

Joint Distribution

Independence:

Let (X, Y) be a two dimensional continuous random variable with joint density $f_{XY}(x, y)$ and marginal density f_X and f_Y , respectively. X and Y are independent if and only if

$$f_{XY}(x, y) = f_X(x)f_Y(y), \quad \text{for all } x \text{ and } y$$

Practice Example: The joint density for (X, Y) is $f_{XY}(x, y) = c$, $8.5 \leq x \leq 10.5, 120 \leq y \leq 240$. Are X and Y independent?

We already learned that the joint density function is $f_{XY}(x, y) = \frac{1}{240}$

The marginal density functions can be calculated as:

$$f_X(x) = \int_{120}^{240} \frac{1}{240} dy = \frac{1}{2}, 8.5 \leq x \leq 10.5$$

$$f_Y(y) = \int_{8.5}^{10.5} \frac{1}{240} dx = \frac{1}{120}, 120 \leq y \leq 240$$

It is true that, $f_{XY}(x, y) = f_X(x)f_Y(y)$, for all x and y

Therefore, X and Y are independent.

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Practice Example: Are (X, Y) with the following joint density independent?

- (1) Answer the question by inspecting the joint probability table
- (2) Answer the question by checking against the independence condition

x\Y	0	1	2	3
0	0.840	0.030	0.020	0.010
1	0.060	0.010	0.008	0.002
2	0.010	0.005	0.004	0.001

From the table, it is easy to find out that:

$$f_X(0) = 0.90; f_Y(0) = 0.91; f_{XY}(0,0) = 0.84$$

$$f_{XY}(0,0) \neq f_X(0) \times f_Y(0)$$

Hence, they are not independent!

Joint Distribution

Expected Value:

Let (X, Y) be a two dimensional continuous random variable with joint density $f_{XY}(x, y)$. Let $H(X, Y)$ be a random variable. The expected value of $H(X, Y)$ is given by:

$$\text{Discrete case: } E[H(X, Y)] = \sum_{all\ x} \sum_{all\ y} H(x, y) f_{XY}(x, y) ; E[XY] = \sum_{all\ x} \sum_{all\ y} xy f_{XY}(x, y)$$

$$\text{Continuous case: } E[H(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} H(x, y) f_{XY}(x, y) dy dx ; E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{XY}(x, y) dy dx$$

Practice: Evaluate **Univariate expectation via the joint density. I.e., Find $E[H(X, Y)]$, when $H(X, Y) = X$; $H(X, Y) = Y$**

Actually,

$$E[X] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f_{XY}(x, y) dy dx = \int_{-\infty}^{\infty} x \int_{-\infty}^{\infty} f_{XY}(x, y) dy dx = \int_{-\infty}^{\infty} x f_X(x) dx$$
$$E[Y] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{XY}(x, y) dy dx = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{XY}(x, y) dx dy = \int_{-\infty}^{\infty} y \int_{-\infty}^{\infty} f_{XY}(x, y) dx dy = \int_{-\infty}^{\infty} y f_Y(y) dy$$

Switched the order of integration

Joint Distribution

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Practice Example: The joint density for (X, Y) is $f_{XY}(x, y) = \frac{1}{240}, 8.5 \leq x \leq 10.5, 120 \leq y \leq 240$.

Find $E[XY]$, $E[X]$, and $E[Y]$

$$E[X] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f_{XY}(x, y) dy dx = \int_{8.5}^{10.5} \int_{120}^{240} \frac{x}{240} dy dx = \int_{8.5}^{10.5} \frac{x}{2} dx = 9.5$$

$$E[Y] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{XY}(x, y) dx dy = \int_{120}^{240} \int_{8.5}^{10.5} \frac{y}{240} dx dy = \int_{120}^{240} \frac{y}{120} dy = 180$$

$$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{XY}(x, y) dy dx = \int_{120}^{240} \int_{8.5}^{10.5} \frac{xy}{240} dx dy = \int_{120}^{240} \frac{19y}{240} dy = 1710$$

Joint Distribution

Expected Value:

Let (X, Y) be a two dimensional continuous random variable with joint density $f_{XY}(x, y)$. Let $H(X, Y)$ be a random variable. The expected value of $H(X, Y)$ is given by:

$$\text{Discrete case: } E[H(X, Y)] = \sum_{\text{all } x} \sum_{\text{all } y} H(x, y) f_{XY}(x, y) ; E[XY] = \sum_{\text{all } x} \sum_{\text{all } y} xy f_{XY}(x, y)$$

$$\text{Continuous case: } E[H(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} H(x, y) f_{XY}(x, y) dy dx ; E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f_{XY}(x, y) dy dx$$

Practice Example: Evaluate $E[X + Y]$ show that $E[X + Y] = E[X] + E[Y]$

$$\begin{aligned} E[X + Y] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (x + y) f_{XY}(x, y) dy dx = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f_{XY}(x, y) dy dx + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y f_{XY}(x, y) dx dy \\ &= \int_{-\infty}^{\infty} x f_X(x) dx + \int_{-\infty}^{\infty} y f_Y(y) dy = E[X] + E[Y] \end{aligned}$$

Switched the order of integration



Joint Distribution

Covariance:

Let X and Y be random variables with means μ_X and μ_Y respectively. The covariance between X and Y , denoted by $Cov(X, Y)$ or σ_{XY} is given by:

$$Cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$$

Can this quantity represent some kind of correlation between X and Y ?

