

将一枚硬币抛掷两次，引进事件： $A_1 = \{\text{第一次出现正面}\}$ ，
 $A_2 = \{\text{第二次出现正面}\}$ ， $A_3 = \{\text{正反面各出现一次}\}$ ， $A_4 = \{\text{正面出现两次}\}$ 。
问：

- (A). A_1, A_2, A_3 相互独立。
- (B). A_2, A_3, A_4 相互独立。
- (C). A_1, A_2, A_3 两两独立。
- (D). A_2, A_3, A_4 两两独立。

Soln.

- 首先 A_1, A_2, A_3 相互独立, 指 $P(A_1 A_2 A_3) = 0$

两两独立. $P(A_1 A_2) = 0, P(A_2 A_3) = 0$ 就不行.

$$P(A_1, A_3) = 0.$$

样本

- 试验的样本空间有四点. $\Omega = \{(正, 正), (正, 反), (反, 正), (反, 反)\}$

- $A_1 \subset A_4$ $A_2 \subset A_4$, 且 A_3, A_4 不相容.

第一次正面 第二次正面

- 依古典概型. $P(A_1) = P(A_2) = P(A_3) = \frac{1}{2}$ $P(A_4) = \frac{1}{4}$

- A_1, A_2 独立. obvious. 第一次和第二次出现正面没关系.

- $A_1 > A_3$ 亦独立, A_3 : 正反面各出现一次
 A_2 → 不一定第-次正面. 反之. 第一次正面
也不一定正反面各出现一次.

- 而 A_1, A_2, A_3 亦非相互独立.

$$P(\underbrace{A_1 A_2}_{\text{两次正面}}) = P(A_1) P(A_2).$$

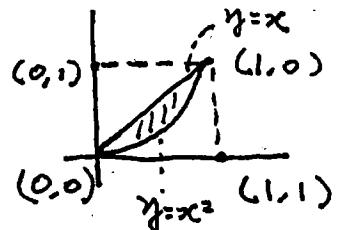
$$P(\underbrace{A_1 A_2 A_3}_{\text{三次正面}}) = 0$$

一正一反

→ 该平面区域 D 是由坐标为 $(0,0), (0,1), (1,0), (1,1)$ 的四个点围成之正方形，今向 D 内随机投入 10 个点，求这 10 个点中，至少有 2 点落于曲线 $y = x^2$ 或直线 $y = x$ 所围成区域 D_1 之概率。

Soln:

- 投点-差T点，视为一个随机试验。入 D_1 的概率。



$$\text{Area of } D_1 = \int_0^1 (x - x^2) dx = \frac{1}{6}$$

$$\text{故每 T 点入 } D_1 \in \text{Prob: } P = \frac{1}{6}$$

- 10次投点，意味着10次试验。

设 B_k = 10次投点中，有k个点落入 D_1 中。
 $\therefore B_k \sim B(10, \frac{1}{6}) =$

- $\{\text{至多有 k 个点入 } D_1\} = \{B_1 \cup B_2 \cup \dots \cup B_{10}\}$

至多不好算，算至多。

$$\{1 - \text{至多有 } k \text{ 点}\} = \{1 - B_0 \cup B_1\}$$

$$= 1 - C_{10}^0 \left(\frac{1}{6}\right)^0 \left(\frac{5}{6}\right)^{10} - C_{10}^1 \left(\frac{1}{6}\right)^1 \left(\frac{5}{6}\right)^9$$

$$\approx 0.52.$$

12页
Prob
P₉₃₅

甲乙二人轮流投篮，游戏规则规定甲先开始，且甲每轮只投一次。
而乙每轮连续投两次，先投中者胜。设甲、乙每次投篮之命中率，
分别为 $P_1 = 0.5$ ， $P_2 = \frac{1}{2}$ 时，甲乙的负概率相同。

Schl:

什么时候甲会赢呢。

$$\begin{aligned} & \text{甲投进} + \underbrace{\text{甲不进} \cdot \text{乙不进} \cdot \text{乙不进} \cdot \text{甲投进}}_{P} + \underbrace{\text{甲不进} \cdot \text{乙不进} \cdot \text{乙不进} \cdot \text{甲不进} \cdot \text{乙不进} \cdot \text{甲投进}}_{(1-P)(1-q)^2(1-q)^2 P} \\ & \quad (1-P)(1-q)^2 \quad (1-P)(1-q)^2 P \end{aligned}$$

$$P\{\text{甲会赢}\} = P + \underbrace{(1-P)(1-q)^2}_{\text{记作 } a} P + a^2 P.$$

$$= P(1+a+a^2+\dots) = P \cdot \frac{1}{1-a}.$$

$$q=0.5 \Rightarrow a=\frac{1}{2}(1-P)$$

甲乙赢的概率相同：

$$P\{\text{甲会赢}\} = 0.5 \cdot 3P \quad P \cdot \frac{1}{1-\frac{1}{2}(1-P)} = 0.5 \quad \Rightarrow P = \frac{3}{7}.$$

Let (X_1, X_2, X_3) be a three-dimensional random vector,
where X_k , $k=1, 2, 3$ are independent geometric r.v.'s
with parameter $1/3$

- a) Calculate $P(X_1=1, X_2=2, X_3=3)$.
- b) Calculate the probability that exactly two X_k 's equals to 2.

What's the probability that it takes more than
5 flips to get a head? (Fair coin $p = \frac{1}{2}$)

$$\Pr(X > 5) = 1 - \Pr(X \leq 5)$$
$$= 1 - (1 - (1 - \frac{1}{2})^5) = \frac{1}{32}.$$

$$1) \quad x = \frac{1}{2} + \frac{2}{2^2} + \frac{3}{2^3} + \dots$$

$$\frac{1}{2}x = \frac{1}{2^2} + \frac{2}{2^3} + \dots$$

↓

$$\frac{1}{2}x = \frac{1}{2} + \frac{1}{2^2} + \frac{1}{2^3} + \dots$$

$$x = 2.$$

$$2) \quad w = \frac{1^2}{2} + \frac{2^2}{2^2} + \frac{3^2}{2^3} + \dots$$

$$\frac{1}{2}w = \frac{1^2}{2^2} + \frac{2^2}{2^3} + \dots$$

↓

$$\frac{1}{2}w = \frac{1^2}{2} + \frac{2^2 - 1^2}{2^2} + \frac{3^2 - 2^2}{2^3} + \dots$$

$$= \frac{1}{2} + \frac{(2+1)}{2^2} + \frac{(2+3)}{2^3} + \dots + \frac{2N-1}{2^N}$$

$$= 2 \cdot \sum_{N=1}^{\infty} \frac{N}{2^N} - \sum_{N=1}^{\infty} \frac{1}{2^N}$$

$$= (2 \cdot 2 - 1) = 3.$$

$$w = 6$$

→ 有个记忆之法。

$$\frac{1}{2} + \frac{1}{2^2} + \dots = 1$$

$\frac{1}{2} + \frac{2}{2^2} + \frac{3}{2^3} + \dots + \frac{N}{2^N} = 2$ 3部分上的自然数之
递增序列，在下面
指教序列的作用下，
作用被削弱。

$$3) \lim_{N \rightarrow \infty} \frac{a_n}{a_{n-1}} = \lim_{n \rightarrow \infty} \frac{N}{2^n} \cdot \frac{2^{n-1}}{N-1} = \frac{1}{2}$$

于是级数收敛。

$$\zeta = \frac{1}{2} + \frac{2}{2^2} + \frac{3}{2^3} + \dots = \sum_{N=1}^{\infty} \frac{N}{2^N} \quad ?$$

$$\sum \frac{N^2}{2^N} \quad ?$$

⋮

$$\sum \frac{N^5}{2^N} \quad ?$$

$$\zeta = \sum_{N=1}^{\infty} \frac{N^m}{2^N} \quad \text{converge?}, = ?$$

$0 < m < \infty$

$\in N$

Plot

$$\zeta \sim m$$

I have a bag containing nine ordinary coins,
and one double headed one; I remove a coin
and flip it three times. It comes up heads
each time. What's the probability that it's the
double header?

贝叶斯 Bayes 公式：

A_1 : 拿出的硬币是 double header

A_2 : single

B: 连掷三次都是 Heads.

$$P(A_1 | B) = \frac{P(B | A_1) \cdot P(A_1)}{P(B)}$$

拿出的硬币是 double
header 的概率, based
on 3 次都是 heads up.
是 prob 为 1.

- $P(A_1)$:
double header vs
single

- $P(B)$ 出现三次 heads vs single
今概公式加分类讨论:

$$P(B) = P(B | A_1)P(A_1) + P(B | A_2)P(A_2).$$

$$= 1 \cdot \frac{1}{10} + \frac{1}{8} \cdot \frac{9}{10} = \frac{17}{80}$$

于是 $P(A_1 | B) = \frac{\frac{1}{10}}{\frac{17}{80}} = \frac{8}{17}.$

Suppose we toss a fair coin, let N denote the number
of tosses, until we get a head (including the final toss).

What's $E(N)$ and $\text{Var}(N)$?

5: → 赤俗称之分类讨论,枚举.
 破币问题的状态,是全概率公式 → 枚举不过是遍历.

第一次得到H 第二次 第三次
 立 H TH TTH
 $\frac{1}{2}$ $\frac{1}{2} \cdot \frac{1}{2}$ $\frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2}$

a) $\bar{N} = \frac{1}{2} \cdot 1 + \frac{1}{2^2} \cdot 2 + \frac{1}{2^3} \cdot 3 + \cdots + \frac{1}{2^n} \cdot N + \cdots$

b) $\frac{1}{2} \bar{N} = \frac{1}{2^2} \cdot 1 + \frac{1}{2^3} \cdot 2 + \cdots$

a-b: $\frac{1}{2} \bar{N} = \frac{1}{2} + \frac{1}{2^2} + \cdots = \frac{\frac{1}{2}}{1 - \frac{1}{2}} = 1 \Rightarrow \bar{N} = 2$

Var(N) = $E(N^2) - [E(N)]^2$

$$E(N^2) = \sum_{N=1}^{\infty} \frac{N^2}{2^N} = 6.$$

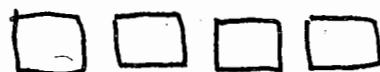
$$Var(N) = 6 - 2^2 = 2.$$

Let's play a game. There're 4 sealed boxes.

There's 100 pounds in one box and the others
are empty. A player can pay X to open a box
and take the content, as many times as they can.

Assuming this is fair game, what is the value of X .

先算平均需几次，才能打开含100个锁子。



第一次中： $\frac{1}{4}$

第二次中： $\frac{3}{4} \cdot \frac{1}{3}$

第三次中
第一次不中
令下的三个锁子，
中有一个

第四次中：

$$\frac{3}{4} \cdot \frac{2}{3} \cdot \frac{1}{2} = \frac{1}{4}$$

第一次
不中
第二次
不中
第三次
中

第四次中： $\frac{3}{4} \cdot \frac{2}{3} \cdot \frac{1}{2} \cdot 1$

$$\bar{n} = 1 \cdot \frac{1}{4} + 2 \cdot \frac{\frac{3}{4} \cdot \frac{1}{3}}{\frac{1}{4}} + 3 \cdot \frac{1}{4} + 4 \cdot \frac{1}{4} = 2.5$$

$$x = \frac{100}{\bar{n}} = 40$$

- We play a game:

I pick a number n from 1 to 100.

If you guess correctly, I pay you \$ n and zero otherwise.

How would you pay to play this game?

这个题的解答有些牵强，不过如下：

抽到那个人当然选择抽的数，越小越好，恨不能每次都是1。

假如若他每次都是1，那么被抓的概率加大。

于是乎，一个较好的 strategy 是，

Well, 这里便是牵强的地方。

抽中 k 的概率是 $\frac{1}{k}$ 。
正比于 $(\text{为什么不 是 } \frac{1}{k^2}?)$

于是抽中 k 的概率是：

$$P(k) = \frac{\frac{1}{k}}{\sum_{k=1}^{100} \frac{1}{k}} \quad (\rightarrow \text{归一化})$$

$$\sum_{k=1}^{100} \frac{1}{k} = 5.1 \text{ 左右}$$

精确值是一个很复杂的分数

$$P(k) \cdot \left(k \cdot \frac{1}{100} + 0 \cdot \frac{99}{100} \right)$$

抽中 k 的概率

猜中，于是得到 k。
猜不中。

全概率公式

$$X = \underbrace{\sum_{k=1}^{100} \frac{1}{k}}_{\text{归一系数} C} \cdot \left(k \cdot \frac{1}{100} + 0 \right) = \frac{1}{C} \approx 0.19$$

归一系数 C

$$C \approx 5.1$$

- Consider the following game. The player tosses a die once only. The payoff is 1 dollar for each dot on upturned face. Assuming a fair die, at what level should you set the ticket price of this game?

Soln:

很容易的一题.

色子随机抛出, 6个面, 则每个面出现之概率为 $\frac{1}{6}$.

每面上的点数为 $1, 2, 3, 4, 5, 6$. 于是 Expectation (or mean) 为

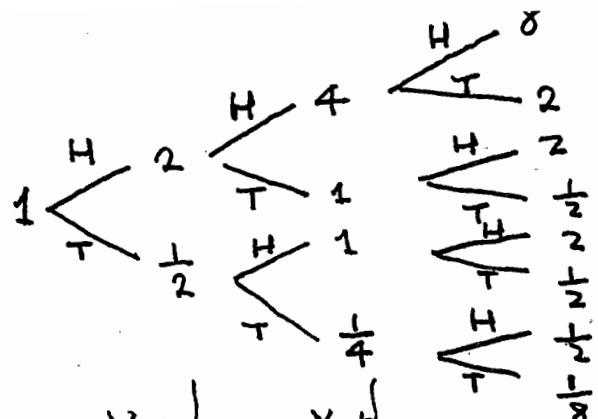
$$E(X) = \frac{1}{6}(1 + 2 + 3 + 4 + 5 + 6)$$

$$= \frac{21}{6} = 3.5$$

平均收益为3.5, 于是票价亦应设为3.5.

Suppose you have a fair coin, and you start with a dollar, and if you toss H, your position doubles, if you toss T, your position halves. What's the expected value of the money you have if you toss infinitely?

S₁:



第一次

$$2 \cdot \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{2}$$

$$= \frac{5}{4}$$

$$4 \cdot \frac{1}{4} + 2 \cdot \frac{1}{4} + \frac{1}{4} \cdot \frac{1}{4}$$

$$= \frac{25}{16} = \left(\frac{5}{4}\right)^2$$

第二次

$$\begin{aligned} & \xrightarrow{\text{第三次}} \left(8 + 6 + \frac{3}{2} + \frac{1}{8}\right) \cdot \frac{1}{8} \\ &= 1 + \frac{3}{4} + \frac{1}{64} + \frac{3}{16} \\ &= \frac{64 + 48 + 1 + 12}{64} = \frac{125}{64} = \left(\frac{5}{4}\right)^3 \end{aligned}$$

于是，牧羊归途中：第n次，是 $(\frac{5}{4})^n$.

