Dissertation Proposal

A study about ordinal patterns transition graphs

Eduarda Tatiane Caetano Chagas Advisor: Heitor S. Ramos

Co-advisor: Alejandro C. Frery

{eduarda.chagas, ramosh}@dcc.ufmg.br, acfrery@laccan.ufal.br

1 Introduction

The main hypothesis underlying this dissertation project can be briefly stated as follows:

Ordinal pattern transition graphs capture intrinsic characteristics of the time series underlying phenomena and capture information about system dynamics in finer granularity.

Therefore, we plan to evaluate the power of characterization and consequently of image classification when we analyze them as ordinal patterns transition graphs.

The following sections motivate, support, and detail this dissertation project. This work was divided as follows: Section 2 exposes the motivations of this study. Section 3 presents the proposed methodology. Section 3.1 reports how the patch linearization process of the images occurs. Section 3.2 describes our technique of ordinal amplitude transition graph weighting by amplitudes. In Section 3.3 we report the Information Theory descriptors used throughout this work. Section 5 shows the expected results with the dissertation proposal presented. Section 6 we present the schedule to be followed. We present our conclusions in Section 7.

2 Motivation

Surface classification and land use are among the most important applications of Synthetic Aperture Radar (SAR) imaging [10], for which supervised and unsupervised classification algorithms have been proposed [13, 3, 16].

Classification techniques rely on the extraction and analysis of features from the data, and from additional information and prior knowledge about the scene, the sensor and the acquisition conditions. Texture is among the features that carries most information and, as such, it is important to characterize it in a quantitative manner.

The texture in SAR images is characterized twofold, namely by the marginal properties of the data [11], and by their spatial structure [14]. In this work we focus on the second approach.

The most widely used approach to obtain textural features from SAR imagery is through co-ocurrence matrices and Haralick's descriptors [17]. Other approaches include the Fourier power spectrum [6], and Markov random fields [4].

In our approach, we chose to linearize the image patches and analyze the resulting 1-D signals as time series. By doing that, we will reduce the dimensionality of the data and map spatial dependence of pixels to the temporal correlation. Then, starting from the ordinal patterns from the time series, we propose a novel weighted graph of pattern transitions that is able to discriminate patterns corresponding to different amplitude levels. Finally, we will use well-known features from information theory to characterize image patches.

This dissertation is, thus, situated in this context, and proposes and analyzes a novel technique for SAR image texture characterization that is based on a modified version of ordinal pattern transition graphs, which is able to discriminate similar patterns with different amplitude levels.

3 Methodology

Figure 1 presents a schematic view of this dissertation proposal. In the following, we present a more detailed discussion of each element.

- Transforming a 2-D patch of data into a time series using a Hilbert Space Filling Curve,
- 2. Building an Ordinal Pattern Transition Graph with weighted edges;
- 3. Obtaining a probability distribution function from this graph;
- 4. Computing the Entropy and Statistical Complexity of this distribution.

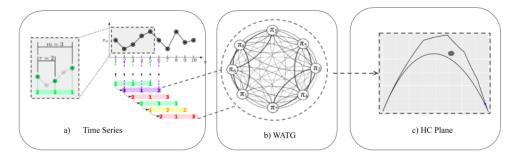


Figure 1. Outline of the technique for the characterization of textures.

3.1 Linearization of image patches

In this step, we perform a data dimensionality reduction by turning 2-D patches into 1-D signals. For this, we opt to do a linearization of the data through techniques that maintain the spatial correlation of the image. This could be accomplished by reading the data by lines, columns or any transformation. In Step 1 we employ a Hilbert Space Filling Curve [9].

Space filling curves were first employed by Nguyen and Quinqueton (1982), to map a texture into a one-dimensional signal. When used as scanning methods of an image, such functions preserve relevant properties of pixel spatial correlation [9].

Assuming an image patch is supported by a $N \times N$ dimension grid, where N is a power of 2, we have the following definition.

