

武汉大学 2021-2022 学年第二学期

《概率论与数理统计B》 期末试题 (A 卷)

一、(12分) 设 A 、 B 为随机事件, $P(A)=0.4, P(A \cup B)=0.7$ 。

(1) 若 A 和 B 互不相容, 求 $P(B)$; (2) 若 A 和 B 相互独立, 求 $P(B)$ 。

二、(12分) 袋中装有 3 个白球 2 个黑球, 每次取一个球, 取后放回并再放入 1 个与取出的球同颜色的球。求

(1) 第三次取出的球是白球的概率;

(2) 如果第三次取出的是白球, 求第一次取出的也是白球的概率。

三、(12分) 设随机变量 X 的分布函数为 $F(x) = \begin{cases} 0 & x < 0 \\ Ax & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$ 。

(1) 求常数 A 的取值范围; (2) 求 $Y = X^2$ 的分布函数。

四、(16分) 若随机向量 (X, Y) 的联合概率密度为 $f(x, y) = \begin{cases} 3x & 0 \leq y \leq x \leq 1 \\ 0 & \text{其他} \end{cases}$ 。

(1) 求 X 和 Y 的边缘概率密度 $f_X(x)$, $f_Y(y)$; (2) 问 X 和 Y 是否独立? (3) 求 $X+2Y$ 的概率密度。

五、(12分) 某产品的合格率 0.9。现抽检 200 个产品, 试分别用切比雪夫不等式和中心极限定理估计合格品数在 175 和 185 之间的概率。

六、(12分) 若 X_1, X_2, \dots, X_n 为来自 $N(\sigma, \mu^2)$ 的样本。 \bar{X} 为样本均值, $Y_i = X_i - \mu$, $Z_i = X_i - \bar{X}, i=1, 2, \dots, n$ 。

(1) 求 $\text{cov}(Z_1, Z_2)$; (2) 求 $\frac{1}{n} \sum_{i=1}^n Y_i^2$ 和 $\frac{1}{n-1} \sum_{i=1}^n Z_i^2$ 的期望和方差。

七、(12分) X_1, \dots, X_n 为总体 X 的样本, X 的密度函数为 $f(x) = \begin{cases} \frac{1}{2\theta}, & -\theta \leq x \leq \theta \\ 0, & \text{其它} \end{cases}$, θ 为未知参数。

求参数 θ 的矩估计量和最大似然估计量, 并判别它们的无偏性。

八、(12分) 轴承内环的锻压零件的平均高度服从正态分布。现从某天生产的产品中抽取 9 只内环, 其平均高度 $\bar{x} = 30.3$ 毫米, 样本标准差 $s = 0.6$ 毫米。正常生产时的零件平均高度为 30 毫米, 标准差不超过 0.5 毫米。试在显著性水平为 $\alpha = 0.05$ 的条件下, 检验这天生产是否正常。

($\Phi(1.96) = 0.975, \Phi(1.65) = 0.95, t_{0.025}(8) = 2.31, t_{0.05}(8) = 1.86, \chi_{0.95}^2(8) = 2.73, \chi_{0.05}^2(8) = 15.51,$

$\chi_{0.95}^2(9) = 3.33, \chi_{0.05}^2(9) = 16.92$)



一. (1) A, B 互不相容, 即 $A \cap B = \emptyset$. $\therefore P(A \cup B) = P(A) + P(B)$
 $P(B) = 0.7 - 0.4 = 0.3$

(2) A, B 相互独立. 故 $P(AB) = P(A)P(B)$ $\therefore P(A \cup B) = P(A) + P(B) - P(AB)$
 $= P(A) + P(B) - P(A)P(B)$

$\therefore 0.7 = 0.4 + P(B) - 0.4P(B) \quad \therefore P(B) = 0.5$

二. A_i : 第 i 次取出球是白球, $i=1, 2, 3$

(1) $P(A_3) = P(A_1 A_2 A_3) + P(A_1 \bar{A}_2 A_3) + P(\bar{A}_1 A_2 A_3) + P(\bar{A}_1 \bar{A}_2 A_3)$