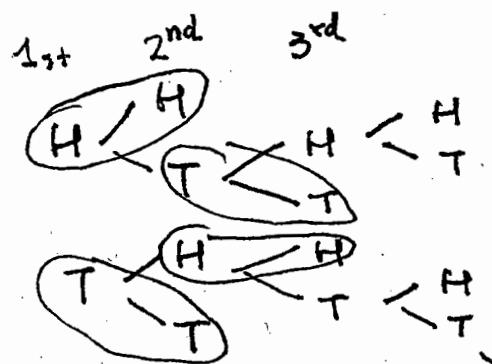
S₂:

第一次，从1变成 $\frac{5}{4}$.

第二次，不视为从 $\frac{5}{4}$ 的起点开始开地，于是为 $(\frac{5}{4}) \cdot \frac{5}{4} = \frac{25}{16}$

第三次为 $(\frac{5}{4})^3$

We play a game, with a fair coin, the game stops when either two heads (H) or tail (T) appear consecutively. What's the expected time until the game stops?



$$2 \cdot \left(\frac{1}{4} \times 2\right) + 3 \times \left(\frac{1}{8}\right) \times 2 + 4 \left(\frac{1}{16} \times 2\right)$$

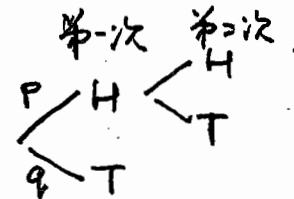
$$\bar{x} =$$

$$x = \sum_{n=2}^{\infty} n \cdot \frac{1}{2^{n-1}} = \sum_{n=2}^{\infty} \left(\frac{n-1}{2^{n-1}} + \frac{1}{2^{n-1}} \right) = \sum_{n=1}^{\infty} \left(\frac{n}{2^n} + \frac{1}{2^n} \right) = 2 + 1 = 3.$$

You toss a biased coin, what is the expected length of time until a head is tossed?
For two consecutive heads?

Soln:

1). 老幼木 = 叉树



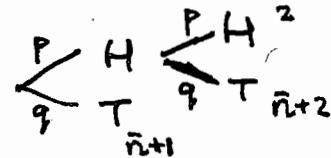
$$\bar{n} = 1 \cdot P + 2 \cdot qP + 3q^2P + \dots$$

$$q\bar{n} = 1 \cdot qP + 2q^2P + \dots$$

$$(1-q)\bar{n} = P \frac{1}{1-q} = 1$$

$$\bar{n} = \frac{1}{P}$$

2).



$$\bar{n} = P^2 2 + q(n+1) + qP(n+2)$$

$$\bar{n}(1 - q - Pq) = 2P^2 + q + 2Pq$$

$$\bar{n} = \frac{2P^2 + q + 2Pq}{1 - q - Pq}$$

$$= \frac{2P^2 + q + 2Pq}{q \cancel{P^2}}$$

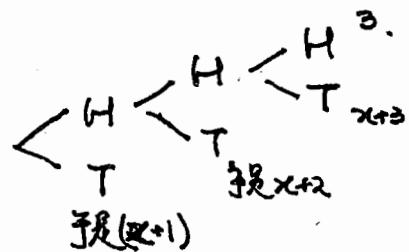
$$= 2 + \frac{q}{P^2} + \frac{2q}{P}$$

For a fair coin, what is the expected number
to get three heads in a row?

Soln:

老办法，一看到这种 consecutive 的题。

先来上二叉树时，抛三次再说。



故 Expected value 为

$$x = \frac{1}{2}(x+1) + \frac{1}{2^2}(x+2) + \frac{1}{2^3}(x+3) + \frac{1}{2^3} \cdot 3.$$

$$= \frac{7}{8}x + \left(\frac{1}{2} + \frac{1}{2} + \frac{3}{8} + \frac{3}{8}\right)$$

$$x = 14$$

平均需要抛 14 次才能得 3 consecutive heads.

~~问题~~ ?

$n \uparrow$ consecutive heads ?

$$x = \sum_{i=1}^n \frac{1}{2^i}(x+i) + \frac{1}{2^n} \cdot n$$

2. Suppose we play a game. I roll a die up to three times.

Each time I roll, you can either take the number showing as dollars, or roll again. What's your expected winning.

Soln:

- 这个题首先要明确 strategy:

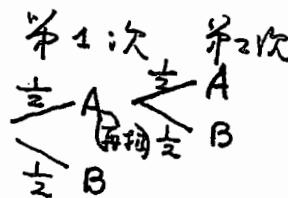
扔完头一个 dice, expected value: $\frac{1}{6}(1+2+3+4+5+6) = 3.5$.

故要是扔到 1, 2, 3, 则再扔一次;

扔到 4, 5, 6, 就从了.

于是 分成两组, $A = E(1, 2, 3) = 2$
 $B = E(4, 5, 6) = 5$.

↑
出现概率均为 $\frac{1}{2}$



$$\begin{aligned} E &= A \cdot \frac{1}{4} + B \cdot \frac{3}{4} \\ &= 2 \cdot \frac{1}{4} + 5 \cdot \frac{3}{4} \\ &= 4.25. \end{aligned}$$

- 31. 如果可以扔三次呢?

R: Based on 前两次的均值 4.25,

扔到 1, 2, 3, 4, 会重扔. 5, 6, R: take.

$$\left(\frac{1}{2}\right) \left(\frac{2}{3}\right) \left(\frac{1}{2}\right)$$

$$\left(\frac{1}{2}\right) \left(\frac{1}{3}\right) \left(\frac{5}{6}\right)$$

$$\left(\frac{1}{2}\right) \left(\frac{5}{6}\right)$$

I take an ordinary-looking coin out of my pocket
and flip it three times. Each time it is a head.

What do you think is the probability that the next
flip is also a head? What if I had flipped the
coin 100 times and each flip was a head?

→ 若是3次投 head, 加上 ordinary looking, 于是这组有3/4是双头 → fair coin.

只有3次

于是下次再投 head 的概率为 $\frac{1}{2}$.

→ 100次 vs 18. 基本不太可能是 fair, 于是, 33% to double-header, 于是
7次投头的 probability 是 1.

若先得 13 further info. 且 P(D) 只是 Double header 的概率为 $\frac{1}{2}$. (P)

$P(\text{Double Header} | 100 \text{ Heads})$

$$= \frac{P(100 \text{ Heads} | \text{Double}) \cdot P(\text{Double})}{P(100 \text{ Heads} | \text{single}) P(S) + P(100 \text{ Heads} | D) P(D)}$$

$$= \frac{1 \cdot P}{\frac{1}{2^{100}} \cdot (1-P) + 1 \cdot P} = \frac{P}{\frac{1}{2^{100}} (1-P) + P} \quad \rightarrow \text{无解作业第1.}$$

You throw a fair coin one million times,

What's the expected # of strings of 6 heads,
followed by 6 tails?

- 其实便是先从 ~~中~~ 中选出 1 个位置放 ~~H~~, 12 个为一组 VS sequence.

(1M-11)

H H H H H H T T T T T T
6 6

共有 1Million -11 个
vacency 可放.

- 送出这个 sequence VS slot + f_b , 再次便是 $6H, 6T$ VS seq.

$$\cancel{1 \times} \cdot \frac{1}{2^{12}} \times (1M-11)$$

- There are $(1\text{Million}-11)$ possible slot for the sequence to occur, in each one of these slots, the possibility is 2^{-12} .

If X, Y totally independent.

$$\mu_X \sigma_X \quad \mu_Y \sigma_Y$$

$$E(X+Y) \quad \text{Var}(X+Y)$$

$$E(X-Y) \quad \text{Var}(X-Y)$$

What if X, Y is not independent?

$$E(X+Y) = \mu_X + \mu_Y$$

$$E(X-Y) = \mu_X - \mu_Y \quad (\text{期望的性质})$$

✓ $\begin{cases} \text{Var}(X+Y) = \sigma_X^2 + \sigma_Y^2 \\ \text{Var}(X-Y) = \sigma_X^2 + \sigma_Y^2 \end{cases}$ 這個條件! $\text{Var}(-Y) = \text{Var}(Y)$

此二式成立條件是 X, Y independent.

$$\text{Var}(X+Y) = \text{Var}(X) + \text{Var}(Y) + 2 \text{Cor}(X, Y)$$

$$\text{Var}(X-Y) = \text{Var}(X) + \text{Var}(Y) - 2 \text{Cor}(X, Y).$$

设随机变量X与Y的概分布分别为

X	0	1
P	$\frac{1}{3}$	$\frac{2}{3}$

Y	-1	0	1
P	$\frac{1}{3}$	$\frac{1}{3}$	$\frac{1}{3}$

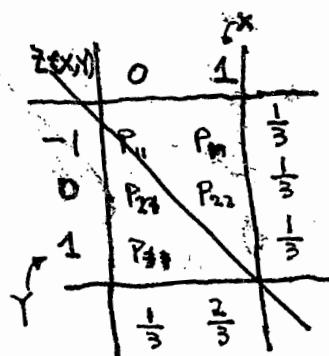
$$\text{且 } P\{X^2 = Y^2\} = 1$$

(I) 求二维随机变量(X, Y)之概率分布.

(II) 求Z=XY之概率分布.

(III) 求X与Y之相关系数.

首先整理清由两个一维构造=13. 边缘分布概率，即为 X 与 Y 本身之概率。



		Y			$P(X^2=Y^2)=1$
		-1	0	1	
X	0	0	b	0	$\frac{1}{3}$
	1	a	0	c	$\frac{2}{3}$
		$\frac{1}{3}$	$\frac{1}{3}$	$\frac{1}{3}$	

三个0因为

$X=0$ 时， Y 不可能取 ± 1

叠加

下表为边缘概率！

于是 $a=b=c=\frac{1}{3}$ ，很容易得到

记住 $\text{Cov}(X, Y) = \underbrace{\mathbb{E}XY}_{0} - \underbrace{\mathbb{E}X \cdot \mathbb{E}Y}_{\frac{2}{3} \cdot 0}$

$Z = XY$ 与 Y 独立同分布。 $\begin{array}{c|ccc} Z & -1 & 0 & 1 \\ \hline \frac{1}{3} & \frac{1}{3} & \frac{1}{3} & \frac{1}{3} \end{array}$

$$\rho_{XY} = \text{Corr}(X, Y) = \frac{\text{Cov}(X, Y)}{\sqrt{\text{Var}(X)} \cdot \sqrt{\text{Var}(Y)}} = 0$$

$$\mathbb{E}X = 0 \cdot \frac{1}{3} + 1 \cdot \frac{2}{3} = \frac{2}{3}$$

可能之取值 取该值之概率

$$\mathbb{E} \text{Var } X = \underbrace{\mathbb{E}X^2}_{0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{2}{3}} - \mathbb{E}X \cdot \mathbb{E}X = \frac{2}{3} - \frac{2}{3} \cdot \frac{2}{3} = \frac{2}{9}$$

$(\mathbb{E}X^2)^2$

$$0^2 \cdot \frac{1}{3} + 1^2 \cdot \frac{2}{3} = \frac{2}{3}$$

方差计算：

各个数据点离开 mean 的距离和，乘以概率。

$$\left[\frac{2}{3} \right]$$

$$\text{Var } X = (0 - \frac{2}{3})^2 \cdot \frac{1}{3} + (1 - \frac{2}{3})^2 \cdot \frac{2}{3} = \frac{6}{27} = \frac{2}{9}$$

Linear regression.

What is linear regression?

In statistics, linear regression is an approach for modeling the relationship between a scalar dependent variable y and one or more explanatory variables denoted x . The case of one explanatory variable is called simple linear regression.

What is an Random variable (r.v.)

Math

What's the difference between a r.v. & a common var like in $x+1=5$

代数系先中之变量.

$x+1=6$, $x \in \mathbb{R}$ take a random value.

If Random variable X , it can take many values, with some probability.

比如:

$Y = \text{sum of seven dice values (facing up)}$.

于是 Y 可以取 7 到 7×6 中的值, (with different p)

随机变量可以取多个值!!

$P(Y=42)$ 求 $Y=42$ 的概率,

求 "sum of seven dice values" $\neq 42$

定义为 Y 后, 你不再需要写这么长了.

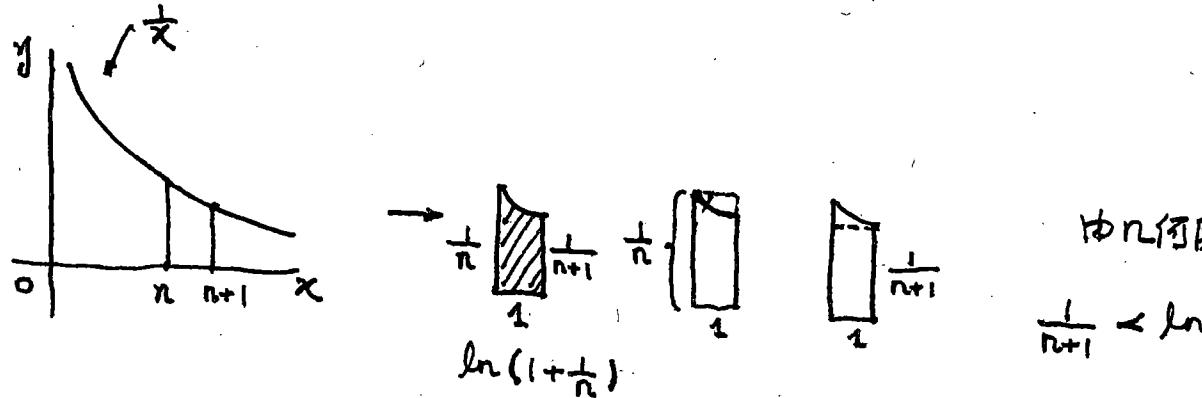
而是简单的多为 $Y=42$ 即可. 希望你很方便.

证明¹⁾ $\forall n \in N$, 有 $\frac{1}{n+1} < \ln(1 + \frac{1}{n}) < \frac{1}{n}$ 成立.

