Lecture 4

Optimal Estimation Theory and Its Applications

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Outline

- Cramér-Rao Bound
- ② Expectation Maximization Method

The Fact

Different performance indexes, for example MSE or MLE, derive different optimal estimators.

Question

If there exists the common performance bound of all estimators?

New Concept

Cramér-Rao Bound

(1) Case 1

If \hat{X} is any unbiased estimate of a deterministic variable X based on the measurement Z, then the covariance of the estimation error $\tilde{X}=X-\hat{X}$ is bounded by

$$P_{\tilde{X}} \geqslant J_F^{-1} \tag{1}$$

where the Fisher information matrix is given by

$$J_F = E\left\{ \left[\frac{\partial \ln f_{Z|X}(z|x)}{\partial x} \right] \left[\frac{\partial \ln f_{Z|X}(z|x)}{\partial x} \right]^T \right\} = -E\left[\frac{\partial^2 \ln f_{Z|X}(z|x)}{\partial x^2} \right]$$
(2)

Equality holds in Eq.1 if and only if

$$\frac{\partial \ln f_{Z|X}(z|x)}{\partial x} = k(x)(x - \hat{X}) \tag{3}$$

It is assumed that $\partial \ln f_{Z|X}/\partial x$ and $\partial^2 \ln f_{Z|X}/\partial x^2$ exist and are absolutely integrable.

An estimate \hat{X} is efficient if it satisfies the Cramér-Rao bound with equality, that is, if Eq.3 holds.

For linear Gaussian measurement model Z=HX+V with V is a zero-mean Gaussian noise with covariance R, the MLE of X and its covariance are

$$\hat{X}_{ML} = (H^T R^{-1} H)^{-1} H^T R^{-1} z \quad P_{ML} = (H^T R^{-1} H)^{-1}$$

$$\hat{X}_{ML} \text{ is unbiased estimate.}$$

$$J_F = -E \left[\frac{\partial^2 \ln f_{Z|X}(z|x)}{\partial x^2} \right] = H^T R^{-1} H$$

$$P_{\mathsf{ML}} = J_F^{-1}$$

Conclusion: The MLE in this case is efficient.

 $\hat{X}_{ML} = (H^T R^{-1} H)^{-1} H^T R^{-1} z$ is unbiased estimate.

 $f_{Z|X}(z|x) = \exp\{-(z - Hx)^T R^{-1}(z - Hx)/2\}/\sqrt{2\pi |R|}$

or

$$\begin{split} & \ln f_{Z|X}(z|x) = -(z-Hx)^T R^{-1}(z-Hx)/2 - \ln(2\pi |R|)/2 \\ & \frac{\partial \ln f_{Z|X}(z|x)}{\partial x} = H^T R^{-1}(z-Hx) = -H^T R^{-1}(Hx-z) \\ & = -(H^T R^{-1}H)x + (H^T R^{-1}H)(H^T R^{-1}H)^{-1}H^T R^{-1}z \\ & = -(H^T R^{-1}H)\left(x - (H^T R^{-1}H)^{-1}H^T R^{-1}z\right) \\ & \text{From the expression of the efficient estimate} & \frac{\partial \ln f_{Z|X}(z|x)}{\partial x} = k(x)(x-\hat{X}), \\ & \text{we find } \hat{X}_{ML} \text{ is the efficient estimate.} \end{split}$$

The ML estimate has several nice properties as the number of independent measurements N goes to infinity

- ① \hat{X}_{ML} converges in probability to the correct value of X as $N \to \infty$. (Consistent)
- ② \hat{X}_{ML} becomes efficient as $N \to \infty$.
- ③ \hat{X}_{ML} becomes Gaussian N(X, PX) as $N \to \infty$.

