0.1 Relation to a More Realistic Model

In this section, we relate our model's stylized treatment of uncertainty to the treatment in a related model with a much more realistic (but much less tractable) structure. Specifically, we use the model in ? (henceforth 'CSTW'), which incorporates transitory and permanent idiosyncratic shocks a la? calibrated to match empirical estimates of the magnitude of such household-level shocks in U.S. data. We use that model to calculate a quantitative relationship between the central measures of uncertainty in the two models, and show how this helps provide an interpretation of the quantitative relationship of uncertainty to precautionary saving in our model.

The CSTW model specifies a household income process commonly used in the modern microfounded consumption literature. Household income y_t is determined by the aggregate wage rate W_t and two idiosyncratic components, the permanent component p_t and the transitory shock ξ_t :

$$y_t = p_t \xi_t W_t \tag{1}$$

The permanent component is subject to a shock ψ :

$$p_t = p_{t-1}\psi_t \tag{2}$$

while the transitory component is:

$$\xi_t = \mu \text{ with probability } \mho_t,$$

= $(1 - \tau_t)\ell\theta_t$ with probability $1 - \mho_t$, (3)

Here μ is the unemployment insurance payment when unemployed. τ_t is the rate of tax collected to pay unemployment benefits, ℓ is time worked per employee. ψ and θ are white noise shocks drawn from log normal distribution and $\mathbb{E}_t \psi_{t+n} = \mathbb{E}_t \theta_{t+n} = 1 \,\forall n > 0$. By changing the value of σ_{ψ}^2 and σ_{θ}^2 , we change the degree of uncertainty faced by households. Further details of the model are described in appendix A.7

CSTW find that in order for the model to generate a plausible distribution of wealth it is necessary to build in some form of *ex ante* hetergogeneity; we follow them in assuming that the time preference rate is the locus of heterogeneity, and in calibrating the mean of the time discount factor to match the U.S. aggregate wealth to income ratio and the the spread of the time discount factor to match the wealth distribution among households in U.S. data.¹

Following CSTW, we set the benchmark annual values of σ_{ψ}^2 and σ_{θ}^2 to be 0.010 and 0.010.² The growth impatience condition $\mathbf{p}_{\dot{\Gamma}}$ of the most patient agent restricts the maximum value of σ_{ψ}^2 we can choose.³

¹The wealth to income ratio and wealth distribution data are obtained from the Survey of Consumer Finances.

²Because CSTW is a quarterly model, these annual figures are translated to the corresponding quantities at the quarterly frequency.

³Under this new model setup, the growth impatience condition needs to be modified to be $\mathbf{p}_{\hat{\Gamma}} = \frac{(\beta R)^{1/\rho} \not p}{\Gamma}$ were $\underline{\Gamma} = \Gamma / \mathbb{E}[\psi^{-1}]$. We take care to make sure this condition is satisfied.

0.2 Translation

Our unemployment shock is a permanent shock to income, so it is natural to interpret alternative values of \mho as proxying for differences in the variance of permanent shocks σ_{ψ}^2 in more realistic models. The difficulty is in knowing quantitatively how to translate alternative values of \mho into the corresponding values of σ_{ψ}^2 .

To help with the translation we have constructed a rough bridge between the two models, as follows. First, we set the parameters that the two models share, G, Ξ, ρ, R, X , to the same values (the values reported in Table ??). And we set the values of parameters unique to the CSTW model to their default values from that paper.

Next, we find a value of \mho at which the two models predict the same ratio of aggregate wealth to income, B. Here, the tractability of our model comes in handy; although the analytical function for B in terms of primitives⁴ is not analytically invertible, it is well-behaved (under appropriate parametrica assumptions) so that finding the \mho for which B matches any particular value is numerically trivial.

For the case in hand, the CSTW model's value is $\bar{B} = 2.56$ (matching the target from a 2010 JEDC symposium organized by ? on solution methods for models of this class), and we calculate that our model generates B = 2.565 for $\mho \approx 0.01167 \equiv \bar{\mho}$.

The last step is to map deviations of the two variables, \mho in the tractable model and σ_{ψ}^2 in the CSTW model, from their benchmark values in such a way that, locally, the mapped change in \mho generates the same change in B as the corresponding change in σ_{ψ}^2 .

