

Bayesian Model Selection

Globular cluster mass profiles

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Testing lowered isothermal models with direct N -body simulations of globular clusters

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LIMPEY

Data:

- 3D Position (x,y,z \rightarrow r)
- 3D Velocity (xy, xz, yz \rightarrow v)

Model parameters:

- m (mass)
- r (half light radius)  scale
- g (truncation energy)  shape
- phi (potential)  gravitational potential
- r_a (anisotropy radius
can \rightarrow infinity, become isotropic)

Assume we have the complete dataset of the cluster's dynamic information (position and velocity)

Goal: Get the estimated of true parameters, posterior distribution
And, Which model we should choose?



use limpey generate dataset

Model Selection

1. fit isotropic data with the **correct** isotropic model
2. fit isotropic data with a more expansive anisotropic model
3. fit anisotropic data with the incorrect isotropic model
4. fit anisotropic data with the **correct** anisotropic model (the anisotropy does not go to 0)

Data X: true parameter {m5r3g1.5phi3.0}

Steps:

1. Assume the models

Model1: isotropic (expected to prefer)

Model2: anisotropic

2. Set up Bayes Method (Why not p-value?)

$$\text{Posterior Prob} = P(\Theta|X) = \frac{P(\Theta)P(X|\Theta)}{P(X)} = \frac{\text{Prior} \cdot \text{Likelihood}}{\text{Evidence}}$$

$$P(X|\Theta) = \prod_i^{N_*} \Lambda_i(X_i|\Theta)$$

N #of stars

$$\Lambda_i = \frac{f(X_i|\Theta)}{\text{Normalizing term}}$$

f distribution function in LIMPEY

$$f(E, J^2) = A \exp\left(-\frac{J^2}{2r_a^2 s^2}\right) E_\gamma\left(g, \frac{\phi(r_t) - E}{s^2}\right)$$

For Model Comparison, We evaluate $K =$

$$\frac{P(M1|X)}{P(M2|X)} = \boxed{\frac{P(X|M1)}{P(X|M2)}} \frac{P(M1)}{P(M2)}$$

A value of $\log K > 1$ means that M1 is more strongly supported by the data than M2

3. Calculate Likelihood/ Choice of Sampling Methods

$$\text{Marginal likelihood} = P(X|M) = \int P(X|\theta, M)P(\theta|M)d\theta = \underline{E_{\text{prior}}[P(X|M)]}$$

Naive Monte Carlo Estimator

$$\hat{p}(X|\theta) = \frac{1}{N} \sum_{i=1}^N p(X|\hat{\theta}_i), \text{ where } \hat{\theta}_i \sim p(\theta) \text{ prior distribution}$$

- Importance Sampling

$$\int p(X|\theta)p(\theta)d\theta = \int p(X|\theta)p(\theta)\frac{g(\theta)}{g(\theta)}d\theta = E_g\left(\frac{p(X|\theta)p(\theta)}{g(\theta)}\right)$$

$$\hat{p} = \frac{1}{N} \sum_i \frac{p(X|\hat{\theta})p(\hat{\theta})}{g(\hat{\theta})}$$

drawing sample from
proposal g instead of prior

importance weight = posterior/ proposal

Application: Estimation of Euler's constant

$$\gamma = \int_1^{\infty} \left(\frac{1}{[x]} - \frac{1}{x} \right) dx$$

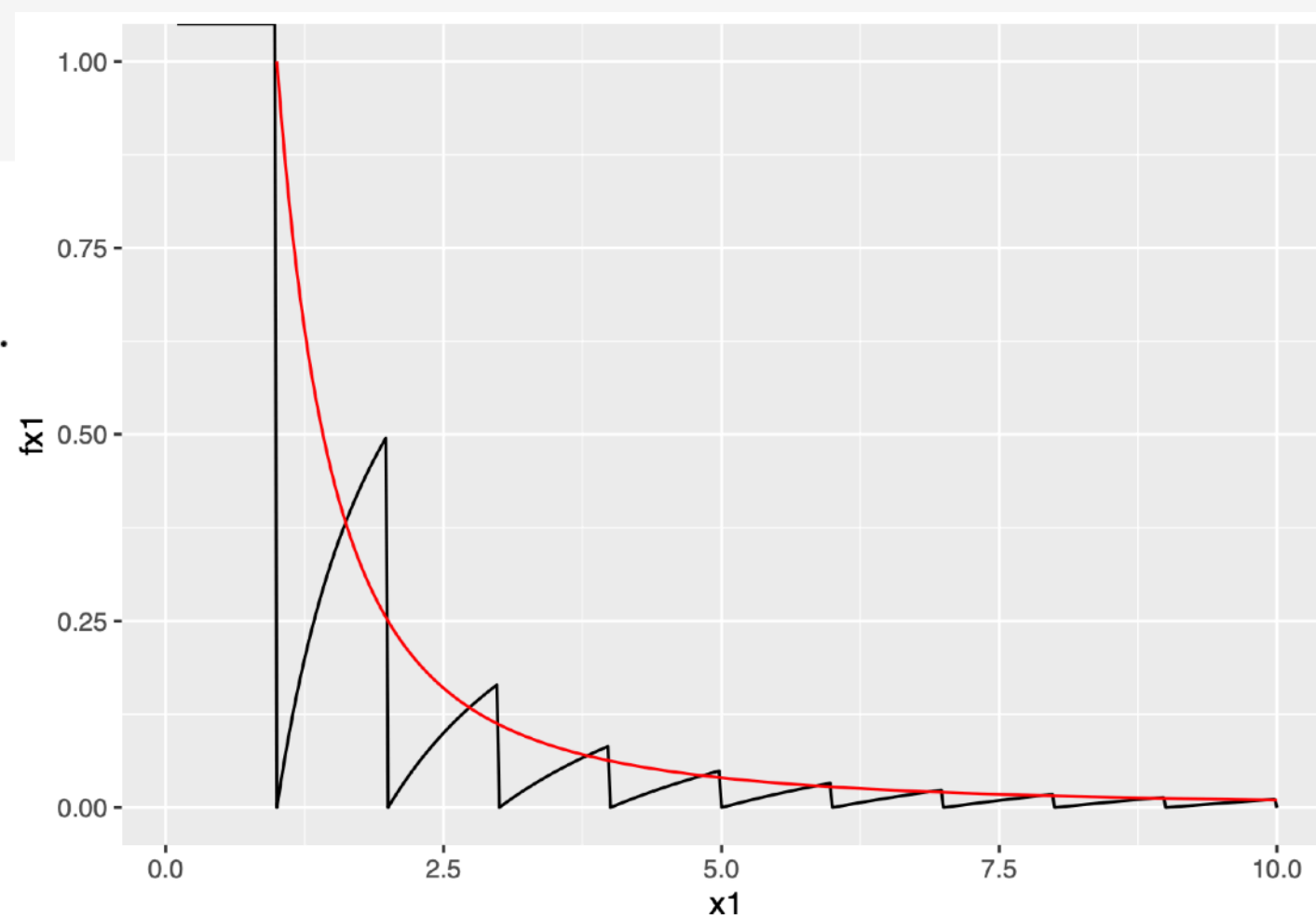
Let importance density $g(x) = x^{-2}$

$$= \int_1^{\infty} x^2 \left(\frac{1}{[x]} - \frac{1}{x} \right) x^{-2} dx$$
$$= \int_1^{\infty} x^2 \left(\frac{1}{[x]} - \frac{1}{x} \right) g(x) dx$$

```
N=10000000
gammalist = c(rep(0,N))
for (i in 1:N){
  unif = runif(1)
  x <- 1/(1-unif)
  gammalist[i] = (x^2)*(1/floor(x)-1/x)
}
mean(gammalist)
```

```
## [1] 0.5769798
```

The true value of Euler constant is 0.57721566...