$= P(A_1) \cdot P(A_2|A_1) P(A_3|A_1 A_2) + P(A_1) P(\bar{A}_2|A_1) P(A_3|A_1 \bar{A}_2) + P(\bar{A}_1) P(A_2|\bar{A}_1) P(A_3|\bar{A}_1 A_2) + P(\bar{A}_1) P(\bar{A}_2|\bar{A}_1) P(A_3|\bar{A}_1 \bar{A}_2)$
 $= \frac{3}{5} \times \frac{4}{6} \times \frac{5}{7} + \frac{3}{5} \times \frac{2}{6} \times \frac{4}{7} + \frac{2}{5} \times \frac{3}{6} \times \frac{4}{7} + \frac{2}{5} \times \frac{3}{6} \times \frac{3}{7} = \frac{3}{5}$

(2) $P(A_1|A_3) = \frac{P(A_1 A_3)}{P(A_3)} = \frac{P(A_1 A_2 A_3) + P(A_1 \bar{A}_2 A_3)}{P(A_3)} = \frac{\frac{3}{5} \times \frac{4}{6} \times \frac{5}{7} + \frac{3}{5} \times \frac{2}{6} \times \frac{4}{7}}{\frac{3}{5}} = \frac{2}{3}$

三. 随机变量 X 分布函数满足单调不减, 右连续. $\lim_{x \rightarrow -\infty} F(x) = 0, \quad \lim_{x \rightarrow +\infty} F(x) = 1$.

(1) $\therefore A > 0$, 且 $F(-) \leq F(1) \quad \therefore 0 < A \leq 1$.

(2) $F_Y(y) \stackrel{P4}{=} P(Y \leq y) = P(X^2 \leq y)$

$\frac{1}{2} y < 0$ 时, $F_Y(y) = 0$

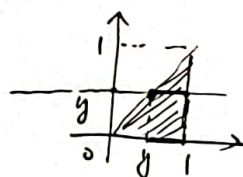
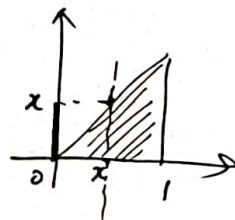
$\frac{1}{2} 0 \leq y < 1$ 时, $F_Y(y) = P(-\sqrt{y} \leq X \leq \sqrt{y}) = P(X \leq \sqrt{y}) - P(X \leq -\sqrt{y})$
 $= F(\sqrt{y}) - F(-\sqrt{y})$

$\frac{1}{2} y \geq 1$ 时, $F_Y(y) = P(X^2 \leq y)$
 $= F(\sqrt{y}) - F(-\sqrt{y}) = 1$

$\therefore F_Y(y) = \begin{cases} 0, & y < 0 \\ A\sqrt{y}, & 0 \leq y < 1 \\ 1, & y \geq 1 \end{cases}$

IV. (1) $f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \begin{cases} \int_0^x 3x dy, & 0 \leq x \leq 1 \\ 0, & \text{其他} \end{cases} = \begin{cases} 3x^2, & 0 \leq x \leq 1 \\ 0, & \text{其他} \end{cases}$

$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \begin{cases} \int_y^1 3x dx, & 0 \leq y \leq 1 \\ 0, & \text{其他} \end{cases} = \begin{cases} \frac{3}{2}(1-y^2), & 0 \leq y \leq 1 \\ 0, & \text{其他} \end{cases}$



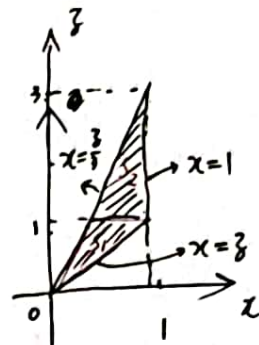
(2) $\therefore f(x, y) \neq f_X(x) f_Y(y)$

$\therefore X, Y$ 不独立.



$$(3) \iint_{\mathbb{R}^2} h(x+y) \cdot f(x,y) dx dy = \int_0^1 \left[\int_0^x h(x+y) \cdot 3x dy \right] dx$$

$$\begin{cases} x+y=z \\ y=0 \end{cases} \Rightarrow y = \frac{z-x}{1} \quad dy = \frac{1}{2} dz \quad \begin{matrix} y=0 \leftrightarrow z=x \\ y=x \leftrightarrow z=3x \end{matrix}$$



$$= \int_0^1 \left[\int_x^{3x} h(z) \cdot 3x \cdot \frac{1}{2} dz \right] dx$$

$$= \iint_{S_1} h(z) \cdot \frac{3}{2} x dz dx + \iint_{S_2} h(z) \cdot \frac{3}{2} x dz dx$$

