Principles of Machine Learning CSCI-B455

Bayesian Decision Theory

- Random experiment, uncertainty
- Sample space S, S is either discrete or continuous
- Event E is a subset of S, which appears with probability P(E)
- P(E): Repeat an experiment many times, count how many times outcome is E.
- $0 \le P(E) \le 1$, P(S) = 1
- E_i and E_j are mutually exclusive if $E_i \cap E_j = \emptyset$
- For exclusive events $P(\bigcup_{i=1}^n E_i) = \sum_{i=1}^n P(E_i)$
- If $E \cap F \neq \emptyset$, then $P(E \cup F) = P(E) + P(F) P(F \cap U)$
- E^c is the complement of event E and $P(E) + P(E^c) = 1$

- Conditional Probability $P(E \mid F)$: Probability of E given that F has occurred, $P(E \mid F) = \frac{P(E \cap F)}{P(F)}$
- $P(F | E) = \frac{P(F \cap E)}{P(E)}$, $P(E \cap F) = P(F \cap E) \Rightarrow P(E | F) \cdot P(F) = P(F | E) \cdot P(E)$
- Bayes' formula $P(E \mid F) = \frac{P(E \cap F)}{P(F)}$
- Mutually exclusive and exhaustive events: $\bigcup_{i=1}^k F_i = S$. Then $P(E) = \sum_{i=1}^n P(E \cap F_i) = \sum_{i=1}^n P(E \mid F_i) \cdot P(F_i)$

•
$$P(F_i|E) = \frac{P(F_i \cap E)}{P(E)} = \frac{P(E|F_i) \cdot P(F_i)}{\sum_{i=1}^n P(E|F_i) \cdot P(F_i)}$$

Bayes' example

• There is a disease that appears with probability one in a million. There is a test that can detect the disease with probability 99% on a person with the disease. However, with 1/1000 probability, the test reports positive on a healthy person. What is the probability that a patient is sick when the test result is positive?

$$P(d = 1) = 1/10^6$$
, $P(d = 0) = 1 - 1/10^6$, $P(d = 1 | t = 1) = ?$ $P(t = 1 | d = 1) = 0.99$, $P(t = 1 | d = 0) = 0.001$

$$P(d=1 | t=1) = \frac{P(t=1 | d=1) \cdot P(d=1)}{P(t=1)} = \frac{P(t=1 | d=1) \cdot P(d=1)}{P(t=1 | d=1) \cdot P(d=1) + P(t=1 | d=0) \cdot P(d=0)}$$

$$\approx 1/10000$$

- Random variable X takes on different values based on the outcome of a random event.
- Probability distribution function of $F(a) = P(X \le a)$, for a real number a

•
$$F(a) = \sum_{\forall x \le a} P(x)$$
, if X is discrete, or $F(a) = \int_{-\infty}^{a} P(x) \, dx$, if X is continuous

• Joint distribution
$$F(x, y) = P(X \le x, Y \le y)$$
, $P(X = x) = \sum_{j} P(x, y_j)$ or $P(X = x) = \int_{-\infty}^{\infty} P(x, y) \, dy$

• Conditional distribution $P(X = x \mid Y = y) = \frac{P(x, y)}{P(y)}$

• Expected value of a random variable X : $E[X] = \begin{cases} \sum_i x_i P(x_i) & \text{if } X \text{ is discrete} \\ \int x P(x) dx & \text{if } X \text{ is continuous} \end{cases}$

•
$$E[aX + b] = aE[X] + b$$
 ; $E[X + Y] = E[X] + E[Y]$

•
$$E[g(X)] = \begin{cases} \sum_i g(x_i) P(x_i) & \text{if } X \text{ is discrete} \\ \int g(x) P(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

• The
$$n$$
th moment of X : $E[X^n] = \begin{cases} \sum_i x_i^n P(x_i) & \text{if } X \text{ is discrete} \\ \int x^n P(x) dx & \text{if } X \text{ is continuous} \end{cases}$