2) 令 $a_n = 1 + \frac{1}{2} + \dots + \frac{1}{n} - \ln n \quad (n=1, 2, 3, \dots)$

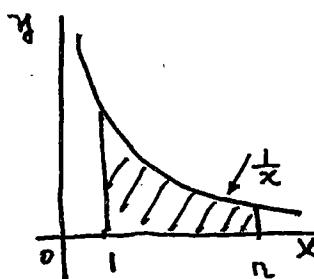
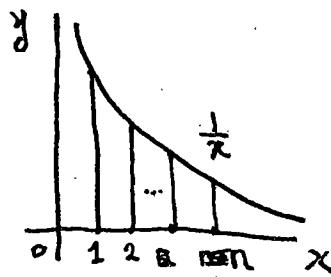
3) 证明数列 $\{a_n\}$ 收敛.

$$1). \text{ 设 } f(x) = \frac{1}{x} \quad \int_n^{n+1} \frac{1}{x} dx = \ln|x| \Big|_n^{n+1} = \ln\left(1 + \frac{1}{n}\right)$$

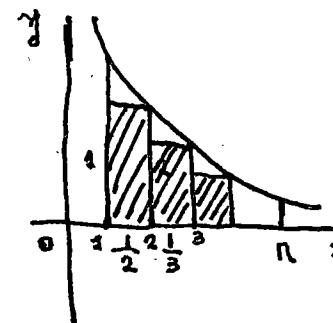


$$\frac{1}{n+1} < \ln\left(1 + \frac{1}{n}\right) < \frac{1}{n}$$

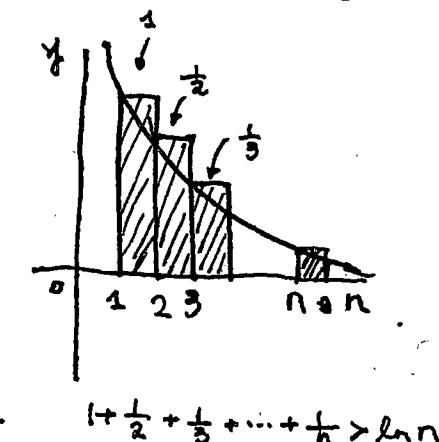
2). 证明 $a_n = 1 - \ln n$ 单调有界.



$$\int_1^n \frac{1}{x} dx = \ln n$$



$$\frac{1}{2} + \frac{1}{3} + \frac{1}{4} + \dots + \frac{1}{n} < \ln n$$



$$1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n} > \ln n$$

a_n

$$0 < 1 + \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{n} - \ln n < 1$$

$$a_n - a_{n-1} = \frac{1}{n} - [\ln n - \ln(n-1)]$$

$$a_{n+1} - a_n = \frac{1}{n+1} - \ln\left(1 + \frac{1}{n}\right) < 0$$

于是 a_n 单调递减且有下界, a_n 收敛.

find the first four ~~power~~ items
of the power series of $\arctan(2x)$

直接求解麻烦. 但 recall $\frac{d}{dx} \arctan(x) = \frac{1}{1+x^2}$

于是 $\frac{d}{dx}(\arctan 2x) = \frac{2}{1+4x^2} = 2 \left[(1 - 4x^2 + (4x^2)^2 - (4x^2)^3 + \dots) \right]$
 $= 2 - 8x^2 + 32x^4 - 128x^6 + \dots$

$\left(\frac{1}{1+x} = 1 - x + x^2 - x^3 \right)$

$$\arctan 2x = \int 2 - 8x^2 + 32x^4 - 128x^6 + \dots$$

$$= 2x - \frac{8}{3}x^3 + \frac{32}{5}x^5 - \frac{128}{7}x^7 + C.$$

$x=0$ 时

$$\arctan 2x = 0$$

于是为前4项.

$C=0$

求積分：

$$\int \ln x \, dx$$

$$\int x^2 e^x \, dx$$

$$\int e^x \cos x \, dx$$

$$\int f(x) g'(x) dx = f(x)g(x) - \int f'(x) g(x) dx.$$

何时用分部积分？ Clue is. 两个 function, g, f , 求导后一个并未复杂，一个简化。

$$x^2 \cdot e^x$$

$$\frac{d}{dx} \downarrow \quad \downarrow \\ 2x \quad e^x$$

(简化) (并未复杂)

$$\int \ln x dx = \ln x \cdot x - \int x d\ln x = \ln x \cdot x - \int x \cdot \frac{1}{x} dx = \ln x - x.$$

$$\text{注2: } d\ln x = \frac{1}{x} dx.$$

$$\int \frac{1}{x} dx = \ln|x|$$

$$\int x^2 e^x dx = \int x^2 de^x = x^2 e^x - \int e^x \cdot 2x dx$$

$$\left(\int e^x x dx = \int x de^x \right) = x e^x - \int e^x dx = x e^x - e^x.$$

$$= x^2 e^x - 2x e^x + e^x + C.$$

然后分步要加绝对值
保证 $\ln x$ 有意义。

$$\int e^x \cos x dx$$

$$= \int e^x \cos x dx = e^x \sin x + \int -\sin x e^x dx$$

$$\rightarrow \int e^x d\cos x = e^x \cos x - \int \cos x e^x dx$$

$$\text{于是 } \int e^x \cos x dx = e^x \cdot \frac{1}{2} (\cos x + \sin x) + C.$$

The average male drinks 2L of water when active outdoors (with a standard deviation of .7L)

You're planning a full day nature trip for 50 men and will bring 110 L of water.

What's the probability that you will run out?

key points 是. 不管最初是什麼 distribution.

i.i.d. 累加 F_0 . 之後 approach standard normal!

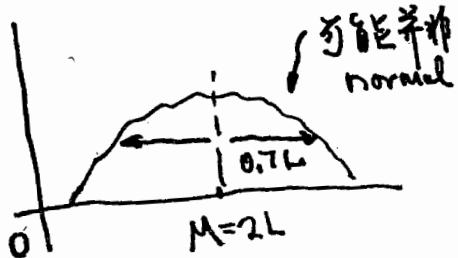
σ : standard deviation.

σ^2 : variance

(variance 是直徑

平均 n^2).

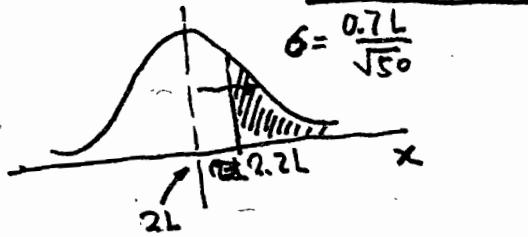
$$\mu = 2L, \sigma = \frac{0.7L}{\sqrt{50}} = 0.099$$



$$P(\text{rain out}) = P(\text{use more than } 110L)$$

$$= P(\text{average use more than } 2.2L)$$

$$= \int_{2.2}^{\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$



或者可以數一下 $n \sigma$.

$$\sigma = 0.099$$

$$\Delta x = 2.2 - 2L = 0.2$$

$$n = \frac{\Delta x}{\sigma} = \frac{0.2}{0.1} = 2$$

26

standard deviation

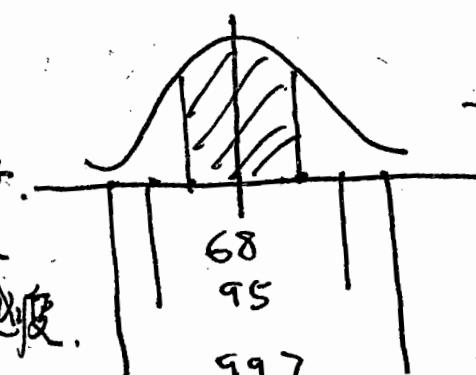


any distribution

i.i.d. 累加 n 次大約怎樣.

然後大約 \rightarrow normal

然後大約這樣.



3.41

If S_t follows a log-normal Brownian motion,
what process does the square of S_t follow?

若 S_t follows a log normal Brownian motion, R. I. T. Stochastic differential eqn.

$$dS_t = \mu S_t dt + \sigma S_t dW_t, \text{ 其中, } \mu > 0, \sigma > 0, W_t: \text{Brownian motion}$$

$$\text{Let } Y_t = S_t^2 \quad \downarrow \text{ito's lemma}$$

$$\begin{aligned} dY_t &= dS_t^2 = 2S_t dS_t + \frac{1}{2} 2(dS_t)^2 \\ &= 2S_t^2 (\mu dt + \sigma dW_t) + \sigma^2 S_t^2 dt \\ &= 2Y_t (\mu dt + \sigma dW_t) + \sigma^2 Y_t dt \\ &= Y_t (2\mu + \sigma^2) dt + Y_t \cdot 2\sigma dW_t \end{aligned}$$

$$\begin{aligned} (dS_t)^2 &= \underbrace{\mu^2 S_t^2 dt^2}_{\rightarrow 0} + 2\mu \sigma S_t^2 dW_t dt \\ &\quad + \sigma^2 S_t^2 dt \end{aligned}$$

于是 Y_t 仍然是 - log-normal, with new parameter.

新 drift $\approx \mu + \sigma^2$

diffusion $\approx 2\sigma$.

Suppose we are doing a random walk on the interval $[0, 1000]$, starting at 80. So with probability $\frac{1}{2}$, this number increases or decreases by one at each step. We stop when one of the boundaries of the interval is reached. What is the probability that this boundary will be 0?

Crucial observation: our location is a martingale.

↳ Expected position at any time in the future is our current position.

Let X_t be the value at time t . It's a martingale.

Let p be the probability of hitting 1000. P[hitting 0] $\leq p \leq 1-p$.

The expected value at stopping is:

$$1000p + 0 \cdot (1-p) = 80$$

$$\therefore p = \frac{8}{100}$$

The probability to hit 0 is thus: $1-p = \frac{92}{100}$.

→ Jump obeys $\sim N(0, 1)$

Everywhere continuous, everywhere not differentiable.

→ The increments

$$W_{t_2} - W_{t_1}, W_{t_3} - W_{t_2}, \dots, W_{t_m} - W_{t_{m-1}}$$

are independent, and the distribution of $W_{t_j} - W_{t_{j-1}}$ is given by

$$W_{t_j} - W_{t_{j-1}} \sim N(0, t_j - t_{j-1})$$

↓ ↓
mean of t_j Variance $\propto \Delta t$.

If $M(x)$ is the cumulative Gaussian function and

$X \sim N(0, 1)$, then what is $E[M(X)]$?

$$\int_{-\infty}^{\infty} e^{-x^2} dx$$

Explain change of measure. Give an example when change of measure is useful in Monte Carlo.

If X is $N(\mu, \sigma^2)$ and $\lambda > 0$, calculate the
values of $E(X^2)$ and $E(\exp(\lambda X))$

$$\rightarrow \text{Var}(x) = (\mathbb{E}(x^2) - [\mathbb{E}(x)]^2)$$

$$\mathbb{E}(x^2) = \sigma^2 + \mu^2$$

$$\begin{aligned}
 \rightarrow \mathbb{E}(\exp(\lambda x)) &= \int_{-\infty}^{+\infty} \exp(\lambda x) \cdot \text{p.d.f. } dx \\
 &= \int_{-\infty}^{+\infty} \exp(\lambda x) \frac{1}{\sqrt{2\pi}\sigma} \exp \frac{(x-\mu)^2}{2\sigma^2} dx \\
 &\quad \left. \begin{array}{l} \text{把 } \lambda x \text{ 提出来, 然后配方.} \\ \text{提 } \lambda. \end{array} \right\} \text{结果是 } z! \\
 &= \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} \exp \frac{x^2 - 2\mu x + \sigma^2 \lambda^2 x + \mu^2}{2\sigma^2} dx \\
 &= \int_{-\infty}^{+\infty} \frac{dx}{\sqrt{2\pi}\sigma} \exp \left(-\frac{x - (\mu + \sigma^2 \lambda)}{2\sigma^2} - \frac{\mu^2}{2\sigma^2} + \frac{(\mu + \sigma^2 \lambda)^2}{2\sigma^2} \right) \\
 &= \exp \left(\mu \lambda + \frac{\sigma^2 \lambda^2}{2} \right)
 \end{aligned}$$

Prove the covariance matrix
is positive definite.

1. There're actually several forms of central Limit theorem.

→ Weekly convergance result in prob.

2. Central limit theorem involves the average of
a sum of independent identically distributed R.V.
with finite second moment.

→ As the # of R.V. increases, the average approaches
a normal distribution with known parameters.

→ Remarkable point: this result is independent of
the starting distribution of the R.V.

3. More formally, X_1, X_2, \dots, X_n be a seq of n independent
and identically distributed R.V. with finite mean μ
and variance σ^2 . Let S_n denote the sum of R.V.

$$S_n = X_1 + X_2 + \dots + X_n, \text{ R.V.}$$

$T_n = \frac{S_n - \mu n}{\sigma \sqrt{n}}$ converges (in distribution to
the standard normal
distribution)

Two independent

Uniform R.V. on $[0, 1]$.

What's the p.d.f. of $X+Y$.

What is the distribution function and density
function of the k th order statistics?

Suppose that three assets A, B, C are such that the correlation coefficient of A & B is 0.9, and the correlation of B & C is 0.8. Is it possible for A & C to have correlation coefficient 0.1 ?

写出 covariance matrix.

$$M = \begin{bmatrix} 1 & 0.9 & 0.8 \\ 0.9 & 1 & 0.1 \\ 0.8 & 0.1 & 1 \end{bmatrix}$$

$$\|I\| = 1 > 0$$

$$\begin{vmatrix} 1 & 0.9 \\ 0.9 & 1 \end{vmatrix} = 0.19 > 0$$

$$|M| = -0.316 < 0$$

非正定. $\therefore M$ can't be covariance.

A plane has one hundred seats and there are exactly one hundred people boarding. They line up at the gate in exactly the same order ~~as~~ as the order of the seats, on the plane, from 1 to 100. However, the first person is Grandma who doesn't see well, so instead of taking the first seat on the plane, she picks a random seat.

Now the rule is, when any other person enters the plane, he or she will try to sit at his own seat.

Eventually If this is not possible, this person chooses one of the free seats randomly. Eventually, everyone boards the plane. What's the probabilities that the last person (number 100) will sit at his assigned seat #?

$\square_1 \square_2 \dots \underbrace{\square_{99}}_{\text{乘客}} \square_{100}$

#1 乘客 take \square_1 , 于是大家全对。#100 亦会 take 第 100 号座位。

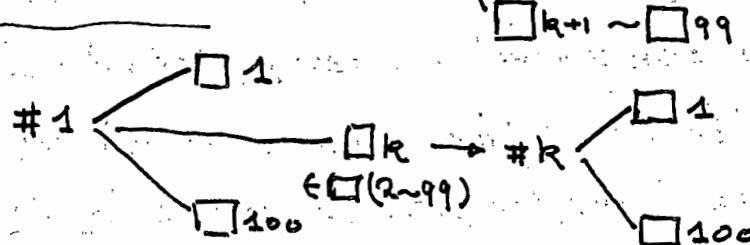
#1 take \square_{100} , 于是 #100 never take his own seat.

#1 take $\square_2 - \square_{99}$ 中的一个. 设为 k.

$\square_2 \sim \square_{k-1}$ 是对的, 划去不再考虑.

#k 剩余的 seats. 有 $\square_1, \square_{k+1} \dots \square_{100}$.

$\blacksquare \quad \#k$ \square_1 大家 $(k+1 \dots 100)$ 全对,
 \square_{100} #100 never takes his own seat.



把整个概率空间分成两份, 左: 最后一个乘客
自己座, 右: 放坐到自己座.

1. 第一个乘客坐 \square_1 , 于是为左空间,
第 $\square_1 \dots \square_{100}$, 不含 \square_1 .

\square_k , 往未及往下看.
 $k \in [2, 99]$

通过以上推得发现, 乘客 100,

只有两种可能, 或者坐在 \square_1 , 或者 \square_{100} .

于是概率为 $\frac{1}{2}$.

| 且达到两种状态
完全 symmetry 的情况.

2. #1 takes \square_k , 于是他在 $\square_1, \square_{k+1}, \dots$
中选择. 依然是左, 右空间一样概率.

3. 最多递归至乘客 $n-1$ 在 1 与 n 间选.

于是始终是 $\frac{1}{2}$.

Given an example of a distribution
with inf variance.

