Supplementary information: A simple explanation for declining temperature sensitivity with warming

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1 A first-hitting-time model of leafout

Our model follows the general understanding of how warm temperatures (forcing) trigger leafout in temperate deciduous trees (Chuine, 2000). We use a first-hitting-time model, which describes the first time a random process hits a threshold, because of its broad applicability and conceptual simplicity. We define leafout day, n_{β} , as the day, n, that cumulative daily temperature, S^n , hits the threshold, β .

We derive the relationship between daily temperature and leafout in two common scenarios. In the first, we take the average daily temperature up until the leafout date. In the second, we take the average daily temperature over a fixed window, such as March 1st to April 30th. In both cases, we discretize time since, although many biological processes depend continuously on time, research typically measures time in discretized units, such as days, weeks, or months.

1.1 Scenario 1: Using average daily temperature until the leafout date

We use the following notation:

 $X_n =$ observed temperature on day n $S_0^n = \sum_{i=0}^n X_i, \text{the cumulative daily temperature from day 0 to day } n$

 $S_0 = \sum_{i=0}^{N} X_i$, the cumulative daily temperature from day 0 to day N

n = day since temperatures start to accumulate, n = 0, 1, ..., N

 $M_0^n = \frac{S_0^n}{n}$, the average daily temperature from day 0 to day n

 $\beta=$ the threshold of interest, $\beta>0,$ (thermal sum required for leafout)

 $n_{\beta} = \underset{n}{\operatorname{argmin}} S_n > \beta$, the first day the cumulative daily temperature passes the threshold (for example, day of year (doy) of leafout).

We model X_n as a Gaussian random walk, $X_n \stackrel{\text{i.i.d}}{\sim} \operatorname{normal}(\alpha_0 + \alpha_1 n, \sigma)$, where $\alpha_0 > 0$ is the average temperature on day n = 0, $\alpha_1 > 0$ is the day-over-day increase in average temperatures, and σ is the standard deviation. This model differs from the traditional Gaussian random walk because of the factor n.

This model has two important consequences:

(1) Leafout time is inversely related to average temperature at leafout time.

Under this model, $M_0^{n_\beta}$ and n_β are inversely proportional. To see why, assume for the moment that the cumulative daily temperature hits the threshold exactly on leafout day. That is, $S_0^{n_\beta} = \beta$. Then

$$M_0^{n_\beta} = \frac{S_0^{n_\beta}}{n_\beta} = \frac{\beta}{n_\beta}$$

rearranging yields

$$n_{\beta} = \frac{\beta}{M_0^{n_{\beta}}}$$

Many global change biology studies use linear regression to quantify the relationship between n_{β} and $M_0^{n_{\beta}}$ (or similar metrics, see Wolkovich et al., 2012; Piao et al., 2017; Keenan et al., 2020, for examples). Regressing n_{β} on $M_0^{n_{\beta}}$ finds a best fit line to the inverse curve, $n_{\beta} = \frac{\beta}{M_0^{n_{\beta}}}$. The relationship is linearized with the logarithm transformation: $\log(n_{\beta}) = \log(\beta) - \log(M_0^{n_{\beta}})$. That is, $\log(n_{\beta})$ is linear in log-average daily temperature with slope -1 and intercept $\log(\beta)$.

(2) The variance of the average temperature may decreases as temperatures rise.

Under the model, the mean and variance of M_0^n are $\mathrm{E}(M_0^n|\alpha_0,\alpha_1)=\frac{1}{n}\sum_{i=0}^n(\alpha_0+\alpha_1i)=\alpha_0+\alpha_1\frac{(n+1)}{2}$ and $\mathrm{Var}(M_0^n|\alpha_0,\alpha_1)=\frac{\sigma^2}{n}$.

