

Basic ARIMA models and stationarity

Time Series for Business

- **"All models are wrong, but some are useful." George E. P. Box**
- Models that give a rule for current or future observations based on past observations.
- We are concerned with a sequence of random variables that may be dependent.
- Our goal is to learn a set of possible models and make sensible model choices.
- Let's start with very simple models and discuss their statistical properties.

- Simplest time series model
 - No trend.
 - No seasonal variations.
 - Independent observations from the same distribution (iid).
- Distributionally, "iid-ness" implies

$$\begin{aligned}f(\mathbf{x}_1, \mathbf{x}_2, \dots, \mathbf{x}_n) &= f(\mathbf{x}_1)f(\mathbf{x}_2) \cdots f(\mathbf{x}_n) \\&= \prod_{i=1}^n f(\mathbf{x}_i).\end{aligned}$$

- Limitation: CANNOT BE USED FOR FORECASTING
- An i.i.d. mean zero Gaussian sequence is called Gaussian white noise: w_1, w_2, \dots, w_t .

An iid sequence

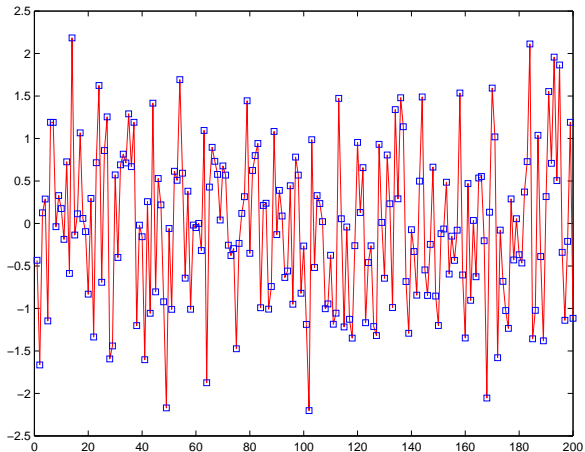


Figure: An iid sequence

Random walk

- How would you model the position of a walk along a straight line?

$$X_t = X_{t-1} + 1$$

- Now imagine a random walk, where you can take a step forward or backward with equal probability.

$$X_t = r_1 + r_2 + \cdots + r_t, \quad t = 1, 2, \dots$$

where r_t is i.i.d.

$$\Pr[r_t = 1] = \frac{1}{2} \quad \text{and} \quad \Pr[r_t = -1] = \frac{1}{2}$$

- This is a simple symmetric random walk.
 - What are the key differences from the i.i.d. sequence
 - What if we take the differencing.

Random walk

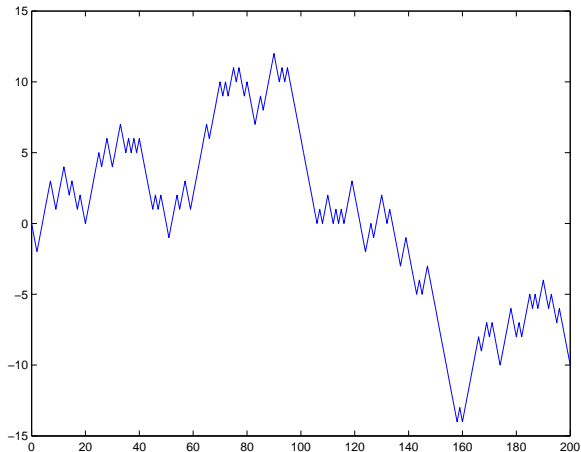


Figure: Simple symmetric random walk

Random walk with a drift

- $X_t - X_{t-1} = r_t$.
- r_t can be any random variables, for instance a Gaussian white noise w_t .
- Adding a drift

$$X_t = \delta + X_{t-1} + w_t.$$

- Assuming the starting position is zero, we have

$$X_t = \delta t + \sum_{i=1}^t w_i.$$

- Random walk acts on the previous step, what if we want to generalize it to several steps before? Note the difference on dependency.