Suppose the random variable X has a probability density function $f_x(x)$.

What density function does $g(x)$ have?

$$P(X \leq x) = F_x(x) = \int_{-\infty}^x f_x(x) dx$$

Let $Y = g(x)$

$$\begin{aligned} \text{p.d.f. of } g(x) &= \frac{d}{dx} P(Y \leq x) \\ &= \frac{d}{dx} P(g(x) \leq x) \\ &= \frac{d}{dx} P(X \leq g^{-1}(x)) \\ &= \frac{d}{dx} F_x(g^{-1}(x)) \\ &= f_x(g^{-1}(x)) \cdot \frac{d}{dx} g^{-1}(x) \end{aligned}$$

Suppose x_1, x_2, \dots, x_n are independent identically distributed from $[0, 1]$,
and uniform on the interval. What's the expected value of the maximum?
What's the expected value of the difference between max & min?

Soln:

$$\text{c.d.f. of } \max\{x_i\} = P\{\max\{x_i\} \leq x\} = P\left\{\bigcap_{i=1}^n x_i \leq x\right\} = \prod_{i=1}^n P\{x_i \leq x\}$$

这一步很关键, 因为 $\max\{x_i\}$ 为取反.

$P\{x_i \leq x\}$ = c.d.f. of $U[0, 1]$

$$P\{x_i \leq x\} = \begin{cases} 0 & x < 0 \\ x^n & x \in [0, 1] \\ 1 & x > 1 \end{cases}$$

$$\frac{x-a}{b-a} = x$$

$$\text{c.d.f.} = \int_0^x p(y) dy \quad \rightarrow \quad p.d.f. = \frac{d}{dx} (\text{c.d.f.})$$

↓
p.d.f.

$$\text{p.d.f. of } \max\{x_i\} = \begin{cases} 0 & x < 0 \\ nx^{n-1} & 0 \leq x \leq 1 \\ 0 & x > 1 \end{cases}$$

$$\begin{aligned} E\{\max\}\} &= \int_0^\infty x \cdot p = \int_0^1 x \cdot x^{n-1} \cdot n \cdot dx \\ &= \frac{n}{n+1} \end{aligned}$$

$\hat{y} = \bar{x}$ 先求 E of $\min\{x_i\}$.

还是构造概率.

$$\begin{aligned} P\{\min\{x_i\} > x\} &= \prod_{i=1}^n P\{x_i > x\} = \prod_{i=1}^n [1 - P\{x_i \leq x\}] = (1-x)^n \\ 1 - P\{\min\{x_i\} \leq x\} & \end{aligned}$$

由 $\forall i, X_i \stackrel{i.i.d.}{\sim} U[0, 1]$, 从而 $P\{x_i \leq x\} = x$

$$\text{c.d.f. } F(M \leq x) = 1 - (1-x)^n$$

$$\begin{aligned} \text{p.d.f. } f(\min\{x_i\}) &= \frac{d}{dx} F = n(1-x)^{n-1} \quad \rightarrow \quad E M = \int_0^\infty x \cdot p \, dx = \int_0^1 x \cdot n(1-x)^{n-1} \, dx \\ &\stackrel{x=1-t}{=} n \int_0^1 (1-t)t^n \, dt = \frac{1}{n+1} \end{aligned}$$

$$\Rightarrow E(M-m) = \frac{n-1}{n+1}$$

What is the distribution function of a uniform
 $U(a, b)$ R.V. ?

E.d.f.

$U(a, b)$

$$\text{P.d.f. } x = \begin{cases} 0 & x \leq a, x = b \\ \frac{1}{b-a} & a < x \leq b \end{cases}$$

p.d.f. density function

c.d.f. distribution function

$$F(x) = \int_{-\infty}^x \frac{1}{b-a} dt = \int_a^x \frac{1}{b-a} dt$$

$$= \frac{1}{b-a} \left[t \right]_a^x = \frac{x-a}{b-a}$$

$$F(x) = \int_{-\infty}^x f(y) dy$$

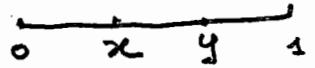
这叫做分布律。

一般用大F.

$$F(x) = \begin{cases} 0 & x < a \\ \frac{x-a}{b-a} & a \leq x \leq b \\ 1 & x > b \end{cases}$$

What's the distribution function of a uniform $U(a, b)$ random variable?

Let $\{x, y\}$ be uniformly distributed on $[0, 1]$ and
separate $[0, 1]$ into 3 pieces, what is the probability
that the 3 pieces can be constructed into a triangle?



不妨设 $y > x$

三段之长度: ~~x~~ , $y - x$,

$$y - x,$$

$$1 - y$$

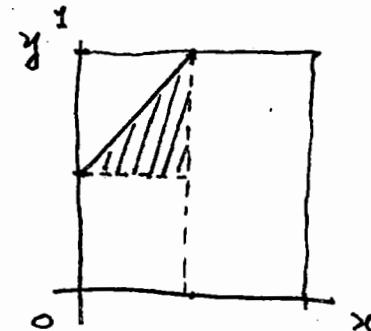
于是有: $x + y - x \geq 1 - y \Rightarrow y \geq \frac{1}{2}$

$$x + 1 - y = y - x \Rightarrow x - y \geq \frac{1}{2} \quad y - x \leq \frac{1}{2}$$

$$y - x + 1 - y \geq x \quad x \leq \frac{1}{2}$$

$$0 \leq y \leq 1$$

$$0 \leq x \leq 1$$



Romeo & Juliet have agreed to meet for a date sometime between 9pm & 10pm. Each of them randomly picks a time within this interval and shows up and wait for the other for fifteen minutes. What's the probability that they will actually meet?

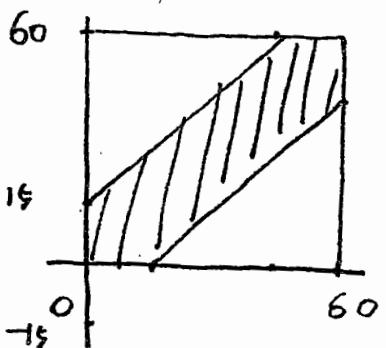
要求，均匀分布，通常看能否化成几何解决。

先转化为分钟，设 Romeo & Juliet 的到达时间分别为 x, y 。
P{ } 有：

$$\begin{aligned} 0 \leq x \leq 60 \\ 0 \leq y \leq 60 \end{aligned} \quad \rightarrow \text{(全集)}$$

in order to meet,

$$|x-y| \leq 15 \Rightarrow \left\{ \begin{array}{l} |y-x| \leq 15 \\ x-15 \leq y \leq x+15 \end{array} \right.$$



$$\frac{60^2 - 45^2}{60^2} = \frac{4^2 - 3^2}{4^2} = \frac{7}{16}$$

You have been captured and blindfolded by pirates,
then placed somewhere on a five-meter-long wooden plank.
Disorientated, each step you take is one meter long but
in a random direction — either towards the sharks
waiting at one end or eventually freedom at the other.

If x (integer) is the distance in meters you start from
the safe end, determine the probability of your
survival as a function of x .

$T = P(\text{Safe end} \mid \text{start at } x)$

設 $x=1$, 則由 $\frac{1}{2}$ 有 $\frac{1}{2}$ 到 2 , $\frac{1}{2}$ 到 3 . go to $f(0)$, $\frac{1}{2}$ go to $f(2)$

$$f(1) = \frac{1}{2} f(0) + \frac{1}{2} f(2)$$

$$f(0) = 1$$

$$f(2) = \frac{1}{2} f(1) + \frac{1}{2} f(3)$$

$$f(3) = \frac{1}{2} f(2) + \frac{1}{2} f(4)$$

$$f(4) = \frac{1}{2} f(5) + \frac{1}{2} f(3)$$

$$\downarrow f(5) = 0$$

$$\Rightarrow f(4) = \frac{1}{2} f(3)$$

$$f(3) = \frac{2}{3} f(2)$$

$$f(2) = \frac{3}{4} f(1)$$

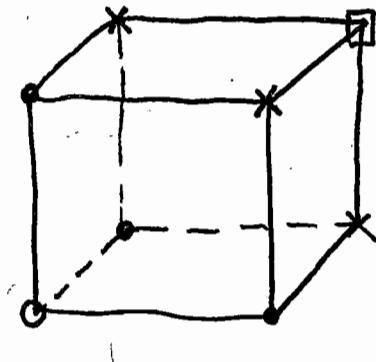
$$f(1) = \frac{4}{5} f(0) = \frac{4}{5}$$

$$\Rightarrow f(1) = \frac{4}{5}, f(2) = \frac{3}{5}, f(3) = \frac{2}{5}, f(4) = \frac{1}{5}$$

$$f(x) = \frac{5-x}{5}$$

Suppose we have an ant travelling on edges of a cube,
going from one vertex to another. The ant never stops,
and it takes him one minutes to go along one edge.

At every vertex, the ant randomly picks one of the
three available edges and starts going along that
edge. What's the expected number of minutes that
it will take the ant to return to that same vertex.



还是暴力打表，反而不求。

从距离为0的点...转回来后，需要的时间 $f(0)$ 。

1

$f(1)$

2

$f(2)$

3

$f(3)$

$$f(0) = f(0) + 1$$

$$\left\{ \begin{array}{l} f(1) = \frac{1}{3} + \frac{2}{3}(f(0) + 1) \\ \downarrow \frac{1}{3} 的概率回到起点。 \end{array} \right.$$

$\frac{2}{3}$ 的概率到距离为2的点。

$$f(2) = \frac{2}{3}(f(1) + 1) + \frac{1}{3}(f(3) + 1)$$

$$f(3) = f(2) + 1$$

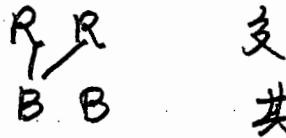
枚举求解。

$$f(1) = 7$$

$$f(0) = 8$$

A drawer contains 2 red socks and 2 black socks.
If you pull out 2 socks at random, what's the
probability they match.

这题容易上手.

Soln 1) 
交叉, Red, black 各一只的情况.
共有 4 种!

$$\frac{\#\{RR, BB\}}{\#\{RR, BB, RB, BR\}} = \frac{1}{3}$$

Soln 2) 先选一只, 比如是 Red. 则剩下 3 只中, 可以选 Black 的概率为

$$\downarrow \quad \frac{1}{2} \cdot \frac{1}{3} = \frac{1}{6}.$$

Black

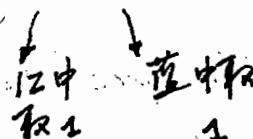
$$\frac{1}{2} \cdot \frac{1}{3}$$

$$\text{总概率为 } \underbrace{\frac{1}{2} \cdot \frac{1}{3} + \frac{1}{2} \cdot \frac{1}{3}}_{(\text{全概率公式})} = \frac{1}{3}$$

Soln 3). 总共的可能性, 4 选 2, 于是是 $C_4^2 = 6$

不一样一只, 于是是 $C_2^1 \cdot C_2^1 = 4$

于是配成对: $\frac{6-4}{6} = \frac{1}{3}$



8 What's the probability that the fourth business day of the month
is Thursday?

此题的**不要** Tricky 处在于, fourth business day!

↑↑

周末, 周日不算 weekday!
business day!

M Tu W Th

1 2 3 4

S M Tu W Th

1 2 3 4 5

7X为 $\frac{3}{7}$

↑仍为 fourth business day!

↓

Fa Sa M Tu W Th

1 2 3 4 5 6

→ A woman has two babies. One of them is a girl.
What's the probability that the other is a boy?

A family has two children whom you know at least one is a girl. What is the probability that both are girls? What if you know that the eldest is a girl? What if you know that they have a girl called Clarissa?

elder younger:

→ B.B. w.p. 0.25

G.G. —

B.G. —

G.B. —

→ At least, one is a girl: 0.75

Both are girls: 0.25

$P(\text{Both are girl} \mid \text{at least one is a girl})$

$$= \frac{0.25}{0.75} = \frac{1}{3}$$

(GG | GG, G.B., G.B.)

三女中能挑出一女

→ If you know eldest is a girl (G.G., G.B.)

Both are girls: G.G.

$$\frac{1}{2}$$

Suppose you are throwing a dart at a circular board.

What's your expected distance from the center?

Suppose you win a dollar, if you hit 10 times in

a row inside a radius of $R/2$, where R is the radius
of the board. You have to pay 10¢ for every try.

If you try 100 times, how much money would you
have lost/made in expectation?

Does your answer change, if you are a professional
and your probability of hitting inside $R/2$ is double
of hitting outside of $R/2$?

→ 平均距离.

$$\bar{r} = \int_0^R \frac{2\pi r dr}{\pi R^2} \cdot r$$



在 \textcircled{R} 上一圆周概率 $\times r$

然后求平均.

$$\bar{r} = \frac{2}{3}R$$

- Try 100次, have to pay $\frac{100 \times 100}{100 \text{ 元}} = 10$ 元.

hit 1次 $R/2$, $P = \frac{1}{4}$

$$10 \times \frac{1}{4} \text{ in a row: } P = \left(\frac{1}{4}\right)^{10} = \frac{1}{2^{20}}$$

$$\text{Try 100 次, 91 \uparrow slot, } 91 \cdot \frac{1}{2^{20}} = 0.86 \times 10^{-4}$$

very bad deal

- 中内圆概率系为外圆两倍, 则

$$\text{hit 10 times in a row: } P = \frac{3^{20}}{3^{20}}$$

$$\text{Try 100 次: } 91 \cdot \frac{3^{20}}{3^{20}} = 0.0274$$

still bad deal.

Suppose $2N$ teams participate in a championship. Once a team loses, it is out. Suppose you know the ranking of teams in advance and you know that a team with higher rank will always win the game. In the beginning, all teams are randomly assigned to play with each other. What would be the probability that in the final, the best teams are playing against the second best?

这题也 easy. 最好的肯定能进 Final.

所以其他队伍进, 这支了. second-best 1号 - 一半的队伍.

N 支队伍. out of $(2N-1)$

$$\text{于是 } P = \frac{N}{2N-1}$$

→ 你得把 best team 去掉
那个 1 去掉.

If I draw two cards from an ordinary deck
with replacement, what's the probability
they are both aces? Without replacement?

Easy one.

With replacement: $\frac{1}{13} \times \frac{1}{13} = \frac{1}{169}$

↓ 52个 slots 中，4个可能有 Ace.

W/O replacement

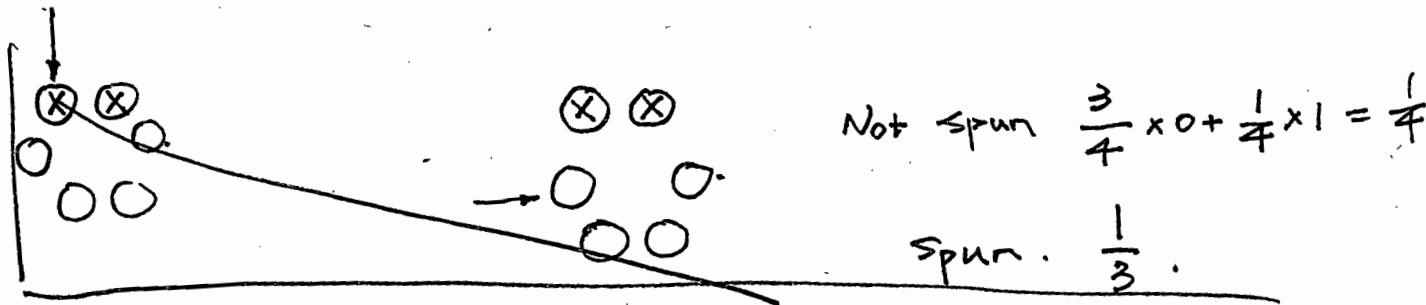
$$\frac{4}{52} \times \frac{3}{51} = \frac{1}{13} \times \frac{1}{17}$$

↓
拿下51张牌中，

3个是 Ace.