By the law of total variance,

$$\begin{aligned} \operatorname{Var}(M_0^n) &= \operatorname{E}(\operatorname{Var}(M_0^n | \alpha_0, \alpha_1)) + \operatorname{Var}(\operatorname{E}(M_0^n | \alpha_0, \alpha_1)) \\ &= \frac{\sigma^2}{n} + \operatorname{Var}(\alpha_0 + \alpha_1 \frac{n+1}{2}) \\ &= \frac{\sigma^2}{n} + \operatorname{Var}(\alpha_0) + \frac{(n+1)^2}{4} \operatorname{Var}(\alpha_1) + (n+1) \operatorname{Cov}(\alpha_0, \alpha_1) \end{aligned}$$

As temperatures rise and leafout date becomes earlier, the variance of the average temperature will decline—provided the variation in temperatures, σ , is sufficiently small.

1.2 Scenario 2: Using average daily temperature over a fixed window

We slightly modify the notation:

 $n = \text{day since temperatures start to accumulate}, n = 0, \dots, a, \dots, b$

 $X_n =$ observed temperature on day n

 $S_a^n = \sum_{i=a}^n X_i$, the cumulative daily temperature from day a to day n

 $M_a^n = \frac{S_a^n}{n-a}$, the average daily temperature from day a to day n

 β = the threshold of interest, $\beta > 0$, (thermal sum required for leafout)

 $n_{\beta} = \underset{n}{\operatorname{argmin}} S_0^n > \beta$, the first day the cumulative daily temperature passes the threshold

(for example, day of year (doy) of leafout).

As before, we model X_n as a Gaussian random walk, $X_n \stackrel{\text{i.i.d}}{\sim} \text{normal} (\alpha_0 + \alpha_1 n, \sigma)$, where $\alpha_0 > 0$ is the average temperature on day n = 0, $\alpha_1 > 0$ is the day-over-day increase in average temperatures, and σ is the standard deviation. We make the additional assumption that $X_n \geq 0$ for all n and $n < n_{\beta} < n_{\beta}$. That is, the cumulative temperature acquired by the plant always increases.

Note that

$$S_a^b \sim \text{normal}\left(\alpha_0(b-a) + \frac{\alpha_1}{2}(b-a)(b+a+1), \sigma\sqrt{b-a}\right)$$

$$M_a^b \sim \text{normal}\left(\alpha_0 + \frac{\alpha_1}{2}(b+a+1), \frac{\sigma}{\sqrt{b-a}}\right)$$

$$S_n^b - S_a^b \sim \text{normal}\left(\alpha_0(b - a - n) + \frac{\alpha_1}{2}((b - n)(b + n + 1) - a(a + 1)), \sigma\sqrt{b + a - n}\right)$$

so that

$$Pr\left(n_{\beta} \leq n \mid M_{a}^{b} = m\right) = Pr\left(n_{\beta} \leq n \mid S_{a}^{b} = (b - a)m\right)$$

$$= Pr\left(S_{0}^{n} \geq \beta \mid S_{a}^{b} = (b - a)m\right)$$

$$= Pr\left(S_{n}^{b} \leq (b - a)m + S_{0}^{a} - \beta\right)$$

$$= Pr\left(S_{n}^{b} - S_{0}^{a} \leq (b - a)m - \beta\right)$$

$$= \Phi\left(\frac{(b - a)m - \beta - [\alpha_{0}(b - a - n) + \frac{\alpha_{1}}{2}((b - n)(b + n + 1) - a(a + 1))]}{\sigma\sqrt{b + a - n}}\right)$$

The distribution of M_a^b shows that consequence (2) above still holds with this model. Consequence (1) no longer holds directly, but will in many situations where average daily temperature until an event correlates strongly with average daily temperature because the window is chosen based, in part, on the expected hitting time (Figs. S1-S2). We note two additional consequences:

(3) The conditional median is quadratic in n:

$$\frac{1}{2} \stackrel{\text{set}}{=} Pr\left(n_{\beta} \le n \mid M_{a}^{b} = m\right)
\Rightarrow 0 = (b-a)m - \beta - \left[\alpha_{0}(b-a-n) + \frac{\alpha_{1}}{2}((b-n)(b+n+1) - a(a+1))\right]
\Rightarrow m = \frac{1}{(a-b)} \left[-\beta - \alpha_{0}(b-a-n) - \frac{\alpha_{1}}{2}((b-n)(b+n+1) - a(a+1))\right]
= \frac{1}{(a-b)} \left[-\beta - \alpha_{0}(b-a) - \frac{\alpha_{1}}{2}(b-a)(b+a+1)\right] + \frac{\alpha_{0} + \frac{\alpha_{1}}{2}}{(a-b)}n + \frac{\frac{\alpha_{1}}{2}}{(a-b)}n^{2}
:= \gamma_{0} + \gamma_{1}n + \gamma_{2}n^{2}$$