Autoregressive and moving average model

- Autoregressive model (AR) is a class of models closely related to random walks.
- It is defined so that the current location is a linear combination of previous locations plus a random term

$$X_t = \phi_1 X_{t-1} + \phi_2 X_{t-2} + \cdots + \phi_p X_{t-p} + w_t.$$

This is an AR(p) model.

- Another related set of models are the moving average models

$$X_t = w_t + \theta_1 w_{t-1} + \cdots + \theta_q w_{t-q}.$$

This is an MA(q) model. It takes sliding window and take a weighted average of the white noises within the window.

Mean and autocovariance functions

- The mean function of $\{X_t\}$ is

$$\mu_X(t) = \mathbb{E}(X_t).$$

- The variance function of $\{X_t\}$ is

$$\sigma_X^2(t) = \mathbb{V}(X_t) = \mathbb{E}[(X_t - \mu_X(t))^2].$$

- The covariance function of $\{X_t\}$ is

$$\begin{aligned}\gamma_X(s, t) &= \text{cov}(X_s, X_t) \\ &= \mathbb{E}[(X_s - \mu_X(s))(X_t - \mu_X(t))].\end{aligned}$$

Mean and autocovariance functions

- The mean function for random walk with drift is

$$\begin{aligned}\mu_t &= E(X_t) = E\left(\delta t + \sum_{i=1}^t w_i\right) \\ &= E(\delta t) + E\left(\sum_{i=1}^t w_i\right) = \delta t + \sum_{i=1}^t E(w_i) = \delta t.\end{aligned}$$

- How about the mean functions for AR(p) and MA(q) model?
- Exercise: variance and autocovariance function for an MA(2) model.
- Autocorrelation function

$$\rho(s, t) = \frac{\gamma(s, t)}{\sqrt{\gamma(s, s)\gamma(t, t)}}.$$

- A time series model for the observed data $\{\mathbf{x}_t\}$ is a specification of the joint distributions (or possibly only the means and covariances) of a sequence of random variables $\{X_t\}$ of which $\{\mathbf{x}_t\}$ is a realization.
- In traditional statistics: Random sampling procedures enable us to obtain replicated observations under identical conditions. Besides, these observations are independent.
- For time series: We have only a single realization at each time point and also dependent over time. More precisely, it is a sample of size one.
- For any inference to be possible, we must recreate some notion of replicability.

- A time series $\{X_t, t = 0, \pm 1, \pm 2, \dots\}$ is said to be stationary if it has statistical properties similar to those of the "time-shifted" series $\{X_{t+h}, t = 0, \pm 1, \pm 2, \dots\}$, for each integer h .

Definition

A time series $\{X_t\}$ is said to be strongly or strictly stationary if the joint density functions depend only on the relative location of the observations, so that

$$f(x_{t_1+h}, x_{t_2+h}, \dots, x_{t_k+h}) = f(x_{t_1}, x_{t_2}, \dots, x_{t_k}),$$

meaning that $(X_{t_1+h}, X_{t_2+h}, \dots, X_{t_k+h})$ and $(X_{t_1}, X_{t_2}, \dots, X_{t_k})$ have the same joint distributions for all h and for all time points $\{t_i\}$.

Stationarity

- Strong stationarity is too strong to be practically useful. Besides, specifying the densities $f(x_{t_1}, x_{t_2}, \dots, x_{t_k})$ is usually very complicated.

Definition

A time series $\{X_t\}$ is weakly stationary if

- (i) $\mu_X(t)$ is independent of t , ie $\mu_X(t) = \mu_X$ for all t and finite.
- (ii) $\sigma_X^2(t)$ is finite.
- (iii) $\gamma_X(t+h, t)$ is independent of t for each h .

- Weak stationarity is also referred to as second-order stationarity, or covariance stationarity. From now on, we shall say stationary to mean weakly stationary.
- Stationarity allows the re-creation of the notion of replicability that is crucial to statistical inference.

Intuition and properties

- Location does not matter - ONLY distance. The sequence consists of identically distributed random variables.
- For autocorrelation function, we have

$$\rho(h) = \frac{\gamma(h)}{\gamma(0)}.$$

- Also when time series is stationary,

$$\gamma(h) = \gamma(-h)$$

and similarly

$$\rho(h) = \rho(-h).$$

- Without stationary, we have little hope of estimating the full $\gamma(s, t)$. With stationarity, we now have many observations that are h apart from one another (when $h \ll T$). Thus inference is possible for stationary processes.

