补充资料:

拟合优度检验的 χ² 统计量推导过程

当检验数据分成r类,需要检验在每一类数据中取值的概率是否为给定的 p_i , $i=1,2,\cdots,r$.对容量为n的数据样本,设其中属于第i类的样品个数为 n_i , $i=1,2,\cdots,r$,($n_1+n_2+\cdots+n_r=n$),证明,当n很大时

$$\chi^2 = \sum_{i=1}^r \frac{(n_i - np_i)^2}{np_i} \sim \chi^2(r-1) \circ$$

证明: $因 n_i \sim b(n, p_i)$, 当 n 很大时, n_i 近似服从正态分布 $n_i \sim N(np_i, np_i(1-p_i))$ 。又因 $i \neq j$ 时,

$$n_i + n_j \sim b(n, p_i + p_j)$$
,

则

$$Var(n_i + n_j) = n(p_i + p_j)(1 - p_i - p_j)$$
,

可得

$$Cov(n_i, n_j) = \frac{Var(n_i + n_j) - Var(n_i) - Var(n_j)}{2},$$

$$= \frac{n(p_i + p_j)(1 - p_i - p_j) - np_i(1 - p_i) - np_j(1 - p_j)}{2} = -np_i p_j,$$

设 $X_i = \frac{n_i - np_i}{\sqrt{np_i}}$,有 X_i 近似服从正态分布,且

$$E(X_i) = \frac{1}{\sqrt{np_i}} E(n_i - np_i) = 0$$
, $Var(X_i) = \frac{1}{np_i} Var(n_i) = \frac{1}{np_i} np_i (1 - p_i) = 1 - p_i$,

$$Cov(X_i, X_j) = \frac{1}{\sqrt{np_i}\sqrt{np_j}}Cov(n_i, n_j) = \frac{1}{\sqrt{np_i}\sqrt{np_j}} \cdot (-np_i p_j) = -\sqrt{p_i p_j},$$

记 $\overrightarrow{X} = (X_1, X_2, \dots, X_r)^T$,有 \overrightarrow{X} 的均值向量和协方差矩阵分别为

$$\overrightarrow{EX} = 0$$
,

$$\operatorname{Cov}(\overrightarrow{X}, \overrightarrow{X}) = \begin{pmatrix} 1 - p_1 & -\sqrt{p_1 p_2} & \cdots & -\sqrt{p_1 p_r} \\ -\sqrt{p_2 p_1} & 1 - p_2 & \cdots & -\sqrt{p_2 p_r} \\ \vdots & \vdots & \ddots & \vdots \\ -\sqrt{p_r p_1} & -\sqrt{p_r p_2} & \cdots & 1 - p_r \end{pmatrix},$$

因

$$D_r = |\lambda E_r - \operatorname{Cov}(\overrightarrow{X}, \overrightarrow{X})| = \begin{vmatrix} \lambda - 1 + p_1 & \sqrt{p_1 p_2} & \cdots & \sqrt{p_1 p_r} \\ \sqrt{p_2 p_1} & \lambda - 1 + p_2 & \cdots & \sqrt{p_2 p_r} \\ \vdots & \vdots & \ddots & \vdots \\ \sqrt{p_r p_1} & \sqrt{p_r p_2} & \cdots & \lambda - 1 + p_r \end{vmatrix}$$

$$= \begin{vmatrix} \lambda - 1 + p_1 & \sqrt{p_1 p_2} & \cdots & \sqrt{p_1 p_r} \\ \sqrt{p_2 p_1} & \lambda - 1 + p_2 & \cdots & \sqrt{p_2 p_r} \\ \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & \cdots & \lambda - 1 \end{vmatrix} + \begin{vmatrix} \lambda - 1 + p_1 & \sqrt{p_1 p_2} & \cdots & \sqrt{p_1 p_r} \\ \sqrt{p_2 p_1} & \lambda - 1 + p_2 & \cdots & \sqrt{p_2 p_r} \\ \vdots & \vdots & \ddots & \vdots \\ \sqrt{p_r p_1} & \sqrt{p_r p_2} & \cdots & p_r \end{vmatrix}$$

$$= (\lambda - 1)D_{r-1} + \sqrt{p_r} \begin{vmatrix} \lambda - 1 + p_1 & \sqrt{p_1 p_2} & \cdots & \sqrt{p_1 p_r} \\ \sqrt{p_2 p_1} & \lambda - 1 + p_2 & \cdots & \sqrt{p_2 p_r} \\ \vdots & \vdots & \ddots & \vdots \\ \sqrt{p_1} & \sqrt{p_2} & \cdots & \sqrt{p_r} \end{vmatrix}$$

$$= (\lambda - 1)D_{r-1} + \sqrt{p_r} \begin{vmatrix} \lambda - 1 & 0 & \cdots & 0 \\ 0 & \lambda - 1 & \cdots & 0 \\ \vdots & \vdots & \ddots & \vdots \\ \sqrt{p_1} & \sqrt{p_2} & \cdots & \sqrt{p_r} \end{vmatrix} = (\lambda - 1)D_{r-1} + p_r(\lambda - 1)^{r-1},$$