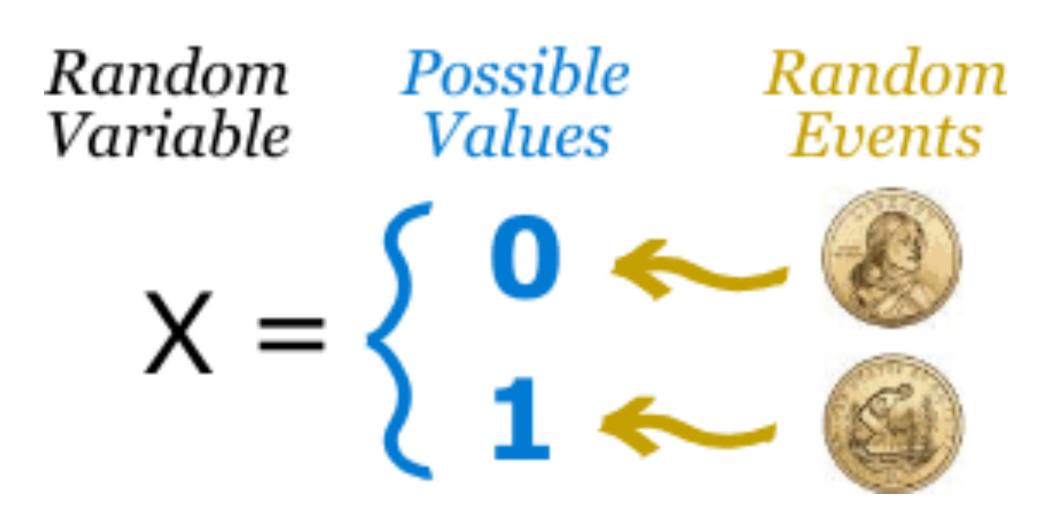
• First moment is $\mu = E[X]$

• Variance of a random variable X measures expected variation of X around $\mu=E[X]$

$$Var(X) = E[(X - \mu)^2] = E[X^2 - 2X\mu + \mu^2] = E[X^2] - 2\mu^2 + \mu^2 = E[X^2] - \mu^2 = \sigma^2$$

- Standard deviation $\sigma = \sqrt{Var(X)}$
- Covariance is the relationship between two random variables $Cov(X, Y) = E[(X \mu_X)(Y \mu_Y)]$
- Correlation is the normalized covariance $Cor(X,Y) = \frac{Cov(X,Y)}{\sqrt{Var(x)Var(Y)}}$, $-1 \leq Corr(X,Y) \leq 1$

Random Variable & Random Process



https://www.mathsisfun.com/data/random-variables.html

- TOSSING A COIN
- The outcome depends on some unobservable parameters (the physics of tossing it).
- It is a random event as the outcome, is uncertain.
- The outcome is indeed observable, and is a random variable, say X.
- X=0 (say heads) with some probability p_0 , and X=1(tails) with probability $p_1=1-p_0$.
- If p_0 (and hence p_1) is known, what would be a good strategy to predict the next?
- Choose the more probable one, to reduce the error, (1-selected.probability).

Random Variable & Random Process

• If more probable one is **not known**, but we have a sequence of previous outcomes, how do we proceed?

We need an estimator, which is easy
$$\hat{p}_0 = \frac{number\ of\ head\ tosses}{total\ number\ of\ tosses}$$
.

Assume, previous outcomes are H,T,T,T,H,H,T,T,T,H. Then $\hat{P}(Heads) = \hat{p}_0 = \frac{4}{10}$

Based on this, the prediction next will be tails.

Notice that, after the next toss, we should update our approximation!

- Revisiting the credit scoring, low-risk and high-risk customers
- The parameters we use X_1, X_2 as the yearly income and savings.
- When a new customer arrives, the bank wants to predict the credit score class.
- The class label of the customer, C, is a **Bernoulli** random variable

- ullet Bernoulli random variable X has two outcomes, Success (1) or Fail (0).
- Success probability shown by P(Success) = P(1) = p
- Failure probability is (1 p)
- E(X) = p and $Var(X) = \sigma^2 = (1 p)^2 \cdot p + (0 p)^2 \cdot (1 p) = p(1 p)$

Choose the class with higher probability for the given input (x_1, x_2)

$$C = 1$$
, if $P(C = 1 | x_1, x_2) > P(C = 0 | x_1, x_2)$, else $C = 0$

The error is
$$E = 1 - max(P(C = 1 | x_1, x_2), P(C = 0 | x_1, x_2))$$

Same story with the coin tossing, but decision depends on two variables.

