the states of the chain represent an assignment of Z
 2. state transitions occur between states that differ in only one $z_{\{d,i\}}$
 3. transition probabilities are based on the **full conditional**:
 - 4.
- 5. where $z^{\{di\}}$ is the set of assignments for Z without the i -th word in document d

$$p(z_{di} = k \mid \mathbf{z}^{\neg di}, \mathbf{x}, \alpha, \beta) = \frac{(\alpha + n_{dz \cdot}^{\neg di}) (\beta + n_{\cdot z w_{di}}^{\neg di}) (W\beta + n_{\cdot z \cdot}^{\neg di})^{-1}}{\sum_{z'=1}^k (\alpha + n_{dz' \cdot}^{\neg di}) (\beta + n_{\cdot z' w_{di}}^{\neg di}) (W\beta + n_{\cdot z' \cdot}^{\neg di})^{-1}}$$

Collapsed Gibbs sampling [Griffths et al. 04]

i	w_i	d_i	iteration	
			1	2
1	MATHEMATICS	1	2	?
2	KNOWLEDGE	1	2	
3	RESEARCH	1	1	
4	WORK	1	2	
5	MATHEMATICS	1	1	
6	RESEARCH	1	2	
7	WORK	1	2	
8	SCIENTIFIC	1	1	
9	MATHEMATICS	1	2	
10	WORK	1	1	
11	SCIENTIFIC	2	1	
12	KNOWLEDGE	2	1	
.	.	.	.	
.	.	.	.	
.	.	.	.	
50	JOY	5	2	

$$p(z_{di} = k \mid \mathbf{z}^{\neg di}, \mathbf{x}, \alpha, \beta) = \frac{(\alpha + n_{dz\cdot}^{\neg di}) (\beta + n_{\cdot zw_{di}}^{\neg di}) (W\beta + n_{\cdot z\cdot}^{\neg di})^{-1}}{\sum_{z'=1}^k (\alpha + n_{dz' \cdot}^{\neg di}) (\beta + n_{\cdot z' w_{di}}^{\neg di}) (W\beta + n_{\cdot z' \cdot}^{\neg di})^{-1}}$$

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.	.	.	.	
.	.	.	.	
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Collapsed Gibbs sampling [Griffths et al. 04]

