# Transparent causal inference for observational epidemiology

Lorenzo Fabbri

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#### Who am I?

- Final-year\* PhD student at the Barcelona Institute for Global Health (ISGlobal).
  - Google Scholar: <a href="https://shorturl.at/IJCU8">https://shorturl.at/IJCU8</a>
  - CV: <a href="https://shorturl.at/N3YeP">https://shorturl.at/N3YeP</a>
  - Code: <a href="https://github.com/lorenzoFabbri">https://github.com/lorenzoFabbri</a>

 Prenatal and childhood exposure to mixtures of non-persistent endocrine disrupting chemicals (EDC) and adolescence neurodevelopment: a triangulation study.

Vrijheid, M., Basagaña, X., Gonzalez, J.R., Jaddoe, V.W., Jensen, G., Keun, H.C., McEachan, R.R., Porcel, J., Siroux, V., Swertz, M.A. and Thomsen, C., 2021. Advancing tools for human early lifecourse exposome research and translation (ATHLETE): Project overview. *Environmental Epidemiology*, *5*(5), p.e166.

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# Approaching the question

#### Some notation

- I will make use of the following notation:
  - A will be our (vector of) exposure(s).
  - W will be our vector of a priori selected confounders.
  - C is an indicator variable for censoring (C = 1 indicating censored).
  - Y will be our outcome.
  - The potential outcomes will be indicated with  $Y^{\{a\}}$ .

#### Approaching the question

1. Start with a well-defined causal question, and define the causal estimand and model (e.g., DAG):

$$\Psi^* = E[Y^{\{a=a^*,c=0\}}] - E[Y^{\{a=a^{nc},c=0\}}]$$

Dang, L.E., Gruber, S., Lee, H., Dahabreh, I.J., Stuart, E.A., Williamson, B.D., Wyss, R., Díaz, I., Ghosh, D., Kıcıman, E. and Alemayehu, D., 2023. A causal roadmap for generating high-quality real-world evidence. *Journal of Clinical and Translational Science*, 7(1), p.e212.

# Intermezzo: causal policies / shift functions

- Binary exposures (presence/absence, treated/untreated):  $Y^{\{a=1\}} Y^{\{a=0\}}$ .
- Categorical exposures:  $Y^{\{a=a_5\}} Y^{\{a=a_0\}}$ .
- Continuous exposures: ?
  - $Y^{\{a=avg(a)\}} Y^{\{a=nc\}}$
  - $Y^{\{a=75^{th}\}} Y^{\{a=25^{th}\}}$ .
  - MTP: if  $a \ge 0.5$  and bw < 70 then  $a^g = a * 0.1$ , else  $a^g = a$ .
  - •
- A zoo of interventions: static, dynamic, MTPs, deterministic, stochastic...

#### Approaching the question

- 1. Start with a well-defined causal question, and define the causal estimand and model (e.g., DAG).
- 2. Consider the **observed** data and **identifiability** conditions. Define the statistical estimand:

$$\Psi = E_W[E[Y|W, A = a^*, C = 0] - E[Y|W, A = a^{nc}, C = 0]]$$

#### Approaching the question

3. Choose the *best* statistical **estimator**:

$$Y \sim A_1 + A_2 + ... + A_p + W$$
Components of mixture *i*

#### To summarize:

- 1. Define your causal estimand.
- 2. Derive the associated statistical estimand.
- 3. Estimate the so-called nuisance functions.
- 4. Estimate your effect of interest.

## How to get the best estimator: g-formula

Remember:  $E[Y|W, A = a^*, C = 0]$ 

$$\Psi^g = \int_A \int_W E[Y|A = a, W = w] f^g(A|W) f(W) dadw$$

Díaz, I., Williams, N., Hoffman, K.L. and Schenck, E.J., 2023. Nonparametric causal effects based on longitudinal modified treatment policies. *Journal of the American Statistical Association*, *118*(542), pp.846-857.

## How to get the best estimator: g-formula

$$\Psi^g = \int_A \int_W E[Y|A = a, W = w(f^g(A|W)) f(W) dadw$$

Díaz, I., Williams, N., Hoffman, K.L. and Schenck, E.J., 2023. Nonparametric causal effects based on longitudinal modified treatment policies. *Journal of the American Statistical Association*, *118*(542), pp.846-857.