$$= \int_0^1 \left[\int_{\frac{z}{3}}^{\frac{z}{2}} h(z) \cdot \frac{3}{2} x dx \right] dz + \int_1^3 \left[\int_{\frac{z}{3}}^1 h(z) \cdot \frac{3}{2} x dx \right] dz$$

$$= \int_0^1 h(z) \cdot \frac{3}{2} \left(\frac{z}{2} \right)^2 dz + \int_1^3 h(z) \left(\frac{z^2}{2} - \frac{z^2}{12} \right) dz$$

$$\therefore f_Z(z) = \begin{cases} \frac{3}{2} z^2, & 0 < z < 1 \\ \frac{3}{2} - \frac{z^2}{2}, & 1 < z < 3 \\ 0, & \text{其他} \end{cases}$$

$$S_1 = \{(x,z) \mid 0 < z < 1, \frac{z}{3} < x < \frac{z}{2}\}$$

$$S_2 = \{(x,z) \mid 1 < z < 3, \frac{z}{3} < x < 1\}$$

五. $\eta_A \sim B(200, 0.9)$

$$E(\eta_A) = np = 200 \times 0.9 = 180$$

$$D(\eta_A) = np(1-p) = 200 \times 0.9 \times 0.1 = 18$$

$$(1) P(175 \leq \eta_A \leq 185) = P(|\eta_A - E(\eta_A)| \leq 5) \geq 1 - \frac{D(\eta_A)}{5^2} = 1 - \frac{18}{25} = \frac{7}{25}$$

$$(2) P(175 \leq \eta_A \leq 185) = P\left(-\frac{5}{\sqrt{18}} \leq \frac{\eta_A - np}{\sqrt{np(1-p)}} \leq \frac{5}{\sqrt{18}}\right) = \Phi\left(\frac{5}{\sqrt{18}}\right) - \Phi\left(-\frac{5}{\sqrt{18}}\right) = 2\Phi\left(\frac{5}{\sqrt{18}}\right) - 1$$

六. (1). $Z_1 = \bar{X}_1 - \bar{X} = \bar{X}_1 - \frac{\bar{X}_1 + \dots + \bar{X}_n}{n} = \frac{n-1}{n} \bar{X}_1 - \frac{1}{n} \bar{X}_2 - \dots - \frac{1}{n} \bar{X}_n$

$$Z_2 = \bar{X}_2 - \bar{X} = \bar{X}_2 - \frac{\bar{X}_1 + \dots + \bar{X}_n}{n} = -\frac{1}{n} \bar{X}_1 + \frac{n-1}{n} \bar{X}_2 - \dots - \frac{1}{n} \bar{X}_n$$

$$\text{cov}(Z_1, Z_2) = \text{cov}\left(\frac{n-1}{n} \bar{X}_1 - \frac{1}{n} \bar{X}_2 - \dots - \frac{1}{n} \bar{X}_n, -\frac{1}{n} \bar{X}_1 + \frac{n-1}{n} \bar{X}_2 - \dots - \frac{1}{n} \bar{X}_n\right)$$

$$\stackrel{\because \bar{X}_1, \bar{X}_n \text{ 独立}}{=} \frac{n-1}{n} \left(-\frac{1}{n}\right) D(\bar{X}_1) + \left(-\frac{1}{n}\right) \cdot \frac{n-1}{n} D(\bar{X}_2) + \left(-\frac{1}{n}\right) \cdot \left(-\frac{1}{n}\right) D(\bar{X}_3) + \dots + \left(-\frac{1}{n}\right) \cdot \left(-\frac{1}{n}\right) D(\bar{X}_n)$$

$$= -\frac{2(n-1)}{n^2} D(\bar{X}) + \frac{n-2}{n^2} D(\bar{X}) = -\frac{1}{n} D(\bar{X}) = -\frac{\sigma^2}{n}$$