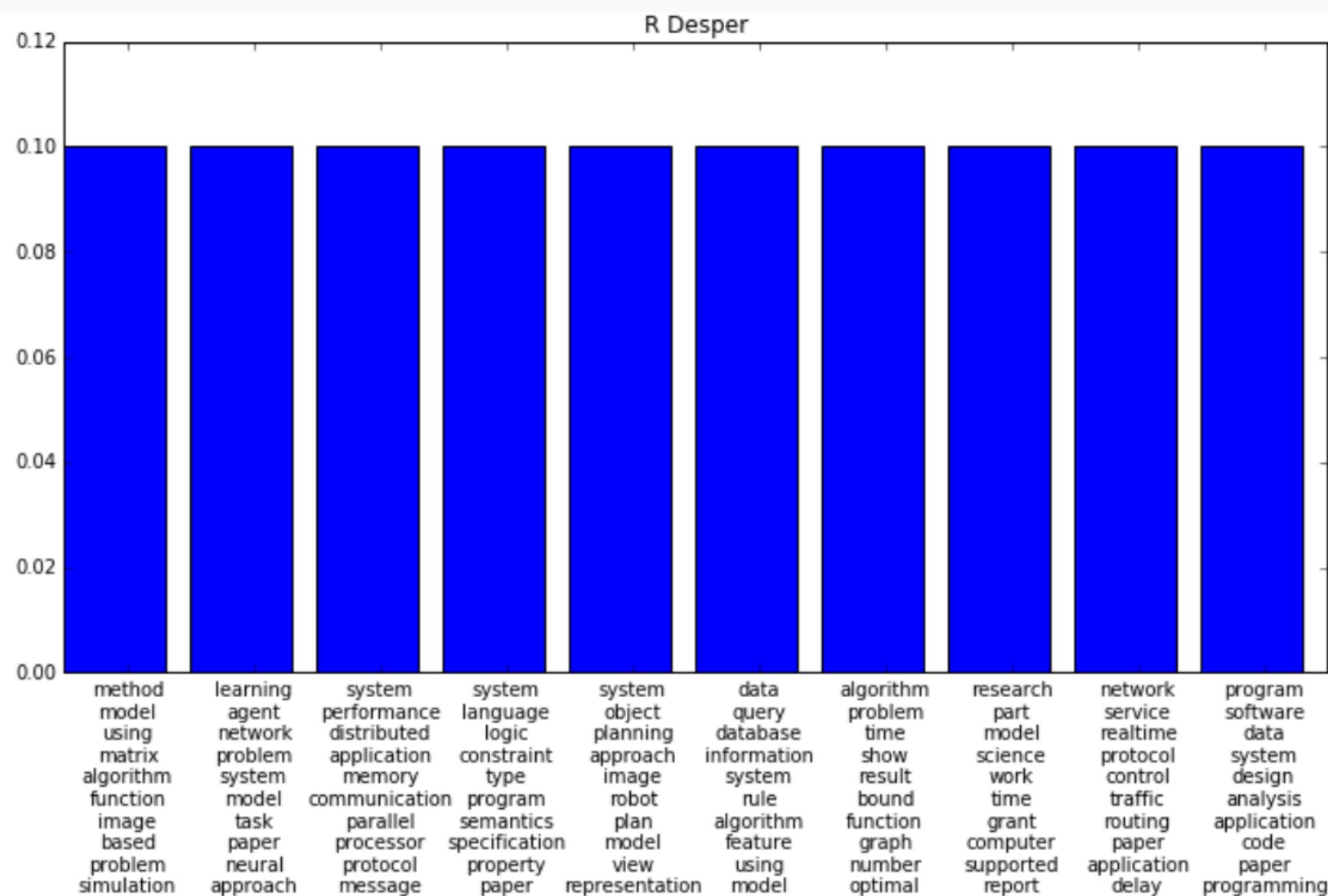
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.	.	.	.	
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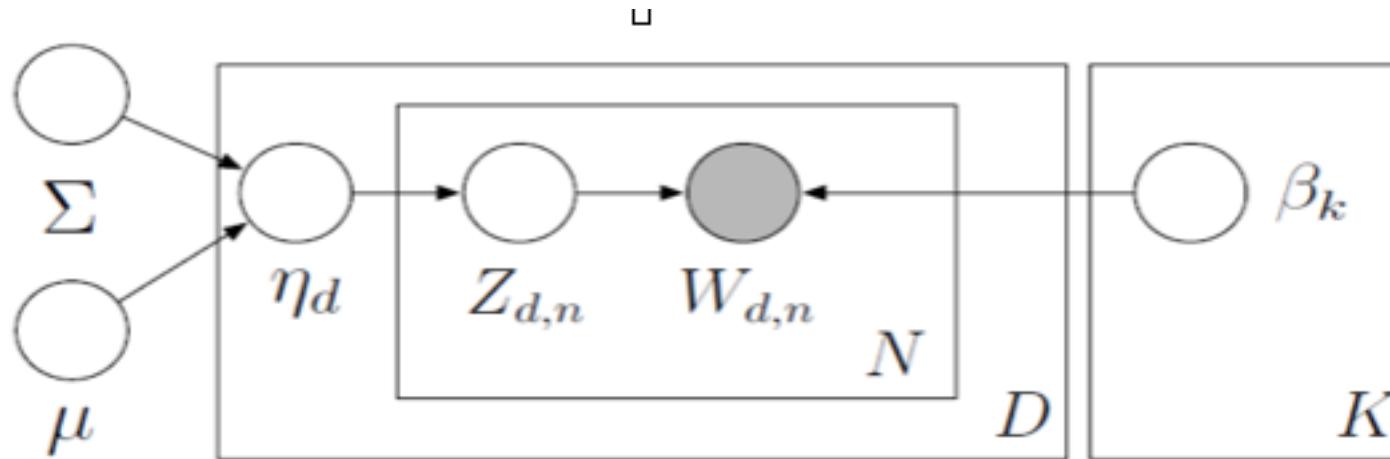
Pros/Cons of Gibbs sampling

- Pros:
 - Ease of implementation
 - Fast iterations (no transcendental functions), fast convergence (at least relative to a full Gibbs sampler)
 - Low memory usage (only require $O(MN)$ storage for the current values of $z_{j,t}$)
- Cons:
 - No obvious parallelization strategy (each iteration depends on previous)
 - Can be difficult to assess convergence