Optimization, get parameter estimates

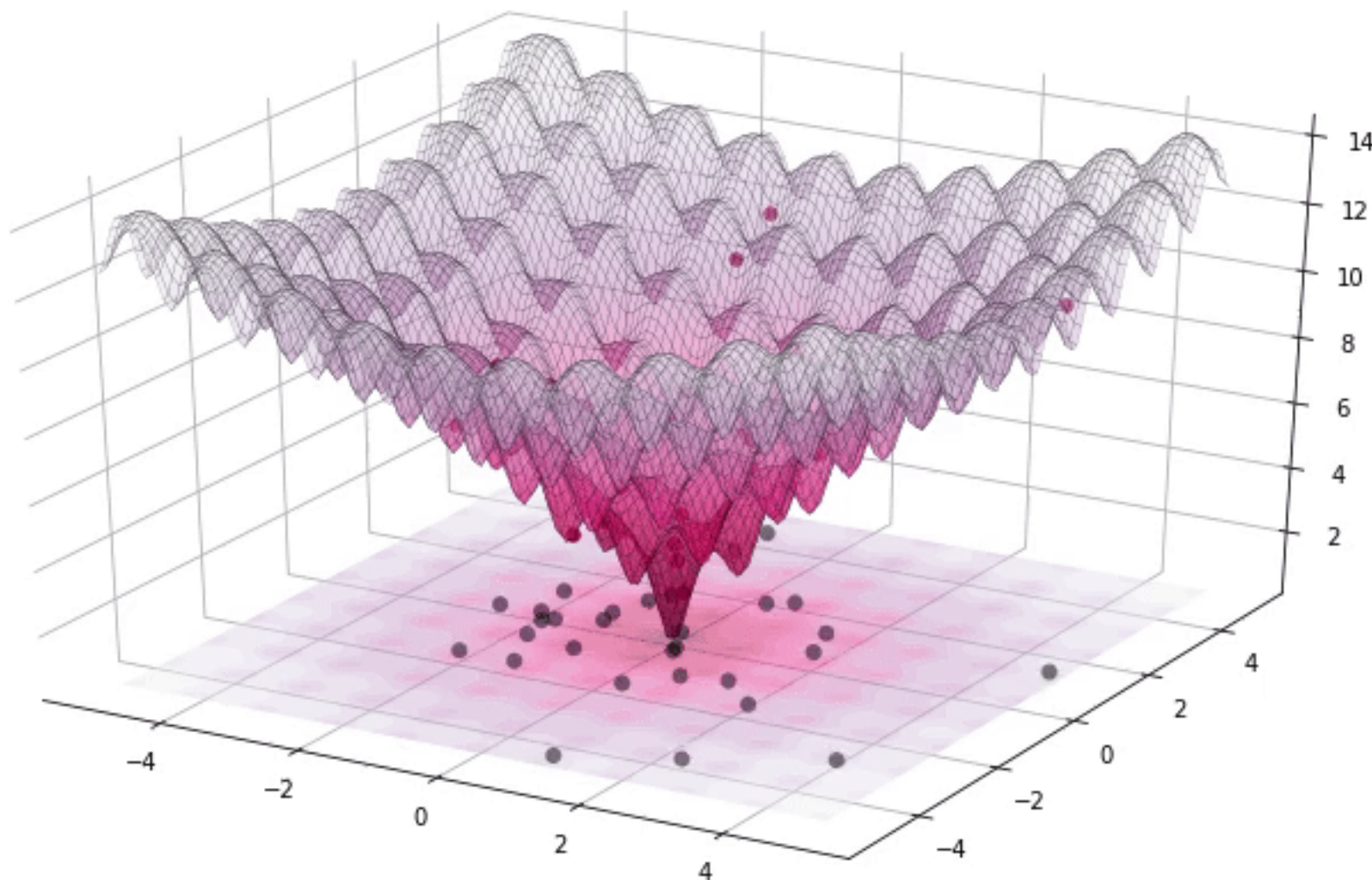
Max posterior =

Min {Negative of the log-posterior function of the input parameters}

There are many optimization methods like gradient/coordinate descent

Here we use **Differential Evolution**

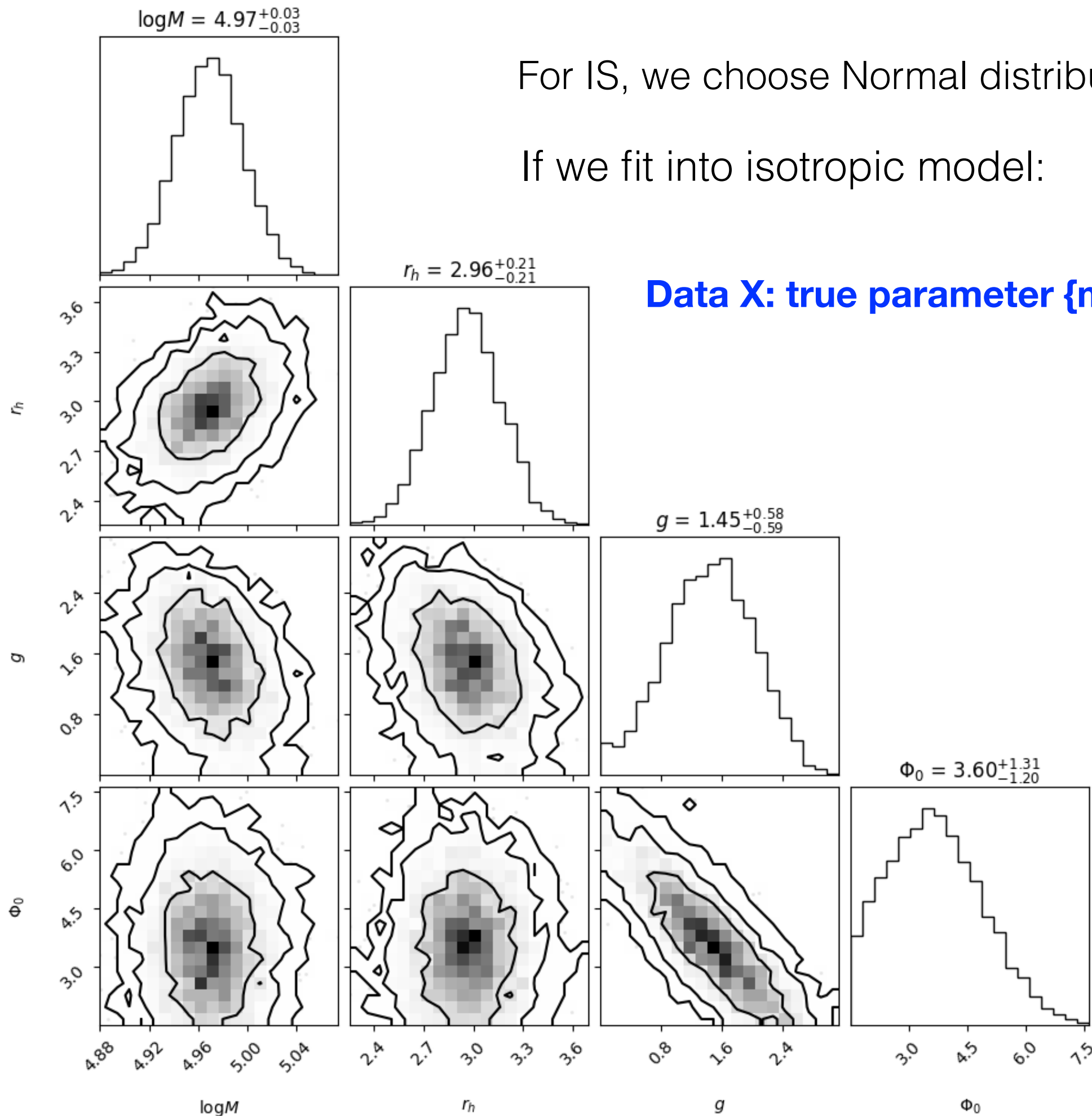
Target function not required to be differential (not even continuous)



