

CS5340 - Lab2 Part2 Report

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Question We suppose that the full observation is generated by a linear-gaussian model.

$$p(x_u | x_{\pi_u}, \theta_u) = \mathcal{N}_u[w_{u0}, \dots, w_{uC}, \sigma_u^2] = \frac{1}{\sqrt{2\pi\sigma_u^2}} \exp \left\{ -0.5 \frac{\left(x_u - \left(\sum_{c \in x_{\pi_u}} w_{uc} x_{uc} + w_{u0} \right) \right)^2}{\sigma_u^2} \right\} \quad (1)$$

To derive the MLE estimates of θ , we can equivalently find the maximum log-likelihood estimate:

$$\begin{aligned} \arg \max_{\theta_u} \sum_{n=1}^N \log p(x_u | x_{\pi_u}, \theta_u | x_{\pi_u}) = \\ \underbrace{\arg \max_{\theta_u} \sum_{n=1}^N \left\{ -\frac{1}{2} \log 2\pi\sigma_u^2 - \frac{1}{2\sigma_u^2} \left(x_{u,n} - \left(\sum_{c \in x_{\pi_u}} w_{uc} x_{uc,n} + w_{u0} \right) \right)^2 \right\}}_{\mathcal{L}} \end{aligned} \quad (2)$$

Compute the MLE estimates of $\theta_u = (W, \sigma_u^2)$.

Derivation of W Let $W = \begin{bmatrix} w_{u0} \\ w_{u1} \\ \vdots \\ w_{uC} \end{bmatrix}$, Take the derivative of \mathcal{L} with respect to w_{u0} :

$$\frac{\partial \mathcal{L}}{\partial w_{u0}} = \sum_{n=1}^N \left\{ \frac{1}{\sigma_u^2} \left(x_{u,n} - \left(\sum_{c \in x_{\pi_u}} w_{uc} x_{uc,n} + w_{u0} \right) \right) \right\} = 0 \quad (3)$$

Take the derivative of \mathcal{L} with respect to $w_{uc}, c = 1, \dots, C$, C is the number of parents of x_u :

$$\frac{\partial \mathcal{L}}{\partial w_{uc}} = \sum_{n=1}^N \left\{ \frac{1}{\sigma_u^2} \left(x_{u,n} - \left(\sum_{c \in x_{\pi_u}} w_{uc} x_{uc,n} + w_{u0} \right) \right) x_{uc,n} \right\} = 0 \quad (4)$$

$$\begin{aligned} \text{Let } Y = \begin{bmatrix} \sum_{n=1}^N x_{u,n} \\ \sum_{n=1}^N x_{u,n} x_{u1,n} \\ \vdots \\ \sum_{n=1}^N x_{u,n} x_{uC,n} \end{bmatrix}, \quad A = \begin{bmatrix} 1 & x_{u1,1} & x_{u2,1} & \cdots & x_{uC,1} \\ 1 & x_{u1,2} & x_{u2,2} & \cdots & x_{uC,2} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_{u1,N} & x_{u2,N} & \cdots & x_{uC,N} \end{bmatrix}, \\ X = A^T A = \begin{bmatrix} N & \sum_{n=1}^N x_{u1,n} & \cdots & \sum_{n=1}^N x_{uC,n} \\ \sum_{n=1}^N x_{u1,n} & \sum_{n=1}^N x_{u1,n}^2 & \cdots & \sum_{n=1}^N x_{u1,n} x_{uC,n} \\ \vdots & \vdots & \ddots & \vdots \\ \sum_{n=1}^N x_{uC,n} & \sum_{n=1}^N x_{uC,n} x_{u1,n} & \cdots & \sum_{n=1}^N x_{uC,n}^2 \end{bmatrix}, \end{aligned}$$

then we can rewrite eq. (3) and eq. (4) as

$$XW = Y$$

Since each line of A is an observation of the parents of x_u , we can assume A is full rank. Thus $X = A^T A$ is also full rank. So the solution is

$$W = X^{-1}Y$$

Derivation of σ_u^2 Take the derivative of \mathcal{L} to σ_u^2 :

$$\frac{\partial \mathcal{L}}{\partial \sigma_u^2} = \sum_{n=1}^N \left\{ -\frac{1}{2\sigma_u^2} + \frac{1}{2\sigma_u^4} \left(x_{u,n} - \left(\sum_{c \in x_{\pi_u}} w_{uc} x_{uc,n} + w_{u0} \right) \right)^2 \right\} = 0 \quad (5)$$

Let $Z = \sum_{n=1}^N \left(x_{u,n} - \left(\sum_{c \in x_{\pi_u}} w_{uc} x_{uc,n} + w_{u0} \right) \right)^2$, then we can rewrite eq. (5) as

$$\frac{N}{2\sigma_u^2} = \frac{Z}{2\sigma_u^4}$$

so the solution is

$$\sigma_u^2 = \frac{Z}{N}$$

Implementation The code is just an implementation of the above derivation.