(4) The conditional mean and variance are sums of negative sigmoids, according to the following identities

$$E\left(n_{\beta} \mid M_{a}^{b} = m\right) = \sum_{n=0}^{\infty} Pr\left(n_{\beta} \ge n \mid M_{a}^{b} = m\right)$$
$$E\left(n_{\beta}^{2} \mid M_{a}^{b} = m\right) = \sum_{n=0}^{\infty} n Pr\left(n_{\beta} \ge n \mid M_{a}^{b} = m\right)$$

1.3 Effect of shifting β on sensitivity

Many biological hypotheses of why plant sensitivities should decline suggest a shifting thermal sum required for leafout, due to declines in over-winter chilling or short photoperiods. However, the first-hitting model suggests that if β (thermal sum for leafout) is changing over time for any reason (increasing or decreasing), the regression coefficient from regressing leafout date on spring-time temperature (commonly referred to as 'sensitivity') should decrease. Using the total law of covariance: sensitivity from changing $\beta = Cov(Y, X)/Var(X)$:

$$Cov(Y,X)/Var(X) = [Cov(E(Y|\beta), E(X)) + E(Cov(Y,X|\beta))]/Var(X)(Cov(Y,X|\beta)/Var(X))$$

[Jonathan, can you work on the text here?] This, however, assumes X and β are independent and $Cov(E(Y|\beta), E(X)) < 0$. The latter is often true because increased mean temperatures (X) lead to earlier leafout dates (Y). The former would be false if mean temperatures (X) are correlated with over-winter chilling, or other covariates that lead to shifts in β . Then, likely $Var(X|\beta) \leq Var(X)$ and then the inequality above could be reversed.

2 Common hypotheses for declining sensitivity

We extended simulations from our first hitting-time model to examine two common biological hypotheses for declining sensitivities, and to understand how easily discernible biological shifts would be from the shifts always present when using linear models to estimate sensitivities.

First, we simulated the most often cited hypothesis (e.g., Fu et al., 2015; Piao et al., 2017): an increasing thermal sum threshold given declines in over-winter chilling. Experiments in controlled environments show that plants require greater thermal sums to budburst or leafout given lower chilling (e.g., Laube et al., 2014; Flynn and Wolkovich, 2018). Connecting chilling to leafout in observational data, however, is difficult as we know little mechanistically about what temperatures determine chilling (but see Rinne et al., 2011; van der Schoot et al., 2014, for recent developments in lab molecular work), and most models of chilling are based on statistical correlations (Erez and Lavee, 1971; Richardson, 1974; Luedeling et al., 2009). Without robust estimates of chilling, teasing out effects of chilling from forcing are likely impossible as the model is non-identifiable when these two factors cannot be separated (ADD: examples from simulations of same leafout date given a few different chilling and forcing requirements).

We further found that under this model—where the thermal sum (β) increases as over-winter chilling declines—would lead to larger estimated sensitivities (see Fig. S3), in contrast with the general expectation in the literature that a larger role for chilling would reduce the estimated sensitivity. The actual effect however, depends on the covariation of X and β (discussed above) and highlights how difficult teasing apart relationships in observational spring phenology may be given correlations across multiple predictors (correlations that are likely exacerbated by climate change). Effects of biological shifts in the drivers of leafout are far more obvious in statistical estimates that include the non-linear nature of plants temperature response compared to linear estimates (see top panel of Fig. S4).