Definition 1. An image scan is a bijective function $f: \mathbb{N} \times \mathbb{N} \to \mathbb{N}$ in the ordered pair set $\{(i,j): 1 \le i, j \le N\}$, which denotes the points in the domain, for the closed range of integers $\{1, \ldots, N^2\}$. A scan rule is $\{f^{-1}(1), \ldots, f^{-1}(N^2)\}$.

This Definition imposes that each pixel is visited only once.

Space filling curves, such as raster-1, raster-2 and Hilbert scanning techniques stipulate a proper function f. Hilbert curves scans an array of pixels of dimension $2^k \times 2^k$, $k \in \mathbb{N}$, never keeping the same direction for more than three consecutive points, as shown in Figure 2. Using the Hilbert curve, we can reduce the data dimensionality by maintaining the spatial dependence information of the analyzed textures.

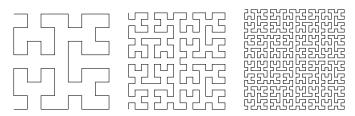


Figure 2. Hilbert space filling curve in areas of: (a) 8×8 , (b) 16×16 and (c) 32×32 .

3.2 Weighted Ordinal Patterns Transition Graph

Step 2 consists of two stages. In the first, the time series is transformed into a sequence of ordinal patterns. In the second, we build a weighted graph describing the transitions between these patterns.

The representation of time series by ordinal patterns was introduced by [1] as a transformation resistant to noise, and invariant to nonlinear monotonic transformations.

Consider $\mathcal{X} \equiv \{x_t\}_{t=1}^T$ a real valued time series of length T.

Let \mathfrak{A}_D (with $D \geq 2$ and $D \in \mathbb{N}$) be the symmetric group of order D! formed by all possible permutation of order D, and the symbol component vector $\boldsymbol{\pi}^{(D)} = (\pi_1, \pi_2, \dots, \pi_D)$ so every element $\boldsymbol{\pi}^{(D)}$ is unique $(\pi_j \neq \pi_k$ for every $j \neq k$). Consider for the time series $\mathcal{X} \equiv \{x_t\}_{t=1}^T$ its

time delay embedding representation, with embedding dimension $D \ge 2$ and time delay $\tau \ge 1$ ($\tau \in \mathbb{N}$, also called "embedding time", "time delay", or "delay"):

$$\mathbf{X}_{t}^{(D,\tau)} = (x_{t}, x_{t+\tau}, \dots, x_{t+(D-1)\tau}), \tag{1}$$

for $t=1,2,\ldots,N$ with $N=T-(D-1)\tau$. Then the vector $\mathbf{X}_t^{(D,\tau)}$ can be mapped to a symbol vector $\boldsymbol{\pi}_t^D \in \mathfrak{A}_D$. This mapping should be defined in a way that preserves the desired relation between the elements $x_t \in \mathbf{X}_t^{(D,\tau)}$, and all $t \in \{1,\ldots,T-(D-1)\tau\}$ that share this pattern (also called "motif") have to mapped to the same $\boldsymbol{\pi}_t^D$.

We define the mapping $\mathbf{X}_t^{(D,\tau)} \mapsto \mathbf{g}_t^D$ by ordering the observations $x_t \in \mathbf{X}_t^{(D,\tau)}$ in increasing order. Consider the time series $\mathcal{X}=(1.8,1.2,3.2,4.8,4.2,4.5,2.3,3.7,1.2,.5)$ depicted in Fig. 3. Assume we are using patterns of length D=5 with unitary time lag $\tau=1$. The code associated to $\mathbf{X}_3^{(5,1)}=(x_3,\ldots,x_7)=(3.2,4.8,4.2,4.5,2.3)$, shown in black, is formed by the indexes in $\pi_3^5=(1,2,3,4,5)$ which sort the elements of $\mathbf{X}_3^{(5,1)}$ in increasing order: 51342. With this, $\widetilde{\pi}_3^5=51342$, and we increase the counting related to this motif in the histogram of all possible patterns of size D=5.

The dash-dot line in Fig. 3 illustrates $\mathbf{X}_1^{(5,2)}$, i.e. the sequence of length D=5 starting at x_1 with lag $\tau=2$. In this case, $\mathbf{X}_1^{(5,2)}=(1.8,3.2,4.2,2.3,1.2)$, and the corresponding motif is $\widetilde{\pi}_1^5=51423$.

