

Introduction to statistical learning

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October 2019

Outline

- 15 hours for this introduction.
- Materials: slides + exercises with R available here
<https://lrouviere.github.io/intro-machine-learning/>
- 4 parts:
 1. Setting for statistical learning
 2. Parametric vs non parametric approaches
 3. Penalized regressions
 4. Trees and random forests
- Prerequisites: basics in probability, statistics (law of large numbers, estimation, bias, variance...) and data mining (linear model, logistic model, linear discriminant analysis...).

Part I

Mathematical setting for SL

Outline

1. Motivations
2. Mathematical framework for statistical learning
3. Some criterion for regression and supervised classification
 - Regression
 - Binary classification
 - Scoring
4. Estimating the risk
5. Overfitting
6. Bibliography

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Statistical learning?

Many "definitions"

1. "... explores way of estimating functional dependency from a given collection of data" [[Vapnik, 2000](#)].
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Statement

- Due to the digital revolution, we are faced with more and more complex data.
- Usual algorithms are not always efficient for these kind of data.

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Statement

- Due to the digital revolution, we are faced with more and more complex data.
- Usual algorithms are not always efficient for these kind of data.
- It is necessary to provide efficient algorithms which (automatically) learn from data.

History - see [[Besse and Laurent,](#)]

Period	Memory	Order of magnitude
1940-70	Byte	$n = 30, p \leq 10$
1970	MB	$n = 500, p \leq 10$
1980	MB	Machine Learning (computer science)
1990	GB	Data-Mining
2000	TB	$p > n$, statistical learning
2010	PB	n and p large, cloud, cluster...
2013	??	Big data
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Computer resources \implies

- **Data Mining** (patterns in large datasets, outliers...).
- **Statistical learning** (algorithms that can **automatically** learn from the data) \implies data decides, not the user!

- Find algorithms that can automatically learn from the data.
- It is not the user who choose both an algorithm and/or the parameters, it is the data which decides.
- But...

Statistical learning

- Find algorithms that can automatically learn from the data.
- It is not the user who choose both an algorithm and/or the parameters, it is the data which decides.
- But...the user should tell to the computer how to do that.

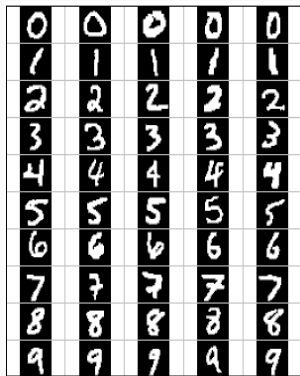
Conclusion

It is necessary to master the basics of machine learning algorithms.

Handwritten recognition

Statistical learning

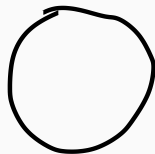
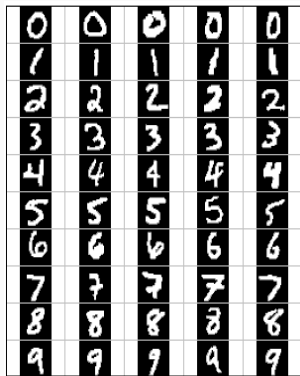
Understand and learn a behavior from **examples**.



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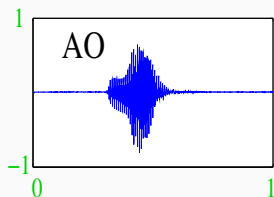
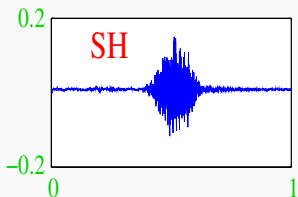
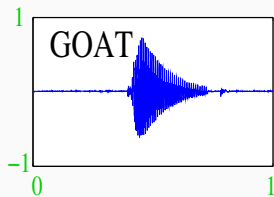
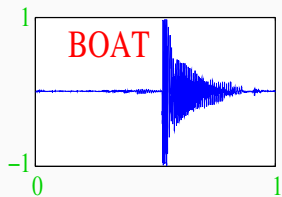
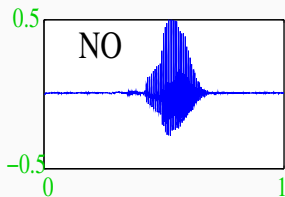
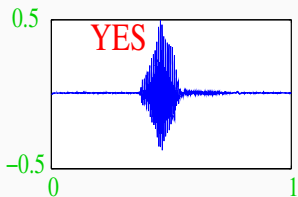
Statistical learning