LDA results

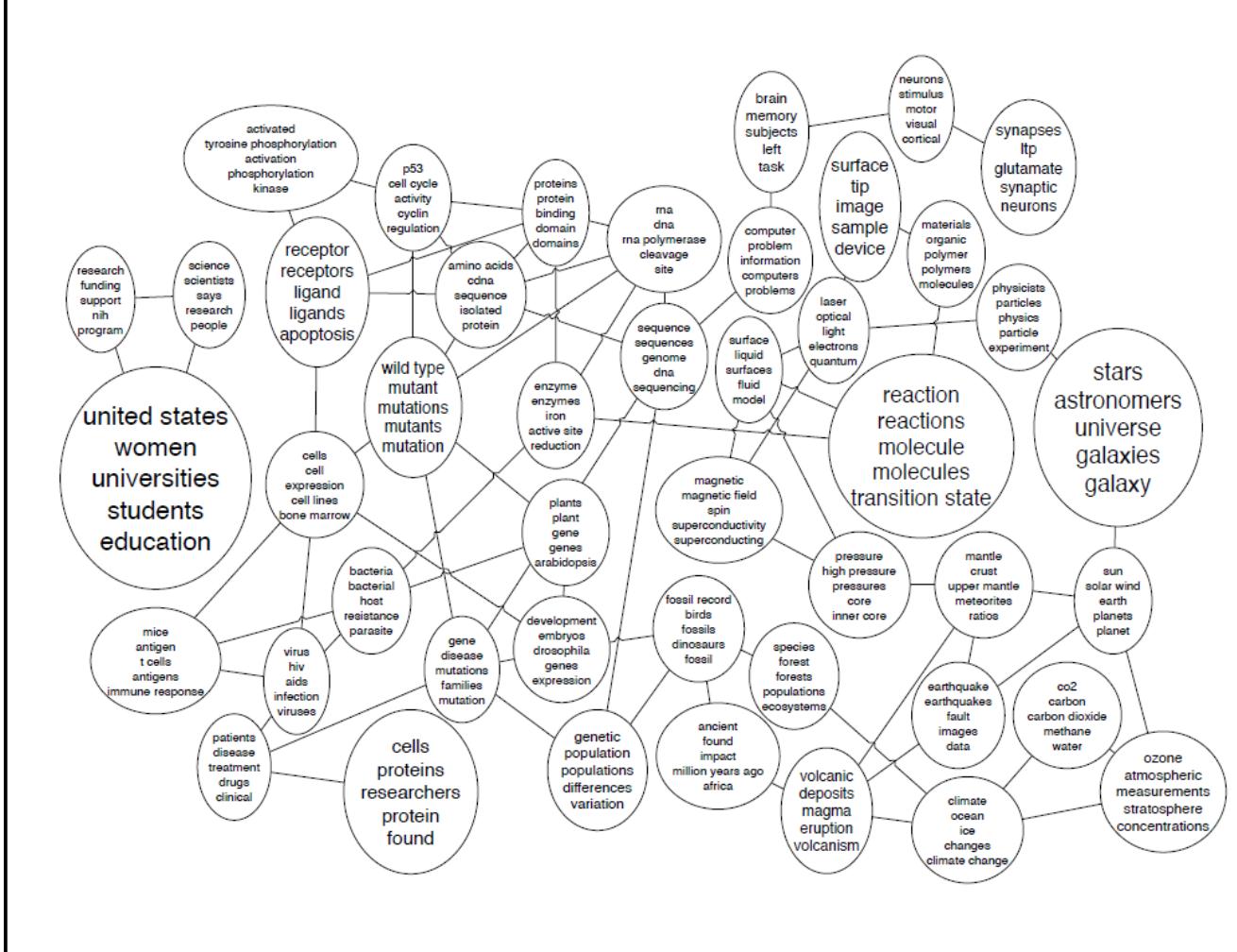


Correlated topic models (CTM) [Blei et al. 2005]