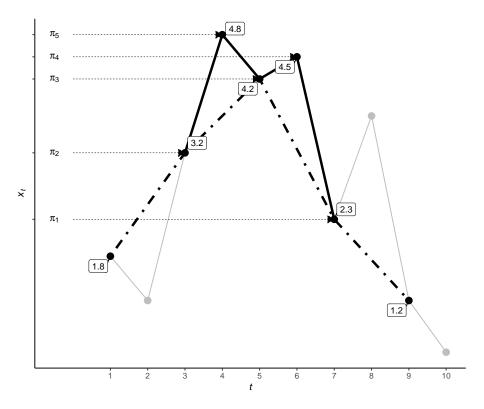


Figure 3. Illustration of the Bandt and Pompe coding

The classical approach consists in analyzing the histogram of these patterns. Alternatively,

one may form an oriented graph with the transitions from $\tilde{\pi}_t^D$ to $\tilde{\pi}_{t+1}^D$. We modify this last approach by assigning weights to the edges related to the absolute difference of the observations. This modification takes into account the scattering properties of the target, and leads to a good characterization of several types of textures.

Denote Π the sequence of symbols obtained by a given series $\mathbf{X}_t^{(D,\tau)}$. The Bandt-Pompe probability distribution is the relative frequency of symbols in the series against D! possible permutations of patterns $\{\widetilde{\pi}_t^D\}_{t=1}^{D!}$:

$$p(\widetilde{\pi}_t^D) = \frac{\#\left\{t : t = 1, \dots, T - (D-1)\tau; \mathbf{X}_t^{(D,\tau)} \text{ is of type } \widetilde{\pi}_t^D\right\}}{T - (D-1)\tau},$$
(2)

that meets the conditions $p(\widetilde{\pi}_t^D) \geq 0$ and $\sum_{i=1}^{D!} p(\widetilde{\pi}_t^D) = 1$.

The Ordinal Pattern Transition Graph G=(V,E) represents the transitions between two consecutive ordinal patterns over time t. The vertices are the patterns, and the edges the transitions between them: $V=\{v_{\widetilde{\pi}_t^D}\}$, and $E=\{(v_{\widetilde{\pi}_t^D},v_{\widetilde{\pi}_{t+1}^D}):v_{\widetilde{\pi}_{t+1}^D},v_{\widetilde{\pi}_{t+1}^D}\in V\}$ [2].

Recent works propose a weighting in the calculation of relative frequencies for ordinal patterns with different amplitude variances, making them contribute differently to the final value of permutation entropy (PE) and thus incorporating amplitude change information within a given set [5]. However, these methods do not consider the amplitude difference present in different time series, weighing them similarly when calculating the final value of their probabilities. Therefore, data with different amplitudes but with similar variance dynamics are not discriminated, losing important information about the system dynamics.

To counterbalance these facts, we propose a modification of the current ordinal pattern transition graph by incorporating meaningful time series information.

Two approaches are considered in relation to the weight of edges in the literature. Some authors employ unweighted edge [12, 7] representing only the existence of such transitions, while others apply the frequency of transitions [15, 18]. The weights $\mathbb{W} = \{w_{v_{\widetilde{\pi}_i^D}, v_{\widetilde{\pi}_j^D}} : v_{\widetilde{\pi}_i^D}, v_{\widetilde{\pi}_j^D} \in V\}$ assigned to each edge describes the chances of transition between two particular patterns $(v_{\widetilde{\pi}_i^D}, v_{\widetilde{\pi}_j^D})$ calculated by their respective relative frequencies, ie:

$$w_{v_{\tilde{\pi}_{i}^{D}, v_{\tilde{\pi}_{j}^{D}}}} = \frac{|\Pi_{\tilde{\pi}_{i}^{D}, \tilde{\pi}_{j}^{D}}|}{T - (D - 1)\tau - 1'}$$
(3)

where $|\Pi_{\widetilde{\pi}_i^D,\widetilde{\pi}_j^D}|$ is the number of transitions from pattern $\widetilde{\pi}_i^D$ to pattern $\widetilde{\pi}_j^D$ and $\sum_{v_{\widetilde{\pi}_i^D},v_{\widetilde{\pi}_j^D}} w_{v_{\widetilde{\pi}_i^D},v_{\widetilde{\pi}_j^D}} = 1$, and the denominator is the number of transitions between sequential patterns in the series of motifs of length $T - (D-1)\tau$.