How to compute $P(C | x_1, x_2)$ by using the training data?

- Prior Probability: P(C = 1), regardless of the input.
- Class likelihood: $P(\mathbf{x} \mid C)$, the probability of \mathbf{x} in class C
- Evidence: $P(\mathbf{x})$, the probability of observing \mathbf{x} regardless of the class.
- Posterior Probability: $P(C | \mathbf{x})$, given input, predict its class, the aim.

$$posterior = \frac{prior \cdot likelihood}{evidence} = \frac{P(C) \cdot P(\mathbf{x} \mid C)}{P(\mathbf{x})} = P(C \mid \mathbf{x})$$

$$P(C = 0 | \mathbf{x}) + P(C = 1 | \mathbf{x}) = 1$$

- Prior Probability: P(C=1), regardless of the input.
- Class likelihood: $P(\mathbf{x} \mid C)$, the probability of \mathbf{x} in class C
- Evidence: $P(\mathbf{x})$, the probable, lity of observing \mathbf{x} regardless of the class.
- Posterior Probability: $P(C | \mathbf{x})$, given input, predict its class, the aim.

$$posterior = \frac{prior \cdot likelihood}{evidence} = \frac{P(C) \cdot P(\mathbf{x} \mid C)}{P(\mathbf{x})} = P(C \mid \mathbf{x})$$

• In case of multiple classes C_1, C_2, \ldots, C_K :

$$P(C_i | \mathbf{x}) = \frac{P(C_i) \cdot P(\mathbf{x} | C_i)}{\sum_{k=1}^{K} P(\mathbf{x} | C_k)}$$

- Wrong decisions (misclassifications) might not be equally costly.
- Compare the consequences of wrong decision on low-risk and high-risk customers.
- We need an adjustment while calculating the error.
- $lpha_i$: classifying input in class C_i
- $\lambda_{i,j}$: incurred loss when input is classified as C_i , while it is C_j .

$$R(\alpha_i | \mathbf{x}) = \sum_{j=1}^K \lambda_{i,j} \cdot P(C_j | \mathbf{x})$$

Choose
$$C_i$$
 when $R(\alpha_i | \mathbf{x}) = \min_k R(\alpha_k | \mathbf{x})$

An example case: Given $\lambda_{1,1}=0,\,\lambda_{2,2}=0,\,\lambda_{1,2}=10,\,\lambda_{2,1}=5$, what would be the optimal decision rule?

Risk when we decide on C_1 :

$$R(\alpha_1 | x) = \lambda_{1,1} \cdot P(C_1 | x) + \lambda_{1,2} \cdot P(C_2 | x) = 0 \cdot P(C_1 | x) + 10 \cdot (1 - P(C_1 | x))$$

Risk when we decide on C_2 :

$$R(\alpha_2 \mid x) = \lambda_{2,1} \cdot P(C_1 \mid x) + \lambda_{2,2} \cdot P(C_2 \mid x) = 5 \cdot P(C_1 \mid x) + 0 \cdot P(C_2 \mid x) = 5P(C_1 \mid x)$$

Choose C_1 when $R(\alpha_1 | x) < R(\alpha_2 | x)$, which means

$$10 \cdot (1 - P(C_1 \mid x)) < 5 \cdot P(C_1 \mid x) \Rightarrow 2 - 2P(C_1 \mid x) < P(C_1 \mid x) \Rightarrow \left| \frac{2}{3} < P(C_1 \mid x) \right|$$

How do we define the case with equal costs of misclassification?