Stochastic convergence: the sequence $Y_1,\,Y_2,\cdots$ of RVs (random variables) converges in probability to $RV\ Y$ if for all $\varepsilon>0$

$$P(||YN - Y|| > \varepsilon) \to 0 \text{ for } N \to \infty$$

(1)Case 2

If \hat{X} is any estimate of a stochastic variable X based on measurement Z, then the covariance of the estimation error $\tilde{X}=X-\hat{X}$ is bounded by

$$P_{\tilde{X}} \geqslant L^{-1} \tag{4}$$

where the Fisher information matrix is given by

$$L = E\left\{ \left[\frac{\partial \ln f_{XZ}(x,z)}{\partial x} \right] \left[\frac{\partial \ln f_{XZ}(x,z)}{\partial x} \right]^T \right\} = -E\left[\frac{\partial^2 \ln f_{XZ}(x,z)}{\partial x^2} \right]$$
(5)

Equality holds in Eq.4 if and only if

$$\frac{\partial \ln f_{XZ}(x,z)}{\partial x} = k(x - \hat{X}) \tag{6}$$

an estimate \hat{X} is efficient if Eq.3 holds.

It is assumed that $\partial \ln f_{Z|X}/\partial x$ and $\partial^2 \ln f_{Z|X}/\partial x^2$ exist and are absolutely integrable with respect to both variables, and .

$$\lim_{x \to +\infty} B(x) f_X(x) = 0, \quad \lim_{x \to -\infty} B(x) f_X(x) = 0 \tag{7}$$

with
$$B(x)\underline{\underline{\underline{\Delta}}}\int_{-\infty}^{+\infty}(x-\hat{X})f_{Z|X}(z|x)\mathrm{d}z$$

To a deterministic variable, its Cramér-Rao Bound

- ullet depends on the likelihood function $f_{Z|X}$
- requires the unbiasedness
- has the efficiency condition about k(x)

$$\frac{\partial \ln f_{Z|X}(z|x)}{\partial x} = k(x)(x - \hat{X})$$

To a random variable, its Cramér-Rao Bound

- ullet depends on the joint PDF f_{XZ}
- requires $\lim_{x \to +\infty} B(x) f_X(x) = 0$, $\lim_{x \to -\infty} B(x) f_X(x) = 0$
- has the efficiency condition about k

$$\frac{\partial \ln f_{XZ}(x,z)}{\partial x} = k(x - \hat{X})$$

To a random variable, we have

$$\begin{split} \frac{\partial \ln f_{XZ}(x,z)}{\partial x} &= \frac{\partial \ln f_{X|Z}(x|z)}{\partial x} + \frac{\partial \ln f_{Z}(z)}{\partial x} = \frac{\partial \ln f_{X|Z}(x|z)}{\partial x} \\ \operatorname{From} &\frac{\partial \ln f_{XZ}(x,z)}{\partial x} = k(x-\hat{X}), \text{ we have} \\ f_{X|Z}(x|z) &= e^{-(k/2)x^Tx + k\hat{X}^Tx + c} \end{split}$$

It is concluded that its efficient estimate necessarily have the Gaussian posterior conditional PDF $f_{X\mid Z}$

MLE in the case of data missing or partly-unobservable data

Consider the observation data Z_{obs} , and partly-unobservable data Z_{mis} . They constitutes the data set $Z=Z_{obs}\cup Z_{mis}$

Here $p\{Z_{mis}|Z_{obs},X\}$ and $p\{Z_{obs},Z_{mis}|X\}$ are known.

Our aim is to obtain the ML estimate, i.e., $\arg\left\{\max_{X}p\{Z_{obs}|X\}\right\}$

based on the available data Z_{obs} and apriori information

$$p\{Z_{mis}|Z_{obs},X\}$$
 and $p\{Z_{obs},Z_{mis}|X\}$.