Consider a deck of 52 cards, ordered such that

$A > K > Q > J > \dots > 2$. I pick one first, then you pick one. What's the probability that my card is larger than yours?



Soln: 利用对称性 $P(\text{我得} \times) = P(\text{我得} \circ) = p$

$$P(\text{Draw}) = \frac{3}{51} = 1 - 2p$$

$$\Rightarrow p = \frac{8}{17}$$

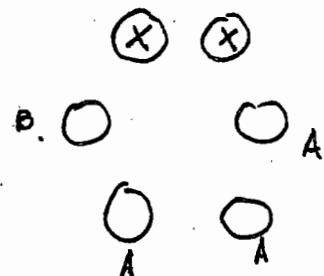
↓ 你第 -3K 亂送.

$\frac{8}{17} = 3K$ 和 $\frac{8}{17} - 3K$ 相等的概率是 同
从拿下的 51 张中, 取到 3K 和 $\frac{8}{17} - 3K$ 相等的 3 张.

You're playing Russian Roulette. There're precisely two bullets in neighbouring chambers of the six shooter revolver. The barrel is spun. The trigger is pulled and the gun does not fire. You're next, do you spin again or pull the trigger?

(12. Russian Roulette you're fire yourself!)

An easy question:



Not spin

因为第一枪没开火，故可能位于空枪的4处，
只有B处，会 fire yourself. \neq

Spin is 1/6.

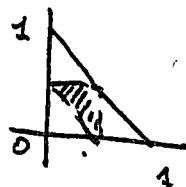
$$\frac{2}{6} = \frac{1}{3} \text{ fire yourself.}$$

Math.

Say I take a rubber band and randomly cut it
into three pieces. What's the probability that
one of the pieces has length greater than $\frac{1}{2}$
of the original circumference of the rubber band.

$$x + y + z = 1$$

$$\Leftrightarrow x + y < 1$$



$$x < \frac{1}{2}$$

$$y < \frac{1}{2}$$

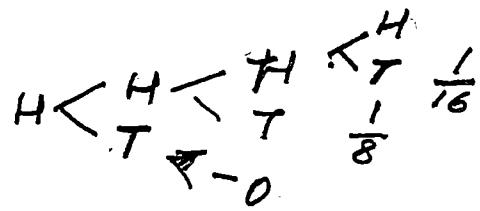
$$z < \frac{1}{2} \Rightarrow x + y > \frac{1}{2}$$

Flip a coin until either HHT or HTT appears.

Is one more likely to appear first?

If so, which one with what probability

HHT

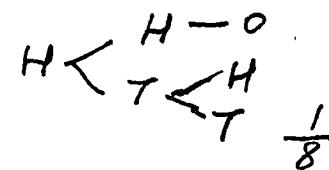


T-T-O

$$\frac{1}{8} + \frac{1}{16} + \frac{1}{32} + \dots$$

$$\frac{\frac{1}{8}}{1 - \frac{1}{2}} = \frac{1}{4}$$

A HTT



T-O

$$x = \frac{3}{8} + \frac{1}{4}(x+2) + \frac{1}{2}(x+1)$$

$$\frac{1}{4}x = \frac{3}{8} + \frac{1}{2} + \frac{1}{2} = \frac{11}{8}$$

$$x = \frac{11}{2}$$

$$P = \frac{1}{x} = \frac{2}{11}$$

$$78 - 12 = 26$$

$$67 + 31 = 98$$

+ 27

$$\frac{21 + 22 + 24 + 31 + 36}{36} = \underline{\underline{3 + \frac{13}{18}}}$$

$$\frac{161}{36}$$

x

	1	2	3	4	5	6	
1	1	2	3	4	5	6	$\frac{3.9}{6}$
2	2	2	3	4	5	6	$\frac{22}{6}$
3	3	3	3	4	5	6	$\frac{24}{6} = 4$
4	4	4	4	4	5	6	$\frac{27}{6}$
5	5	5	—	—	—	6	$\frac{31}{6}$
6	6	—	—	—	—	6	6

You have five coins. One is double-headed.

Pick one coin at random without looking and throw it 5 times.

Suppose the out are five heads, what is the probability
that the coin picked is double-headed one?

$$\begin{aligned}
 P(\text{Double headed} \mid 5 \text{ heads}) &= \frac{P(5 \text{ heads}) \cdot P(\text{two headed})}{P(5 \text{ heads}) \cdot P(\text{One headed}) + P(5 \text{ heads}) P(\text{two headed})} \\
 &= \frac{1 \cdot \frac{1}{5}}{\frac{1}{32} \cdot \frac{4}{5} + 1 \cdot \frac{1}{5}} \\
 &= \frac{8}{9}.
 \end{aligned}$$

You're playing a game where the player gets to draw number 1-100 out of hat, replace and redraw as many times as they want, with their final number being how many dollars they win from the game. Each "redraw" cost extra \$1.

How much would you charge someone to play this game?

假設最後一粒你拿到 x , 便停止.

也就是說 x 態度等於你 之後 的贏利期望.
下一粒.

$$E^* = -1 + \frac{x-1}{100} E + \frac{\sum_{t=x}^{100} t}{100}$$

$$2(101-x)E = (101-x)(100+x) - 200$$

$$\rightarrow E = \frac{(101-x)(100+x) - 200}{2(101-x)}$$

This is maximized by $101 - 10\sqrt{2} \doteq 86.86$.

Thus the best value for x is either 86 or 87.

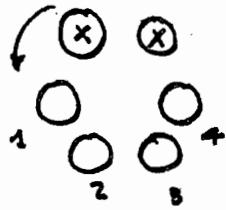
These values give $E \doteq 86.33$ or $E \doteq 86.37$.

So $x=87$ is better, and the value of the game is $1209/14$.

Russian roulette,

4 blank 2 bullets, all in a row.

If someone shoots a blank next to you,
would you take another shot or spin.



shoot blank,

1, 2, 3, 4 中选一个.

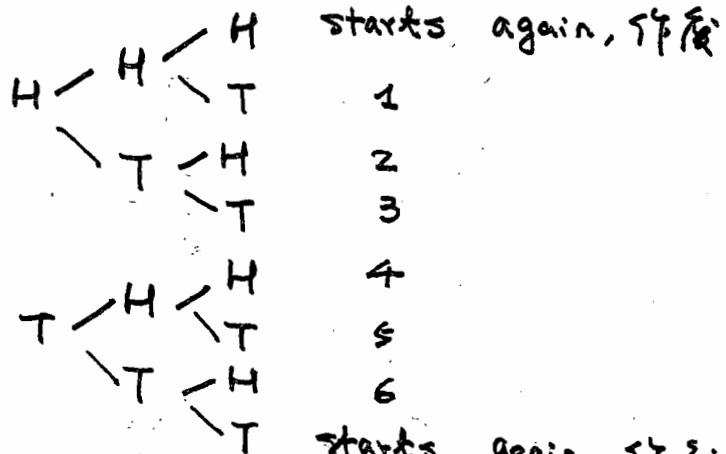
如果是 1, 再打公挂.

2, 3, 4 先打

不 spin, $\frac{3}{4}$ 几率 survive

spin, $\frac{2}{3}$ 几率 survive

故, 不 spin.



starts again, 57歲

1

2

3

4

5

6

starts again, 57歲

鲁羽A游戏，一副去掉大小王的扑克只有52张牌。

庄家把牌一排一张一张翻开，每翻开一张给你一块钱。

一直到翻开一个A为止（这张也给你一块钱）。问，

+块一次你不玩不玩。

假设有 X 张，翻出 A，
于是 X 是一个随机变量。

$$P(X=x) = \left(\frac{16}{17}\right)^{x-1} \left(\frac{1}{17}\right)^x \quad x=1, 2, \dots, 49$$

49, 50, 51, 52

然后求期望。

$$\mathbb{E}X = \sum_{x=1}^{49} x P(X=x)$$

计算很麻烦！

也可以这样想：

M 张红牌， N 张蓝牌，随机翻牌，第一次翻出蓝牌的期望长度。

可以想成：

红牌 M 张



N 张蓝牌插进来，求这一串之平均长度。

N 张蓝牌将 M 张红牌分成 $N+1$ 段，每段之平均长度是 $\frac{M}{N+1}$ 。

↑
单指红牌。

每段最后加一个空格。

平均长度加一。

故第一次翻出蓝牌之期望长度是：

$$1 + \frac{M}{N+1}$$

回到原问题，期望长度是： $1 + \frac{48}{5} = 10.6$

Math.

一个骰子6面分别是1到6。

问你平均要投多少次，才能让每个数字都投到过。

$$1 + \frac{1}{2} + \frac{1}{3} + \frac{1}{4} + \frac{1}{5} + \frac{1}{6}$$

$$= \frac{60+30+20+15+12+10}{60} = \frac{147}{60}$$

- 掷第一次可以得到一个数，第二次掷出不同数之概率是 $\frac{5}{6}$ ，第三次不同数之概率是 $\frac{4}{6}$ 。
 故掷出三个数之期望次数
 是 $\frac{6}{5}$ 。
 期望次数是 $\frac{6}{5}$ 。

类推可得：

$$1 + \frac{6}{5} + \frac{6}{4} + \frac{6}{3} + \frac{6}{2} + \frac{6}{1} = 14.7.$$

Suppose you have a perfectly round disk.

You put three legs on this disk to form a table.

Suppose the legs are perfectly perpendicular to the disk
and attached to the disk firmly.

What's the probability that the table will fall?

当三枚针在同一个半圆时，余例

于是回到保皮书问题， n 个点，随机落到圆周上，

问同时落到一个半圆上之概率是？

一共 N 个点，固定一个点，从这点开始，

顺逆时针各有一个半圆，其它 $N-1$ 个点

都落在同一个（比如顺时针）半圆上的概率是 $(\frac{1}{2})^{N-1}$.

总共 N 个点，所以总概率是

$$\frac{N}{2^{N-1}}$$

一个口袋有10个红色球，20个蓝色球，30个绿色球。

你随机地把球一个一个取出来。问红色球最先被拿走的概率？

12球最先被拿走，即10个12色球被取出时，
口袋内至少还有一个蓝色球和一个绿色球的概率。

如果把这个排列反过来，就是先看见12球前，
先看见蓝球和绿球。共有两种情况。

1. 先蓝后绿再12。

先蓝概率是 $\frac{1}{3}$ ，绿红两色中先绿再12之概率为 $\frac{3}{4}$ 。

(先蓝后其他蓝球不影响绿色之概率)。

$$\text{故第一种情况之概率是 } \frac{1}{3} \times \frac{3}{4} = \frac{1}{4}$$

2. 绿蓝而后再12。

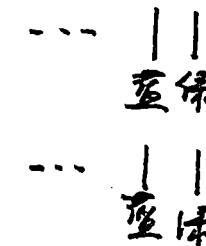
先绿 $\frac{2}{3}$ ，蓝12两色中先蓝再12之概率是 $\frac{2}{3}$

第二种情况之概率是：

$$\frac{1}{2} \times \frac{2}{3} = \frac{1}{3}$$

$$\text{故总概率是 } \frac{1}{3} + \frac{1}{4} = \frac{7}{12}.$$

12 红 绿
10 蓝 30



仔细看一下上面之解法可知，
红蓝绿数量不等，比例要以最简化。

We're playing a game, where we start at 0,
and flip a coin until someone wins.

Math

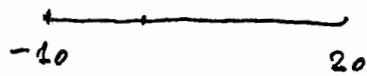
Every head that is flipped adds +1 to the number,
and every tail adds -1.

If the number reaches -10, the interviewer wins
20, you win

What's the probability the interviewer wins?

The first one is $\frac{1}{3}$, think of it like a random walk.

though you don't need



Interviewer wins $\in \mathbb{R} \nsubseteq \frac{1}{3}$.

Suppose you, player A, B are in a tennis tournament.

Player A has a better chance of beating you than player B does.

If you would like to maximize your chance of winning two games in a row,
which sequence you prefer?

ABA or BAB ?

If you want to win two game,
you must win the middle game,
so you should play the weaker player in the middle: ABA.

You have 4 coins, you throw them in the air, for
every head that lands, you get 1 dollar.

How much would you pay to play this game?

$$\rightarrow \frac{1}{16} \times 0 + \frac{4}{16} \times (1) + \frac{6}{16} \times (2) + \frac{4}{16} \times (3) + \frac{1}{16} \times (4) = \$2$$

→ 明显更优先解决！

binomial distribution (每投 coin 是 head 还是 tail 的概率)

$$E(X) = n p = 4 \times 0.5 = 2$$

When A predicts raining, the chance of raining is 60%.

When B — raining, _____ is also 60%.

If both A and B predict to rain (assume they're independent)

What's the chance of raining?

$P(A \text{ wrong})$

→ 两种情况：

- 孩子下而打喷嚏
- 孩子不下却打喷嚏。

若没有下雨， P 是 both A & B are wrong.

$$P(\text{not rain}) = P(\text{both A & B are wrong}) = P(A \text{ wrong}) P(B \text{ wrong})$$

A & B independent

$$= 0.4 \times 0.4 = 0.16.$$

3

$$P(\text{rain}) = 1 - P(\text{not rain}) = 0.84.$$

↓ 这是 A, B 完全 independent 的情况

↓ 若 A 与 B 正相关。

至少也要 0.6.

可是 - range (0.6, 0.84)

- 60% OR, confidence interval 只是而已。

把 a, b 看成一个 coin flip 而而已。

- 这是一个典型的 boost 问题,

multipredictor, how to sum up.

其实还是大于 0.6.

- If A and B are two suckers, then nobody knows the chance of raining. If A and B are two gods, then 60% for sure. generally 60% +/- delta%.

→ depends on the ability of the predictor.

$$P(A \cup B) = 1 - (A \cap B) = 1 - (1 - 0.6) * (1 - 0.6) = 0.84$$

Find positive integers x, y, z
and $x^x + y^y = z^z$.

MP.

Math problem.

- 简单一下都快问倒人快.

- 一道奥赛送命题, 猜几个特殊解, 而后证明大于那些数的都不行, 很搞毛了.

- $1^1 2^2 3^3 4^4 5^5 6^6$

1 4 27 256 3125 46656.

明显无解!

假设 $z > x > y > 0$

$$\begin{aligned} z^x &= (x+1)^{x+1} = \underbrace{(x+1) \cdot (x+1) \cdot \cdots \cdot (x+1)}_{x+1 \text{个}} \\ &= x^{x+1} + (x+1) \cdot x^x + \cdots > 2x^x \geq x^x + y^y \end{aligned}$$

P2.32.