- Let's first note that with stationarity, the (true) mean is constant. We can therefore estimate the mean using the sample mean

$$\bar{x} = \frac{\sum_{t=1}^T x_t}{T}.$$

- Now, let's look at the sample autocovariance function.

$$\hat{\gamma}(h) = \frac{\sum_{t=1}^{T-h} (x_{t+h} - \bar{x})(x_t - \bar{x})}{T}.$$

for $h = 0, 1, \dots, T - 1$.

- To get the sample autocorrelation we simply scale by the variance.

$$\hat{\rho}(h) = \frac{\hat{\gamma}(h)}{\hat{\gamma}(0)}$$

- An important property for the sample autocorrelation function is that when the true model is white noise $\hat{\rho}(h)(h = 1, 2, \dots, H)$ is approximately normally distributed with zero mean and standard deviation of $1/\sqrt{T}$. Since by central limit theorem,

$$\frac{\bar{x} - \mu}{\sigma/\sqrt{T}}$$

is approximately normal for large T .

Examples of stationary processes

- **The iid noise process**

If $\{X_t\}$ is an iid noise process, we write $\{X_t\} \sim \text{IID}(0, \sigma^2)$

We also have,

$$\gamma_X(t+h, t) = \begin{cases} \sigma^2, & \text{if } h = 0; \\ 0, & \text{if } h \neq 0. \end{cases}$$

- **The white noise process** Let $\{X_t\}$ be a sequence of

- Uncorrelated random variables, ie $\gamma_X(h) = 0$ for $h \neq 0$.
- Each variable having zero mean, ie $\mathbb{E}(X_t) = 0$.
- Each variable having finite variance, ie $\mathbb{V}(X_t) = \sigma^2 < \infty$.

Such a sequence is referred to as white noise (with mean 0 and variance σ^2), and indicated by $\{X_t\} \sim \text{WN}(0, \sigma^2)$.

Examples of stationary processes

- $\{X_t\} \sim \text{WN}(0, \sigma^2)$ is clearly a stationary process.
- The covariance function of $\{X_t\} \sim \text{WN}(0, \sigma^2)$ is the same as that of $\text{IID}(0, \sigma^2)$, namely

$$\gamma_X(t+h, t) = \begin{cases} \sigma^2, & \text{if } h = 0; \\ 0, & \text{if } h \neq 0. \end{cases}$$

- **The random walk process**

$$X_t = w_1 + w_2 + \cdots + w_t$$

where $w_t \sim \text{WN}(0, \sigma^2)$.

- $E[X_t] = ?$ and $V[X_t] = ?$
- If $s > t$ then $\text{cov}(Z_s, X_t) = ?$ $\gamma_X(t+h, t) = ?$
- Is the series $\{X_t\}$ stationary?

First order AR process

- A series $\{X_t\}$ is a first-order autoregressive or AR(1) process if

$$X_t = \phi X_{t-1} + w_t \quad t = 0, \pm 1, \pm 2, \dots$$

where

- $\{w_t\} \sim \text{WN}(0, \sigma^2)$,
 - $|\phi| < 1$
 - w_t is uncorrelated with X_s for each $s < t$.
- It is easy to show that the autocovariance function (ACVF) is

$$\gamma_X(h) = \sigma^2 \frac{\phi^{|h|}}{1 - \phi^2}, \quad h = 0, \pm 1, \pm 2, \dots$$

The autocorrelation function (ACF) of an AR(1) is given by

$$\rho_X(h) = \phi^{|h|} \quad h = 0, \pm 1, \pm 2, \dots$$

First order AR process

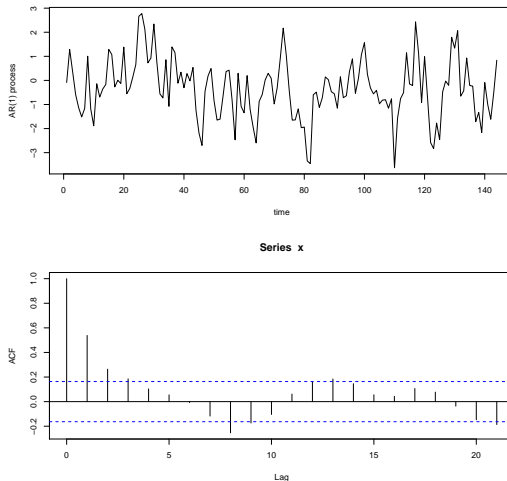


Figure: AR(1) with $\phi = 0.5$.

