

PRINCIPAL COMPONENT ANALYSIS

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Let x be a zero-mean random variable. Suppose we want the direction w such that the projection of x along this direction has maximum variance:

$$\max_w \text{var}(w'x) \quad \text{st.} \quad \|w\| = 1. \quad (1)$$

We have

$$\text{var}(w'x) = E w' x x' w \quad (2)$$

$$= w' \Sigma w. \quad (3)$$

The Lagrangian is

$$L = w' \Sigma w + \lambda(w'w - 1). \quad (4)$$

The stationary condition is

$$\frac{\partial L}{\partial w} = 2\Sigma w - 2\lambda w = 0, \quad (5)$$

$$\Sigma w = \lambda w. \quad (6)$$

Thus w is an eigenvector of Σ . Since

$$w' \Sigma w = w'(\lambda w) = \lambda, \quad (7)$$

the direction with maximum variance is the largest eigenvector. This procedure can be iterated to get the second largest variance projection (orthogonal to the first one), and so on. For a set of data points, we use the ML estimate of the covariance matrix.

1 High-Dimensional Data

Let X be the $d \times n$ matrix of high-dimensional points ($d > n$). Instead of explicitly estimating the covariance matrix, we use the following trick:

$$\left(\frac{1}{n} X X'\right) W = W \Lambda, \quad (8)$$

$$\left(\frac{1}{n} X' X\right) X' W = X' W \Lambda, \quad (9)$$

$$\left(\frac{1}{n} X' X\right) V = V \Lambda. \quad (10)$$

$$\frac{1}{n}X(X'X)F = XFV, \quad (11)$$

$$\left(\frac{1}{n}XX'\right)(XF) = (XF)V, \quad (12)$$

so XF are the eigenvectors of $\frac{1}{n}XX'$. Since

$$\frac{1}{n}X'XF = FV, \quad (13)$$

$$(F'X')(XF) = nF'FV \quad (14)$$

$$(XF)'(XF) = nV, \quad (15)$$

we have $E = \frac{1}{\sqrt{n}}XFV^{-1/2}$.

2 Maximum Information Preservation

Let

$$y = w'x, \quad (16)$$

where $\|w\| = 1$.

We want to maximize $I(X; Y)$ or $H(Y)$ under the constraint on w :

$$J = 1/2(\log(2\pi e)^D + \log |Cy|) + aw'w. \quad (17)$$

$$Cy = E[w'xx'w] = w'E[xx']w = w'Cw. \quad (18)$$

Thus we maximize

$$L = \log |w'Cw| + aw'w \quad dL/dw = 2Cw/|w'Cw| + 2awCw = a|w'Cw|w \quad (19)$$

Thus, w is an eigenvector of the covariance matrix C of x , and the maximum occurs with the eigenvector with the largest eigenvalue.

3 Maximum Likelihood

The data matrix is modeled as linear combinations of a small set of basis vectors plus noise.

$$X = UV + Y \quad (20)$$

X is $D \times N$, where each column is a datum. U is $D \times K$, where each column is a factor. V is $K \times N$, where each column is a vector of coefficients. Y is $D \times N$ noise.

Assuming normal noise with equal variance for all points, ML estimation of UV gives a least squares cost function.

$$J = (X - UV)^2. \quad (21)$$

The stationary conditions are

$$dJ/dU = (X - UV)V' = 0 \quad dJ/dV = U'(X - UV) = 0 \quad XV' = UVV'U'X = U'UV \quad (22)$$

To make the solution unique against rotations and scalings of U and V , we constrain $U'U = I$ and $V'V$ diagonal:

$$V = U'XXX'U = SU1/NXX'U = S/NU \quad (23)$$

The basis functions are the eigenvectors of the covariance matrix, assuming X is zero mean .

4 Generalized PCA

To generalize PCA to other data the noise distribution can be changed and a link function added:

$$X \sim p(f(g(A)h(B))). \quad (24)$$