Our proposal, henceforth referred to as Weighted Amplitude Transition Graph (WATG), incorporates the absolute difference between the observations that produced the patterns.

First, each \mathcal{X} time series is scaled to [0,1], since we are interested in a metric able to compare data sets:

$$\frac{x_i - x_{\min}}{x_{\max} - x_{\min}} \longmapsto x_i, \tag{4}$$

where x_{min} and x_{max} are, respectively, the minimum and maximum values of the series.

Each $\mathbf{X}_{t}^{(D,\tau)}$ vector is associated with a weight β_{t} that measures the largest difference between its elements:

$$\beta_t = \max\{x_i - x_i\},\tag{5}$$

where $x_i, x_j \in \mathbf{X}_t^{(D,\tau)}$.

Traditionally, the transition graph assigns uniform weight to each transition between patterns and normalizes the result obtained by dividing by the total transitions. In this modification, the $w_{v_{\overline{\pi}_i^D},v_{\overline{\pi}_j^D}}$ weights assigned to each edge depict the amplitude difference observed in the transition. So we have that:

$$w_{v_{\widetilde{\pi}_{i}^{D}},v_{\widetilde{\pi}_{j}^{D}}} = \sum_{i:\{\mathbf{X}_{t}^{(D,\tau)} \mapsto \widetilde{\pi}_{i}^{D}\}} \sum_{j:\{\mathbf{X}_{t}^{(D,\tau)} \mapsto \widetilde{\pi}_{j}^{D}\}} |\beta_{i} - \beta_{j}|.$$

$$(6)$$

Thus, the probability distribution taken from the weighted amplitude transition graph is given as follows:

$$\begin{cases} \lambda_{v_{\widetilde{\pi}_{i}^{D}}, v_{\widetilde{\pi}_{j}^{D}}} = 1, & \text{if } (v_{\widetilde{\pi}_{i}^{D}}, v_{\widetilde{\pi}_{j}^{D}}) \in E \\ \lambda_{v_{\widetilde{\pi}_{i}^{D}}, v_{\widetilde{\pi}_{j}^{D}}} = 0, & \text{otherwise.} \end{cases}$$

$$(7)$$

$$p(\widetilde{\pi}_{i}^{D}, \widetilde{\pi}_{j}^{D}) = \frac{\lambda_{v_{\widetilde{\pi}_{i}^{D}}, v_{\widetilde{\pi}_{j}^{D}}} \cdot w_{v_{\widetilde{\pi}_{i}^{D}}, v_{\widetilde{\pi}_{j}^{D}}}}{\sum_{v_{\widetilde{\pi}_{a}^{D}}, v_{\widetilde{\pi}_{b}^{D}}} w_{v_{\widetilde{\pi}_{a}^{D}}, v_{\widetilde{\pi}_{b}^{D}}}}.$$
(8)

Note that the conditions $p(\widetilde{\pi}_i^D, \widetilde{\pi}_i^D) \geq 0$ e $\sum_{\widetilde{\pi}_i^D, \widetilde{\pi}_i^D} p(\widetilde{\pi}_i^D, \widetilde{\pi}_i^D) = 1$ are satisfied.

Thus, series with uniform amplitudes have edges with probability of occurrence well distributed along the graph, while those with large peaks have edges with probability of occurrence much higher than the others.

3.3 Information-Theoretic Descriptors

Entropy measures the disorder or unpredictability of a system characterized by a probability measure \mathbb{P} .