Second, we simulated an alternative hypothesis where warming causes the thermal sum to be reached before a required daylength threshold and plants then leaf out on the first day the daylength threshold is met (CITES). At its extreme this model produces the same leafout day (the day when the required daylength is met) across different temperatures and thus does produce a smaller estimated sensitivity. Estimated sensitivities, however, from a linear model do not necessarily decline depending on exact parameter values. Again this shift appears more obvious from statistical estimates that include the non-linear nature of plants temperature response compared to linear estimates (see Fig. S5).

3 Results using long-term empirical data from PEP725

To examine how estimated sensitivities shift over time, we selected sites of two common European tree species (silver birch, Betula pendula, and European beech, Fagus sylvatica) that have long-term observational data of leafout, through the Pan European Phenology Project (PEP725, Templ et al., 2018). We used a European-wide gridded climate dataset (E-OBS, Cornes et al., 2018) to extract daily minimum and maximum temperature for the grid cells where observations of leafout for these two species were available. We used sites with complete leafout data across both our 10-year (and 20-year) windows to avoid possible confounding effects of shifting sites over time (see Tables S2-S3 for numbers of sites per species x window).

Our estimates of temperature sensitivity from a linear model using untransformed variables show

a decline in sensitivity with recent warming for *Betula pendula* over 10 and 20-year windows, but no decline for *Fagus sylvatica*; using logged variables estimates appeared more similar over time or sometimes suggested an increase in sensitivity (see Figs. S6-S7, Tables S2-S3). This shift in estimated sensitivity when regressing with untransformed versus logged variables suggests the declining estimates with untransformed variables may not be caused by changes in the underlying mechanisms of leafout (i.e., reduced winter chilling) and driven instead by using linear regression for a non-linear process. This hypothesis is supported further by large declines in variance of leafout in recent decades.

Shifts in variance provide another hurdle to robust estimates of temperature sensitivity. Previous work has highlighted how shifting temperature variance (over space and/or time) could lead to shifting estimates of temperature sensitivities (Keenan et al., 2020), but our results stress that variance in both leafout and temperature are shifting. If both shift in step, estimates would not be impacted by changes in temperature variance, but our results suggest variance in temperature—for these data—has declined more than variance in leafout, though both have declined substantially in recent decades (Tables S2-S3).

Estimated sensitivities for the empirical data (PEP725) using logged variables are far lower than the value obtained in our simulations (-1). This likely results from a contrast between our simulations—where we can accurately define the temperature plants experience and the temporal window that drives leafout—and our empirical data, where we do not know how measured temperatures translate into the temperatures that plants accumulate and where we have no clear method to define the relevant temporal window (Güsewell et al., 2017).

These results highlight how the acceleration of biological time due to climate change requires researchers to clarify their assumptions. Expecting temperature sensitivity to remain constant as temperatures rise assumes the relationship between response and temperature is proportional. But the underlying biological processes suggest this relationship is seldom proportional, or even linear. In fact, when our model holds, declining sensitivity with rising temperatures should be the null hypothesis of any analysis of temperature sensitivity based on linear regression or similar methods.

References

Chuine, I. 2000. A unified model for budburst of trees. Journal of Theoretical Biology 207:337–347.

Cornes, R. C., G. van der Schrier, E. J. van den Besselaar, and P. D. Jones. 2018. An ensemble version of the E-OBS temperature and precipitation data sets. Journal of Geophysical Research: Atmospheres 123:9391–9409.

Erez, A., and S. Lavee. 1971. Effect of climatic conditions on dormancy development of peach buds. i. temperature. Amer Soc Hort Sci J.