Understand and learn a behavior from **examples**.



What is the number? 0, 1, 2...?

Speech recognition



Ozone prediction

- During one year, we have measured **ozone concentration** in a city (V4) ;
- Other meteorological variables are available (temperature, nebulosity, wind...).

```
> head(Ozone)
```

	V1	V2	V3	V4	V5	V6	V7	V8	V9	V10	V11	V12	V13
1	1	1	4	3	5480	8	20	NA	NA	5000	-15	30.56	200
2	1	2	5	3	5660	6	NA	38	NA	NA	-14	NA	300
3	1	3	6	3	5710	4	28	40	NA	2693	-25	47.66	250
4	1	4	7	5	5700	3	37	45	NA	590	-24	55.04	100
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Question

Can we **explain and predict** ozone concentration for tomorrow given meteorological predictions?

Spam detection

- For 4 601 emails, we have identified 1813 **spams**.
- In addition to this class label there are **57 variables** indicating the frequency of some words and characters in the e-mail.

```
> spam[1:5,c(1:8,58)]  
  make address  all num3d  our over remove internet type  
1 0.00    0.64 0.64    0 0.32 0.00   0.00    0.00 spam  
2 0.21    0.28 0.50    0 0.14 0.28   0.21    0.07 spam  
3 0.06    0.00 0.71    0 1.23 0.19   0.19    0.12 spam  
4 0.00    0.00 0.00    0 0.63 0.00   0.31    0.63 spam  
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Question

From these informations, can we **automatically detect** if a new e-mail is (or not) a spam?

Supervised vs unsupervised learning

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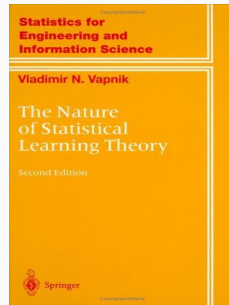
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Wide range of applications

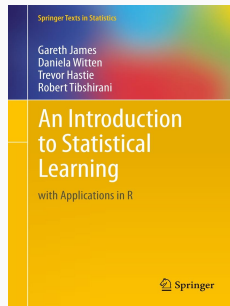
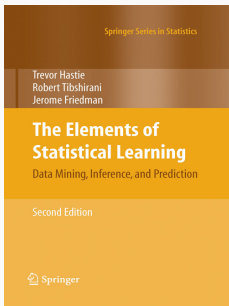
finance, economy, marketing, biology, medecine...

References

- Reference book: [Vapnik, 2000]



The Elements of Statistical Learning [[Hastie et al., 2009](#), [James et al., 2015](#)]



- Available (with datasets, R commands...) at:
<https://web.stanford.edu/~hastie/ElemStatLearn/>
<http://www-bcf.usc.edu/~gareth/ISL/>
- This course is largely based on these two books.

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Regression vs supervised classification

- **Input/output data:** $d_n = (x_1, y_1), \dots, (x_n, y_n)$ where $x_i \in \mathcal{X}$ are the inputs $y_i \in \mathcal{Y}$ the outputs.

Goal

1. **Explain** connections between inputs x_i and outputs y_i ;
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Vocabulary

- When the output \mathcal{Y} is continuous, we are faced with a **regression** problem.
- When the output is categorical ($\text{Card}(\mathcal{Y})$ finite), it is a **supervised classification** problem.

Examples

- Most of the presented problems are **supervised learning** problems: we have to predict an output y by inputs x :

y_i	x_i	
Number	picture	Super. Class.
Word	curve	Super. Class.
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Remark

