

MATH357 - Statistics

Based on lectures from Winter 2025 by Prof. Abbas Khalili.
Notes by Louis Meunier

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§1 REVIEW OF PROBABILITY

↪ **Definition 1.1** (Measurable Space, Probability Space): We work with a set Ω = sample space = {outcomes}, and a σ -algebra \mathcal{F} , which is a collection of subsets of Ω containing Ω and closed under taking complements and countable unions. The tuple (Ω, \mathcal{F}) is called *measurable space*.

We call a nonnegative function $P : \mathcal{F} \rightarrow \mathbb{R}$ defined on a measurable space a *probability function* if $P(\Omega) = 1$ and if $\{E_n\} \subseteq \mathcal{F}$ a disjoint collection of subsets of Ω , then $P(\bigcup_{n \geq 1} E_n) = \sum_{n \geq 1} P(E_n)$. We call the tuple (Ω, \mathcal{F}, P) a *probability space*.

↪ **Definition 1.2** (Random Variables): Fix a probability space (Ω, \mathcal{F}, P) . A Borel-measurable function $X : \Omega \rightarrow \mathbb{R}$ (namely, $X^{-1}(B) \in \mathcal{F}$ for every $B \in \mathfrak{B}(\mathbb{R})$) is called a *random variable* on \mathcal{F} .

- *Probability distribution*: X induces a probability distribution on $\mathfrak{B}(\mathbb{R})$ given by $P(X \in B)$
- *Cumulative distribution function (CDF)*:

$$F_X(x) := P(X \leq x).$$

Note that $F(-\infty) = 0, F(+\infty) = 1$ and F right-continuous.

We say X *discrete* if there exists a countable set $S := \{x_1, x_2, \dots\} \subset \mathbb{R}$, called the *support* of X , such that $P(X \in S) = 1$. Putting $p_i := P(X = x_i)$, then $\{p_i : i \geq 1\}$ is called the *probability mass function* (PMF) of X , and the CDF of X is given by

$$P(X \leq x) = \sum_{i: x_i \leq x} p_i.$$

We say X *continuous* if there is a nonnegative function f , called the *probability distribution function* (PDF) of X such that $F(x) = \int_{-\infty}^x f(t) dt$ for every $x \in \mathbb{R}$. Then,

- $\forall B \in \mathfrak{B}(\mathbb{R}), P(X \in B) = \int_B f(t) dt$
- $F'(x) = f(x)$
- $\int_{-\infty}^{\infty} f(x) dx = 1$

If $X : \Omega \rightarrow \mathbb{R}$ a random variable and $g : \mathbb{R} \rightarrow \mathbb{R}$ a Borel-measurable function, then $Y := g(X) : \Omega \rightarrow \mathbb{R}$ also a random variable.

↪ **Definition 1.3** (Moments): Let X be a discrete/random variable with pmf/pdf f and support S . Then, if $\sum_{x \in S} |x| f(x) / \int_S |x| f(x) dx < \infty$, then we say the first moment/mean of X exists, and define

$$\mu_X = \mathbb{E}[X] = \begin{cases} \sum_{x \in S} x f(x) \\ \int_S x f(x) dx \end{cases}.$$

Let $g : \mathbb{R} \rightarrow \mathbb{R}$ be a Borel-measurable function. Then, we have

$$\mathbb{E}[g(X)] = \begin{cases} \sum_{x \in S} g(x) f(x) \\ \int_S g(x) f(x) dx \end{cases}.$$

Taking $g(x) = |x|^k$ gives the so-called “ k th absolute moments”, and $g(x) = x^k$ gives the ordinary “ k th moments”. Notice that $\mathbb{E}[\cdot]$ is linear in its argument.

For $k \geq 1$, if μ exists, define the central moments

$$\mu_k := \mathbb{E}[(X - \mu)^k],$$

where they exist.

↪ **Definition 1.4** (Moment Generating Function (mgf)): If X a r.v., the mgf of X is given by

$$M(t) := \mathbb{E}[e^{tX}],$$

if it exists for $t \in (-h, h)$, $h > 0$. Then, $M^{(n)}(0) = \mathbb{E}[X^n]$.

↪ **Definition 1.5** (Multiple Random Variable): $X = (X_1, \dots, X_n) : \Omega \rightarrow \mathbb{R}^n$ a random vector if $X^{-1}(I) \in \mathcal{F}$ for every $I \in \mathfrak{B}_{\mathbb{R}^n}$. (It suffices to check for “rectangles” $I = (-\infty, a_1] \times \dots \times (-\infty, a_n]$, as before.)