For IS, we choose Normal distribution as proposal

If we fit into isotropic model:

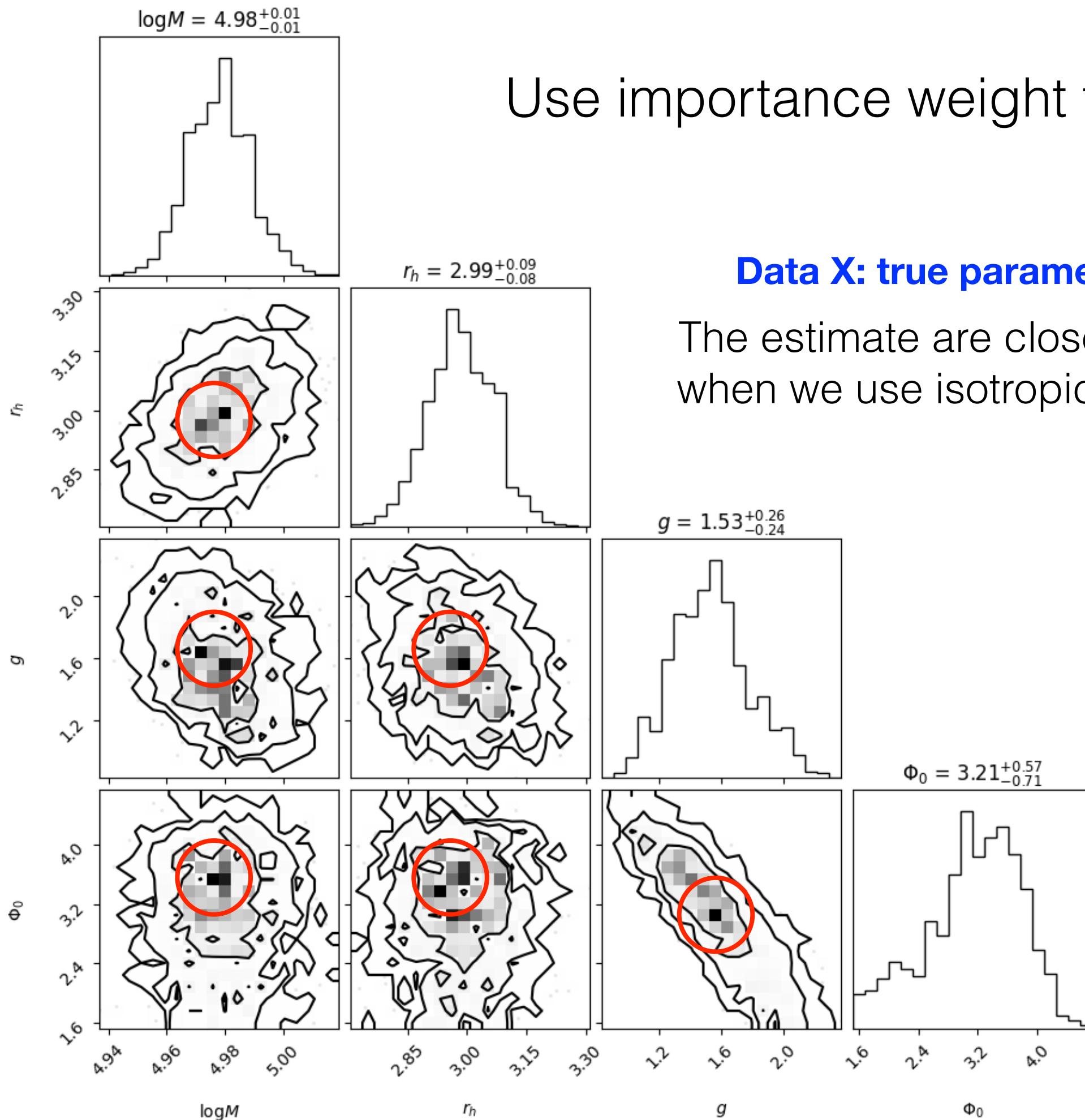
Data X: true parameter {m5r3g1.5phi3.0}



Use importance weight to resample:

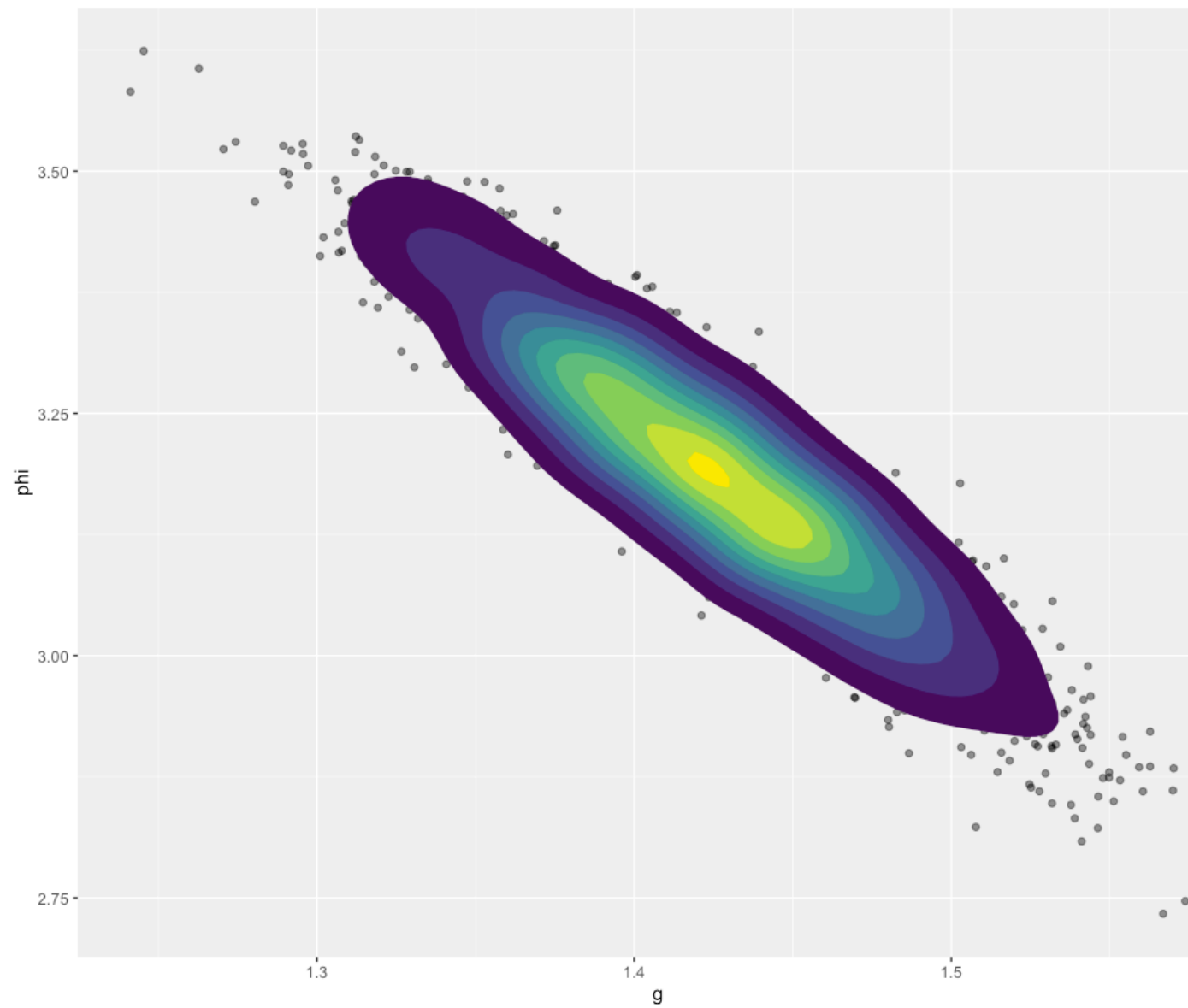
Data X: true parameter {m5r3g1.5phi3.0}

The estimate are close to true parameter when we use isotropic model

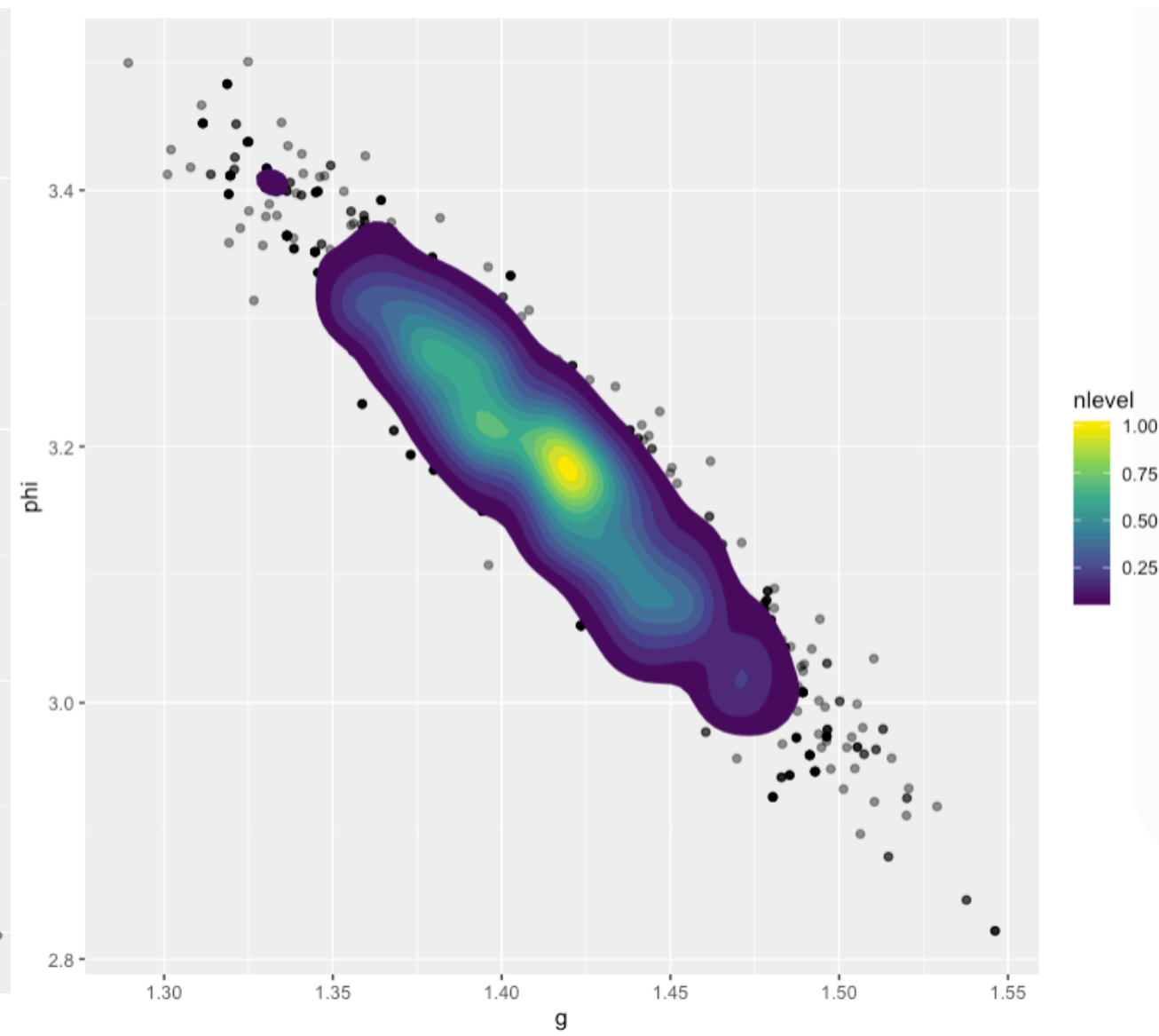


g vs phi

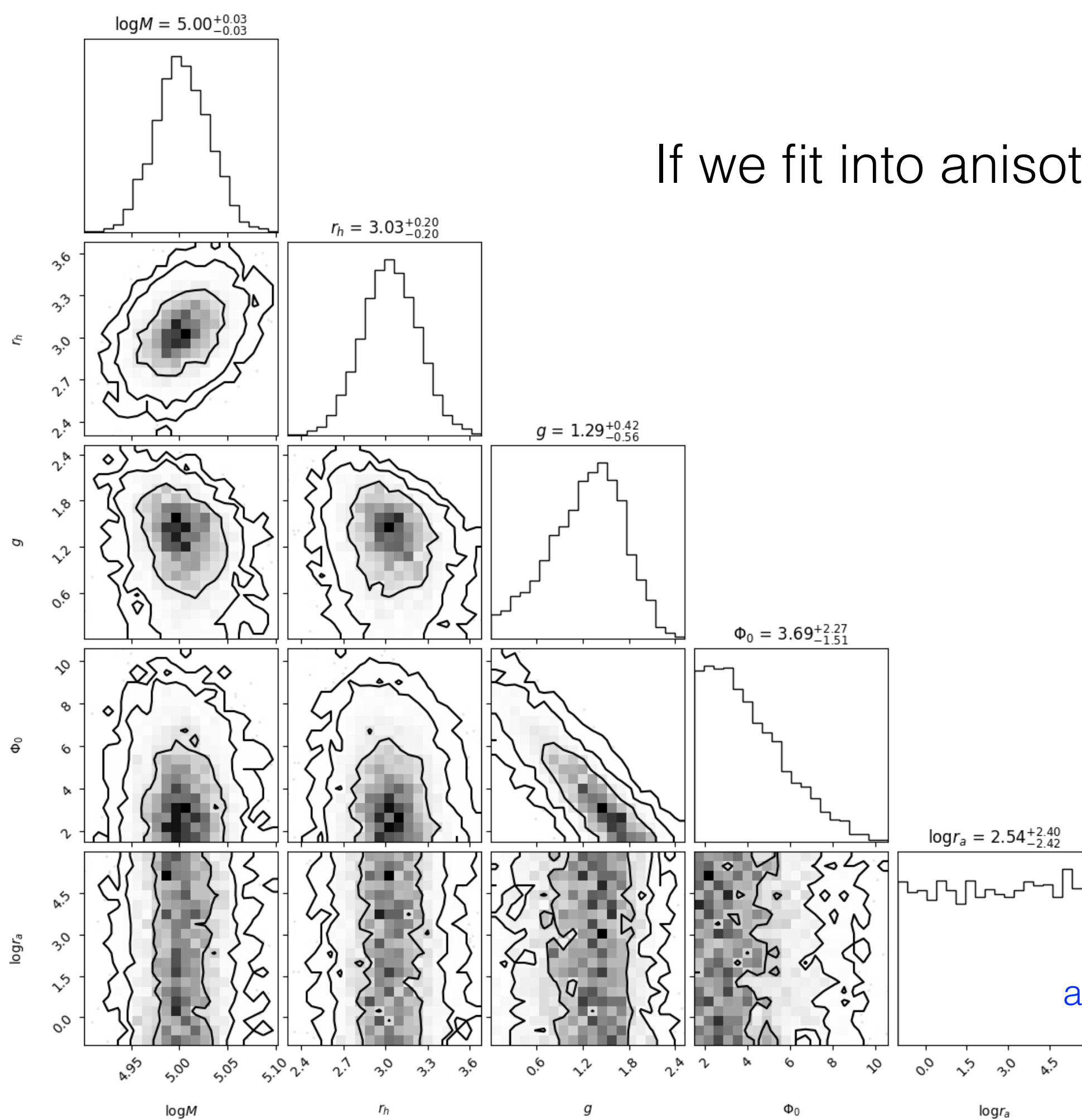