Denote $L(X) = \ln p\{Z_{obs}|X\}$. To obtain the iterative estimate, we define the estimate after the k - th iteration as $\hat{X}^{(k)}$, and have $L(X) - L(\hat{X}^{(k)}) = \ln \frac{p\{Z_{obs}|X\}}{p\{Z_{obs}|\hat{X}^{(k)}\}} = \ln \sum_{Z_{mis}} \frac{p\{Z_{obs}, Z_{mis}|X\}}{p\{Z_{obs}|\hat{X}^{(k)}\}}$ $= \ln \sum_{Z_{mis}} \frac{p\{Z_{obs}|Z_{mis}, X\}p\{Z_{mis}|X\}}{p\{Z_{obs}|\hat{X}^{(k)}\}}$

If the operator \ln and \sum can exchange the order, then the computation will be easily implemented.

$$L(X) - L(\hat{X}^{(k)}) = \ln \sum_{Z_{mis}} \frac{p\{Z_{obs}|Z_{mis}, X\}p\{Z_{mis}|X\}p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\}}{p\{Z_{obs}|\hat{X}^{(k)}\}p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\}}$$

From the Jensen Inequality $\sum\limits_{j}\lambda_{j}=1\Rightarrow\ln(\sum\limits_{j}\lambda_{j}y_{j})\geqslant\sum\limits_{j}\lambda_{j}\ln y_{j}$ and the fact that $\sum\limits_{Z_{mis}}p\{Z_{mis}|Z_{obs},\hat{X}^{(k)}\}=1$, we have

$$L(X) - L(\hat{X}^{(k)}) \geqslant$$

$$\sum_{Z_{mis}} p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\} \ln \left[\frac{p\{Z_{mis}|X\}p\{Z_{obs}|Z_{mis}, X\}}{p\{Z_{obs}|\hat{X}^{(k)}\}p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\}} \right]$$

In other words, the operators exchange their order via the Jensen Inequality.

$$L(X) \geqslant L(\hat{X}^{(k)}) + \sum_{Z_{mis}} p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\} \ln \left[\frac{p\{Z_{mis}|X\}p\{Z_{obs}|Z_{mis}, X\}}{p\{Z_{obs}|\hat{X}^{(k)}\}p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\}} \right]$$

Through conditional probability integral (called conditional mean), the unobservable data is smoothed.

Given the initial estimate value $\hat{X}^{(0)}$, iteratively find the maximum point of $\hat{X}^{(k+1)} = \arg\max_X l(X,\hat{X}^{(k)})$ until $\hat{X}^{(k+1)} = \hat{X}^{(k)}$

We have

$$\begin{split} \hat{X}^{(k+1)} &= \arg\max_{X} \\ &\left[L(\hat{X}^{(k)}) + \sum_{Z_{mis}} p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\} \ln\left[\frac{p\{Z_{mis}|X\}p\{Z_{obs}|Z_{mis}, X\}}{p\{Z_{obs}|\hat{X}^{(k)}\}p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\}} \right] \right] \\ &= \arg\max_{X} \left[\sum_{Z_{mis}} p\{Z_{mis}|Z_{obs}, \hat{X}^{(k)}\} \ln\left[p\{Z_{mis}|X\}p\{Z_{obs}|Z_{mis}, X\} \right] \right] \\ &= \arg\max_{X} \left[E_{Z_{mis}|Z_{obs}, \hat{X}^{(k)}} \ln\left[p\{Z_{obs}, Z_{mis}|X\} \right] \right] \end{split}$$

It is proved that such iterative computation always converge.

Remark: $p\{Z_{mis}|Z_{obs},X\}=p\{Z_{mis},Z_{obs}|X\}/p\{Z_{obs}|X\}$ is known

The implementation of EM algorithm includes two steps

① Expectation (E-Step)

$$Q(X|\hat{X}^{(k+1)}) = E_{Z_{mis}|Z_{obs},\hat{X}^{(k)}} \ln [p\{Z_{obs}, Z_{mis}|X\}]$$

② Maximization(M-Step)

$$\hat{X}^{(k+1)} = \arg\max_{X} \left[Q(X|\hat{X}^{(k+1)}) \right]$$

McLachlan G J, Krishnan T. **The EM Algorithm and Extension**. New York: Wiley,1997