某地区青年血压服从 $N(110, 122)$

求这一青年血压 x .

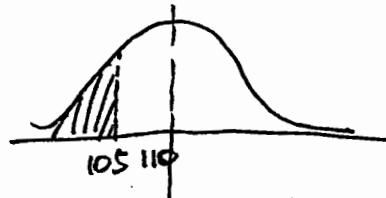
求 1) $P(x \leq 105)$,

$$P(100 \leq x \leq 120)$$

2) 确定 x 小于 x , 使

$$P(x > x) \leq 0.05$$

$$1) \frac{X-110}{\sqrt{122}} \sim N(0,1).$$



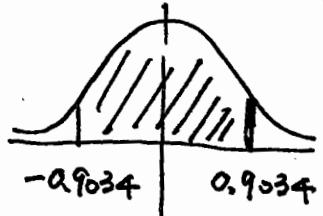
$$a. \frac{105-110}{\sqrt{122}} = -0.453.$$

$$\text{P}(\text{normcdf}(-0.453)) = 0.3254$$

$$\text{P}(X \leq 105) = 0.3254$$

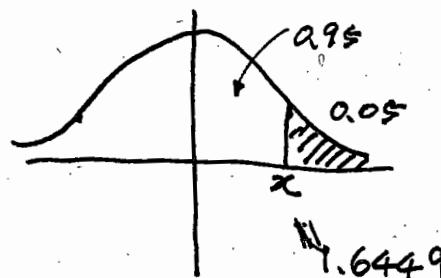
$$b. \frac{100-110}{\sqrt{122}} = -0.9034$$

$$\frac{120-110}{\sqrt{122}} = 0.9034$$



$$\text{P}(100 < X \leq 120) = 0.6347$$

2)



$$\text{* } \frac{x-110}{\sqrt{122}} = 1.6449 \quad x = 128.169$$

$$P_{2,30} \quad X \sim N(5, 4)$$

$$\text{f}^{\wedge} a \quad 1) \quad P(X < a) = 0.90$$

$$\Rightarrow P(|X - 5| > a) = 0.01$$

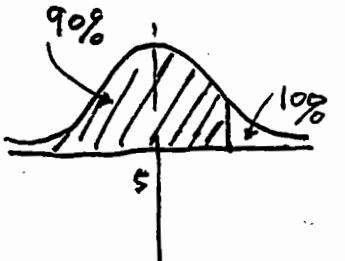
$$N(\mu, \sigma^2)$$

$$\begin{matrix} \uparrow \\ 5 \\ \uparrow \\ 4 \end{matrix}$$

$$\sigma = 2$$

$$P(X < a) = \int_{-\infty}^a \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx.$$

1)



$$X \sim N(\mu, \sigma^2)$$

$$\frac{X-\mu}{\sigma} \sim N(0,1) \quad \frac{X-5}{2} \sim N(0,1)$$

$$\text{norminv}(0.90) = 1.2816$$

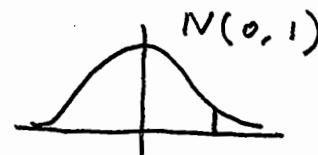
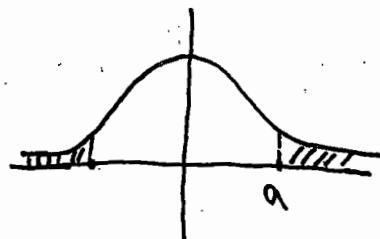
$$N(0,1) : x = 1.2816$$



$$\frac{X-5}{2} = 1.2816$$

$$X = 5 + 2.5632 = \underline{\underline{7.5632}}$$

$$2). P(|X-5| > a) = 0.01$$



$$\text{norminv}(0.995) = 2.5758$$

$$\frac{|X-5|}{2} \sim N(0,1)$$

$$P\left(\frac{|X-5|}{2} > 2.5758\right) = 0.01$$

5.14

$$a = \underline{\underline{4.5516}}$$

思路是通过 norminv
求标准正态的 quantile
(即 cdf 的 inv)

P2.24.

设某车间有同类型设备100台。

各台独立工作。

每台设备处于故障状态的概率为0.01。

又假定各设备的故障可由一名维修人员修理。

求此车间应配备多少维修人员。

为保证设备发生故障而不能及时修理的
概率小于0.01。

EX-112-11R
设各处故障之概率为 p , $P(X \geq k)$ 为 0.01

$$P(X \geq k) = \sum_{l=0}^{100} C_{100}^l p^l (1-p)^{100-l}$$
$$= 1 - \sum_{l=0}^k C_{100}^l p^l (1-p)^{100-l} = 0.01.$$

$$l=0 \quad C_{100}^0 p^0 (1-p)^{100} = 0.3660$$

$$l=1 \quad C_{100}^1 p^1 (1-p)^{99} = 0.3697$$

$$l=2 \quad C_{100}^2 p^2 (1-p)^{98} = 0.1849$$

$$l=3 \quad C_{100}^3 p^3 (1-p)^{97} = 0.0610$$

$$l=4 \quad C_{100}^4 p^4 (1-p)^{96} = 0.0149$$

$$k=4.41$$

$$P(X \geq 4) = 1 - P(X=0,1,2,3,4) = 0.0035 < 0.01$$

P_{2.22}

设某类电子元件寿命的分布密度为：

$$f(x) = \begin{cases} \frac{\alpha}{x^2} & x > 100 \text{ 小时} \\ 0 & x \leq 100 \end{cases}$$

1). $\alpha = ?$

2). 一台设备中要使用三个这种元件。

问在开始运行 150 小时中，三个元件中
至少有一个要替换之概率为？

3). 开始使用后三个元件今至替换之概率。

$$4) \int_{-\infty}^{+\infty} f(x) dx = 1$$

$$\int_{100}^{+\infty} \frac{a}{x^2} dx = 1$$

$$1 = a \cdot \left(-\frac{1}{x}\right) \Big|_{100}^{+\infty}$$

$$\frac{1}{a} = \frac{1}{100} - 0 \Rightarrow a = 100$$

单个是指数分布
三个是二项分布

$$2). P(X \leq 150) = \int_{-\infty}^{150} f(x) dx$$

$$= \int_{100}^{150} \frac{100}{x^2} dx$$

$$= 100 \cdot \frac{1}{x} \Big|_{150}^{100} = 100 \left(\frac{1}{100} - \frac{1}{150} \right) = 1 - \frac{2}{3} = \frac{1}{3}$$

$$P(\text{至少有1个}) = 1 - P(\text{全没有}) = 1 - \left(\frac{2}{3}\right)^3 = \frac{27-8}{27} = \frac{19}{27}$$

$$3) \text{全3个} \quad P(3个) = \frac{1}{3} \times \frac{1}{3} \times \frac{1}{3} = \frac{1}{27}$$

P_{2.16}

一电话交换台每分钟收到呼喚之次数
服从于参数为 λ 的泊松分布. 求

- ① 每分钟恰有 8 次呼喚之概率.
- ② 每分钟呼喚次数大于 10 之概率.

$$P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

$$\lambda=4, k=8$$

$$P(X=8) = e^{-4} \frac{4^8}{8!} = 0.0298.$$

$$P(X>10) = 1 - P(X=1, 2, 3, \dots, 10)$$

P_{2,15} — 一辆小汽车行驶，每天大量汽车通过，
每

行驶辆车在一天车时数内事故率为 0.0001.

在某时段内有 1000 辆汽车通过。

问事故次数不小于 2 次概率。

事故次数服从泊松分布

$$P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

1000辆汽车

$$\lambda = 0.0001 \times 1000 = 0.1$$

$$P(X=0) = e^{-0.1} \frac{0.1^0}{0!} = e^{-0.1}$$

$$P(X=1) = e^{-0.1} \frac{0.1^1}{1!} = 0.1 e^{-0.1}$$

~~P(x>2)~~ $P(\text{不小于} z) = 1 - P(\text{不大于} z)$

$$= 1 - P(X=0) - P(X=1)$$

$$= 1 - e^{-0.1} (1 + 0.1)$$

$$= \cancel{0.1953} \\ 0.0047$$

甲乙二人投篮，命中概率为 0.6, 0.7. 今各投 3 次。

求

1) 二人投中概率之和

2) 甲比乙投中次数多之概率

$$\begin{aligned}1. P[\bar{Y}] &= P(Y_0, Z_0) + P(Y_1, Z_1) + P(Y_2, Z_2) + P(Y_3, Z_3) \\&= 0.4^3 \cdot 0.3^3 + C_3^1 0.6 \cdot 0.4^2 \cdot C_3^1 0.7 \cdot 0.3^2 + C_3^2 0.6^2 \cdot 0.4 \cdot C_3^2 0.7^2 \cdot 0.3 + 0.6^3 \cdot 0.7^3 \\&= 0.3208\end{aligned}$$

$$\begin{aligned}2. P[Y] &= P(Y_1, Z_0) + P(Y_2, Z_0) + P(Y_3, Z_0) \\&\quad + P(Y_2, Z_1) + P(Y_3, Z_1) \\&\quad + P(Y_3, Z_2)\end{aligned}$$

三人能各得一份筹码。

各人能得出概率为 $\frac{1}{5}$, $\frac{1}{3}$, $\frac{1}{4}$.

问三人中至少一人能将此筹码得出之概率。

$$P(\text{至少一人}) = 1 - P(\text{一人没有}) \\ = 1 - \frac{4}{5} \times \frac{2}{3} \times \frac{3}{4} = \frac{3}{5}.$$

第 P1.26.

甲乙两盒放棋子

甲 60枚，40短

乙 20枚，10短

任取一盒，从中任取一枚。

发现是长棋子。

求此枚来自甲盒概率。

$$P(\text{亲自甲盒} | \text{长中快检}) = \frac{P(\text{甲} \& \text{快})}{P(\text{长中快检})} = \frac{\frac{1}{2} \times \frac{3}{5}}{\frac{1}{2} \times \frac{3}{5} + \frac{1}{2} \times \frac{2}{3}} = \frac{\frac{3}{10}}{\frac{3}{10} + \frac{1}{3}} = \frac{9}{19}$$

$\begin{matrix} \text{甲} \\ \text{长} \end{matrix}$ $\begin{matrix} \text{快} \\ \text{中快检} \end{matrix}$ $\begin{matrix} 1 \\ \text{乙} \end{matrix}$ $\begin{matrix} 1 \\ \text{中快检} \end{matrix}$

高
1.25

已知男人中 5% 色盲.

女人中 0.25%.

今从男女人数相等之人群中随机选一人.

恰好是色盲. 求此人男性的概率多少.

$$P(\text{男} | \text{色音}) = \frac{P(\text{男} \& \text{色音})}{P(\text{色音})} = \frac{5\% \times \frac{1}{2}}{P(\text{色音} | \text{男}) P(\text{男}) + P(\text{色音} | \text{女}) P(\text{女})}$$
$$= \frac{5\% \times \frac{1}{2}}{5\% \times \frac{1}{2} + 0.25\% \times \frac{1}{2}} = \frac{20}{21}$$

P_高
1.24

甲乙两袋.

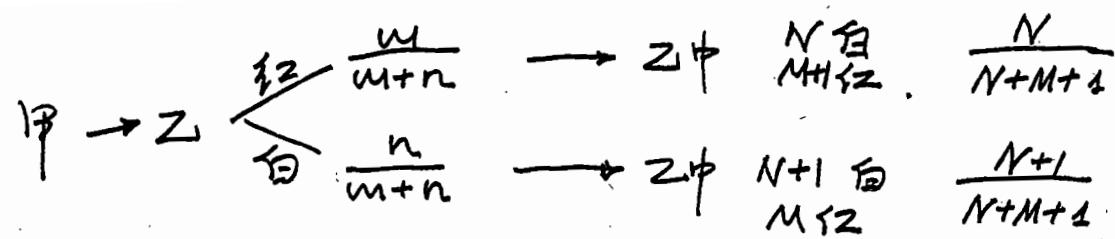
甲袋中 N 只白球, M 只红球

乙袋中 N 只白球, M 只红球

今先从甲袋中取一只球入乙.

再从乙中取一?

问可取到白球概率.



$$P_{tot} = \frac{mN + n(N+1)}{(m+n)(M+N+1)}$$

2. 袋中 10 粒，9 白 1 红。

10 人依次于袋中各取一粒。

每人取完后不再放回。

问：第一人，第二人，直至最后一个取得红球者为多少。

约 1/10

与你取之顺序无关

P_{1.18}

掷两骰子，已知两点之和为7.

求其中一枚为1点之概率.

(1, 6), (2, 5), (3, 4), (4, 3) (5, 2) (6, 1)

$\frac{1}{3}$.

P.17

基油漆公司发出17桶漆，其中白10桶，黑4桶，123桶。

物流过程中所有标签脱落。

发货人随已将这些油漆发给顾客。

问一个退货4桶白，3黑，2红之顾客。

核所返颜色得货之损坏为多少。

$$\frac{C_{10}^4 \times C_7^3 \times C_3^2}{C_{17}^9}$$

1.15

将3个球随机放入4个杯子.

第

求杯中球的最大数分别为1, 2, 3之概率

最大数为3，三只放入一个杯子。

$$\frac{1}{4} \cdot \frac{1}{4} \cdot \frac{1}{4} = \frac{1}{64}$$

↓
第一只球
4个杯子中选一个。

最大数为2。

$$C_3^2 \cdot 4^1 = 3 \times 4 \times 3$$

①两只球
不在一只杯。 两只球在
哪只杯。 全球在
哪只杯。
哪只杯。

总选择 4^3 ， $P = \frac{3 \times 4 \times 3}{64} = \frac{9}{16}$

最大数1。

大球都在不同杯。

$$\frac{4 \times 3 \times 2}{4 \times 4 \times 4} = \frac{3}{8}$$

P_{1.12}

有 1500 个产品中有 400 个次品，
1100 个正品。现从中取 200 个，

- 1) 求恰有 90 个次品的概率
- 2) 求至少有 2 个次品概率。

$$1972-2., 次品概率 p = \frac{400}{15} = \frac{4}{15},$$

$$\text{正品. } q = 1 - p = \frac{11}{15}.$$

含有 90 次品

$$C_{200}^{90} p^{90} q^{110}.$$

可能不对.

$$\checkmark \frac{C_{400}^{90} C_{1100}^{110}}{C_{1500}^{200}}$$

$$\text{至少有两个次品} = 1 - 0 \text{ 次品} - 1 \text{ 次品}.$$

$$= 1 - q^{200} - C_{200}^1 p^1 q^{199}.$$

葛金榮

P1.10

房间里有10个人，分别佩戴1号到10号胸章。

你选3人，纪录其胸章，~~并~~

1) 最小号码为5号胸章。

2). 最大号码为5号胸章。

2). 任取3个，先作阶配制

$$\begin{aligned} C_{10}^3 &= \frac{10 \times 9 \times 8}{3 \times 2 \times 1} \\ &= \frac{720}{6} = 120. \end{aligned}$$

(注: 不是 $10 \times 9 \times 8$.)