First order MA process

- A series $\{X_t\}$ is a first-order moving average or MA(1) process if

$$X_t = w_t + \theta w_{t-1}, \quad t = 0, \pm 1, \pm 2, \dots$$

where

- $\{w_t\} \sim \text{WN}(0, \sigma^2)$
 - θ is a real-valued constant.
- It is easy to show that the autocovariance function (ACVF) is

$$\gamma_X(t+h, t) = \begin{cases} \sigma^2(1 + \theta^2), & \text{if } h = 0; \\ \sigma^2\theta, & \text{if } h = \pm 1; \\ 0, & \text{if } |h| > 1 \end{cases}$$

- Clearly, an MA(1) process is a stationary process.
- Also, the autocorrelation function (ACF) is

$$\rho_X(t+h, t) = \rho_X(h) = \begin{cases} 1, & \text{if } h = 0; \\ \theta/(1 + \theta^2), & \text{if } h = \pm 1; \\ 0, & \text{if } |h| > 1, \end{cases}$$

First order MA process

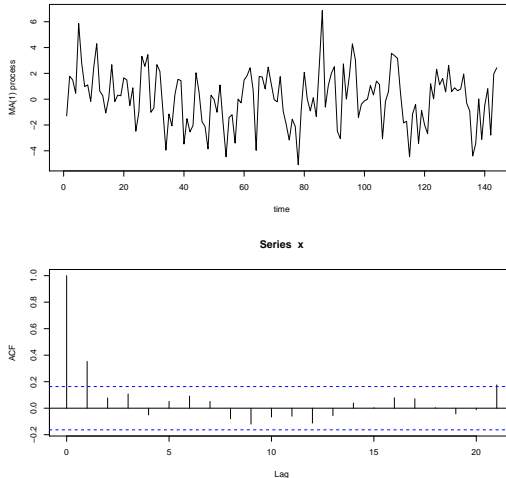


Figure: MA(1) with $\theta = 2$.

A Test for IID noise using the sample (ACF)

- For iid noise with finite variance, we have, for $h \neq 0$,

$$\hat{\rho}(h) \sim N(0, \frac{1}{T})$$

- Steps of the diagnostic for iid noise
 - Plot the lag h versus $\hat{\rho}(h)$.
 - Draw two horizontal lines at $\pm 1.96/\sqrt{T}$. *These two lines are drawn automatically in R*
 - You should have about 95% of the the values of $\{\hat{\rho}(h) : h = 1, 2, \dots\}$ within the lines *if the noise is indeed iid.*

A Test for IID noise using the sample (ACF)

Which of the two depicts an IID noise?

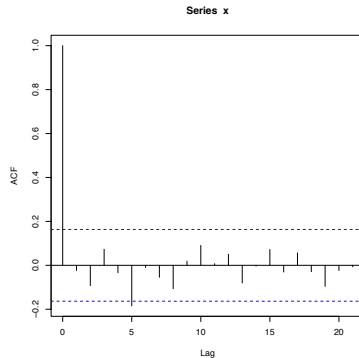
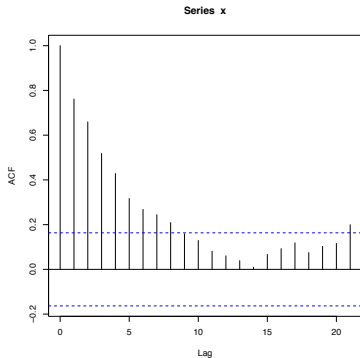


Figure: Which is more likely i.i.d?