



Due December 2, 2021

HW 8

1. Let X_1, \ldots, X_n represent a random sample from a distribution with pdf

$$f(x;\theta) = \frac{x}{\theta}e^{-\frac{x^2}{2\theta}}, \qquad x > 0$$

a) Show that $E(X^2) = 2\theta$.

Hint: Do a change of variable $y = x^2$ in the integral, and then use the pdf of the Gamma distribution.

- b) Show that $\hat{\theta} = \frac{\sum_{i=1}^{n} X_i^2}{2n}$ is an unbiased estimator of θ .
- 2. From Devore's textbook:

p. 348: Q10

p. 350: Q20

/. a)
$$E[x'] = \int_{0}^{\infty} x' \stackrel{\times}{\partial} \cdot e^{\frac{2\pi}{3}} dx = dt$$
 $e[x'] = \int_{0}^{\infty} -2\pi t e^{\frac{2\pi}{3}} - 2\pi t e^{\frac{2\pi}{3}} dt = dt$
 $e[x'] = \int_{0}^{\infty} -2\pi t e^{\frac{2\pi}{3}} - 2\pi t e^{\frac{2\pi}{3}} dt = dt$
 $e[x'] = 2\pi t e^{\frac{2\pi}{3}} - 2\pi t e^{\frac{2\pi}{3}} -$

2.
$$\alpha$$
) $E(\bar{x}) = \mu$ $V(\bar{x}) = \frac{6^2}{n}$ $V(x) = E(x^2) - [E(x)]^2$

$$= \sum E(x^2) = Vx + [E(x)]^2$$

$$= \frac{6^2}{n} + \mu^2 \neq \mu$$
So \bar{x} is not an imbiased estimator for μ^2 .
$$\Rightarrow \lambda_0 = \frac{6^2}{n} + \mu^2 - k_0^2 = \frac{6^2}{n} + \mu^2 - k_0^2$$

3. a)
$$\hat{p} = \frac{\times + \sqrt{34}}{n + \sqrt{4n}}$$

$$L(\hat{p}) - \hat{p} = L(\frac{\times + \sqrt{4n}}{n + \sqrt{4n}}) - \hat{p} = \frac{np + \sqrt{4n}}{n + \sqrt{4n}} - \hat{p}$$

$$= \frac{n\sqrt{4n}(\frac{1}{2} - p) + n(\frac{1}{2} - p)}{n^2 - n} = \frac{\sqrt{4n}(\frac{1}{2} - p)}{n + \sqrt{4n}} = \frac{\sqrt{4n}(\frac{$$

Elx-Ks] = M2 So Th+ W2 - K6 = K2

 $=7\frac{6^2}{h}=k6^2=>k=\frac{1}{h}$

$$= \frac{dn(\frac{1}{2}-p)+(p-\frac{1}{2})}{n-1} = \frac{(1-nn)(p-\frac{1}{2})}{n-1}$$

$$= \frac{(4n-1)(\frac{1}{2}-p)}{(4n-1)(4n+1)} = \frac{1}{2n+1}$$

$$Var(p) = Var(\frac{x+pyy}{n+n}) = \frac{1}{(n+n)}(np(1-p))$$

$$MSE = Var(p)+E(p)-p)^{2}$$

$$= \frac{np(1-p)}{n+n}(np(1-p))+\frac{1}{(n+n)}(np(1-p))$$

$$= \frac{np(1-p)}{n+n}+\frac{1}{n+n$$