Let $\mathbb{P} = \{p_{(\widetilde{\pi}_1^D, \widetilde{\pi}_1^D)}, p_{(\widetilde{\pi}_1^D, \widetilde{\pi}_2^D)}, \dots, p_{(\widetilde{\pi}_{D!}^D, \widetilde{\pi}_{D!}^D)}\} = \{p_1, \dots, p_{D!^2}\}$ be the probability distribution taken from the time series weighted amplitude transition graph \mathbb{X} . The normalized Shannon entropy is given by:

$$H(\mathbb{P}) = -\frac{1}{2\log D!} \sum_{\ell=1}^{D!^2} p_{\ell} \log p_{\ell}.$$
 (9)

The ability of the entropy to capture system properties is limited, so it is necessary to use it in conjunction with other descriptors to perform a more complete analysis. Other interesting measures are distances between the \mathbb{P} probability function and a probability measure that describes

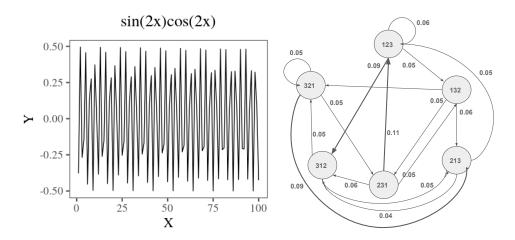


Figure 4. Example of (a) the first 100 samples of the $sin(2x) \cdot cos(2x)$ series and (b) the WATG for the series with D=3 and $\tau=1$

a non-informative process, typically the uniform distribution.

The Jensen-Shannon distance to the uniform distribution $\mathbb{U} = (\frac{1}{D!^2}, \dots, \frac{1}{D!^2})$ is a measure of how similar the underlying dynamics are to a process without information; it is calculated as:

$$Q'(\mathbb{P}, \mathbb{U}) = \sum_{\ell=1}^{D!^2} \left(p_\ell \log \frac{p_\ell}{u_\ell} + u_\ell \log \frac{u_\ell}{p_\ell} \right). \tag{10}$$

This quantity is also called "disequilibrium." The normalized disequilibrium is $Q = Q' / \max\{Q'\}$.

Conversely to entropy, statistical complexity seeks to find interaction and dependence structures among the elements of a given series, being an extremely important factor in the study of dynamic systems.

The Statistical Complexity is then defined as [8]:

$$C(\mathbb{P}, \mathbb{U}) = H(\mathbb{P})Q(\mathbb{P}, \mathbb{U}). \tag{11}$$

In our analysis, each time series can then be described by a point $(H(\mathbb{P}), C(\mathbb{P}, \mathbb{U}))$. The set of all pairs $(H(\mathbb{P}), C(\mathbb{P}, \mathbb{U}))$ for any time series described by patterns of length D lies in a compact subset of \mathbb{R}^2 : the Entropy-Complexity plane.

Through such a tool it is possible to discover the nature of the series, determining if it corresponds to a chaotic (or other deterministic dynamics) or stochastic sequences.

4 Textural classification of SAR regions

Widely used in recognizing geographical features and patterns, synthetic aperture radar (SAR) images are rich in texture information. For this analysis, we will use the HH intensity band of three different region SAR images with uninhabited aerial vehicle (UAVSAR) SAR data, which

provide fully polarimetric SAR observations:

- Sierra del Lacandon National Park, Guatemala (acquired on April 10, 2015)¹;
- Cape Canaveral Ocean Regions (acquired on September 22, 2016);
- Urban area of the city of Munich, Germany (acquired on June 5, 2015)².

will be used in this work 160 samples of size 128×128 , with the following configuration: 40 samples from Guatemalan forest regions; 80 samples from the oceanic regions of Cape Canaveral, where they were divided into two types that have distinct contrast information; and 40 samples of urban regions of the city of Munich. Figure 5 shows examples of each of them.

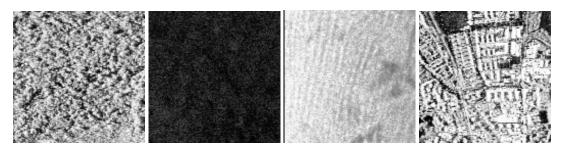


Figure 5. Types of regions analyzed: Forest, Sea Type 1, Sea Type 2, and Urban.