这样子虽然有重复, 是~~单~~次.

A B C
B A C
是同一组.

1) 最小号码为5.

$$\begin{array}{c} C_5^1 \quad 1 \quad C_5^2 \\ \diagdown \quad \uparrow \quad \uparrow \\ \text{号码为5} \quad 6-10 \\ \text{中选出5个?} \end{array}$$

6 7 8 9 10

$$\begin{array}{l} \text{最小号码为5} \\ P = \frac{1}{120} = \cancel{\frac{1}{12}} = 25\% \\ \frac{1}{12} \end{array}$$

2) 最大号码为5.

$$\begin{array}{c} C_5^1 \quad 1 \quad C_4^2 \\ \diagup \quad \uparrow \quad \uparrow \\ \text{号码为5} \quad 1-4 \\ \boxed{\begin{array}{l} \text{不重要} \\ \text{对次序 does not} \\ \text{matter} \end{array}} \\ \text{中选出5个?} \end{array}$$

$$P = \frac{1}{120} = \cancel{\frac{1}{120}} = 25\%$$

Chapter 4 Probability Theory, soln.

1

- Coin toss game.

A toss n

B toss n.

Card game

Probability

A bigger x

equal $1-2x$

smaller x

$$\begin{matrix} \text{两张一样} \\ \downarrow \\ P(\text{equal}) = \frac{1}{17} \end{matrix}$$

A先随便选一张

在剩余的51张中，B抽到

与A同样的可能性是3张

所以是

$$\frac{3}{51} = \frac{1}{17}$$

1. Let X, Y be r.v. with joint p.d.f. $f_{X,Y}(x,y)$,¹⁾ show
that pdf of $X+Y$

$$f_{X+Y}(z) = \int_{-\infty}^{+\infty} f_{X,Y}(x, z-x) dx$$

2) if X, Y are independent,

write pdf of $X+Y$ in terms of $f_X(x)$, $f_Y(y)$

Marginal probability density functions.

Soln:

$$1) F_Z(z) = P(Z \leq z) = P(X+Y \leq z)$$

$$\begin{aligned} &= \iint_{x+y \leq z} f_{X,Y}(x,y) dx dy \\ &= \int_{-\infty}^{+\infty} dx \int_{-\infty}^{z-x} dy f(x,y) \end{aligned}$$

) = 重积分
over region $x+y \leq z$

$$\begin{aligned} f_Z(z) &= \frac{d}{dz} F_Z(z) = \int_{-\infty}^{+\infty} dx \left[\frac{d}{dz} \underbrace{\int_{-\infty}^{z-x} dy f(x,y)}_{\text{支限重积分}} \right] \\ &= \int_{-\infty}^{+\infty} dx f_{X,Y}(x, z-x) \end{aligned}$$

$$2). 若 X, Y 独立, 则 f_{X,Y}(x,y) = f_X(x) f_Y(y)$$

$$f_Z(z) = \int_{-\infty}^{+\infty} dx \left[\frac{d}{dz} \underbrace{\int_{-\infty}^{z-x} dy f_X(x) f_Y(y)}_{\text{独立且为常数}} \right]$$

$$\begin{aligned} &= \int_{-\infty}^{+\infty} f_X(x) dx \cdot \left[\frac{d}{dz} \int_{-\infty}^{z-x} dy f_Y(y) \right] \\ &= \int_{-\infty}^{+\infty} dx f_X(x) f_Y(z-x) \end{aligned}$$

对称性

↓ 令 $u=x$, 则 $u=z-u$ 都是 dummy var

$$= \int_{-\infty}^{+\infty} f_X(u) f_Y(z-u) du$$

← 其实是 $f_X(x) f_Y(y)$
之差分

$\xrightarrow{\text{对称性}}$

$$f_Z(z) = \int_{-\infty}^{+\infty} f_Y(u) f_X(z-u) du$$

as well.

2. Prove the following version of Chebyshov's inequality:

Let r.v. X be such that

$$E(|X - EX|^k) < \infty \text{ for some fixed } k \geq 1,$$

then it holds for all $t > 0$

$$P(|X - EX| \geq t) \leq \frac{E(|X - EX|^k)}{t^k}$$

Soln:

$$\begin{aligned} P(|X - \mu| \geq t) &= \int_{|x-\mu| \geq t} dF_x(x) \leq \int_{|x-\mu| \geq t} \frac{|x-\mu|^k}{t^k} dF_x(x) \leq \int_{|x-\mu| \geq t} \frac{|x-\mu|^k}{t^k} dF_x(x) \\ &\quad + \int_{|x-\mu| < t} \frac{|x-\mu|^k}{t^k} dF_x(x) \\ &\quad \text{由 } |x-\mu| \geq t > 0 \\ &\quad \text{及 } \frac{|x-\mu|}{t} \text{ 是 } x \geq t \text{ 时的数.} \end{aligned}$$

3. X and Y are two indep. r.v., they have Poisson dist.

with parameters $\lambda > 0$, and $\mu > 0$, respectively.

Show $X+Y$ is also distributed poission with parameter $\lambda+\mu$.

sohn:

Let $Z = X + Y$, $Z = k$ if $X + Y = k$. 考以下事件之并

$$\{Z = k\} = \left\{ \begin{array}{l} X=0 \\ Y=k \end{array} \right\} \cup \left\{ \begin{array}{l} X=1 \\ Y=k-1 \end{array} \right\} \cup \dots \cup \left\{ \begin{array}{l} X=k \\ Y=0 \end{array} \right\}$$

$$P(Z = k) = \sum_{i=0}^k P(X=i \text{ & } Y=k-i)$$

$\downarrow X, Y$ 独立

$$= \sum_{i=0}^k P(X=i) \cdot P(Y=k-i)$$

$$= \sum_{i=0}^k \frac{\lambda^i}{i!} e^{-\lambda} \cdot \frac{\mu^{k-i}}{(k-i)!} e^{-\mu}$$

$$= \frac{e^{-(\lambda+\mu)}}{k!} \underbrace{\sum_{i=0}^k \frac{\lambda^i}{i!} \frac{\mu^{k-i}}{(k-i)!}}_{III} \cdot k!$$

$\swarrow X, Y$ 相互独立.

$$P(X=i) = e^{-\lambda} \frac{\lambda^i}{i!}$$

$$\sum_{i=0}^k C_k^i \lambda^i \mu^{k-i}$$

$\underbrace{\quad}_{III} = 2$ 项定理

$$(\lambda+\mu)^k$$

$$= (\lambda+\mu)^k \frac{e^{-(\lambda+\mu)}}{k!}$$

∴ $Z = (X+Y) \sim \text{Poisson}(\lambda+\mu)$

4. Let X have normal dist. $N(\mu, \sigma^2)$, Let $Y = e^X$,

compute the moment of $E(Y^k)$ for an arbitrary $k \in \mathbb{N}$.

Soln:

by definition, mgf

$$\phi(t) = \phi_x(t) = E(e^{tx}) \xrightarrow{\text{本題 } Y=e^X} E(Y^t)$$

令 $t=k$, \bar{x} 是 normal dist. 由 $\phi(t) = \exp(\mu t + \frac{\sigma^2 t^2}{2})$

$$E(Y^k) = \phi(k) = \exp\left(\mu k + \frac{k^2 \sigma^2}{2}\right)$$

5.

X_1, \dots, X_n are i.i.d. r.v. with $E(X_i) = \mu$, $\text{Var}(X_i) = \sigma^2 < \infty$

Compute the expected value of the following r.v.

$$\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2$$

$$\downarrow \equiv \frac{1}{n} \sum_{i=1}^n X_i$$

Soln:

$$\mathbb{E} \left[\frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 \right]$$

Expectation
of $\frac{1}{n} \sum_{i=1}^n$

$$= \frac{1}{n} \sum_{i=1}^n \mathbb{E} (X_i - \bar{X}_n)^2$$

$$\downarrow \{ EY_i = X_i - \bar{X}_n \}$$

$$\mathbb{E} Y_i^2 = DY_i + (\mathbb{E} Y_i)^2$$



$$\mathbb{E} Y_i = \mathbb{E}(X_i - \bar{X})$$

$$DY_i = D(X_i - \bar{X})$$

$\downarrow \{ EY_i = X_i - \bar{X} \}$

$$= EX_i - E(\bar{X})$$

$$= D \left[X_i - \frac{1}{n} X_i - \left(\frac{1}{n} X_1 + \dots + \frac{1}{n} X_n \right) \right] = \mu - \frac{n\mu}{n} = 0$$

$$= \left(\frac{n-1}{n} \right)^2 DX_i + \frac{1}{n^2} (DX_1 + DX_2 + \dots + DX_n)$$

$\uparrow \{ EX_i \}$

$$= \left(\frac{n-1}{n} \right)^2 \sigma^2 + \frac{1}{n^2} (n-1) \sigma^2$$

$$= \frac{n(n-1)}{n^2} \sigma^2$$

$$= \frac{n-1}{n} \sigma^2$$

于是 $\mathbb{E} Y_i^2 = \frac{n-1}{n} \sigma^2$

Soln

$$= \frac{1}{n} \sum_{i=1}^n \frac{n-1}{n} \sigma^2 = \frac{n-1}{n} \sigma^2$$

一本 500 页的书中共有 100 个印刷错误。

每页错误个数近似服从 Poisson 分布。

- 1) 随机取一页, 求这一页上错误不少于 2 个之概率。
- 2) 随机取 4 页, 求这 4 页上错误不少于 5 个之概率。

首先，两只随机分布之叠加仍是随机分布

$$X_1 \sim P(\lambda_1), X_2 \sim P(\lambda_2)$$

$$X = X_1 + X_2 \sim P(\lambda_1 + \lambda_2)$$

- $X \sim P(\lambda)$, 由 $P(X=k) \stackrel{\text{定义}}{=} P_k = e^{-\lambda} \frac{\lambda^k}{k!}$

$$E(X) = \lambda, D(X) = \lambda$$

- 设每页之错误分布 $X \sim P(\lambda)$

500页之错误分布 $Y \sim P(500\lambda)$ ("每页"累加起来)

$$EY = 500\lambda = 100 \Rightarrow \lambda = 0.2.$$

→ $X \sim P(0.2) \quad P(X=k) = e^{-0.2} \frac{0.2^k}{k!}$

- 随机取一页, 其上错误不少于2个之概率,

$$P = 1 - P_0 - P_1$$

$$= 1 - e^{-0.2} \frac{0.2^0}{0!} - e^{-0.2} \frac{0.2^1}{1!} = 0.0175$$

∴ $X \sim P(4\lambda) = P(0.8)$.

- 随机取4页, 错误不少于5个之概率

$$P = 1 - \underbrace{P_0 + P_1 + P_2 + P_3 + P_4}_{\text{少于5个之概率}}$$

少于5个之概率

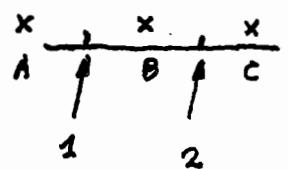
Jason throws two darts at a board, aiming for the center.

Prob
stat.

second lands farther from the center than the first.

If Jason throws a third dart, what's the probability
that the third is farther than the first?

Assume Jason's skillfulness is constant.



其实很简单.

1, 2 把 dartboard 分成 3 块.

$$\frac{\#\{B, C\}}{\#\{A, B, C\}} = \frac{2}{3}$$

How many people do we need in class to make the probability that two people have the same birthday more than $\frac{1}{2}$?

Prob
Stat.

$P(\text{at least two people have the same birthday})$

$= 1 - P(\text{no people have the same birthday})$

$P(\text{at least two}) > \frac{1}{2} \Rightarrow P(\text{no people have the same ..}) < \frac{1}{2}$



$$\frac{365}{365} \cdot \frac{364}{365} \cdot \frac{363}{365} \cdot \frac{365+1-n}{365} < \frac{1}{2}$$



$$n = 23.$$

Six letters were to be placed in envelopes for posting.

Unfortunately, the letters were dropped before putting in their envelopes, and they were placed in at random.

- What's the probability that none is in the correct envelope?
 - _____ exactly one _____.

Let A be the event that letter A is in its correct envelope, and
similarly B is the event that B is in its _____, then.

$$P(A) = \frac{1}{5} \quad (\text{5个信封中选一个A})$$

$$P(A \text{ and } B) = \frac{1}{5} \cdot \frac{1}{4}$$

A: 5选1 B: 全下4个中选1.

至少有1封对:

$$P(A \text{ or } B \text{ or } C \text{ or } D \text{ or } E)$$

$$\begin{aligned} &= P(A) + P(B) + P(C) + P(D) + P(E) \\ &\quad - P(A \text{ and } B) - P(B \text{ and } C) - \dots \\ &\quad + P(A \text{ and } B \text{ and } C) + P(B \text{ and } C \text{ and } D) + \dots \\ &\quad - P(A \text{ and } B \text{ and } C \text{ and } D) - P(\dots) - \dots \\ &\quad + P(A \text{ and } B \text{ and } C \text{ and } D \text{ and } E) \end{aligned}$$

$$= 5 \times \left(\frac{1}{5}\right)$$

$$- C_5^2 \frac{1}{5} \times \frac{1}{4}$$

$$+ C_5^3 \frac{1}{5} \times \frac{1}{4} \times \frac{1}{3}$$

$$- C_5^4 \frac{1}{5} \times \frac{1}{4} \times \frac{1}{3} \times \frac{1}{2}$$

$$+ 1 \cdot \frac{1}{5} \times \frac{1}{4} \times \frac{1}{3} \times \frac{1}{2} \times 1$$

$$\begin{aligned} &= 1 - \frac{5 \times 4}{2} \times \frac{1}{5 \times 4} + \frac{5 \cdot 4 \cdot 3}{3 \cdot 2 \cdot 1} \cdot \frac{1}{5 \cdot 4 \cdot 3} - \frac{5 \cdot 4 \cdot 3 \cdot 2}{4 \cdot 3 \cdot 2 \cdot 1} \cdot \frac{1}{5 \cdot 4 \cdot 3 \cdot 2} \\ &\quad + \frac{1}{5!} \end{aligned}$$

$$= 1 - \frac{1}{2!} + \frac{1}{3!} - \frac{1}{4!} + \frac{1}{5!} = \frac{120 - 60 + 20 - \frac{5}{24} + \frac{1}{120}}{120} = \frac{76}{120}$$

$$P(\text{全不对}) = 1 - P(\text{至少一封对}) = 1 - \frac{76}{120} = \frac{44}{120} = \frac{11}{30}$$



$$\begin{aligned} &P(A \text{ or } B \text{ or } C) \\ &= P(A) + P(B) + P(C) \\ &\quad - (P(A \cap B) + P(B \cap C) + P(C \cap A)) \\ &\quad + P(A \cap B \cap C) \end{aligned}$$

部分用 induction.

n封信全错:

$$-\frac{1}{2!} + \frac{1}{3!} - \frac{1}{4!} + \cdots + (-1)^n \frac{1}{n!}$$

$$e^{+x} = 1 + f'(0)x + \frac{f''(0)}{2!}x^2 + \cdots = 1 + x + \frac{x^2}{2!} + \frac{x^3}{3!} + \cdots$$

↑ 在 $x=0$ 处展开

$$e^{-1} = 1 - 1 + \frac{1}{2!} - \frac{1}{3!} + \cdots (-1)^n \frac{1}{n!} = 0.3678.$$

6封信全错:

$$\frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} - \frac{1}{5!} + \frac{1}{6!} = \frac{53}{144}$$

6封信对 1 封:

imagine the letter laid out in a row, and the envelopes in a matching row.

$$\frac{\binom{1}{6} \underbrace{\text{6封全不对}}_{\frac{11}{30} \times 5!} \text{arrangement}}{\text{送出一封对的.}} = \frac{6 \times \frac{11}{30} \times 120}{720} = \frac{11}{30}$$

所有6封之可能排法 $6!$

A deck of cards with numbers on $1, 2, \dots, N$

Draw a card randomly from the deck,
keep it at hand, then draw another one;
if the new one is larger than the largest
number in hand, keep it, otherwise, discard it.
Stop when you have the card N .