IS



IS with weight



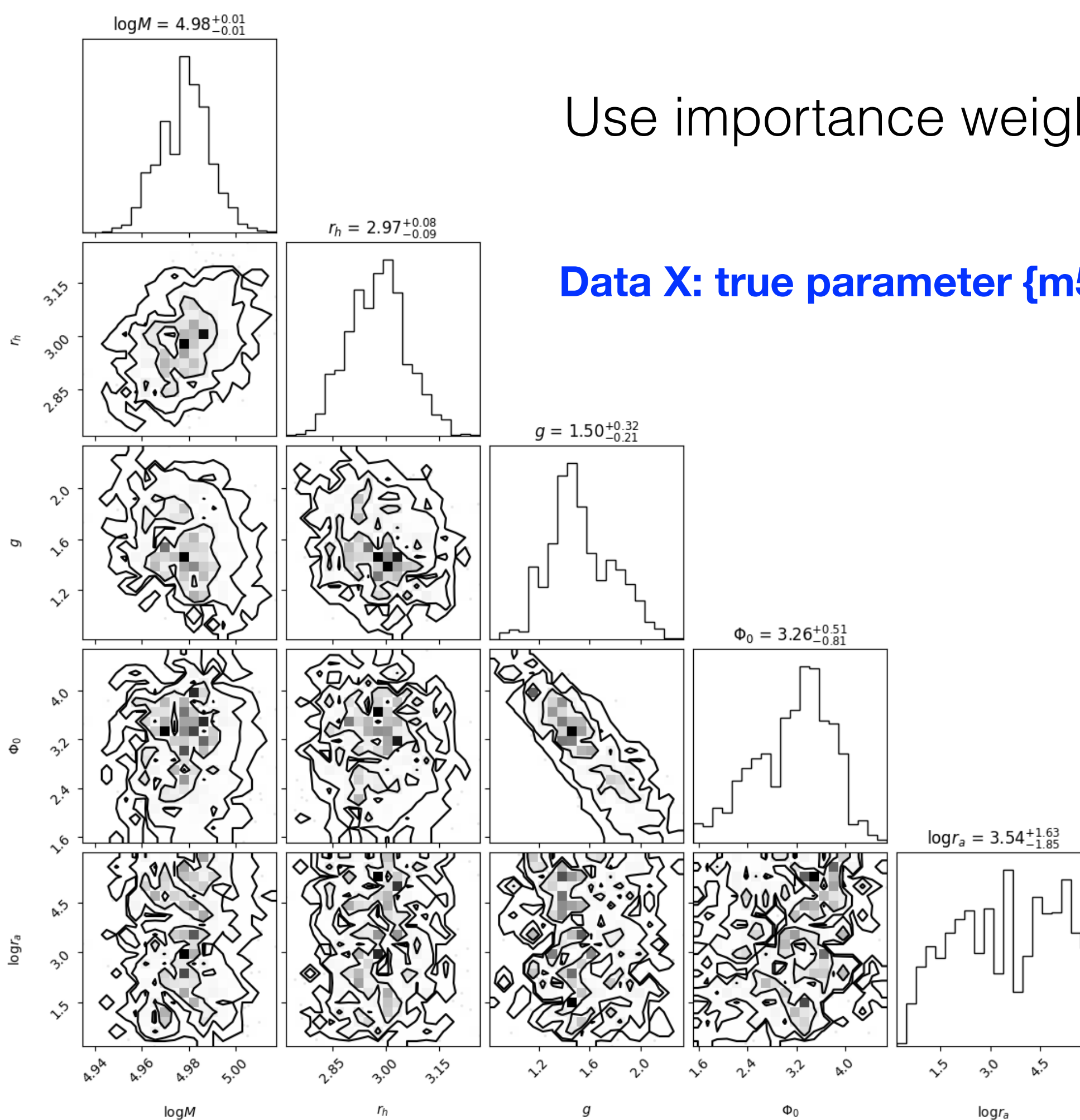
If we fit into anisotropic model



additional parameter r_a

Use importance weight to resample

Data X: true parameter {m5r3g1.5phi3.0}



what's the estimate for r_a ?