则由递推公式可得

$$\begin{split} D_r &= (\lambda - 1)D_{r-1} + p_r(\lambda - 1)^{r-1} = (\lambda - 1)[(\lambda - 1)D_{r-2} + p_{r-1}(\lambda - 1)^{r-2}] + p_r(\lambda - 1)^{r-1} \\ &= (\lambda - 1)^2 D_{r-2} + p_{r-1}(\lambda - 1)^{r-1} + p_r(\lambda - 1)^{r-1} \\ &= (\lambda - 1)^3 D_{r-3} + p_{r-2}(\lambda - 1)^{r-1} + p_{r-1}(\lambda - 1)^{r-1} + p_r(\lambda - 1)^{r-1} \\ &= (\lambda - 1)^r + p_1(\lambda - 1)^{r-1} + \dots + p_{r-1}(\lambda - 1)^{r-1} + p_r(\lambda - 1)^{r-1} \\ &= (\lambda - 1)^{r-1}(\lambda - 1 + p_1 + \dots + p_{r-1} + p_r) = \lambda(\lambda - 1)^{r-1} \,, \end{split}$$

令

$$D_r = |\lambda E_r - \text{Cov}(\overrightarrow{X}, \overrightarrow{X})| = \lambda(\lambda - 1)^{r-1} = 0$$
,

可得协方差矩阵 $Cov(\vec{X}, \vec{X})$ 的特征值为

$$\lambda_1=\lambda_2=\cdots=\lambda_{r-1}=1$$
 , $\lambda_r=0$,

有 $Cov(\vec{X}, \vec{X})$ 正交相似于对角阵

$$\Lambda = \begin{pmatrix} 1 & 0 & \cdots & 0 & 0 \\ 0 & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & \cdots & 1 & 0 \\ 0 & 0 & \cdots & 0 & 0 \end{pmatrix},$$

即存在正交阵C, 使得

$$C^{-1}\operatorname{Cov}(\overrightarrow{X},\overrightarrow{X})C = C^{T}\operatorname{Cov}(\overrightarrow{X},\overrightarrow{X})C = \Lambda$$
,

作正交变换

$$\overrightarrow{X} = C\overrightarrow{Y} = C(Y_1, Y_2, \dots, Y_r)^T,$$

有 $\vec{Y} = C^T \vec{X}$,则 \vec{Y} 的均值向量和协方差矩阵分别为

$$\overrightarrow{EY} = \overrightarrow{C}^T E \overrightarrow{X} = \overrightarrow{C}^T \cdot 0 = 0$$
,

$$Cov(\vec{Y}, \vec{Y}) = Cov(C^T \vec{X}, C^T \vec{X}) = C^T Cov(\vec{X}, \vec{X})C = \Lambda = \begin{pmatrix} 1 & 0 & \cdots & 0 & 0 \\ 0 & 1 & \cdots & 0 & 0 \\ \vdots & \vdots & \ddots & \vdots & \vdots \\ 0 & 0 & \cdots & 1 & 0 \\ 0 & 0 & \cdots & 0 & 0 \end{pmatrix},$$

则 $Y_1, Y_2, \dots, Y_{r-1}, Y_r$ 相互独立,且 Y_1, Y_2, \dots, Y_{r-1} 都服从正态分布 N(0,1) ,而 $EY_r = 0$, $Cov(Y_r) = 0$,即 Y_r 几 乎必然等于 0 ,因

$$\chi^{2} = \sum_{i=1}^{r} \frac{(n_{i} - np_{i})^{2}}{np_{i}} = \sum_{i=1}^{r} X_{i}^{2} = \overrightarrow{X}^{T} \overrightarrow{X} = \overrightarrow{Y}^{T} C^{T} C \overrightarrow{Y} = \overrightarrow{Y}^{T} \overrightarrow{Y} = \sum_{i=1}^{r} Y_{i}^{2} = \sum_{i=1}^{r-1} Y_{i}^{2}, \quad a.s. \ .$$

故

$$\chi^2 = \sum_{i=1}^r \frac{(n_i - np_i)^2}{np_i} \stackrel{\sim}{\sim} \chi^2(r-1) \ .$$