• Assume $\lambda_{i,j}=0$ if i=j, and $\lambda_{i,j}=1$ for all $i\neq j$. Then:

$$R(\alpha_i | \mathbf{x}) = \sum_{k=1}^K \lambda_{i,k} \cdot P(C_k | \mathbf{x}) = \sum_{k \neq i} P(C_k | \mathbf{x}) = 1 - P(C_i | \mathbf{x})$$

Therefore, choosing the maximum posterior probability, guarantees the minimum risk, when misclassification costs are equal.

Notice that this is actually very rare in practice!

How do we define the case with equal costs of misclassification?

If $\lambda_{1,1}=0,\,\lambda_{2,2}=0,\,\lambda_{1,2}=\lambda_{2,1}=w$, what would be the optimal decision rule?

Risk when we decide on C_1 :

$$R(\alpha_1 \mid x) = \lambda_{1,1} \cdot P(C_1 \mid x) + \lambda_{1,2} \cdot P(C_2 \mid x) = 0 \cdot P(C_1 \mid x) + w \cdot (1 - P(C_1 \mid x))$$

Risk when we decide on C_2 :

$$R(\alpha_2 \mid x) = \lambda_{2,1} \cdot P(C_1 \mid x) + \lambda_{2,2} \cdot P(C_2 \mid x) = w \cdot P(C_1 \mid x) + 0 \cdot P(C_2 \mid x) = wP(C_1 \mid x)$$

Choose C_1 when $R(\alpha_1 | x) < R(\alpha_2 | x)$, which means

$$w \cdot (1 - P(C_1 | x)) < w \cdot P(C_1 | x) \Rightarrow 1 - P(C_1 | x) < P(C_1 | x) \Rightarrow \frac{1}{2} < P(C_1 | x)$$

What if the cost of misclassification is extremely high?

- When the computed risk is greater than a threshold, more complex systems will handle the input, even maybe manual.
- We introduce an additional decision α_{K+1} as **REJECT**.

The loss function
$$\lambda_{i,k}$$
 is then
$$\lambda_{i,k} = \begin{cases} 0 & \text{if } i = k \\ \lambda & \text{if } i = K+1 \\ 1 & \text{otherwise} \end{cases}$$

$$R(\alpha_{K+1} | \mathbf{x}) = \sum_{k=1}^{K} \lambda \cdot P(C_k | \mathbf{x}) = \lambda$$

$$R(\alpha_i | \mathbf{x}) = \sum_{k \neq i} \lambda_{i,k} \cdot P(C_k | \mathbf{x}) = 1 - P(C_i | \mathbf{x})$$

$$1 - P(C_i | \mathbf{x}) < \lambda \rightarrow P(C_i | \mathbf{x}) > 1 - \lambda$$

Choose C_i if $R(\alpha_i|\mathbf{x}) < R(\alpha_i|\mathbf{x})$, for i=1,2,...,K,K+1, which means $P(C_i|\mathbf{x}) > 1-\lambda$

What if the cost of misclassification is extremely high?

If $\lambda_{1,1}=0, \, \lambda_{2,2}=0, \, \lambda_{1,2}=10, \, \lambda_{2,1}=5$ and $\lambda_{r,1}=\lambda_{r,2}=1$, what would be the optimal decision rule?

Risk when we decide on
$$C_1$$
: $R(\alpha_1 | x) = \lambda_{1,1} \cdot P(C_1 | x) + \lambda_{1,2} \cdot P(C_2 | x) = 0 \cdot P(C_1 | x) + 10 \cdot (1 - P(C_1 | x))$

Risk when we decide on
$$C_2$$
: $R(\alpha_2 | x) = \lambda_{2,1} \cdot P(C_1 | x) + \lambda_{2,2} \cdot P(C_2 | x) = 5 \cdot P(C_1 | x) + 0 \cdot P(C_2 | x) = 5P(C_1 | x)$

Risk when we decide on reject: $R(\alpha_r|x) = \lambda_{r,1} \cdot P(C_1|x) + \lambda_{r,2} \cdot P(C_2|x) = 1 \cdot P(C_1|x) + 1 \cdot (1 - P(C_1|x)) = 1$

