

Mean and Variance of Integrated Conductance in a Continuous-Time Markov Chain

1 Setup

Let X_t be a continuous-time Markov chain on a finite state space $\{1, \dots, K\}$ with generator $Q \in \mathbb{R}^{K \times K}$. We write the state-dependent conductance as a vector $\gamma = (\gamma_1, \dots, \gamma_K)^\top$ and let $\Gamma = \text{diag}(\gamma)$.

Given a time horizon $t > 0$ and a sample path of the chain, we define the *integrated conductance*

$$A(t) = \int_0^t \gamma(X_s) ds. \quad (1)$$

We are interested in the first two moments of $A(t)$, and in particular its mean and variance, conditional on the chain starting in state i and ending in state j :

$$X_0 = i, \quad X_t = j.$$

Note that the interval-averaged conductance used in the main text is

$$\bar{\gamma}_{i \rightarrow j}(t) = \frac{1}{t} \mathbb{E}[A(t) | X_0 = i, X_t = j], \quad (2)$$

so moments of $\bar{\gamma}_{i \rightarrow j}(t)$ follow directly from moments of $A(t)$ by dividing by powers of t .

2 Tilted generator and Feynman–Kac representation

Define the *tilted generator*

$$Q(\alpha) = Q + \alpha \Gamma, \quad \alpha \in \mathbb{R}, \quad (3)$$

and the corresponding transition matrix

$$P^{(\alpha)}(t) = e^{Q(\alpha)t}. \quad (4)$$

A standard Feynman–Kac argument for continuous-time Markov chains yields

$$P_{ij}^{(\alpha)}(t) = \mathbb{E}\left[\exp\left(\alpha \int_0^t \gamma(X_s) ds\right) \mathbf{1}_{\{X_t=j\}} \mid X_0 = i\right]. \quad (5)$$

We abbreviate $A(t) = \int_0^t \gamma(X_s) ds$ and write $P(t) = P^{(0)}(t) = e^{Qt}$.

Differentiating with respect to α at $\alpha = 0$ gives

$$U_{ij}(t) := \frac{\partial}{\partial \alpha} P_{ij}^{(\alpha)}(t) \Big|_{\alpha=0} = \mathbb{E}[A(t) \mathbf{1}_{\{X_t=j\}} \mid X_0 = i], \quad (6)$$

$$W_{ij}(t) := \frac{\partial^2}{\partial \alpha^2} P_{ij}^{(\alpha)}(t) \Big|_{\alpha=0} = \mathbb{E}[A^2(t) \mathbf{1}_{\{X_t=j\}} \mid X_0 = i]. \quad (7)$$

3 Conditional mean and variance of integrated conductance

By conditioning on the final state, we obtain

$$\mathbb{E}[A(t) \mid X_0 = i, X_t = j] = \frac{U_{ij}(t)}{P_{ij}(t)}, \quad (8)$$

$$\mathbb{E}[A^2(t) \mid X_0 = i, X_t = j] = \frac{W_{ij}(t)}{P_{ij}(t)}. \quad (9)$$

Therefore the conditional variance of the integrated conductance is

$$\text{Var}(A(t) \mid X_0 = i, X_t = j) = \frac{W_{ij}(t)}{P_{ij}(t)} - \left(\frac{U_{ij}(t)}{P_{ij}(t)}\right)^2. \quad (10)$$

In terms of the interval-averaged conductance $\bar{\gamma}_{i \rightarrow j}(t) = A(t)/t$, we have

$$\mathbb{E}[\bar{\gamma}_{i \rightarrow j}(t)] = \frac{1}{t} \frac{U_{ij}(t)}{P_{ij}(t)}, \quad (11)$$

$$\text{Var}(\bar{\gamma}_{i \rightarrow j}(t)) = \frac{1}{t^2} \left[\frac{W_{ij}(t)}{P_{ij}(t)} - \left(\frac{U_{ij}(t)}{P_{ij}(t)}\right)^2 \right]. \quad (12)$$

4 Block-matrix representation

The derivatives in (6)–(7) can be computed without explicit diagonalization of Q using matrix exponentials of block matrices.

4.1 First derivative via a $2K \times 2K$ block matrix

Consider the augmented generator

$$\mathcal{Q}_2 = \begin{pmatrix} Q & \Gamma \\ 0 & Q \end{pmatrix} \in \mathbb{R}^{2K \times 2K}. \quad (13)$$

Its exponential has the block form

$$e^{\mathcal{Q}_2 t} = \begin{pmatrix} e^{Qt} & F_1(t) \\ 0 & e^{Qt} \end{pmatrix}, \quad (14)$$

where

$$F_1(t) = \int_0^t e^{Qu} \Gamma e^{Q(t-u)} du = \frac{\partial}{\partial \alpha} e^{(Q+\alpha\Gamma)t} \Big|_{\alpha=0}. \quad (15)$$

Identifying $F_1(t)$ with the first derivative of the tilted semigroup, we have

$$U(t) = F_1(t), \quad (16)$$

so the mean of $A(t)$ conditional on (i, j) can be written as

$$\mathbb{E}[A(t) \mid X_0 = i, X_t = j] = \frac{F_{1,ij}(t)}{[e^{Qt}]_{ij}}. \quad (17)$$

4.2 Second derivative via a $3K \times 3K$ block matrix

To access the second derivative simultaneously, define the $3K \times 3K$ block matrix

$$\mathcal{Q}_3 = \begin{pmatrix} Q & \Gamma & 0 \\ 0 & Q & \Gamma \\ 0 & 0 & Q \end{pmatrix}. \quad (18)$$

Its exponential has the form

$$e^{\mathcal{Q}_3 t} = \begin{pmatrix} e^{Qt} & F_1(t) & F_2(t) \\ 0 & e^{Qt} & F_1(t) \\ 0 & 0 & e^{Qt} \end{pmatrix}, \quad (19)$$

where $F_1(t)$ is as above, and

$$F_2(t) = \int_0^t ds_1 \int_0^{s_1} ds_2 e^{Qs_2} \Gamma e^{Q(s_1-s_2)} \Gamma e^{Q(t-s_1)}. \quad (20)$$

The second derivative of the tilted semigroup can be written as the symmetric double integral

$$W(t) = \frac{\partial^2}{\partial \alpha^2} e^{(Q+\alpha\Gamma)t} \Big|_{\alpha=0} = \int_0^t \int_0^t e^{Qs_<} \Gamma e^{Q(s_>-s_<)} \Gamma e^{Q(t-s_>)} ds_1 ds_2, \quad (21)$$

where $s_< = \min(s_1, s_2)$ and $s_> = \max(s_1, s_2)$. Since the square $[0, t] \times [0, t]$ splits into two identical triangular regions $s_2 < s_1$ and $s_1 < s_2$, we obtain

$$W(t) = 2F_2(t). \quad (22)$$

Combining this with (8)–(10), we find

$$\mathbb{E}[A(t) | X_0 = i, X_t = j] = \frac{F_{1,ij}(t)}{[e^{Qt}]_{ij}}, \quad (23)$$

$$\mathbb{E}[A^2(t) | X_0 = i, X_t = j] = \frac{2F_{2,ij}(t)}{[e^{Qt}]_{ij}}, \quad (24)$$

$$\text{Var}(A(t) | X_0 = i, X_t = j) = \frac{2F_{2,ij}(t)}{[e^{Qt}]_{ij}} - \left(\frac{F_{1,ij}(t)}{[e^{Qt}]_{ij}} \right)^2. \quad (25)$$