Since the symbolization process is invariant to monotonous transformations and resistant to contamination effects, contrast changes are not capable of causing changes in the final results obtained by the descriptors. Thus, the different types of oceanic regions considered in this work were studied as a single more general class.

Figure 6 shows examples of forest, sea and urban samples as time series, after the linearization process.

5 Expected results

The main contributions already achieved are: (i) a time series discrimination technique with different amplitude information and (ii) a methodology for characterization of regions in SAR images, which proved to be adequate for the proposed data set. In addition, we have preliminary results from quantitative analysis, indicating that such method has a good performance in SAR texture classification.

Following the evaluation of this new metric, two studies are planned: the impact analysis of data imputation methods, and the prediction and simulation of events, both using transition graphs.

 $^{^{1}} https://uavsar.jpl.nasa.gov/cgi-bin/product.pl?jobName=Lacand_30202_15043_006_150410_L090_CX_01\#dados$

https://uavsar.jpl.nasa.gov/cgi-bin/product.pl?jobName=munich_19417_15088_002_150605_L090_CX_01#

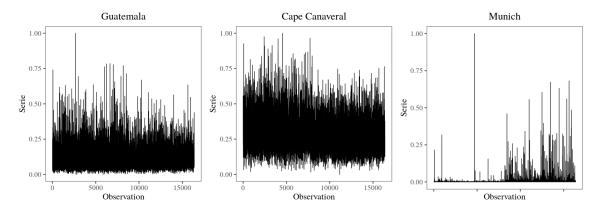


Figure 6. Analysis of the amplitude of the different types of regions: (a) Oceanic Region; (b) Forest Regions and (c) Urban Regions

Exploratory work to study the power of reconstruction of time series data imputation techniques is now in its early stages. We hope that with the use of transition graphs we will be able to contrast the impacts caused by the different approaches and be able to measure which one can most effectively describe the original series, for this we will use the kullback-leibler divergence.

Finally, we plan to conduct a study of ordinal patterns with transition graphs and Markov Chains. The purpose is to combine the evidence of the time series, its ordinal patterns, and the transition graph to make simulations and predictions about the series. One of the innovations of this work is to address one of the limitations of ordinal pattern-based time series analysis techniques that so far do not allow forecasting or simulations.

6 Schedule and checkpoints

This work is in its first year (second semester). As mentioned earlier, the initial research steps have already been taken. The first submission to the conference in this context was made at the Latin American GRSS & ISPRS Remote Sensing Conference, on the proposal for characterization of SAR images.

In the next steps, we plan: (1) conduct a literature review of current SAR image texture classification models; (2)study and propose an amplitude-based time series classification model; (3) investigate the characterization of transition graphs using kullback-leibler divergence; (4) study and design the transition graph prediction and event simulation algorithm; (5) publish the scientific results obtained with this work; and (6) write and present this dissertation.

An estimated schedule for the aforementioned steps, over a period of twelve, is as follows.

7 Final remarks

This proposal presents the dissertation project entitled "A study about ordinal patterns transition graphs". We argue that the information extracted from the ordinal transition graphs gives us

	Jan	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1	Χ											
2	Χ	X										
3			X	X	X							
4						X	X	X				
5									X	X	X	X
6									X	X	X	X

Table 1. Monthly agenda for 2020

more details about the dynamics of the system that governs the time series, being an important characterization tool.

Therefore, the main expected contributions to this work are: (i) the proposal of a methodology of characterization of time series with different amplitudes, (ii) the evaluation of the method in relation to the others proposed in the literature, (iii) the proposed algorithm for evaluating the imputation techniques given, (iv) the proposal of a time series simulation and prediction algorithm based on transition graphs and Markov chains.

We present the key elements needed to achieve the above contributions.