#Effective sample size
 (“exchange rate”
 btw dependent & independent samples)

$$n_{\text{eff}} \equiv \frac{(\sum_{i=1}^n w_i)^2}{\sum_{i=1}^n w_i^2}$$

Model 1: 305.5455 Total 2500 samples

Model 2: 113.3407 (proposal distribution not good enough)

$$\text{Bayes Factor } \log K = \frac{P(M1|X)}{P(M2|X)} = \frac{P(X|M1)}{P(X|M2)} \frac{P(M1)}{P(M2)} = 2.41$$

$\log_{10} K$	K	Strength of evidence
0 to 1/2	1 to 3.2	Not worth more than a bare mention
1/2 to 1	3.2 to 10	Substantial
1 to 2	10 to 100	Strong
> 2	> 100	Decisive

As expected, we prefer Model 1

Markov Chain Monte Carlo (MCMC)

Markov Chain: Time homogenous property

$$\Pr(X_{n+1} = x \mid X_1 = x_1, X_2 = x_2, \dots, X_n = x_n) = \Pr(X_{n+1} = x \mid X_n = x_n)$$

MCMC: By constructing a Markov Chain such that it will converge to its desired stationary distribution

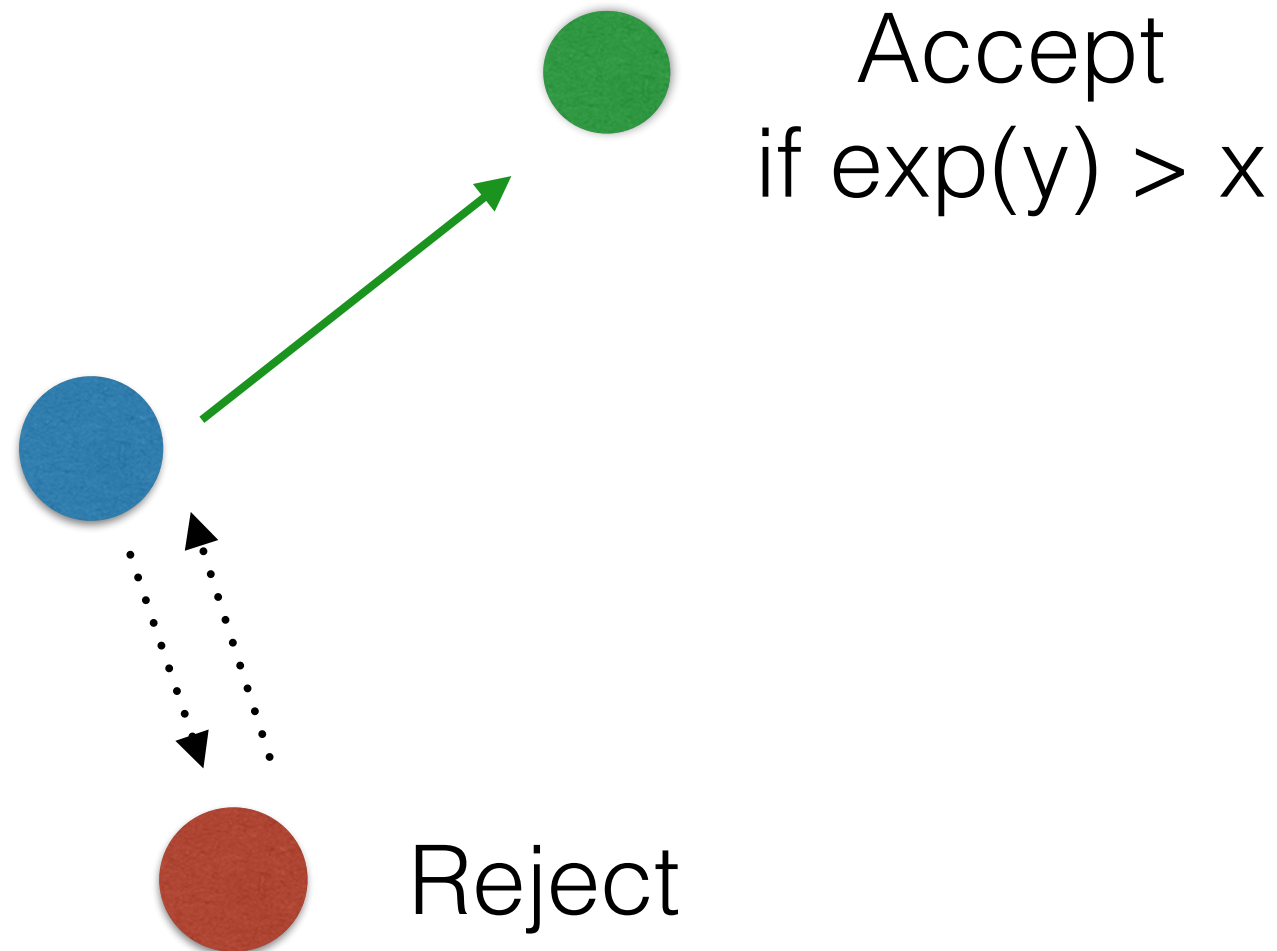
Ex: Metropolis Hasting Algorithm

In Bayesian, the samples generated by MCMC can be used to evaluate the integral over high dimensional variable

Same Goal: Calculate the likelihood

1. Calculate difference of logs of likelihood*prior between current and new proposed points = y
2. Generate $X \sim \text{Unif}[0,1]$ independently

Detailed Options:
Thinning
Finite Adaptive MCMC



if return 0 probability from prior
if return $-\text{Inf}$ from the log likelihood
if $\exp(y) < x$

Data X: true parameter {m5r3g1.5phi3.0}

100 chain, each chain 500 samples

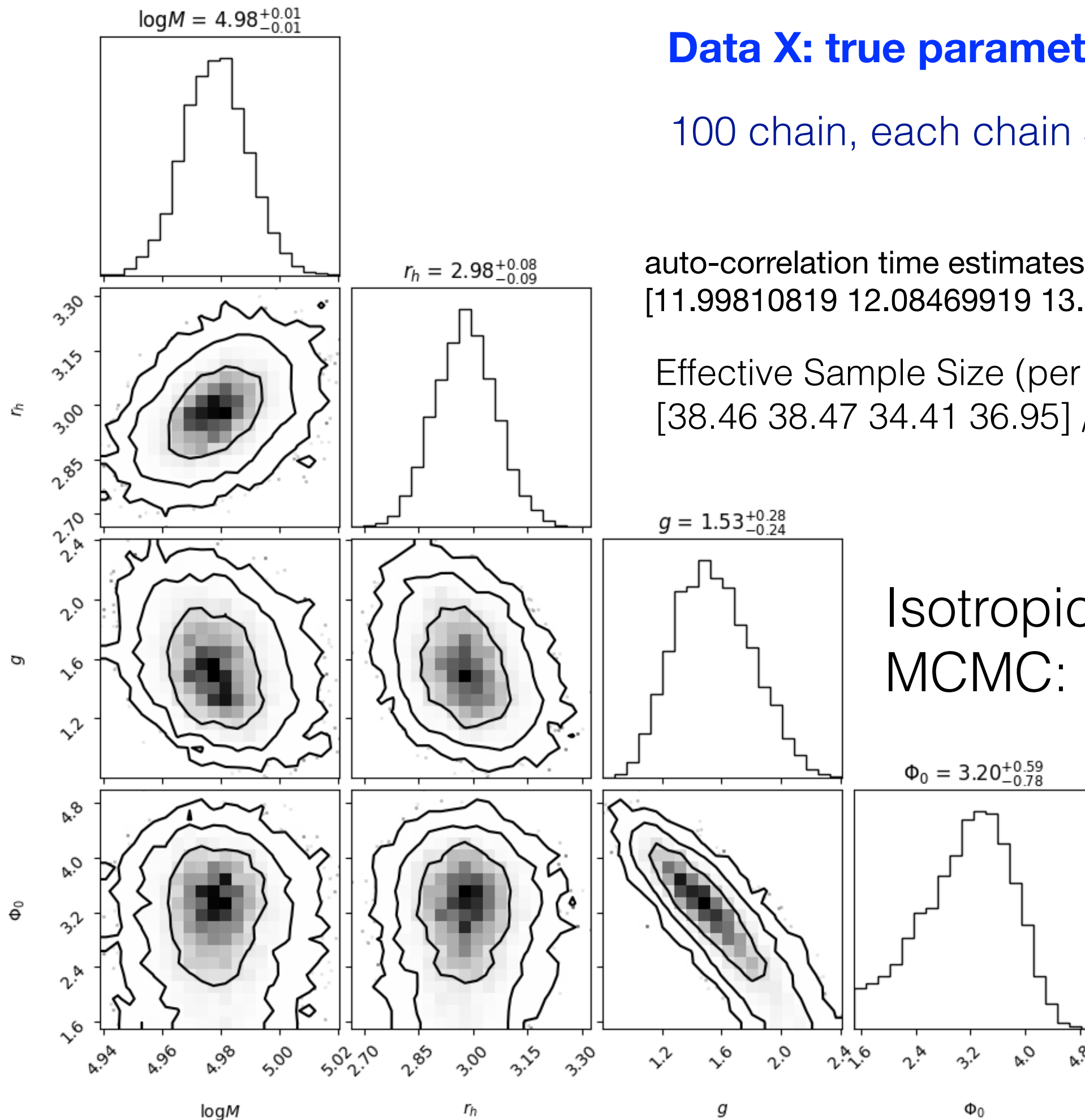
auto-correlation time estimates:

[11.99810819 12.08469919 13.53103676 12.53233813]

Effective Sample Size (per chain)

[38.46 38.47 34.41 36.95] /500 samples

Isotropic Model
MCMC: Good Estimate



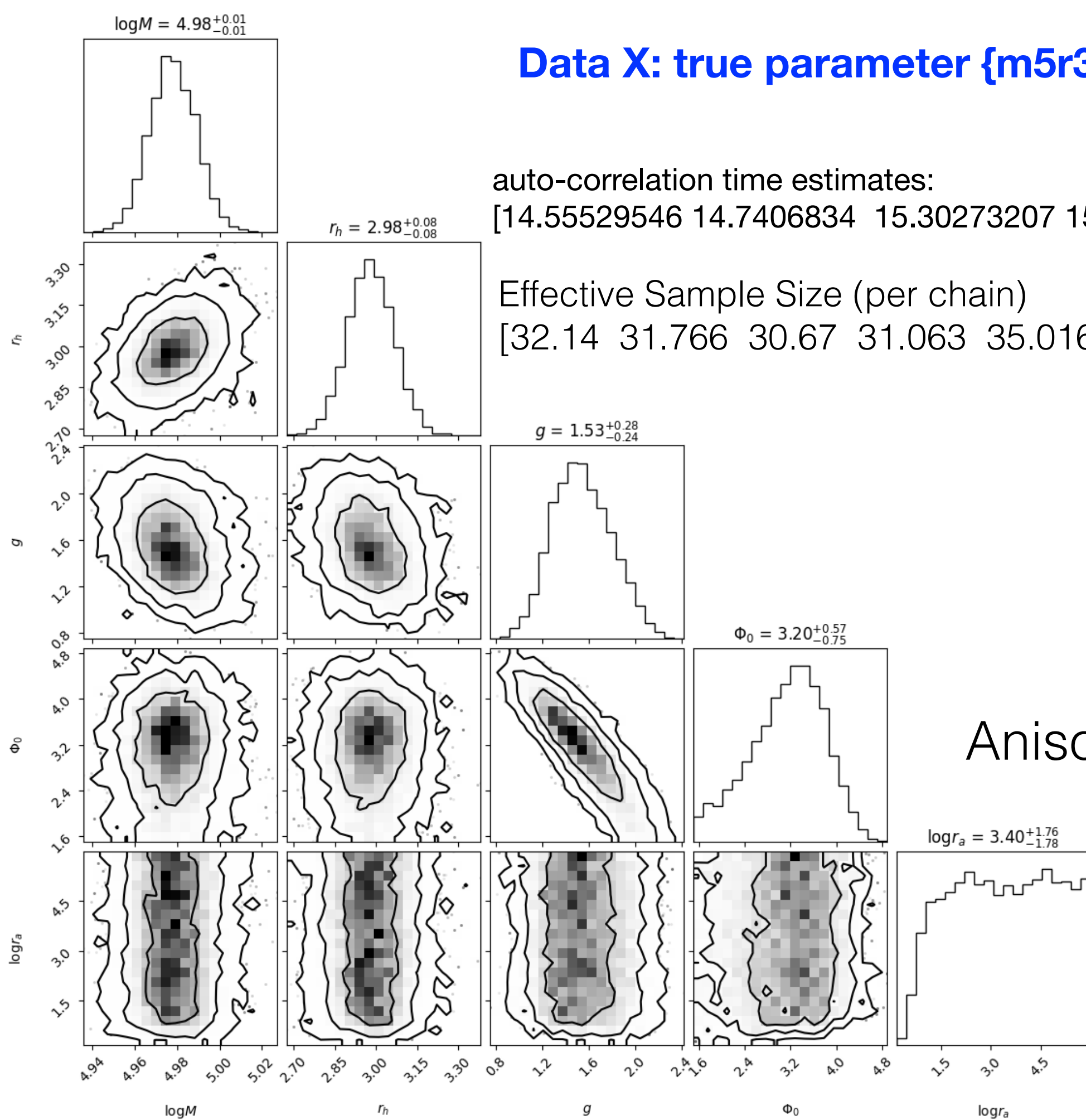
Data X: true parameter {m5r3g1.5phi3.0}

auto-correlation time estimates:

[14.55529546 14.7406834 15.30273207 15.09608124 13.27906024]

Effective Sample Size (per chain)