- Flynn, D. F. B., and E. M. Wolkovich. 2018. Temperature and photoperiod drive spring phenology across all species in a temperate forest community. New Phytologist 219:1353–1362.
- Fu, Y. S. H., H. F. Zhao, S. L. Piao, M. Peaucelle, S. S. Peng, G. Y. Zhou, P. Ciais, M. T. Huang, A. Menzel, J. P. Uelas, Y. Song, Y. Vitasse, Z. Z. Zeng, and I. A. Janssens. 2015. Declining global warming effects on the phenology of spring leaf unfolding. Nature 526:104–107.
- Güsewell, S., R. Furrer, R. Gehrig, and B. Pietragalla. 2017. Changes in temperature sensitivity of spring phenology with recent climate warming in Switzerland are related to shifts of the preseason. Global Change Biology 23:5189–5202.
- Keenan, T. F., A. D. Richardson, and K. Hufkens. 2020. On quantifying the apparent temperature sensitivity of plant phenology. New Phytologist 225:1033–1040.
- Laube, J., T. H. Sparks, N. Estrella, J. Höfler, D. P. Ankerst, and A. Menzel. 2014. Chilling outweighs photoperiod in preventing precocious spring development. Global Change Biology 20:170–182.
- Luedeling, E., M. H. Zhang, G. McGranahan, and C. Leslie. 2009. Validation of winter chill models using historic records of walnut phenology. Agricultural and Forest Meteorology 149:1854–1864.
- Piao, S., Z. Liu, T. Wang, S. Peng, P. Ciais, M. Huang, A. Ahlstrom, J. F. Burkhart, F. Chevallier, I. A. Janssens, et al. 2017. Weakening temperature control on the interannual variations of spring carbon uptake across northern lands. Nature Climate Change 7:359.
- Richardson, E. 1974. A model for estimating the completion of rest for 'Redhaven' and 'Elberta' peach trees. HortScience 9:331–332.
- Rinne, P. L. H., A. Welling, J. Vahala, L. Ripel, R. Ruonala, J. Kangasjarvi, and C. van der Schoot. 2011. Chilling of dormant buds hyperinduces FLOWERING LOCUS T and recruits GA-Inducible 1,3-beta-Glucanases to reopen signal conduits and release dormancy in *Populus*. Plant Cell 23:130–146.
- Templ, B., E. Koch, K. Bolmgren, M. Ungersböck, A. Paul, H. Scheifinger, T. Rutishauser, M. Busto, F.-M. Chmielewski, L. Hájková, S. Hodzić, F. Kaspar, B. Pietragalla, R. Romero-Fresneda, A. Tolvanen, V. Vučetič, K. Zimmermann, and A. Zust. 2018. Pan European Phenological database (PEP725): a single point of access for European data. International Journal of Biometeorology 62:1109–1113.
- van der Schoot, C., L. K. Paul, and P. L. H. Rinne. 2014. The embryonic shoot: a lifeline through winter. Journal of Experimental Botany 65:1699–1712.
- Wolkovich, E. M., B. I. Cook, J. M. Allen, T. M. Crimmins, J. L. Betancourt, S. E. Travers, S. Pau, J. Regetz, T. J. Davies, N. J. B. Kraft, T. R. Ault, K. Bolmgren, S. J. Mazer, G. J. McCabe, B. J. McGill, C. Parmesan, N. Salamin, M. D. Schwartz, and E. E. Cleland. 2012. Warming experiments underpredict plant phenological responses to climate change. Nature 485:494–497.

4 Tables

Table S1: Number of sites for each species used in the study (*Betula pendula* and *Fagus sylvatica*) using the PEP725 over 10 and 20-year windows.

species	n sites						
	(1950-1960)	(2000-2010)	(1950-1970)	(1970-1990)	(1990-2010)		
Betula pendula	17	17	45	45	45		
Fagus sylvatica	24	24	47	47	47		
Quercus robur	20	20	43	43	43		

Table S2: Climate and phenology statistics for two species (Betula pendula, Fagus sylvatica, across 45 and 47 sites respectively) from the PEP725 data across all sites with continuous data from 1950-1960 and 2000-2010. ST is spring temperature from 1 March to 30 April, ST.leafout is temperature 30 days before leafout, and GDD is growing degree days 30 days before leafout. Slope represents the estimated sensitivity using untransformed leafout and ST, while log-slope represents the estimated sensitivity using log(leafout) and log(ST). We calculated all metrics for for each species x site x 10 year period before taking mean or variance estimates. See also Fig. S6.