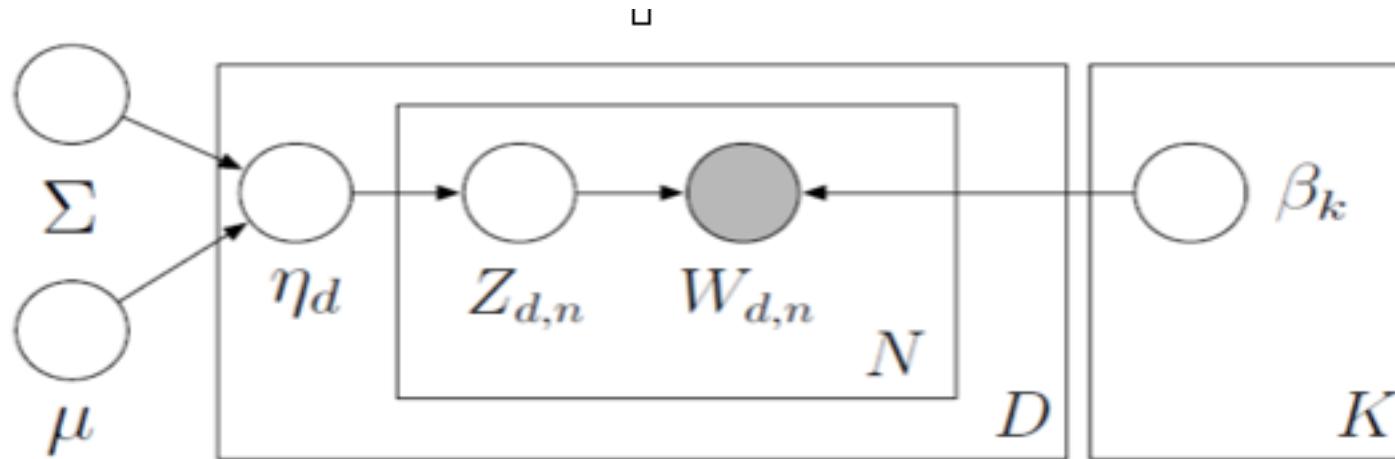


- Draw topics from a logistic normal, where topic occurrence can exhibit correlations

Correlated topic models (CTM) [Blei et al. 2005]

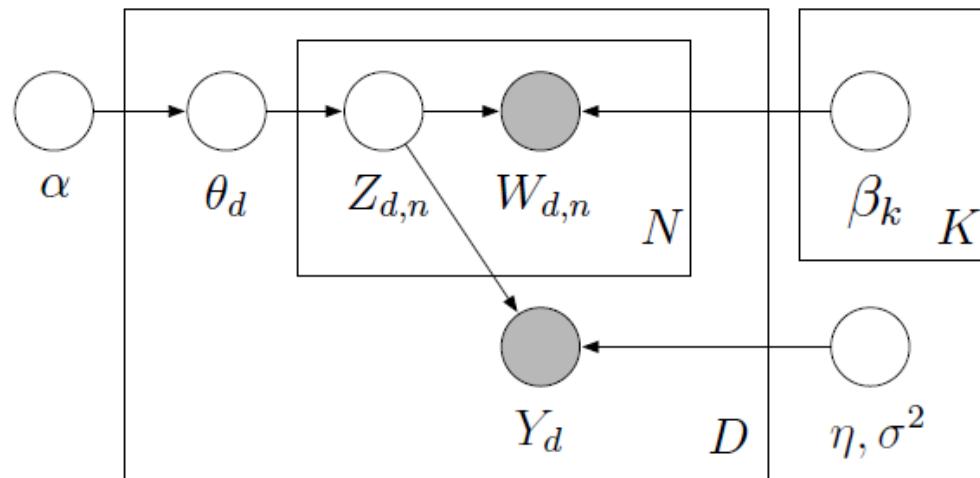


Correlated topic models (CTM) [Blei et al. 2005]