To choose
$$C_1$$
, $R(\alpha_1 | x) < R(\alpha_r | x)$ $10 - 10P(C_1 | x) < 1 \Rightarrow P(C_1 | x) > \frac{9}{10}$

To choose
$$C_2$$
, $R(\alpha_2 | x) < R(\alpha_r | x)$ 5 · $P(C_1 | x) < 1 \Rightarrow P(C_1 | x) < \frac{1}{5}$

$$\frac{1}{5}$$
 $\frac{9}{10}$ C_2 REJECT C_1

Discriminant Functions

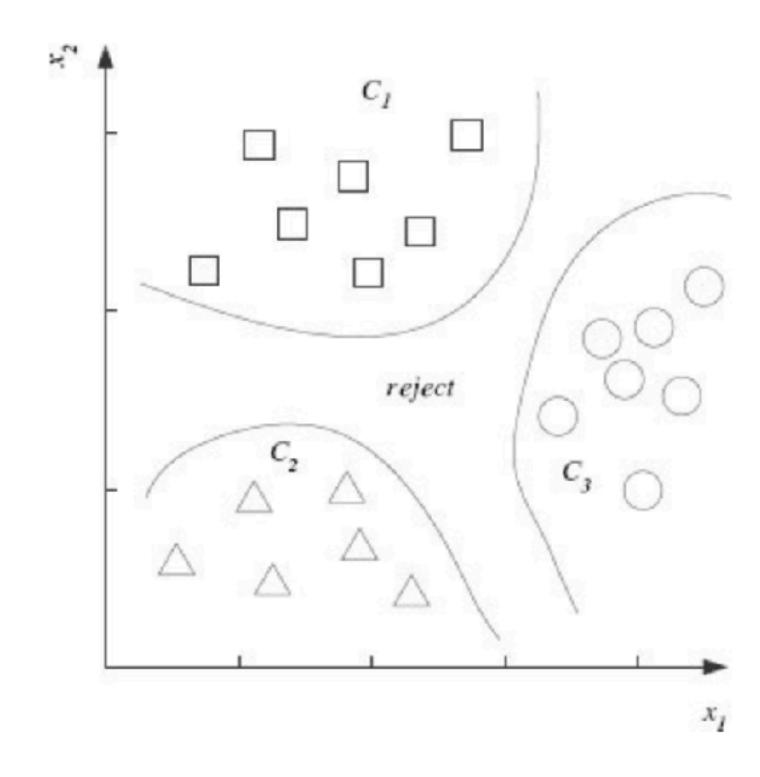
Discriminant functions can be used as classifiers: Choose

class
$$C_i$$
, if $g_i(\mathbf{x}) = \max_k g_k(\mathbf{x})$, for $g_i(\mathbf{x}) = -R(\alpha_i | \mathbf{x})$

Assuming 0/1 *loss* function, $g_i(\mathbf{x}) = -R(\alpha_i | \mathbf{x}) = P(C_i | \mathbf{x})$

Even by neglecting the common denominator $p(\mathbf{x})$,

$$g_i(\mathbf{x}) = P(C_i | \mathbf{x}) = P(C_i)P(\mathbf{x} | C_i)$$



Discriminant Functions

An example: Let the likelihood ratio be
$$\mathscr{E} = \frac{P(x \mid C_1)}{P(x \mid C_2)}$$
.

If we define discriminant function as
$$g(x) = \frac{P(C_1|x)}{P(C_2|x)}$$

$$g(x) = \frac{P(C_1|x)}{P(C_2|x)} = \frac{P(x|C_1) \cdot P(C_1)/P(x)}{P(x|C_2) \cdot P(C_2)} = \frac{P(x|C_1) \cdot P(C_1)/P(x)}{P(x|C_2) \cdot P(C_2)} = \ell \cdot \frac{P(C_1)}{P(C_2)}.$$

Notice that if $P(C_1) = P(C_2)$, likelihood ratio becomes directly the discriminant.

Reading Material

• Chapter 3, excluding 3.5 Association Rules