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$$IF_{1} = \int_{A} \int_{W} \frac{I(A = a, W = w)}{(f(A = a, W = w))} \{Y - E[Y|A, W]\} f^{g}(A|W) f(W) dadw$$
$$= \frac{f^{g}(A|W)}{f(A|W)} \{Y - E[Y|A, W]\}.$$

$$\Psi^g = \int_A \int_W E[Y|A = a, W = w] f^g(A|W) f(W) dadw$$

$$IF_{2} = \int_{A} \int_{W} E[Y|A, W] f^{g}(A|W) \{I(W = w) - f(W = w)\} dadw$$

$$= \int_{A} E[Y|A, W] f^{g}(A|W) da - \Psi$$

$$= \int_{A} E[Y|A^{*}, W) f(A|W) f^{+}(A|W) da - \Psi$$

$$\Psi^g = \int_A \int_W E[Y|A = a, W = w] f^g(A|W) f(W) dadw$$

$$EIF = \frac{f^{g}(A|W)}{f(A|W)} \{Y - E[Y|A, W]\} + E[Y|A^{g}, W] - \Psi$$

• One possible way to use the EIF to estimate the parameter of interest is by **plugging in** estimates of the individual components and averaging over the sample (AIPW).

# How to get the *best* estimator: nuisance functions

• 
$$\Psi_{AIPW}^{g} = avg(\widehat{EIF})$$
  
=  $\frac{1}{n} \sum_{i} \frac{f^{g}(A_{i}|W_{i})}{f(A_{i}|W_{i})} \{Y_{i} - E[Y_{i}|A_{i}, W_{i}]\} + E[Y_{i}|A_{i}^{g}, W_{i}]$ 

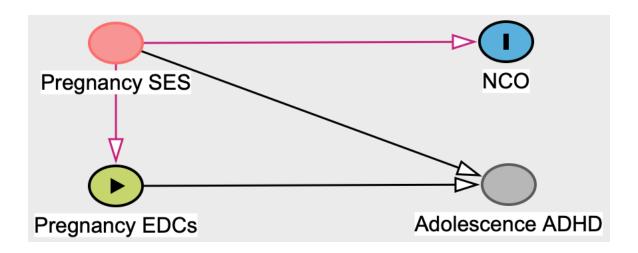
- We combine 2 nuisance functions (double-robustness): the propensity score and the outcome model.
- For my project, probably using Imtp R package: <a href="https://github.com/nt-williams/Imtp">https://github.com/nt-williams/Imtp</a>. Works also with multiple exposures. Implements parametric g-computation, IPW, TML, and SDR estimators. Static, dynamic, and MT policies.

#### Why that?

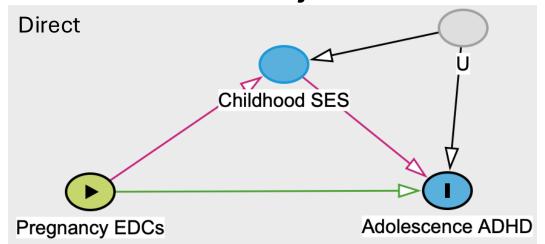
- Literature on mixtures (environmental epidemiology): quantile g-computation, BKMR, BWQS...
  - Parametric models.
  - No causal "background".
  - Discretization of exposures that are continuous in nature.

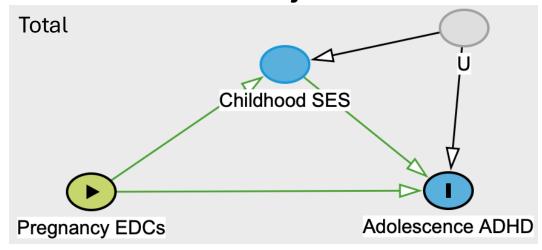
- Observational studies are subject to bias (many).
- We can strengthen our conclusions by combining different data sources and methods:
  - Negative controls (exposure and/or outcome).
  - Cross-cohort comparisons.
  - Genetic variants (MR), although different estimand.
  - •

- Observational studies are subject to bias (many).
- We can strengthen our conclusions by combining different data sources and methods:
  - Negative control outcome, assuming "SES" is main culprit.



• **Direct or total effects**: "...when there is an unmeasured common cause of the intermediate and the outcome, associations adjusted for the intermediate are subject to bias."





What if there is no arrow between EDCs and SES?

VanderWeele, T.J., Mumford, S.L. and Schisterman, E.F., 2012. Conditioning on intermediates in perinatal epidemiology. *Epidemiology*, 23(1), pp.1-9.

#### Thank you for your attention!

You can contact me at <a href="lorenzo.fabbri@isglobal.org">lorenzo.fabbri@isglobal.org</a>.