

习题6.4 第14 题第(1) 小题的判断无偏性两种方法:

14. 设 X_1, X_2, \dots, X_n 为独立同分布变量, $0 < \theta < 1$,

$$P\{X_1 = -1\} = \frac{1-\theta}{2}, \quad P\{X_1 = 0\} = \frac{1}{2}, \quad P\{X_1 = 1\} = \frac{\theta}{2}.$$

(1) 求 θ 的MLE $\hat{\theta}_1$, 并问 $\hat{\theta}_1$ 是否无偏。

解: 总体 X 概率函数为

$$p(x; \theta) = \left(\frac{1-\theta}{2}\right)^{\frac{x(x-1)}{2}} \left(\frac{1}{2}\right)^{-(x+1)(x-1)} \left(\frac{\theta}{2}\right)^{\frac{x(x+1)}{2}} = \frac{1}{2}(1-\theta)^{\frac{x^2-x}{2}} \theta^{\frac{x^2+x}{2}}, \quad x = -1, 0, 1,$$

$$\text{则似然函数 } L(\theta) = \prod_{i=1}^n \frac{1}{2}(1-\theta)^{\frac{x_i^2-x_i}{2}} \theta^{\frac{x_i^2+x_i}{2}} = \frac{1}{2^n}(1-\theta)^{\frac{1}{2}\left(\sum_{i=1}^n x_i^2 - \sum_{i=1}^n x_i\right)} \theta^{\frac{1}{2}\left(\sum_{i=1}^n x_i^2 + \sum_{i=1}^n x_i\right)},$$

$$\text{有 } \ln L(\theta) = \frac{1}{2}\left(\sum_{i=1}^n x_i^2 - \sum_{i=1}^n x_i\right) \ln(1-\theta) + \frac{1}{2}\left(\sum_{i=1}^n x_i^2 + \sum_{i=1}^n x_i\right) \ln \theta - n \ln 2,$$

$$\text{令 } \frac{d \ln L(\theta)}{d\theta} = \frac{1}{2}\left(\sum_{i=1}^n x_i^2 - \sum_{i=1}^n x_i\right) \cdot \frac{-1}{1-\theta} + \frac{1}{2}\left(\sum_{i=1}^n x_i^2 + \sum_{i=1}^n x_i\right) \cdot \frac{1}{\theta} = 0,$$

$$\text{得 } \theta = \frac{\sum_{i=1}^n x_i^2 + \sum_{i=1}^n x_i}{2 \sum_{i=1}^n x_i^2} = \frac{1}{2} + \frac{\sum_{i=1}^n x_i}{2 \sum_{i=1}^n x_i^2},$$

$$\text{故 } \theta \text{ 的MLE } \hat{\theta}_1 = \frac{1}{2} + \frac{\sum_{i=1}^n X_i}{2 \sum_{i=1}^n X_i^2};$$

方法一: 设 X_1, X_2, \dots, X_n 中取值 $-1, 0, 1$ 分别有 n_{-1}, n_0, n_1 次, 有 $n_{-1} + n_0 + n_1 = n$,

$$\text{则 } \sum_{i=1}^n X_i^2 = n_1 + n_{-1}, \quad \sum_{i=1}^n X_i = n_1 - n_{-1},$$

$$\text{即 } \theta \text{ 的MLE } \hat{\theta}_1 = \frac{\sum_{i=1}^n X_i^2 + \sum_{i=1}^n X_i}{2 \sum_{i=1}^n X_i^2} = \frac{n_1}{n_1 + n_{-1}};$$

$$\text{因 } E(\hat{\theta}_1) = E\left(\frac{n_1}{n_{-1} + n_1}\right) = E\left[E\left(\frac{n_1}{n_{-1} + n_1} \middle| n_{-1} + n_1\right)\right] = E\left[\frac{1}{n_{-1} + n_1} E(n_1 | n_{-1} + n_1)\right],$$

$$\text{且 } P\{X = 1 | X = -1 \text{ 或 } X = 1\} = \frac{P\{X = 1\}}{P\{X = -1 \text{ 或 } X = 1\}} = \frac{\frac{\theta}{2}}{\frac{1-\theta}{2} + \frac{\theta}{2}} = \theta,$$

则在 $n_{-1} + n_1 = m$ 的条件下, n_1 服从二项分布 $b(m, \theta)$, $E(n_1 | n_{-1} + n_1 = m) = m\theta$,

