

# Homework2

Qinglin Li, 5110309074

## Problem 1

Since every person have only two hands, the game continue  $n$  rounds. The final state must be some cycles.

Before the  $n$ -th round, the state must be some lines(a single without self-loop is also considered to be a line) with some cycles.

In every round, we either change a line into a cycle or combine two lines into one line.

Let  $X_n$  be the number of cycles when there are  $n$  lines remaining.

Let  $A_n$  be the event changing a line into cycle when there are  $n$  lines remaining.

$$\Pr(A_n) = \frac{n}{\binom{2n}{2}} = \frac{1}{2n-1}$$

$$X_n = \begin{cases} X_{n-1} + 1 & \Pr = \frac{1}{2n-1} \\ X_{n-1} & \Pr = \frac{2n-2}{2n-1} \end{cases} \quad (1)$$

$$E[X_n] = \frac{1}{2n-1} (1 + E[n-1]) + \frac{2n-2}{2n-1} E[n-1] = E[n-1] + \frac{1}{2n-1}$$

$$E[X_n] = \sum_{i=1}^n \frac{1}{2i-1}$$

## Problem 2

Let  $X_i$  be the number of balls in each bins in uniformly random case

Let  $Y_i$  be the number of balls in each bins in 2-choice paradigm case

Obviously, in paradigm 1-3, the answer is  $E[\max_i(X_i + Y_i)]$

$$E[\max_i (X_i + Y_i)] \leq E[\max_i X_i] + E[\max_i Y_i]$$

Thus

$$E[\max_i (X_i + Y_i)] = O\left(\frac{\ln \frac{n}{2}}{\ln \ln \frac{n}{2}} + \ln \ln \frac{n}{2}\right) = O\left(\frac{\ln n}{\ln \ln n}\right)$$

$$E[\max_i (X_i + Y_i)] \geq E[\max_i X_i]$$

Thus

$$E[\max_i (X_i + Y_i)] = \Omega\left(\frac{\ln \frac{n}{2}}{\ln \ln \frac{n}{2}}\right) = \Omega\left(\frac{\ln n}{\ln \ln n}\right)$$

So the answer is  $\Theta\left(\frac{\ln n}{\ln \ln n}\right)$

### Problem 3

Let  $H_i$  be Number of Heads - Number of Tails

$$E[|H_i|] = \sum_{x>0} x \Pr(H_i = x) - \sum_{x<0} x \Pr(H_i = x) = 2 \sum_{x>0} x \Pr(H_i = x)$$

Let  $i = 2n$

$$E[|H_i|] = 4 \sum_{y=0}^n y \frac{\binom{2n}{n+y}}{2^{2n}} = 2^{2-2n} \sum_{y=0}^n y \binom{2n}{n+y}$$

Let  $S_1 = \sum_{y=0}^n y \binom{2n}{n+y} = 0 \cdot \binom{2n}{n} + \dots + n \cdot \binom{2n}{2n}$

Let  $S_2 = \sum_{y=0}^n y \binom{2n}{y} = n \cdot \binom{2n}{n} + \dots + 0 \cdot \binom{2n}{0} = n \cdot \binom{2n}{n} + \dots + 0 \cdot \binom{2n}{2n}$

$$S_1 + S_2 = n \sum_{y=n}^{2n} \binom{2n}{y} = \frac{n}{2} \left( 2^{2n} + \binom{2n}{n} \right)$$

$$S_2 = \sum_{y=0}^n y \binom{2n}{y} = \sum_{y=0}^n 2n \binom{2n-1}{y-1} = 2n \sum_{y=0}^{n-1} \binom{2n-1}{y} = n \cdot 2^{2n-1}$$

$$S_1 = (S_1 + S_2) - S_2 = \frac{n}{2} \binom{2n}{n}$$

$$E[|H_i|] = 2^{2-2n} S_1 = 2^{2-2n} \frac{(2n)!n}{n!n!}$$

By Stirling's approximation

$$E[|H_i|] \approx 2^{2-2n} \frac{\sqrt{4\pi n} \left(\frac{2n}{e}\right)^{2n} n}{2\pi n \left(\frac{n}{e}\right)^{2n}} = 4\sqrt{\frac{n}{\pi}} = \Theta(\sqrt{n})$$

$$\begin{aligned} & \Pr(H \geq k) \\ &= \Pr(H \geq k | H_n \geq k) \cdot \Pr(H_n \geq k) + \Pr(H \geq k | H_n < k) \cdot \Pr(H_n < k) \\ &= \Pr(H_n \geq k) + \Pr(H \geq k | H_n < k) \cdot \Pr(H_n < k) \end{aligned}$$

Since  $H_n$  is either greater or less than  $H_{n-1}$  by 1, by reflection principle

$$\Pr(H \geq k | H_n < k) \cdot \Pr(H_n < k) = \Pr(H_n \geq k + 1)$$

Thus

$$\Pr(H \geq k) = \Pr(H_n \geq k) + \Pr(H_n \geq k + 1)$$

$$\begin{aligned} & \Pr(H = k) \\ &= \Pr(H \geq k) - \Pr(H \geq k + 1) \\ &= \Pr(H_n \geq k) + \Pr(H_n \geq k + 1) - \Pr(H_n \geq k + 1) - \Pr(H_n \geq k + 2) \\ &= \Pr(H_n = k) + \Pr(H_n = k + 1) \end{aligned}$$

$$\begin{aligned} E[H_i] &= \sum_{k \geq 0} k(\Pr[H_n = k] + \Pr[H_n = k + 1]) \\ &= \sum_{k \geq 1} (2k - 1) \Pr[H_n = k] \\ &= \Theta(E[H_n]) \\ &= \Theta(\sqrt{n}) \end{aligned}$$

## Problem 4

1. Let  $Z = \frac{X - \mu_X}{\sigma_X}$

$$\Pr(X - \mu_X \geq t\sigma_X) = \Pr(Z \geq t) = \Pr\left(Z + \frac{1}{t} \geq t + \frac{1}{t}\right) \leq \Pr\left(\left(Z + \frac{1}{t}\right)^2 \geq \left(t + \frac{1}{t}\right)^2\right)$$

By Markov's inequality

$$\Pr(X - \mu_X \geq t\sigma_X) \leq \frac{E\left[\left(Z + \frac{1}{t}\right)^2\right]}{\left(t + \frac{1}{t}\right)^2}$$

$$E[Z^2] = \frac{E[(X - \mu_X)^2]}{E[Var[X]]} = 1$$

$$E[Z] = \frac{E[X - \mu_X]}{E[Var[X]]} = 0$$

Thus

$$\Pr(X - \mu_X \geq t\sigma_X) \leq \frac{1 + \frac{1}{t^2}}{\left(t + \frac{1}{t}\right)^2} = \frac{1}{1 + t^2}$$

2.

$$\Pr(|X - \mu_X| \geq t\sigma_X) = \Pr((X - \mu_X)^2 \geq t^2 Var[X]) = \Pr\left(\frac{(X - \mu_X)^2}{Var[X]} + 1 \geq t^2 + 1\right)$$

By Markov's inequality

$$\Pr(|X - \mu_X| \geq t\sigma_X) \leq \frac{E\left[\frac{(X - \mu_X)^2}{Var[X]} + 1\right]}{t^2 + 1} = \frac{2}{t^2 + 1}$$