		mean (ST)				var (ST)					slope			log-slope		
years	species	30	45	60	mean	30	45	60	var	GDD	30	45	60	30	45	60
					ST.lo				(lo)							
1950-1960	Betula	6.3	5.2	5.6	7.0	2.0	2.6	3.4	110.5	71.7	3.3	-	-	0.20	-	-
												2.1	4.3		0.09	0.17
2000-2010	Betula	6.1	4.9	6.6	6.8	3.7	2.4	1.2	47.0	64.6	-	0.5	-	0.00	0.02	-
											0.1		3.6			0.22
1950-1960	Fagus	6.3	5.3	5.6	7.5	1.9	2.6	3.3	71.9	83.8	2.0	-	-	0.12	-	-
												0.9	2.8		0.05	0.11
2000-2010	Fagus	6.2	5.0	6.7	7.7	3.8	2.4	1.2	38.3	86.7	-	1.2	-	-	0.06	-
											0.7		3.4	0.03		0.20
1950-1960	Quercus	6.4	5.4	5.7	8.9	2.0	2.5	3.3	87.7	119.3	1.7	-	-	0.09	-	-
												1.2	3.0		0.05	0.11
2000-2010	Quercus	6.3	5.1	6.8	8.7	3.7	2.4	1.2	51.0	113.3	-	0.8	-	-	0.04	-
											0.4		4.1	0.01		0.24

Table S3: Climate and phenology statistics for two species ($Betula\ pendula$, $Fagus\ sylvatica$, across 17 and 24 sites respectively) from the PEP725 data across all sites with continuous data from 1950-2010. ST is spring temperature from 1 March to 30 April, ST.leafout is temperature 30 days before leafout, and GDD is growing degree days 30 days before leafout. Slope represents the estimated sensitivity using untransformed leafout and ST, while log-slope represents the estimated sensitivity using log(leafout) and log(ST). We calculated all metrics for for each species x site x 20 year period before taking mean or variance estimates. See also Fig. S7.

		mean (ST)		var (ST)						slope		log-slope				
years	species	30	45	60	mean	30	45	60	var	GDD	30	45	60	30	45	60
					ST.lo				(lo)							
1950-1970	Betula	6.4	4.9	5.8	7.1	3.7	2.7	2.6	79.9	72.5	1.1	1	ı	0.08	-	-
												1.0	4.3		0.03	0.19
1970-1990	Betula	6.4	5.4	5.9	7.2	2.2	2.9	1.3	104.8	72.2	-	-	-	-	-	-
											0.0	2.0	6.1	0.02	0.07	0.33
1990-2010	Betula	5.8	5.3	6.8	6.7	2.1	2.7	0.9	36.2	60.0	-	0.0	-	-	0.00	-
											1.2		3.3	0.07		0.21
1950-1970	Fagus	6.1	4.7	5.6	7.6	3.8	2.8	2.7	63.4	86.0	1.0	-	-	0.05	0.00	-
												0.2	3.1			0.12
1970-1990	Fagus	6.2	5.2	5.6	7.5	2.3	3.0	1.3	56.2	81.3	-	-	-	-	-	-
											0.2	1.3	2.5	0.01	0.04	0.16
1990-2010	Fagus	5.5	5.2	6.7	7.5	2.2	2.8	1.0	32.8	79.9	-	0.1	-	-	0.01	-
											0.6		2.8	0.03		0.15
1950-1970	Fagus	6.2	4.9	5.7	9.2	3.7	2.7	2.6	78.0	127.8	1.3	-	-	0.07	-	-
												0.7	3.7		0.02	0.14
1970-1990	Fagus	6.3	5.4	5.8	8.6	2.3	2.9	1.2	57.6	109.1	0.2	-	-	0.00	-	-
												1.2	3.1		0.04	0.15
1990-2010	Fagus	5.7	5.3	6.8	8.6	2.1	2.7	1.0	70.6	112.3	-	-	-	-	0.00	-
											1.0	0.0	3.5	0.06		0.19