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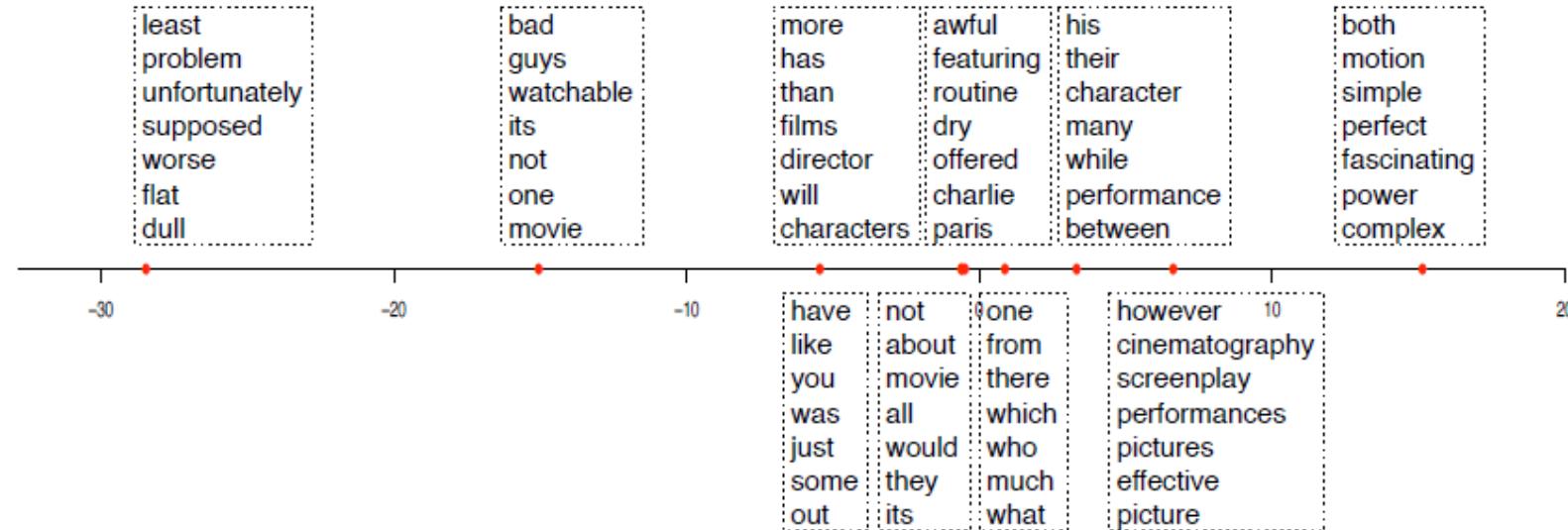
Supervised LDA [Blei & McAuliffe 07]



- ① Draw topic proportions $\theta | \alpha \sim \text{Dir}(\alpha)$.
- ② For each word
 - Draw topic assignment $z_n | \theta \sim \text{Mult}(\theta)$.
 - Draw word $w_n | z_n, \beta_{1:K} \sim \text{Mult}(\beta_{z_n})$.
- ③ Draw response variable $y | z_{1:N}, \eta, \sigma^2 \sim N(\eta^\top \bar{z}, \sigma^2)$, where

$$\bar{z} = (1/N) \sum_{n=1}^N z_n.$$

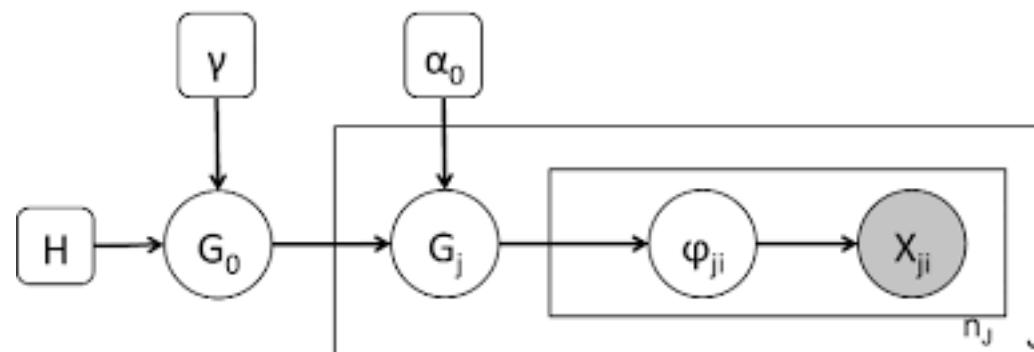
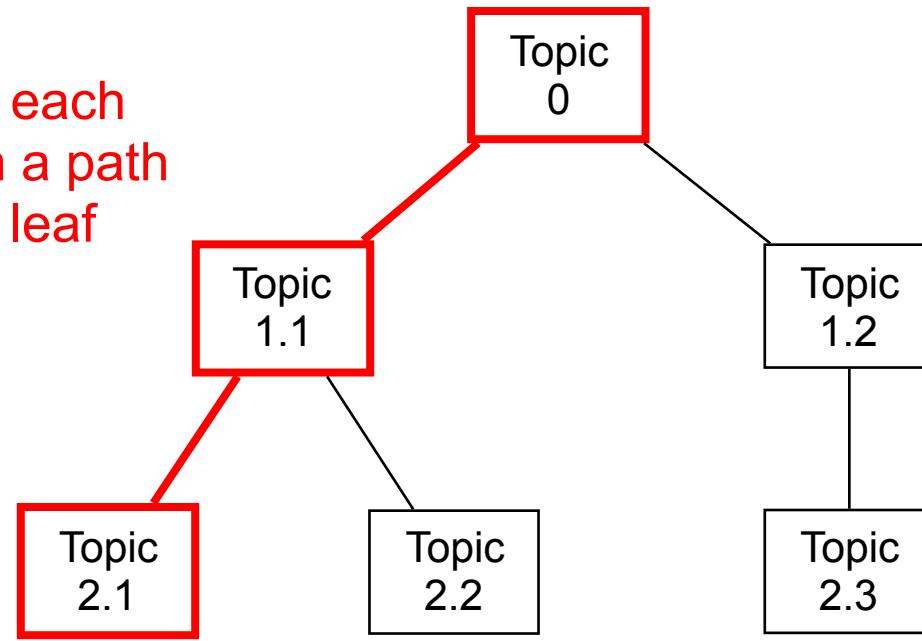
Supervised LDA [Blei & McAuliffe 07]