[32.14 31.766 30.67 31.063 35.016] /500 samples

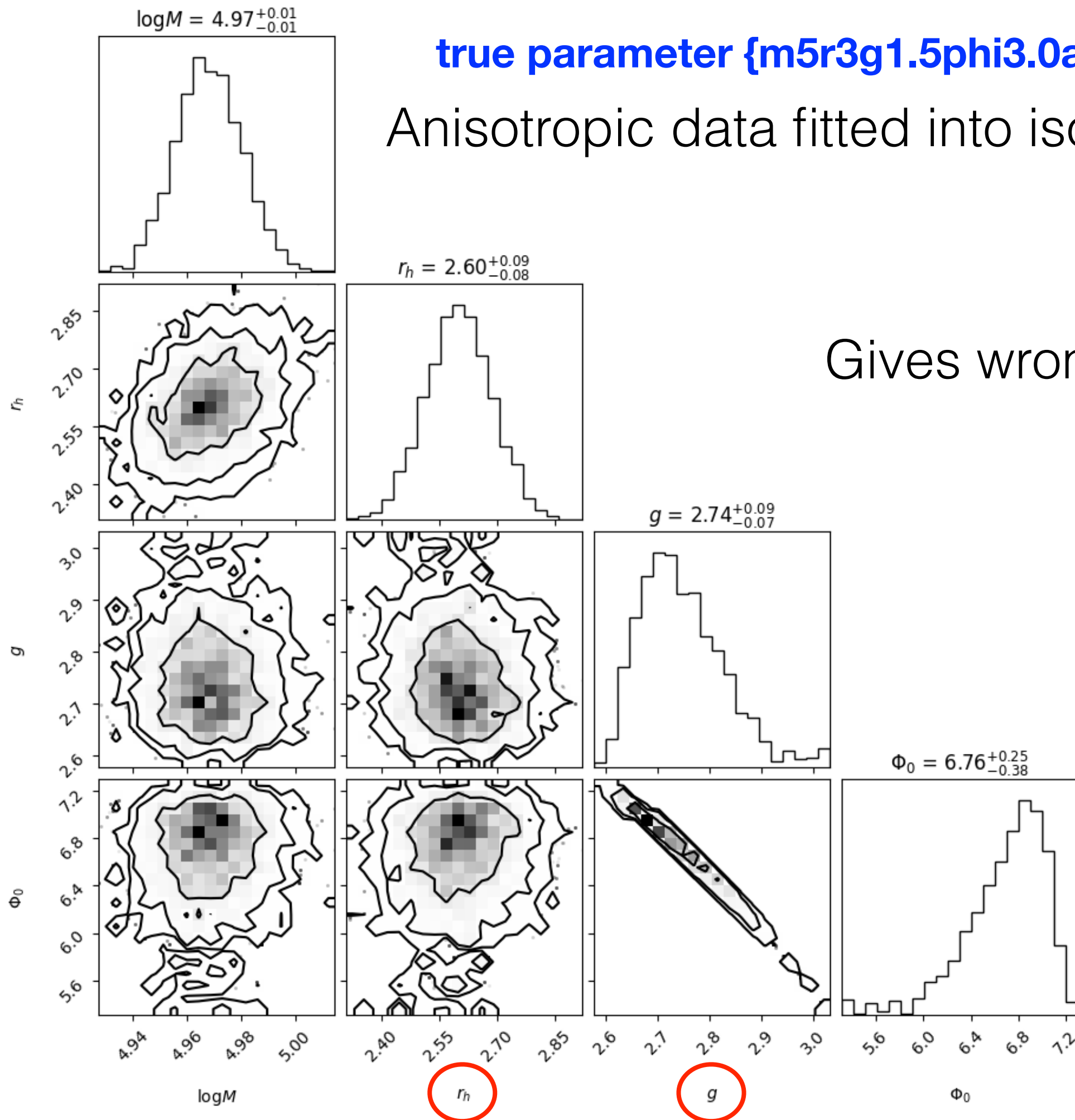


Anisotropic Model

true parameter {m5r3g1.5phi3.0a0.8}

Anisotropic data fitted into isotropic model

Gives wrong estimate

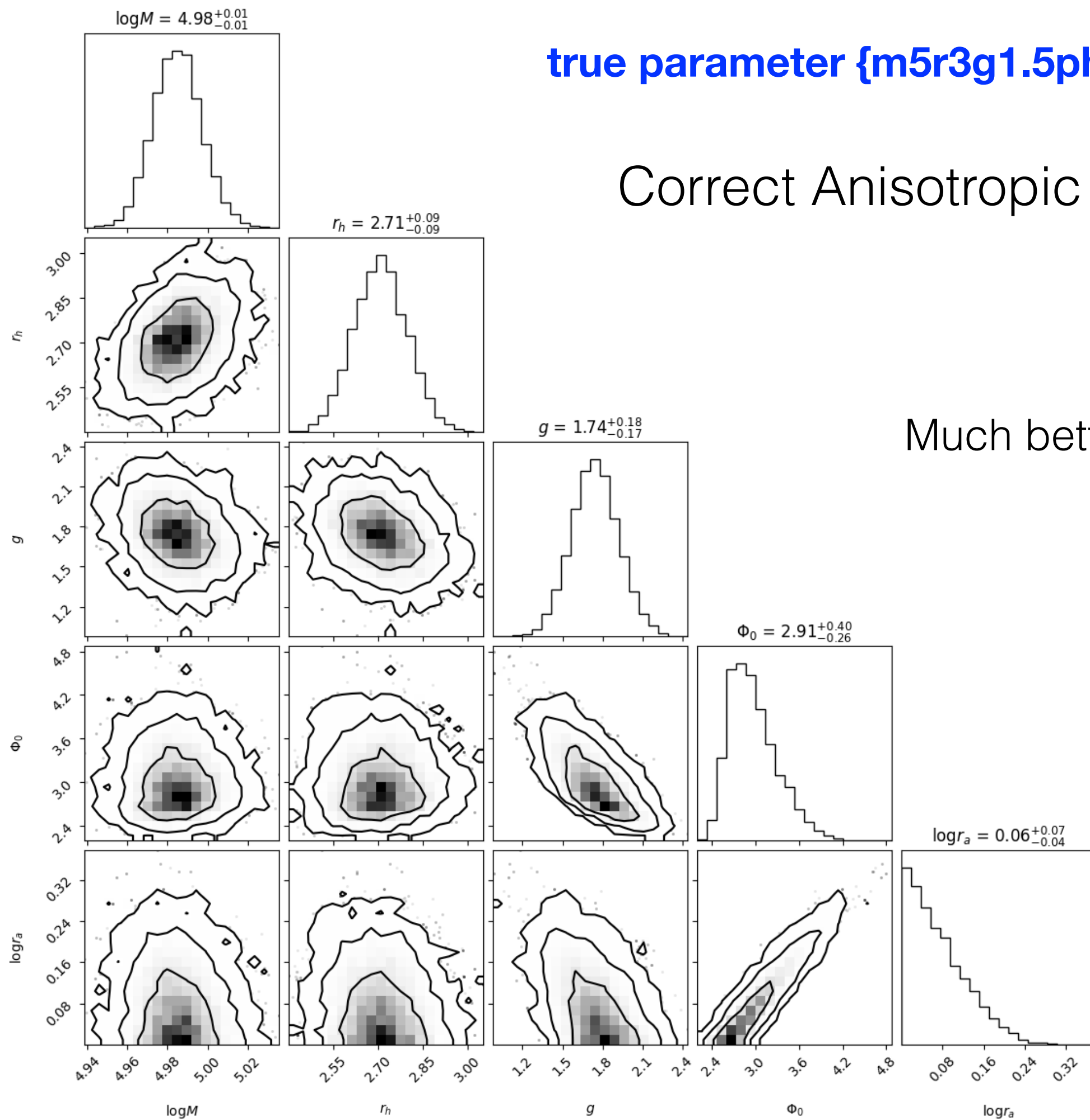


Missing r_a

true parameter {m5r3g1.5phi3.0a0.8}

Correct Anisotropic Model

Much better estimates



Next Step:

1. Try other sampling methods, such as Bridge Sampling, Nested Sampling
2. Consider the case when we don't have complete dataset (Projection/ missing positions)
3. Error Measurements/ Uncertainty in the model