Let F be the CDF of X , and let $A \subseteq \{1, \dots, n\}$, enumerating A by $\{i_1, \dots, i_k\}$. Then, the CDF of the subvector $X_A = (X_{i_1}, \dots, X_{i_k})$ is given by

$$F_{X_A}(x_{i_1}, \dots, x_{i_k}) = \lim_{\substack{x_{i_j} \rightarrow \infty, \\ i_j \in \mathcal{I} \setminus A}} F(x_1, \dots, x_n).$$

In particular, the marginal distribution of X_i is given by

$$F_{X_i}(x) = \lim_{x_1, \dots, x_{i-1}, x_{i+1}, \dots, x_n \rightarrow +\infty} F(x_1, \dots, x, \dots, x_n).$$

Let $g : \mathbb{R}^n \rightarrow \mathbb{R}$ measurable. Then,

$$\mathbb{E}[g(X_1, \dots, X_n)] = \begin{cases} \sum_{(x_1, \dots, x_n)} g(x_1, \dots, x_n) f(x_1, \dots, x_n) \\ \int \dots \int g(x_1, \dots, x_n) f(x_1, \dots, x_n) dx_1 \dots dx_n \end{cases}.$$

We have the notion of a joint mgf,

$$M(t_1, \dots, t_n) = \mathbb{E}\left[e^{\sum_{i=1}^n t_i X_i}\right],$$

if it exists for $0 < \left(\sum_{i=1}^n t_i^2\right)^{\frac{1}{2}} < h$ for some $h > 0$. Notice that $M(0, \dots, 0, t_i, 0, \dots, 0)$ is equal to the mgf of X_i .

↪ **Definition 1.6** (Conditional Probability): Let (X_1, \dots, X_n) a random vector. Let $\mathcal{I} = \{1, \dots, n\}$ and A, B disjoint subsets of \mathcal{I} with $k := |A|, h := |B|$. Write $X_A = (X_{i_1}, \dots, X_{i_k})^t$, similar for B . Then, the conditional probability of A given B is given by

$$f_{X_A|X_B}(x_a|x_b) := f_{X_A|X_B=x_B}(x_A) = \frac{f_{X_A, X_B}(x_a, x_b)}{f_{X_B}(x_b)},$$

provided the denominator is nonzero. Sometimes we have information about conditional probabilities but not the main probability function; we have that

$$f(x_1, \dots, x_n) = f(x_1)f(x_2|x_1)f(x_3|x_1, x_2)\dots f(x_n|x_1, \dots, x_{n-1}),$$

which follows from expanding the previous definition and observing the cancellation.

Let $X = (X_1, \dots, X_n) \sim F$. We say X_1, \dots, X_n (mutually) independent and write $\prod_{i=1}^n X_i$ if

$$F(x_1, \dots, x_n) = \prod_{i=1}^n F_{X_i}(x_i),$$

where F_{X_i} the marginal cdf of X_i . Equivalently,

$$\begin{aligned} \prod_{i=1}^n X_i &\Leftrightarrow f(x_1, \dots, x_n) = \prod_{i=1}^n f_{X_i}(x_i) \\ &\Leftrightarrow P(X_1 \in B_1, \dots, X_n \in B_n) = \prod_{i=1}^n P(X_i \in B_i) \quad \forall B_i \in \mathfrak{B}_{\mathbb{R}} \\ &\Leftrightarrow M_X(t_1, \dots, t_n) = \prod_{i=1}^n M_{X_i}(t_i). \end{aligned}$$

If X, Y are two random variables with cdfs F_X, F_Y such that $F_X(z) = F_Y(z)$ for every z , we say X, Y identically distributed and write $X \stackrel{d}{=} Y$ (note that X need not equal Y pointwise). If X_1, \dots, X_n a collection of random variables that are independent and identically distributed with common cdf F , we write $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} F$.

Further, define the covariance, correlation of two random variables X, Y respectively:

$$\text{Cov}(X, Y) := \sigma_{X,Y} := \mathbb{E}[(X - \mathbb{E}[X])(Y - \mathbb{E}[Y])] = \mathbb{E}[XY] - \mu_X \mu_Y, \quad \rho_{X,Y} := \frac{\sigma_{XY}}{\sigma_X \sigma_Y},$$

if $\mathbb{E}[|X - \mathbb{E}[X]| |Y - \mathbb{E}[Y]|] < \infty$.