5 Figures

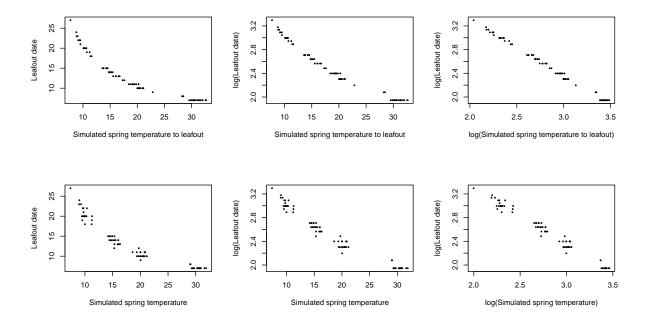


Figure S1: Simulated leafout as a function of temperature across different temperatures highlights non-linearity of process. Here we simulated sets of data where leafout constantly occurs at 200 growing degree days (thermal sum of mean daily temperatures with 0°C as base temperature) across mean temperatures of 10, 15, 20 and 30°C (constant SD of 4), we calculated estimated mean temperature until leafout date (top row) or across a fixed window (bottom row, similar to estimates of 'spring temperature'). While within any small temperature range the relationship may appear linear, its non-linear relationship becomes clear across the greater range shown here (left). Taking the log of both leafout and temperature (right) linearizes the relationship.

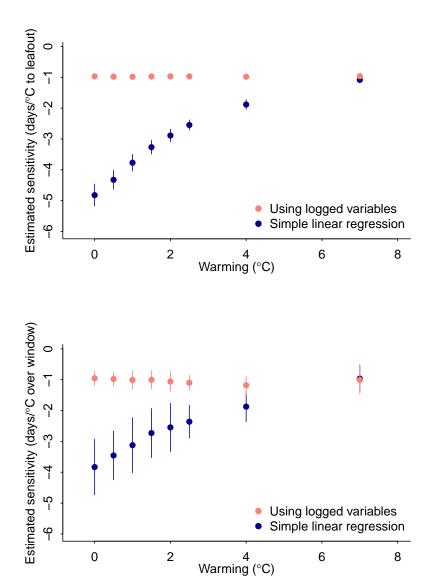


Figure S2: A simple model generates declining sensitivities with warming. We show declines in estimated sensitivities with warming from simulations (top: using average temperature until leafout, bottom: using a fixed window) with no underlying change in the biological process when sensitivities were estimated with simple linear regression ("Simple linear regression"). This decline disappears using regression on logged predictor and response variables ("Using logged variables").

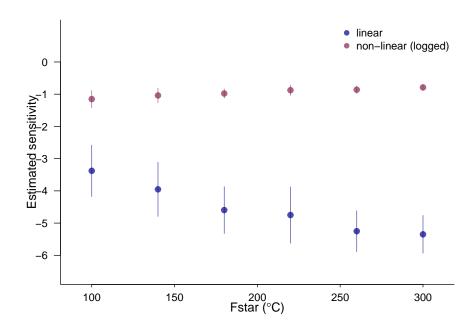
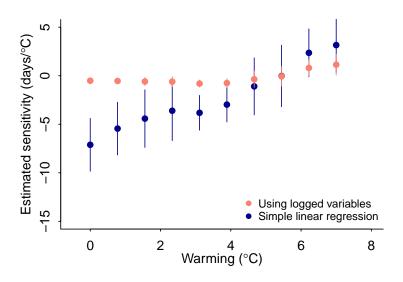


Figure S3: Simulated leafout as a function of required thermal sum for leafout. Here we simulated sets of data where leafout occurs at varying thermal sums (sum of mean daily temperatures with 0° C as base temperature).



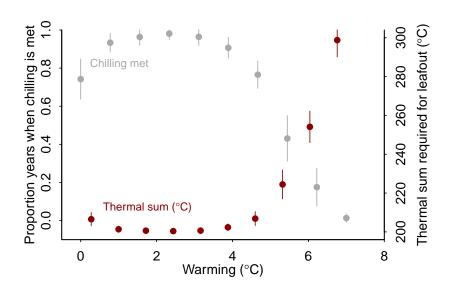


Figure S4: Simulated leafout as a function of temperature across different levels of warming with shifts in underlying biology through lower chilling. Here we simulated sets of data where leafout occurs at a thermal sum of 200 (sum of mean daily temperatures with 0°C as base temperature) when chilling is met, and requires a higher thermal sum when chilling is not met. We show estimated sensitivities in the top panel, and the shifting cues in the bottom panel.

