

Q3. This is an illustration of convergence rate of methods that are similar to Monte

Carlo method, consider the estimation of $z = \pi = 3.1415 \dots$ via the estimator $Z \stackrel{\text{def}}{=} 4\mathbb{1}\{U_1^2 + U_2^2 < 1\}$, where U_1, U_2 are independent uniform(0, 1) r.v.'s. Here $\mathbb{1}\{U_1^2 + U_2^2 < 1\}$ is a Bernoulli r.v. with mean $\pi/4$, so the variance of Z is $4^2(\pi/4)(1 - \pi/4) \approx 2.70$. Thus, to get the leading 3 in π correct w.p. 95% we need $R \approx 1.96^2 2.70 / 0.5^2 = 42$, whereas for the next 1 we need R 100 times more. \square

Q5.

(a) Test for significance ($H_0 : \mu_d = 0$)

Letting $d_i = y_i - z_i$, $\bar{d} = 1.80$, $S_d = 3.60$

$$t_0 = 1.80 / (3.60 / \sqrt{4}) = 1.0$$

For $\alpha = 0.05$, $t_{3,0.025} = 3.18$

Since $|t_0| < 3.18$, do not reject the null hypothesis.

(b) Sample size needed for $\beta \leq 0.20$

$$\delta = 2/3.60 = 0.556$$

For $\alpha = 0.05$, $\beta \leq 0.20$ and $\delta = 0.556$

$n = 30$ observations.

Q2.

For this problem the true optimal solution can be computed analytically: $x^* = 2.611$ years, giving an expected cost of \$11,586. This solution is obtained by minimizing the expected cost, which can be written as

$$2000x + \int_0^\infty 20000 I(y \leq 1) \frac{e^{-y/x}}{x} dx$$

where I is the indicator function.

Q1. Let X represent a process $S - T$ such that S has a distribution that complements the exponential arrival times such that X is normally distributed with zero mean and standard deviation σ . Then:

$$W_{n+1} - w = \phi(W_n - w) + X_n$$

$$W_{n+1} - w = \phi^{n+1}(W_0 - w) + \sum_{j=0}^n \phi^{n-j} X_j$$

$$E(W_{n+1}) = w + \phi^{n+1}(E(W_0) - w) \rightarrow w \text{ as } n \rightarrow \infty$$

$$\text{Var}(W_{n+1}) = \phi^{2(n+1)} + \sigma^2 \sum_{j=0}^n \phi^{2j} \rightarrow \sigma^2 / (1 - \phi^2) \text{ as } n \rightarrow \infty$$

The first two lines in the above can be obtained by induction.

Q4. In the calculation below, substitute epsilon = 3 and t = 2.821 for 9 degrees of freedom and at 1% confidence level obtained as 0.01/(5-1). Also use S12, S23, S24, S25 in the calculations below.

\bar{Y}_i	1	2	3	4	5
	8.9464	6.5448	7.1387	9.5497	11.1213

S_{ij}^2	1	2	3	4	5
1		0.0251	0.0375	0.0330	0.0620
2			0.0028	0.0104	0.0226
3				0.0051	0.0104
4					0.0046

$$t_{0.01/(5-1), (10-1)} = 2.821$$

$$\bar{Y}_{.1} = 8.9464 > \bar{Y}_{.2} + t\sqrt{S_{12}^2/10} = 6.6861$$

$$\bar{Y}_{.3} = 7.1387 > \bar{Y}_{.2} + t\sqrt{S_{13}^2/10} = 6.71755$$

$$\bar{Y}_{.4} = 9.5497 > \bar{Y}_{.2} + t\sqrt{S_{14}^2/10} = 6.70685$$

$$\bar{Y}_{.5} = 11.1213 > \bar{Y}_{.2} + t\sqrt{S_{15}^2/10} = 6.76693$$

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~~Thus, there was adequate data to select the best, policy 2, with 99% confidence.~~ $\hat{S}^2 = \max_{i \neq j} S_{ij}^2 = 0.0620$. The seconde-stage sample size,

$$R = \max \left(R_0, \left\lceil \frac{t^2 \hat{S}^2}{\epsilon^2} \right\rceil \right) = \max \left(10, \left\lceil \frac{(2.821^2)(0.0620)}{3^2} \right\rceil \right) = 10$$

Thus, 10 replication is sufficient to make statistical comparisons.

Since $\min_{i=1}^5 \{\bar{Y}_{.i}\} = \bar{Y}_{.2}$, there was adequate data to conclude that policy 2 has the least expected cost per day with 99% confidence.

However, the conclusion remains as above. The answer does not change.