↪ **Theorem 1.1**: If X_1, \dots, X_n independent and $g_1, \dots, g_n : \mathbb{R} \rightarrow \mathbb{R}$ borel-measurable functions, then $g_1(X_1), \dots, g_n(X_n)$ also independent.

↪ **Definition 1.7** (Conditional Expectation): Let X, Y be random variables and $g : \mathbb{R} \rightarrow \mathbb{R}$ a borel-measurable function. We define the following notions:

$$\mathbb{E}[g(X)|Y = y] = \begin{cases} \sum_{x \in S_X} g(x)f(x|y) & \text{discrete} \\ \int_{S_X} g(x)f(x|y) dx & \text{cnts} \end{cases}.$$

$$\text{Var}(X|Y = y) = \mathbb{E}[X^2|Y = y] - \mathbb{E}^2[X|Y = y].$$

↪ **Theorem 1.2**: If $\mathbb{E}[g(X)]$ exists, then $\mathbb{E}[g(X)] = \mathbb{E}[\mathbb{E}[g(X)|Y]]$, where the first nested \mathbb{E} is with respect to x , the second y .

↪ **Theorem 1.3**: If $\mathbb{E}[X^2] < \infty$, then $\text{Var}(X) = \text{Var}(\mathbb{E}[X|Y]) + \mathbb{E}[\text{Var}(X|Y)]$. In particular, $\text{Var}(X) \geq \text{Var}(\mathbb{E}[X|Y])$.

§2 STATISTICS

↪ **Definition 2.1** (Inference): We consider some population with some characteristic we wish to study. We can model this characteristic as a random variable $X \sim F$. In general, we don't have access to F , but wish to take samples from our population to make inferences about its properties.

(1) *Parametric inference*: in this setting, we assume we know the functional form of X up to some parameter, $\theta \in \Theta \subset \mathbb{R}^d$, where Θ our "parameter space". Namely, we know $X \sim F_\theta \in \mathcal{F} := \{F_\theta \mid \theta \in \Theta\}$.

(2) *Non-parametric inference*: in this setting we know nothing about F itself, except perhaps that F continuous, discrete, etc.

Other types exist. We'll focus on these two.

↪ **Definition 2.2** (Random Sample): Let $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} F$. Then X_1, \dots, X_n called a *random sample* of the population.

We also call X_i the "pre-experimental data" (to be observed) and x_i the "post-experimental data" (been observed).

↪ **Definition 2.3** (Statistics): Let $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} F$ where X_i a d -dimensional random vector. Let

$$T : \underbrace{\mathbb{R}^d \times \mathbb{R}^d \times \dots \times \mathbb{R}^d}_{n\text{-fold}} \rightarrow \mathbb{R}^k$$

be a borel-measurable function. Then, $T(X_1, \dots, X_n)$ is called a *statistic*, provided it does not depend on any unknown.

⊗ **Example 2.1:** $\bar{X}_n := \frac{1}{n} \sum_{i=1}^n X_i$ (the “sample mean”) and $S_n^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2$, (the “sample variance”) are both typical statistics.

↪ **Theorem 2.1:** Let $x_1, \dots, x_n \in \mathbb{R}$, then

- (a) $\operatorname{argmin}_{\alpha \in \mathbb{R}} \left\{ \sum_{i=1}^n (x_i - \alpha)^2 \right\} = \bar{x}_n$;
- (b) $\sum_{i=1}^n (x_i - \bar{x}_n)^2 = \sum_{i=1}^n (x_i^2) - n\bar{x}_n^2$;
- (c) $\sum_{i=1}^n (x_i - \bar{x}_n) = 0$.

↪ **Theorem 2.2:** Let $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} F$, and $g : \mathbb{R} \rightarrow \mathbb{R}$ borel-measurable such that $\operatorname{Var}(g(X)) < \infty$. Then,

- (a) $\mathbb{E} \left[\sum_{i=1}^n g(X_i) \right] = n \mathbb{E}[g(X_1)]$;
- (b) $\operatorname{Var} \left(\sum_{i=1}^n g(X_i) \right) = n \operatorname{Var}(X_1)$.

↪ **Theorem 2.3:** Let $X_1, \dots, X_n \stackrel{\text{iid}}{\sim} F$ with $\sigma^2 < \infty$, then

1. $\mathbb{E}[\bar{X}_n] = \mu$, $\operatorname{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$, $\mathbb{E}[S_n^2] = \sigma^2$.
2. If $M_{X_1}(t)$ exists in some neighborhood of 0, then $M_{\bar{X}_n}(t) = M_{X_1}\left(\frac{t}{n}\right)^n$, where it exists.