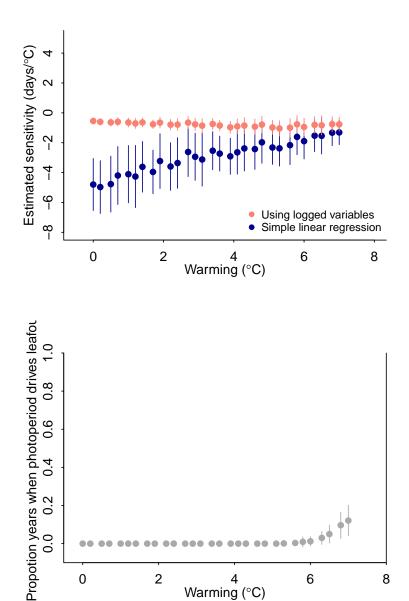


Figure S5: Simulated leafout as a function of temperature across different levels of warming with shifts in underlying biology through a daylength threshold. Here we simulated sets of data where leafout occurs at a thermal sum of 200 (sum of mean daily temperatures with 0° C as base temperature) as long as the daylength of that day is ≥ 12 hours; otherwise leafout occurs on the day when daylength is 12 hours. We show estimated sensitivities in the top panel, and the shifting cues in the bottom panel.

4 Warming (°C)

6

8

2

0

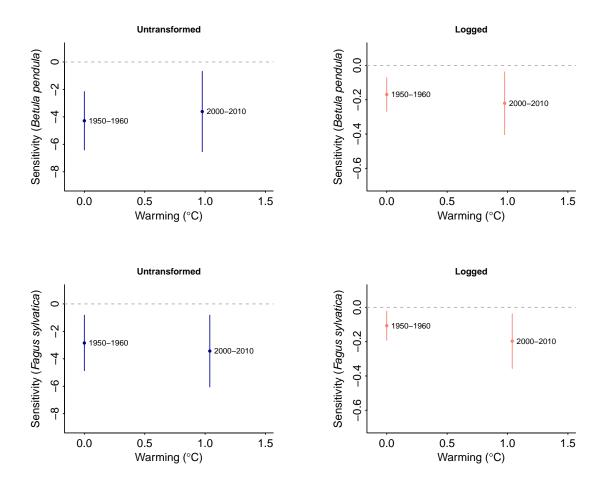


Figure S6: Sensitivities from PEP725 data using 10 year windows of data for two species (top - Betula pendula, bottom - Fagus sylvatica; all lines show 78% confidence intervals from linear regressions). Amounts of warming are calculated relative to 1950-1960 and we used only sites with leafout data in all years shown here. See Table S2 for further details.

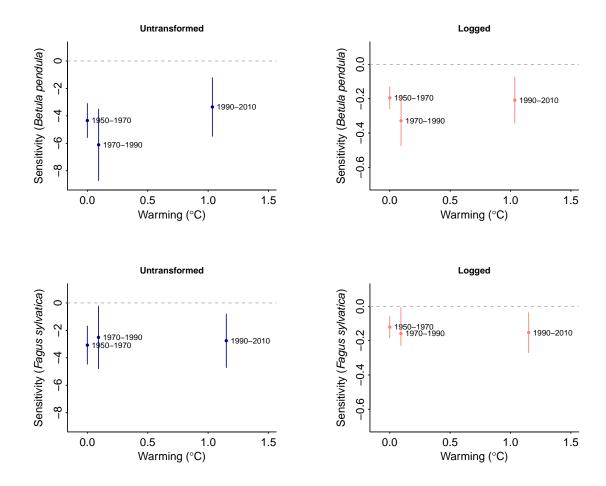


Figure S7: Sensitivities from PEP725 data using 20 year windows of data for two species (top - Betula pendula, bottom - Fagus sylvatica; all lines show 78% confidence intervals from linear regressions). Amounts of warming are calculated relative to 1950-1970 and we used only sites with leafout data in all years shown here. See Table S3 for further details.