It turns out that the relationships between steady-state net worth and our principal measures of uncertainty are not far from linear in either model; see 1 and 2 for plots of the relation between B and each model's central measure of uncertainty (\mathcal{O} for the tractable model; σ_{ψ}^2 for CSTW), over the range from half the benchmark value to the full benchmark value. Because it is computationally expensive to calculate, the results from the CSTW model are presented at a set of points sufficient to reveal the shape of the underlying continuous function; because it is analytic, we can plot the function for the tractable model exactly.

Numerically, for a one-half reduction in the magnitude of the uncertainty measure, the two models predict a similar decline in the size of ne worth: In the tractable case, to a value about 0.71 times as large is the benchmark, and for the CSTW case to about 0.65 times the original value.

The two figures are encouragingly similar to each other over the corresponding ranges of their chief uncertainty measures, so we are comfortable in approximating the first order relationships in the vicinity of the target \bar{B} by

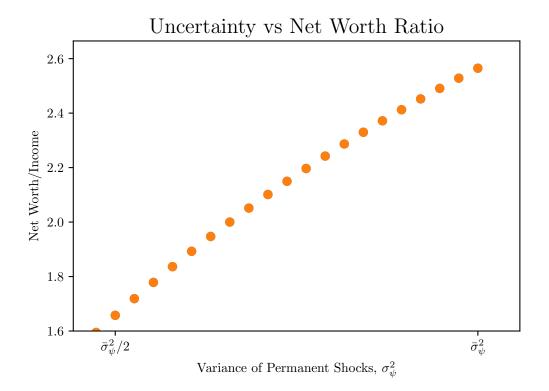
$$B_{\text{tract}}(\mho) \approx \bar{B} + (\mho - \bar{\mho})\zeta_{\text{tract}}$$
 (5)

 $B = \left(\frac{(\mho\Xi\mathsf{G})(1-\alpha)}{\Xi\mathsf{G} - \mathcal{D}\mathbf{\hat{p}}}\right) \left[\frac{\Gamma}{\mathsf{R}} - \frac{1}{2-\Lambda} + \kappa^u \left(\frac{1 + \mathbf{p}_{\Gamma}^{-\rho} - 1}{\mho}\right)^{-\rho}\right]^{-1} \tag{4}$

Figure 1 Tractable Model

Uncertainty vs Net Worth Ratio B 2.6 2.4 2.2 2.0 1.8 $\overline{\overline{0}}/2$ $\overline{\overline{0}} = 1.012$

Figure 2 CSTW Model



in the tractable model and

$$B_{\text{cstw}}(\sigma_{vb}^2) \approx \bar{B} + (\sigma_{vb}^2 - \bar{\sigma}_{vb}^2)\zeta_{\text{cstw}}$$
 (6)

in the CSTW model.

Under these assumptions, there is some ν such that

$$\nu \left(\frac{dB_{\text{tract}}^e}{d\mho} \right) \approx \left(\frac{dB_{\text{cstw}}^e}{d\sigma_{\psi}^2} \right)$$
 (7)

$$\nu \zeta_{\text{tract}} \approx \zeta_{\text{cstw}}$$

$$\nu \approx \zeta_{\text{cstw}} / \zeta_{\text{tract}}$$
(8)

$$\nu \approx \zeta_{\rm cstw}/\zeta_{\rm tract}$$
 (9)

When we undertake this exercise, we obtain a value of $\nu \approx 0.254$. That is, the conclusion is that a rough interpretation of a one unit change in \mho is that it is equivalent to a change of about (1/4) as much in the measurable quantity σ_{ψ}^2 .

There is at present little evidence on the size of ζ_{cstw} . But recent years have seen a growing number of estimates of statistics like σ_{ψ}^2 across countries. Studies comparing small open economies with well-measured data both on saving rates and proxies for σ_{ψ}^2 should be able to measure an empirical counterpart to ζ_{cstw} , if variations in this statistic are in fact large enough to contribute importantly to the difference in saving rates across countries (as, for example, the IMF seems to believe is true [cite IMF advice to China to bolster its social safety net to encourage consumption.