Question:

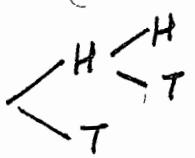
What is the expectation of numbers of
cards at hand.

We can use conditional probability on the first number drawn.

Let the result be $f(N)$

$$f(N) = \frac{1}{N} + \dots + \frac{1}{1}$$

You have a fair coin, calculate the expected number N
of throws that you needed to find two HH at the end
of the sequence?



$$n = \frac{1}{2}(n+1) + \frac{1}{2} \cdot 2 + \frac{1}{2}(n+2)$$

$$n = 6.$$

Prob.

一个青蛙，equal chance to jump to right or left.

向右跳，2分钟后果青蛙将回到原地。

向左跳，0.7的几率于4分钟后果青蛙将回到原地

0.3 —— 于3分钟后果~~青蛙~~
跳出木箱。

假设x次跳出木

$$x = \frac{1}{2}(x+2) + \frac{1}{2} \times 0.7 \times (x+4) + \frac{1}{2} \times 0.3 \times 3$$

$$x = 19$$

Chapter 4. Probability Theory, Questions.

- Coin toss game.

Two gamblers are playing a coin toss game. Gambler A has $(n+1)$ fair coins; B has n fair coins. What is the probability that A will have more heads than B if both flip all their coins.

- Card game.

A casino offers a simple card game. There are 52 cards in a deck with 4 cards for each value 2, 3, 4, 5, 6, 7, 8, 9, 10, J, Q, K, A. Each time the cards are thoroughly shuffled (so each card has equal probability of being selected). You pick up a card from the deck and the dealer picks another one without replacement. If you have a larger number, you win; if the numbers are equal or yours is smaller, the house wins — as in all other casinos, the house always has better odds of winning. What is your probability of winning?

- Drunken passenger

A line of 100 airline passengers are waiting to board a plane. They each hold a ticket to one of 100 seats on the flights. For convenience, let's say the n -th passenger in line has a ticket for the seat number n . Being drunk, the first person in line picks a random seat (equally likely for each seat). All of the other passengers are sober, and will go to their proper seats unless it's already occupied; In that case, they will randomly choose a free seat. You're person number 100. What is the probability that you end up in your seat (i.e., seat #100)

- N points on a circle

Given N points drawn randomly on the circumference of a circle, what is the probability that they are all within a semicircle?

Start

药厂制剂车间用自动装瓶机封装药液，
机器工作正常时，每瓶药液净重500克。某日

随机取710瓶，称得为：504, 498, 496, 487,
509, 476, 482, 510, 469, 472. 7问此机器
是否正常工作。

- 500克是什么？总体均值

假设有何？总体均值等于500.

若不能否定，则理解为装瓶机正常

假检验分三步.

1. 提出原假设：总体期望 = 500.

2. 建立检验统计量：

样本均值: 500

$$\sqrt{n-1} \frac{\bar{x} - \mu_0}{s} \sim t(n-1)$$

10个样本 样本方差

统计量 T

即统计量 T 服从自由度为 $n-1$ 的 t -分布，且不带未知参数，
它可作为判断 H_0 的检验统计量，这种检验方法，称为
 t 检验法。

原理：

- 每一次抽样，可视为一个随机变量，R.V.

此 R.V. 与总体 ~~同~~ 同分布。

- 10次抽样，可视为 10 个不同之 R.V.

→ 独立同分布
i.i.d.

- 则 \bar{X} 是把这 10 个随机变量相加再除以 n
求平均。

- S 本身也是随机变量

通过样本方差和差商之公式，若原总体符合正态

则 T 符合 t 分布。

→ t, S^2 相互独立 R.V.

$t \sim N(0, 1)$, $S^2 \sim \chi^2(n)$

$\therefore t = \frac{\bar{X} - \mu_0}{\sqrt{\frac{s^2}{n}}} \sim t$ 分布
(自由度为 n)

$$f_{t^2}(x) = f_{\chi^2/n}(x) = \frac{\frac{1}{2}(\frac{n+1}{2})}{\sqrt{\pi n}} \cdot \left(1 + \frac{x^2}{n}\right)^{-\frac{n+1}{2}}$$

T 究极可视为阶乘之推广

1. The sample space of two children is given by

$$\Omega = \{(b, b), (b, g), (g, b), (g, g)\}$$
 $P - \text{at least one boy}, P = \text{at least one girl}.$

Since Ms Jackson is invited, she has at least one son,

Let B be the event that at least one of the children is a boy,
and A be the event that both children are boys.

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(\{(b, b)\})}{P(\{(b, b), (g, b), (b, g)\})} = \frac{1}{3}$$

2. 另一个可能是男 ($\frac{1}{2}$)

女 — 女 ($\frac{1}{2}$)

故两个男孩的概率为 $\frac{1}{2}$.

Boys and girls:

1. A company is holding a dinner for working mothers with at least one son. Ms. Jackson, a mother with two children, is invited. What's the probability that both children are boys?
2. Your new colleague, Ms. Parker is known to have two children. If you see her walking with one of her children and that child is a boy, what is the probability that both children are boys?

一共 52 张牌中取 5 张, 顺序 does not matter.

共有 C_{52}^5 种.

4 张 - 1 张 $C_{13}^1 \cdot C_{48}^1$
13 组
 $A \dots K$ ↓
剩下 48 张.

$$P = \frac{13 \times 48}{C_{52}^5}$$

Full house.

$C_3^1 C_{12}^1 C_4^3 C_4^2$
13 组
~~双三~~
1 组 3 张
剩下 12 组
中 2 张
1 组 2 张

$$P = \frac{13 \times 12 \times 4 \times 6}{C_{52}^5}$$

Two pairs.

$C_3^1 C_{12}^1 C_4^2 C_4^2 C_{44}^1$
剩下 3 张

$$P = \frac{13 \times 12 \times 6 \times 6 \times 44}{C_{52}^5}$$

Texas holdem.

得到43K-2J1S
full house (33K-2J1S-2J)

两对
的概率分别是多少？

$$我 > A \quad p$$

$$我 = A \quad 1 - 2p = \frac{3}{5}$$

$$我 < A \quad p.$$

$$1 - 2p = \frac{1}{7} \quad p = \frac{8}{7}$$

A casino offers a simple card game.

52 张牌，你和 dealer 各抽 3 张，无放回。

若你的牌大，你赢。

牌等或你的牌小，庄家赢。

问：你赌赢的概率多少。

soln:

A 和 B 都只扔 n 个时：

$$A > B \quad p$$

$$A = B \quad 1 - 2p$$

$$A < B \quad p$$

($A > B$ 之概率
等于 $A < B$)

by symmetry)

不管怎样 A 依然大于 B

$\frac{1}{2}$ + 概率 A 为 head,

从而 $A > B$

不管怎样

A 不大于 B.

(最多相等).

B + B n 个, A + B n+1 个.

p.

$$\frac{1}{2} (1 - 2p)$$

0

→ 都 放而扔完后.

$A > B$ 之概率:

$$p + \frac{1}{2} (1 - 2p) = \frac{1}{2}.$$

Two gamblers are playing a coin toss game,
Gambler A has $(n+1)$ fair coins; B has n
fair coins. What's the probability that A
will have more heads than B if both flip
all their coins?

Prove s^2 is the unbiased
estimator of σ^2

Stats

s^2 : sample variance.

$$\text{Pf.} \quad E(S^2) = E\left(\frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1}\right) = \sigma^2$$

Soln.

Let X_1, X_2, \dots, X_n be n independent observations from a population with mean μ and var σ^2 .

- Pf. S^2 为 X_i 有

$$E(X_i) = \mu, \quad \text{Var}(X_i) = \sigma^2$$

$$- \left| \begin{aligned} \text{Var}(X) &= E(X^2) - (EX)^2 \\ \text{Rif. } E(X^2) &= \sigma^2 + \mu^2 \end{aligned} \right.$$

$$\left| \begin{aligned} \text{Var}(\bar{X}) &= E(\bar{X}^2) - [E(\bar{X})]^2 \\ &\quad \xrightarrow{\text{Sample mean, 是 R.V.}} \\ \text{有: } E(\bar{X}^2) &= \frac{\sigma^2}{n} + \mu^2 \end{aligned} \right.$$

$$\begin{aligned} &E[\sum(X_i - \bar{X})^2] \\ &= E[\sum(X_i^2 - 2X_i\bar{X} + \bar{X}^2)] \\ &= E(\sum X_i^2 - \sum 2X_i\bar{X} + \sum \bar{X}^2) \\ &\quad \xrightarrow{\begin{array}{l} \bar{X} \text{ 是常数,} \\ \text{是 } X_i \text{ 的函数, } \bar{X} = \frac{\sum X_i}{n} \end{array}} \\ &= E(\sum X_i^2 - 2\bar{X} \sum X_i + n\bar{X}^2) \\ &\quad \xrightarrow{n\bar{X}} \end{aligned}$$

$$\begin{aligned} &= E(\sum X_i^2 - n\bar{X}^2) \\ &= \sum E(X_i^2) - nE(\bar{X}^2) \\ &= \sum (\sigma^2 + \mu^2) - n\left(\frac{\sigma^2}{n} + \mu^2\right) \\ &= n\sigma^2 + n\mu^2 - \sigma^2 - n\mu^2 = (n-1)\sigma^2 \end{aligned}$$

Pf.

$$E(S^2) = \sigma^2$$

$$\text{其中 } S^2 = \frac{\sum_{i=1}^n (X_i - \bar{X})^2}{n-1},$$

是 σ^2 之无偏估计。

Let play a game,

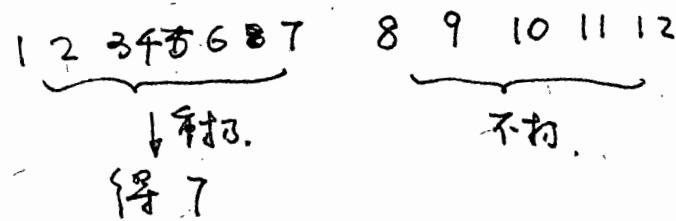
I give you a 12-sided die, and will pay you whatever the die lands on. If you're unhappy with the roll, you can choose to roll another two 6-sided dice, and I will pay you the sum of the two dice. How much are you willing to pay to play this game?

$$\frac{1+12}{2} = 6.5$$

7, 8, 9, 10, 11, 12

前面2, 6面 die; $\frac{1+6}{2} = 3.5 \quad 3.5 \times 2 = 7.$

于是低于7的都算中奖.



$$\frac{7 \times 7 + \frac{8+12}{2} \times 5}{12} = \frac{49 + 50}{12} = 8.25.$$

Chapter 4 Probability Theory, soln.

1

- Coin toss game.

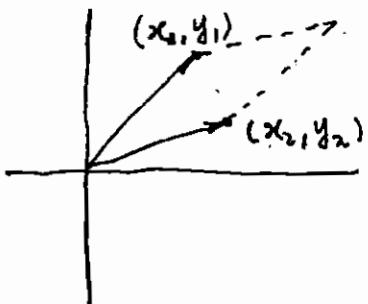
- A toss n & B toss n .

$$\begin{array}{ccc}
 \text{Heads of A} & > & \text{Heads of B} & x \\
 & = & & 1-2x \\
 & < & & x \quad \text{by symmetry.}
 \end{array}$$

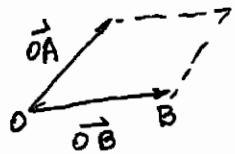
- 再抛一次，A是heads + B是tails:

$$x + (1-2x) \cdot \frac{1}{2} + 0 = x + \frac{1}{2} - x = \frac{1}{2}$$

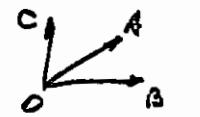
↓ ↓ ↗
 n次时比B大。 n次时与B相等。 n次时比B大。
 n+1次不管什么。 n+1次时，若是head， 则不管什么。
 然后比B大。 则比B大，概率是 $\frac{1}{2}$ 。 还是比B大。



如何迅速算出图示平行四边形之面积?

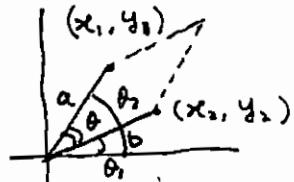


由两个向量之叉积，大小为二者围成△之面积。
方向垂直于AOB平面。



在二维的情形下，

$$|\vec{OA} \times \vec{OB}| = \begin{vmatrix} x_1 & y_1 \\ x_2 & y_2 \end{vmatrix}$$



证明如下。 $\text{Area} = a \cdot b \cdot \sin\theta$.

$$\sin\theta = \sin(\theta_2 - \theta_1)$$

$$= \sin\theta_2 \cos\theta_1 - \cos\theta_2 \sin\theta_1$$

$$= \frac{y_2}{b} \frac{x_1}{a} - \frac{x_2}{b} \frac{y_1}{a}$$

$$\text{Area} = y_2 x_1 - x_2 y_1 = \begin{vmatrix} x_1 & y_1 \\ x_2 & y_2 \end{vmatrix}$$

$$\rightarrow \int_{-\infty}^{+\infty} e^{-x^2} dx.$$

两种办法求解.

1. 令 $a = \int_{-\infty}^{+\infty} e^{-x^2} dx$

R: $a^2 = (\int_{-\infty}^{+\infty} e^{-x^2} dx)^2 \quad \text{→ 换元}$

$$= \int_{-\infty}^{+\infty} e^{-x^2} dx \cdot \int_{-\infty}^{+\infty} e^{-y^2} dy$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} e^{-(x^2+y^2)} dx dy \quad \text{累次积分化为重积分}$$

{ 极坐标.

2. $\int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2}{2\sigma^2}} = 1 \quad (*)$

$$a = \int_0^{+\infty} e^{-x^2}$$

$\sqrt{\pi}$. 仅讨论 $x > 0$.

$$\left\{ \mu=0, 2\sigma^2=1, \sigma=\frac{1}{\sqrt{2}} \right.$$

* \Rightarrow 化为

$$\frac{1}{\sqrt{\pi}} \int_0^{+\infty} e^{-x^2} = 1$$

于是 $a = \sqrt{\pi}$