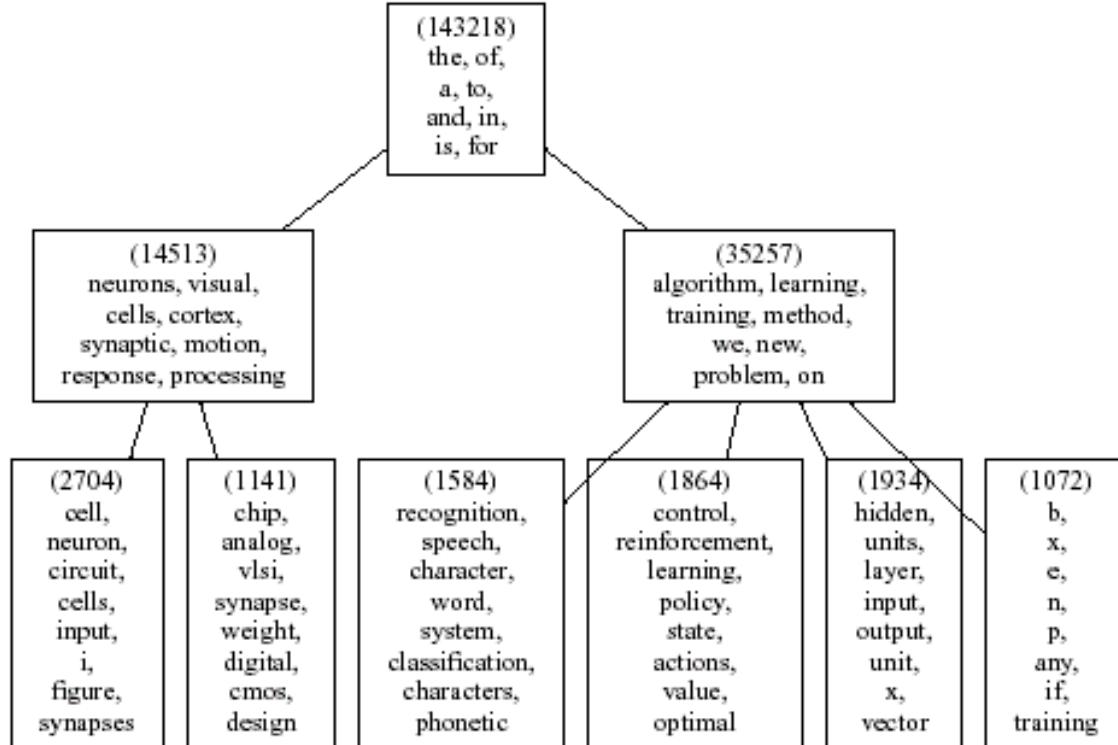
- 10-topic sLDA model on movie reviews (Pang and Lee, 2005).
- Response: number of stars associated with each review

Learning topic hierarchies

The topics in each document form a path from root to leaf



Learning topic hierarchies



Today's lecture

- Latent semantic indexing
- Continue on topic model
 - Bayesian inference of topic model
 - Variational inference for LDA
 - Gibbs sampling, Markov chain Monte-Carlo