可得 $E(n_1 | n_{-1} + n_1) = (n_{-1} + n_1)\theta$,

$$\text{故 } E(\hat{\theta}_1) = E\left[\frac{1}{n_{-1} + n_1} E(n_1 | n_{-1} + n_1)\right] = E(\theta) = \theta, \quad \hat{\theta}_1 \text{ 是 } \theta \text{ 的无偏估计};$$

$$\begin{aligned}
\text{方法二: } E(\hat{\theta}_1) &= \frac{1}{2} + \frac{1}{2} E \left(\frac{\sum_{j=1}^n X_j}{\sum_{i=1}^n X_i^2} \right) = \frac{1}{2} + \frac{1}{2} \sum_{j=1}^n E \left(\frac{X_j}{\sum_{i=1}^n X_i^2} \right) = \frac{1}{2} + \frac{n}{2} E \left(\frac{X_1}{X_1^2 + \sum_{i=2}^n X_i^2} \right) \\
&= \frac{1}{2} + \frac{n}{2} E \left[E \left(\frac{X_1}{X_1^2 + \sum_{i=2}^n X_i^2} \middle| \sum_{i=2}^n X_i^2 \right) \right] \\
&= \frac{1}{2} + \frac{n}{2} E \left(\frac{-1}{1 + \sum_{i=2}^n X_i^2} \cdot \frac{1-\theta}{2} + \frac{0}{0 + \sum_{i=2}^n X_i^2} \cdot \frac{1}{2} + \frac{1}{1 + \sum_{i=2}^n X_i^2} \cdot \frac{\theta}{2} \right) \\
&= \frac{1}{2} + \frac{n}{2} \left(\theta - \frac{1}{2} \right) E \left(\frac{1}{1 + \sum_{i=2}^n X_i^2} \right),
\end{aligned}$$

因 $\sum_{i=2}^n X_i^2$ 是 X_2, \dots, X_n 中取值 -1 或 1 的次数, 则 $\sum_{i=2}^n X_i^2$ 服从二项分布 $b\left(n-1, \frac{1}{2}\right)$,

$$\text{即 } P\left\{\sum_{i=2}^n X_i^2 = k\right\} = C_{n-1}^k \cdot \left(\frac{1}{2}\right)^k \left(\frac{1}{2}\right)^{n-1-k} = \frac{(n-1)!}{k!(n-1-k)!} \cdot \left(\frac{1}{2}\right)^{n-1}, \quad k=0, 1, \dots, n-1,$$

$$\begin{aligned}
\text{则 } E(\hat{\theta}_1) &= \frac{1}{2} + \frac{n}{2} \left(\theta - \frac{1}{2} \right) \cdot \sum_{k=0}^{n-1} \frac{1}{1+k} \cdot \frac{(n-1)!}{k!(n-1-k)!} \cdot \left(\frac{1}{2}\right)^{n-1} \\
&= \frac{1}{2} + \frac{n}{2} \left(\theta - \frac{1}{2} \right) \cdot \left(\frac{1}{2}\right)^{n-1} \cdot \sum_{k=0}^{n-1} \frac{(n-1)!}{(k+1)!(n-1-k)!} = \frac{1}{2} + \frac{n}{2^n} \left(\theta - \frac{1}{2} \right) \cdot \sum_{k=0}^{n-1} \frac{1}{n} \cdot C_n^{k+1} \\
&= \frac{1}{2} + \frac{n}{2^n} \left(\theta - \frac{1}{2} \right) \cdot \frac{1}{n} (2^n - 1) = \frac{1}{2} + \left(\theta - \frac{1}{2} \right) \cdot \left(1 - \frac{1}{2^n} \right) = \frac{1}{2^{n+1}} + \theta \left(1 - \frac{1}{2^n} \right),
\end{aligned}$$

故 $\hat{\theta}_1$ 不是 θ 的无偏估计;

两种方法结论不同, 问题在于当所有样品全部取值 0 时, 有 $\sum_{i=1}^n X_i^2 = n_{-1} + n_1 = 0$, 此时

$$\hat{\theta}_1 = \frac{\sum_{i=1}^n X_i^2 + \sum_{i=1}^n X_i}{2 \sum_{i=1}^n X_i^2} = \frac{n_1}{n_{-1} + n_1} \text{ 无定义, 正确的做法应该是在 } \sum_{i=1}^n X_i^2 = n_{-1} + n_1 \neq 0 \text{ 的条件下, 计算}$$

$\hat{\theta}_1$ 的数学期望。事实上在 $\sum_{i=1}^n X_i^2 = n_{-1} + n_1 \neq 0$ 的条件下, $\hat{\theta}_1$ 的数学期望就是 θ 。

方法一中相当于当 $\sum_{i=1}^n X_i^2 = n_{-1} + n_1 = 0$ 时将 $\hat{\theta}_1$ 取为 θ , 所以最后的答案就是 θ ; 在方法二

中相当于当 $\sum_{i=1}^n X_i^2 = n_{-1} + n_1 = 0$ 时将 $\hat{\theta}_1$ 取为 $\frac{1}{2}$, 且 $\sum_{i=1}^n X_i^2 = n_{-1} + n_1 = 0$ 即所有样品全部取值 0 的

概率为 $\frac{1}{2^n}$ ，所以最后的答案就是 $\frac{1}{2} \cdot \frac{1}{2^n} + \theta \cdot \left(1 - \frac{1}{2^n}\right) = \frac{1}{2^{n+1}} + \theta \left(1 - \frac{1}{2^n}\right)$ 。可见方法一的结论是正确的。