Prove that $Cov(X, Y) = E[XY] - E[X]E[Y]$

$$\begin{aligned} Cov(X, Y) &= E[(X - \mu_X)(Y - \mu_Y)] \\ &= E[XY - \mu_Y X - \mu_X Y + \mu_X \mu_Y] \\ &= E[XY] - \mu_Y E[X] - \mu_X E[Y] + \mu_X \mu_Y \\ &= E[XY] - E[Y]E[X] - E[X]E[Y] + E[X]E[Y] \\ &= E[XY] - E[X]E[Y] \end{aligned}$$

Evaluate $Cov(X, Y)$ when $X = Y$.

Can $Cov(X, Y) < 0$?

Joint Distribution

Prove that $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)$.

We already learned that When X and Y are independent, $\text{Var}(X + Y) = \text{Var}(X) + \text{Var}(Y)$

Proof:

$$\begin{aligned}\text{Var}(X + Y) &= E[(X + Y)^2] - (E[X + Y])^2 \\ &= E[X^2 + 2XY + Y^2] - (E[X] + E[Y])^2 \\ &= E[X^2] + 2E[XY] + E[Y^2] - (E[X])^2 - 2E[X]E[Y] - (E[Y])^2 \\ &= E[X^2] - (E[X])^2 + E[Y^2] - (E[Y])^2 + 2E[XY] - 2E[X]E[Y] \\ &= \text{Var}(X) + \text{Var}(Y) + 2\text{Cov}(X, Y)\end{aligned}$$

Joint Distribution

Theorem: Let (X, Y) be a two dimensional continuous random variable with joint density $f_{XY}(x, y)$. If X and Y are independent then $E[XY] = E[X]E[Y]$

Proof:

It is known that if X and Y are independent, we have $f_{XY}(x, y) = f_X(x)f_Y(y)$

Then,

$$\begin{aligned} E[XY] &= \iint_{-\infty}^{\infty} xy f_{XY}(x, y) dy dx = \iint_{-\infty}^{\infty} xy f_X(x) f_Y(y) dy dx \\ &= \int_{-\infty}^{\infty} x f_X(x) \int_{-\infty}^{\infty} y f_Y(y) dy dx = \int_{-\infty}^{\infty} x f_X(x) E[Y] dx = E[Y] \int_{-\infty}^{\infty} x f_X(x) dx = E[Y]E[X] \end{aligned}$$

Joint Distribution

Theorem: Let (X, Y) be a two dimensional continuous random variable with joint density $f_{XY}(x, y)$. If X and Y are independent then $E[XY] = E[X]E[Y]$

From this theorem, we have

$$X \text{ and } Y \text{ are independent} \Rightarrow \text{Cov}(X, Y) = 0$$

But, $\text{Cov}(X, Y) = 0 \nRightarrow X \text{ and } Y \text{ are independent !!!}$

Here is an example:

Let $X \sim N(0, 1)$. We have, $m_X(t) = e^{\frac{t^2}{2}}$

Then,

$$E[X] = \frac{d}{dt} m_X(t)_{t=0} = t e^{\frac{t^2}{2}} \Big|_{t=0} = 0; E[X^2] = \frac{d^2}{dt^2} m_X(t)_{t=0} = \left[t^2 e^{\frac{t^2}{2}} + e^{\frac{t^2}{2}} \right]_{t=0} = 1;$$
$$E[X^3] = \frac{d^3}{dt^3} m_X(t)_{t=0} = \left[t^3 e^{\frac{t^2}{2}} + 3t e^{\frac{t^2}{2}} \right]_{t=0} = 0$$

Let $Y = X^2$. X and Y are related. But,

$$\text{Cov}(X, Y) = E[XY] - E[X]E[Y] = E[X^3] - E[X]E[X^2] = 0$$

Joint Distribution

Covariance:

Let X and Y be random variables with means μ_X and μ_Y respectively. The covariance between X and Y , denoted by $Cov(X, Y)$ or σ_{XY} is given by $Cov(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] = E[XY] - E[X]E[Y]$

Practice Example: Find $Cov(X_1, Y_1)$ and $Cov(X_2, Y_2)$. Can you tell the correlation of X and Y before the calculation?

$X_1 \backslash Y_1$	0	1	2
0	0.40	0.006	0.004
1	0.006	0.30	0.004
2	0.004	0.006	0.27

$$\begin{aligned}E[X_1 Y_1] &= 1.4 \\E[X_1] &= 0.87 \\E[Y_1] &= 0.868 \\Cov(X_1, Y_1) &= 0.645\end{aligned}$$

$X_2 \backslash Y_2$	0	1	2
0	0.1	0.1	0.1
1	0.13	0.13	0.14
2	0.1	0.1	0.1

$$\begin{aligned}E[X_2 Y_2] &= 1.01 \\E[X_2] &= 1.0 \\E[Y_2] &= 1.01 \\Cov(X_2, Y_2) &= 0\end{aligned}$$