Adjustment and diagnosis

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Box-cox transformation

- Variation changes with the level of the series.
- Box-cox transformation:

$$w_t = \begin{cases} \log(y_t) & \text{if } \lambda = 0; \\ (y_t^{\lambda} - 1)/\lambda & \text{otherwise.} \end{cases}$$

- The logarithm is natural log.
- ullet Chose the proper λ

```
#BoxCox.lambda() choose a value of lambda for you.
lambda <- BoxCox.lambda(elec) # = 0.27
plot(BoxCox(elec,lambda))</pre>
```



Box-cox transformation

 Once a transformation is chosen, we need to forecast the transformed data

$$y_t = \begin{cases} \exp(w_t) & \text{if } \lambda = 0; \\ (\lambda w_t + 1)^{1/\lambda} & \text{otherwise.} \end{cases}$$

- If some $y_t \le 0$, no power transformation is possible unless all observations are adjusted by adding a constant to all values.
- Choose a simple value of λ .
- Transformations sometimes make little difference to the forecasts but have a large effect on prediction intervals.

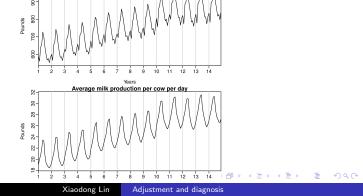


Calendar adjustments

 Some variation seen in seasonal data may be due to simple calendar effects

Monthly milk production per cow

 Calendar variation between the months due to different numbers of days in each month



Calendar adjustments

- Seasonal pattern becomes much cleaner after adjustment.
- Similar adjustments can be done on number of trading days/month ect.
- Population adjustments
- Inflation adjustments: consumer price index(CPI). $x_t = y_t/z_t \times z_{2000}$ gives the adjusted price at year 2000 dollar.

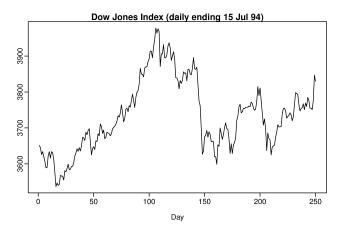
R code

```
plot(log(elec), ylab="Transformed electricity demand",
 xlab="Year", main="Transformed monthly electricity demand")
title(main="Log",line=-1)
# The BoxCox.lambda() function chooses a value of lambda.
lambda <- BoxCox.lambda(elec) # = 0.27</pre>
plot(BoxCox(elec,lambda))
monthdays \leftarrow rep(c(31,28,31,30,31,30,31,30,31,30,31),14)
monthdays [26 + (4*12)*(0:2)] < -29
par(mfrow=c(2,1))
plot(milk, main="Monthly milk production per cow",
  ylab="Pounds",xlab="Years")
plot(milk/monthdays,
main="Average milk production per cow per day",
  vlab="Pounds", xlab="Years")
```

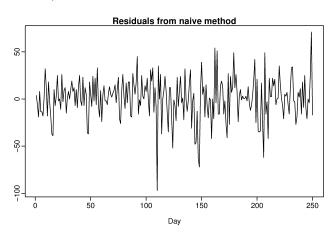
Residual diagnostics

- A residual in forecasting is $\epsilon_t = y_t \hat{y}_t$.
- The residuals are uncorrelated.
- The residuals have zero mean.
- We also need to examine the variance and distribution of the residuals

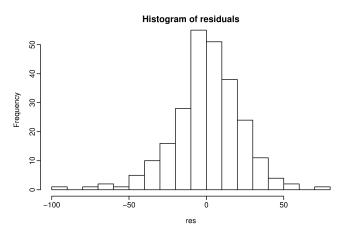
The daily DJ index for 250 days.



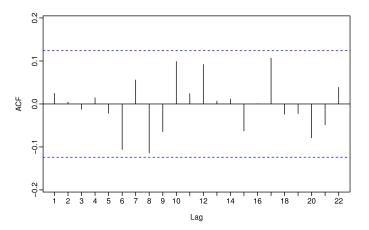
Residuals computed from the naive method



Histogram of the residuals



The ACF plot



R code

```
dj2 \leftarrow window(dj, end=250)
plot(dj2, main="Dow Jones Index (daily ending 15 Jul 94)",
  ylab="", xlab="Day")
res <- residuals(naive(dj2))
plot(res, main="Residuals from naive method",
  vlab="", xlab="Day")
Acf(res, main="ACF of residuals")
hist(res, nclass="FD", main="Histogram of residuals")
# lag=h and fitdf=K
> Box.test(res, lag=10, fitdf=0)
        Box-Pierce test
X-squared = 10.6425, df = 10, p-value = 0.385
> Box.test(res,lag=10, fitdf=0, type="Lj")
        Box-Ljung test
X-squared = 11.0729, df = 10, p-value = 0.3507
```