a note on beyesian linear regression

general form:

$$y_i = eta^T x_i + \epsilon_i$$

now assume y_i comes from $N(eta^T x_i, \sigma^2)$

compute the likelihood given assumption that ϵ_i are independent:

$$egin{aligned} P(y|X,eta,\sigma^2) &= \prod_{i=1}^n P(y_i|x_i,eta,\sigma^2) \ &= \prod_{i=1}^n rac{1}{\sqrt{2\pi}\sigma} exp\left\{rac{1}{2\sigma^2}(y_i-eta^Tx_i)^2
ight\} \ &= (2\pi\sigma^2)^{-n/2} exp\left\{-rac{1}{2\sigma^2}\sum_{i=1}^n (y_i-eta^Tx_i)^2
ight\} \end{aligned}$$

note that,

$$egin{aligned} \sum_{i=1}^n (y_i - eta^T x_i)^2 &= (y - Xeta)^T (y - Xeta) \ &= y^T y - 2eta^T X^T y + eta^T X^T Xeta \end{aligned}$$

then we will compute the conditional likelihood $P(\beta|y,x,\sigma^2)$. Suppose we have known the value of σ^2 , and prior distribution of β comes from $N_k(\beta_0,\Sigma_0)$. Then $P(\beta)=$

$$((2\pi)^k |\Sigma_0|)^{-1/2} exp\left\{-rac{1}{2}(eta-eta_0)^T \Sigma_0^{-1} (eta-eta_0)
ight\}.$$

According to bayes throey,

$$\begin{split} &P(\beta|y,X,\sigma^{2})\\ &\propto P(\beta)P(y|X,\beta,\sigma^{2})\\ &= ((2\pi)^{k}|\Sigma_{0}|)^{-1/2}exp\left\{-\frac{1}{2}(\beta-\beta_{0})^{T}\Sigma_{0}^{-1}(\beta-\beta_{0})\right\}\\ &\quad *(2\pi\sigma^{2})^{-n/2}exp\left\{-\frac{1}{2\sigma^{2}}\left(y^{T}y-2\beta^{T}X^{T}y+\beta^{T}X^{T}X\beta\right)\right\}\\ &\propto exp\left\{-\frac{1}{2}\left(\beta^{T}\Sigma_{0}^{-1}\beta_{0}-\beta^{T}\Sigma_{0}^{-1}\beta_{0}-\beta_{0}^{T}\Sigma_{0}^{-1}\beta+\beta_{0}^{T}\Sigma_{0}^{-1}\beta_{0}\right)\right\}\\ &\quad *exp\left\{-\frac{1}{2\sigma^{2}}\left(-2\beta^{T}X^{T}y+\beta^{T}X^{T}X\beta\right)\right\}\\ &\propto exp\left\{\beta^{T}\Sigma_{0}^{-1}\beta_{0}-\frac{1}{2}\beta^{T}\Sigma_{0}^{-1}\beta+\beta^{T}X^{T}y/\sigma^{2}-\frac{1}{2}\beta^{T}X^{T}X\beta/\sigma^{2}\right\}\\ &= exp\left\{\beta^{T}\left(\Sigma_{0}^{-1}\beta_{0}+X^{T}y/\sigma^{2}\right)-\frac{1}{2}\beta^{T}\left(\Sigma_{0}^{-1}+X^{T}X/\sigma^{2}\right)\beta\right\} \end{split}$$

this is proportional to a multivariate normal distribution(to be proved later), that means:

$$eta|y,X,\sigma^2 \sim N(A_{\sigma^2}^{-1}B_{\sigma^2},A_{\sigma^2}^{-1})$$

where
$$A_{\sigma^2}=\Sigma_0^{-1}+X^TX/\sigma^2$$
 , $B_{\sigma^2}=\Sigma_0^{-1}eta_0+X^Ty/\sigma^2$

Do something similar, now we have known the value of β . Let $\gamma=\frac{1}{\sigma^2}$ and prior distribution of γ comes from $gamma\left(v_0/2,v_0\sigma^2/2\right)$. Recall that the pdf for a gamma(x|a,b) is

$$f(x|a,b) = rac{b^a x^{a-1} e^{-bx}}{\Gamma(a)}$$

so the prior distribution of γ is

$$P(\gamma) \propto (v_0 \sigma_0^2/2)^{v_0/2} \gamma^{rac{v_0}{2}-1} exp\left\{-\gamma v_0 \sigma^2/2
ight\}$$

the likelihood is, as mentioned before,

$$egin{aligned} P(y|X,eta,\gamma=1/\sigma^2) &= (2\pi\sigma^2)^{-n/2}exp\left\{-rac{1}{2\sigma^2}\left(y^Ty-2eta^TX^Ty+eta^TX^TXeta
ight)
ight\} \ &\propto \gamma^{n/2}exp\left\{-rac{\gamma}{2}\left(y^Ty-2eta^TX^Ty+eta^TX^TXeta
ight)
ight\} \end{aligned}$$

Again using bayesian theory,

$$egin{aligned} P(\gamma|y,X,eta) &\propto P(\gamma)P(y|X,eta,\gamma) \ &\propto \gamma^{n/2}exp\left\{-rac{\gamma}{2}\left(y^Ty-2eta^TX^Ty+eta^TX^TXeta
ight)
ight\}*\gamma^{rac{v_0}{2}-1}exp\left\{-\gamma v_0\sigma^2/2
ight\} \ &= \gamma^{rac{1}{2}(v_0+n)-1}exp\left\{-rac{\gamma}{2}\left(v_0\sigma_0^2+y^Ty-2eta^TX^Ty+eta^TX^TXeta
ight)
ight\} \end{aligned}$$

that is

$$\gamma|y,X,eta\sim gamma\left(rac{v_0+n}{2},rac{1}{2}\left(v_0\sigma_0^2+y^Ty-2eta^TX^Ty+eta^TX^TXeta
ight)
ight)$$

which is equivalent to

$$\sigma^2|y,X,eta \sim inverse-gamma\left(rac{v_0+n}{2},rac{1}{2}\left(v_0\sigma_0^2+SSR(eta)
ight)
ight)$$

where
$$SSR(\beta) = y^Ty - 2\beta^TX^Ty + \beta^TX^TX\beta$$

after we have the posterior distribution, we do the Gibbs sampling.

STEP 1: sample γ_0 fron $P(\gamma)$, $\sigma_0^2=1/\gamma_0$

STEP 2: sample eta_t from $N(A_{\sigma_t^2}^{-1}B_{\sigma_t^2},A_{\sigma_t^2}^{-1})$

STEP 3: sample σ_{t+1}^2 from $inverse-gamma(rac{v_0+n}{2},rac{1}{2}(v_0\sigma_0^2++SSR(eta_t)))$, t=t+1

STEP 4: return to STEP 2, t=0 to 10000