(或者)

$$\text{cov}(Z_1, Z_2) = \text{cov}(\bar{X}_1 - \bar{X}, \bar{X}_2 - \bar{X})$$

$$= \text{cov}(\bar{X}_1, \bar{X}_2) - \text{cov}(\bar{X}_1, \bar{X}) - \text{cov}(\bar{X}_2, \bar{X}) + \text{cov}(\bar{X}, \bar{X})$$

$$= -\text{cov}\left(\bar{X}_1, \frac{\bar{X}_1}{n} + \dots + \frac{\bar{X}_n}{n}\right) - \text{cov}\left(\bar{X}_2, \frac{\bar{X}_1}{n} + \dots + \frac{\bar{X}_n}{n}\right) + D(\bar{X}) = -\frac{1}{n} \text{cov}(\bar{X}_1, \bar{X}_1) - \frac{1}{n} \text{cov}(\bar{X}_2, \bar{X}_2) + D(\bar{X})$$

$$= -\frac{2}{n} D(\bar{X}) + D(\bar{X}) = -\frac{1}{n} D(\bar{X}) = -\frac{\sigma^2}{n}$$



$$(2) \frac{1}{n-1} \sum_{i=1}^n Z_i^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = S^2$$

由命题 1.21 知 $E(S^2) = D(\bar{X}) = \sigma^2$

或者由命题 13.1(2) 知 $\frac{(n-1)S^2}{\sigma^2} \sim \chi^2_{(n-1)}$

$$\begin{aligned} \therefore E\left(\frac{(n-1)S^2}{\sigma^2}\right) &= n-1 & \Rightarrow & \frac{n-1}{\sigma^2} E(S^2) = n-1 & \Rightarrow & E(S^2) = \sigma^2 \\ D\left(\frac{(n-1)S^2}{\sigma^2}\right) &= 2(n-1) & \Rightarrow & \frac{(n-1)^2}{\sigma^4} D(S^2) = 2(n-1) & \Rightarrow & D(S^2) = \frac{2\sigma^4}{n-1} \end{aligned}$$

$$\frac{1}{n} \sum_{i=1}^n Y_i^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \mu)^2$$

$\because X_i \sim N(\mu, \sigma^2) \xrightarrow{\text{P75}} \frac{X_i - \mu}{\sigma} \sim N(0, 1)$ 又 X_1, \dots, X_n 独立

$$\therefore \left(\frac{X_1 - \mu}{\sigma}\right)^2 + \dots + \left(\frac{X_n - \mu}{\sigma}\right)^2 \sim \chi^2_{(n)}$$

$$\therefore E\left(\frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2\right) = n \quad \Rightarrow \quad E\left(\sum_{i=1}^n (X_i - \mu)^2\right) = n\sigma^2$$

$$D\left(\frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \mu)^2\right) = 2n \quad \Rightarrow \quad D\left(\sum_{i=1}^n (X_i - \mu)^2\right) = 2n\sigma^4$$

$$\therefore E\left(\frac{1}{n} \sum_{i=1}^n Y_i^2\right) = \frac{1}{n} E\left(\sum_{i=1}^n (X_i - \mu)^2\right) = \frac{1}{n} \cdot n\sigma^2 = \sigma^2$$

$$D\left(\frac{1}{n} \sum_{i=1}^n Y_i^2\right) = \frac{1}{n^2} D\left(\sum_{i=1}^n (X_i - \mu)^2\right) = \frac{1}{n^2} \cdot 2n\sigma^4 = \frac{2\sigma^4}{n}$$

七. 矩估计: $\because E(X) = 0 \therefore$ 对 $|X|$ 作矩估计.

$$\frac{|X_1| + |X_2| + \dots + |X_n|}{n} = E(|X|)$$

$$E(|X|) = \int_{-\infty}^{+\infty} |x| \cdot f(x) dx = \int_{-\theta}^0 |x| \cdot \frac{1}{2\theta} dx + \int_0^{\theta} x \cdot \frac{1}{\theta} dx = \frac{\theta}{2}$$

$$\therefore \frac{\theta}{2} = \frac{|X_1| + \dots + |X_n|}{n} \Rightarrow \hat{\theta}_1 = \frac{2}{n} (|X_1| + \dots + |X_n|)$$

最大似然估计:

$$L(\theta) = \prod_{i=1}^n f(x_i) = f(x_1) f(x_2) \dots f(x_n) = \begin{cases} \frac{1}{(2\theta)^n}, & -\theta \leq x_1, \dots, x_n \leq \theta \\ 0, & \text{其他} \end{cases}$$

要 $L(\theta)$ 取值越大, 即 θ 越小越好.