References

- [1] C. Bandt and B. Pompe. Permutation entropy: A natural complexity measure for time series. *Physical review letters*, 88:174102, 2002.
- [2] J. Borges, H. Ramos, R. Mini, O. A. Rosso, A. C. Frery, and A. A. F. Loureiro. Learning and distinguishing time series dynamics via ordinal patterns transition graphs. *Applied Mathematics and Computation*, 362:UNSP 124554, 2019.
- [3] K. S. Chen, W. P. Huang, D. H. Tsay, and F. Amar. Classification of multifrequency polarimetric sar imagery using a dynamic learning neural network. *IEEE Transactions on Geoscience and Remote Sensing*, 34(3):814–820, 1996.
- [4] H. Deng and D. A. Clausi. Unsupervised segmentation of synthetic aperture radar sea ice imagery using a novel markov random field model. *IEEE Transactions on Geoscience and Remote* Sensing, 43:528–538, 2005.
- [5] B. H. Fadlallah, B. Chen, A. Keil, and J. C. Príncipe. Weighted-permutation entropy: a complexity measure for time series incorporating amplitude information. *Physical review. E, Statistical, nonlinear, and soft matter physics,* 87 2:022911, 2013.
- [6] J. B. Florindo and O. M. Bruno. Fractal descriptors based on Fourier spectrum applied to texture analysis. *Physica A: Statistical Mechanics and its Applications*, 391(20):4909–4922, 2012.

- [7] C. W. Kulp, J. M. Chobot, H. R. Freitas, and G. D. Sprechini. Using ordinal partition transition networks to analyze ecg data. *Chaos: An Interdisciplinary Journal of Nonlinear Science*, 26(7):073114, 2016.
- [8] P. W. Lamberti, M. T. Martín, A. Plastino, and O. A. Rosso. Intensive entropic non-triviality measure. *Physica A: Statistical Mechanics and its Applications*, 334(1–2):119–131, 2004.
- [9] J.-H. Lee and Y.-C. Hsueh. Texture classification method using multiple space filling curves. *Pattern Recognition Letters*, 15(12):1241–1244, 1994.
- [10] J.-S. Lee, M. R. Grunes, E. Pottier, and L. Ferro-Famil. Unsupervised terrain classification preserving polarimetric scattering characteristics. *IEEE Transactions on Geoscience and Remote Sensing*, 42(4):722–731, 2004.
- [11] A. J. Luckman, A. C. Frery, C. C. F. Yanasse, and G. B. Groom. Texture in airborne SAR imagery of tropical forest and its relationship to forest regeneration stage. *International Journal of Remote Sensing*, 18(6):1333–1349, 1997.
- [12] M. McCullough, M. Small, T. Stemler, and H. H.-C. Iu. Time lagged ordinal partition networks for capturing dynamics of continuous dynamical systems. *Chaos: An Interdisciplinary Journal of Nonlinear Science*, 25(5):053101, 2015.
- [13] D. Ratha, A. Bhattacharya, and A. C. Frery. Unsupervised classification of polsar data using a scattering similarity measure derived from a geodesic distance. *IEEE Geoscience and Remote Sensing Letters*, 15(1):151–155, 2018.
- [14] J. V. Soares, C. D. Rennó, A. R. Formaggio, C. C. F. Yanasse, and A. C. Frery. An investigation of the selection of texture features for crop discrimination using SAR imagery. *Remote Sensing of Environment*, 59(2):234–247, 1997. Spaceborne Imaging Radar Mission.
- [15] T. Sorrentino, C. Quintero-Quiroz, A. Aragoneses, M. Torrent, and C. Masoller. Effects of periodic forcing on the temporally correlated spikes of a semiconductor laser with feedback. *Optics Express*, 23, 2015.
- [16] J. J. van Zyl and C. F. Burnette. Bayesian classification of polarimetric sar images using adaptive a priori probabilities. *International Journal of Remote Sensing*, 13(5):835–840, 1992.
- [17] H. Zakeri, F. Yamazaki, and W. Liu. Texture analysis and land cover classification of tehran using polarimetric synthetic aperture radar imagery. *Applied Sciences*, 7(12):452, apr 2017.
- [18] J. Zhang, J. Zhou, M. Tang, H. Guo, M. Small, and Y. Zou. Constructing ordinal partition transition networks from multivariate time series. In *Scientific Reports*, 2017.