$$\text{又 } -\theta \leq x_1, \dots, x_n \leq \theta \Leftrightarrow |x_1|, \dots, |x_n| \leq \theta$$

$$\therefore \hat{\theta} = \max\{|x_1|, \dots, |x_n|\}$$

$$\therefore \hat{\theta}_2 = \max\{|x_1|, |x_2|, \dots, |x_n|\}$$

$$\because E(|X|) = \frac{\theta}{2} \quad \therefore E(\hat{\theta}_1) = E\left[\frac{2}{n} (|X_1| + \dots + |X_n|)\right] = \frac{2}{n} [E(|X_1|) + \dots + E(|X_n|)] = \frac{2}{n} \cdot \left(\frac{\theta}{2} + \dots + \frac{\theta}{2}\right) = \theta$$

$\therefore \hat{\theta}_1$ 是无偏估计



$$X \sim f(x) = \begin{cases} \frac{1}{2\theta}, & 0 \leq x \leq \theta \\ 0, & \text{其他} \end{cases}$$

$$F_{|X|}(x) \stackrel{P41}{=} P(|X| \leq x) = \begin{cases} 0, & x < 0 \\ P(X \leq x) - P(X < -x), & x \geq 0 \end{cases} = \begin{cases} 0, & x < 0 \\ F_X(x) - F_X(-x), & x \geq 0 \end{cases} \quad \text{或者} \quad = \begin{cases} 0, & x < 0 \\ \frac{x}{\theta}, & 0 \leq x \leq \theta \\ 1, & x > \theta \end{cases}$$

$$\therefore f_{|X|}(x) = F'_{|X|}(x) = \begin{cases} 0, & x < 0 \\ f_X(x) + f_X(-x), & x > 0 \end{cases} = \begin{cases} \frac{1}{\theta}, & 0 < x < \theta \\ 0, & \text{其他} \end{cases}$$

由 P26 (3.7.1)

$$F_{\hat{\theta}_2}(x) = F_{|X|}^n(x) = \begin{cases} 0, & x < 0 \\ \frac{x^n}{\theta^n}, & 0 \leq x \leq \theta \\ 1, & x > \theta \end{cases} \quad \therefore f_{\hat{\theta}_2}(x) = \begin{cases} \frac{n x^{n-1}}{\theta^n}, & 0 < x < \theta \\ 0, & \text{其他} \end{cases}$$

$$\therefore E(\hat{\theta}_2) = \int_{-\infty}^{+\infty} x \cdot f_{\hat{\theta}_2}(x) dx = \int_0^{\theta} x \cdot \frac{n x^{n-1}}{\theta^n} dx = \frac{n}{n+1} \theta \neq \theta \quad \therefore \hat{\theta}_2 \text{ 不是无偏估计}$$

八. 正第2作时 $\mu=30$, $\sigma \leq 0.5$, 故需时两者都加以检验
对 μ 作检验.

$$H_0: \mu=30, \quad H_1: \mu \neq 30$$

$$\text{检验统计量} \quad \frac{\bar{X} - 30}{S/\sqrt{n}} \sim t(n-1)$$

$$\text{拒绝域} \quad W = \left\{ \left| \frac{\bar{X} - 30}{S/\sqrt{n}} \right| > t_{\frac{\alpha}{2}}(n-1) \right\} = \left\{ \left| \frac{\bar{X} - 30}{S/\sqrt{n}} \right| > 2.31 \right\}$$

$$\because \left| \frac{\bar{X} - 30}{S/\sqrt{n}} \right| = \left| \frac{30.3 - 30}{0.6/\sqrt{9}} \right| = 1.5 < 2.31 \quad \therefore \text{不在拒绝域中. 故接受 } H_0, \text{ 即 } \mu=30$$

再对 σ 作检验.

$$H_0: \sigma \leq 0.5, \quad H_1: \sigma > 0.5$$

$$\text{检验统计量: } \frac{(n-1)S^2}{0.5^2} \sim \chi^2(n-1)$$

$$\text{拒绝域} \quad W = \left\{ \frac{(n-1)S^2}{0.5^2} > \chi^2_{\alpha}(n-1) \right\} = \left\{ \frac{(n-1)S^2}{0.5^2} > 15.51 \right\}$$

$$\because \frac{(n-1)S^2}{0.5^2} = \frac{8 \times 0.6^2}{0.5^2} = 11.52 < 15.51 \quad \therefore \text{不在拒绝域中. 故接受 } H_0, \text{ 即 } \sigma \leq 0.5$$

\therefore 这天生产正常.