- One **output** y_i .
- Wide range of **input objects** x_i (continuous, categorical, curves, pictures...).

Mathematical framework (begin)

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Interpretation

$\ell(y, y')$ measure the **cost (error)** between **one** prediction y' and **one** observation y .

Statistical framework

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- **Technical problem**: this function is random \implies (very) **difficult to minimize**.

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Risk of a machine

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- Such a function f^* (if it exists) is called the optimal machine for the cost function ℓ .

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- **Interpretation**: the risk of f_n comes closer to the optimal risk as n grows.

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Important conclusion

In practice, it is crucial to choose a relevant cost function for the problem we are faced.

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- Indeed, $\forall f : \mathcal{X} \rightarrow \mathbb{R}$, we have

$$\mathcal{R}(f^\star) = \mathbf{E}[(Y - f^\star(X))^2] \leq \mathbf{E}[(Y - f(X))^2] = \mathcal{R}(f).$$

- **Problem:** f^\star is unknown in practice. We have to find an estimate $f_n(x) = f_n(x, \mathcal{D}_n)$ such that $f_n(x) \approx f^\star(x)$.

Universal consistency

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Definition

f_n is **universally consistent** if

$$\lim_{n \rightarrow +\infty} \mathcal{R}(f_n) = \mathcal{R}(f^\star)$$

for any distribution of (X, Y) .

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- For any classification rule f ,

$$\mathcal{R}(f^\star) = \mathbf{P}(f^\star(X) \neq Y) \leq \mathbf{P}(f(X) \neq Y) = \mathcal{R}(f).$$

Universal consistency

- **Problem:** f^\star is unknown in practice. We have to find $f_n(x) = f_n(x, \mathcal{D}_n)$ such that $f_n(x) \approx f^\star(x)$.

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\implies See **Exercise 1 - IML0**.

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1. Motivations
2. Mathematical framework for statistical learning
3. Some criterion for regression and supervised classification
 - Regression
 - Binary classification
 - Scoring
4. Estimating the risk
5. Overfitting
6. Bibliography

Scoring function

- Always in **binary classification** ($\mathcal{Y} = \{-1, 1\}$).
- But... instead of a classification rule $f : \mathcal{X} \rightarrow \{-1, 1\}$, we want **to find a function** $S : \mathcal{X} \rightarrow \mathbb{R}$ such that

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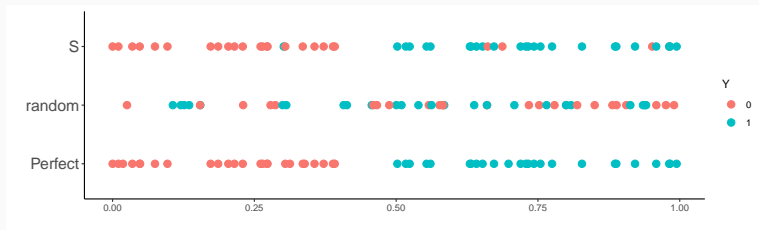
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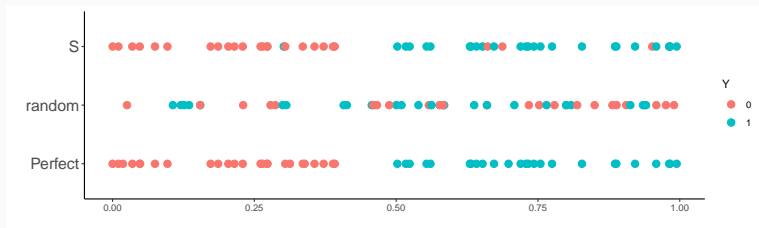
$S(x)$

- Such a function is a **score function**: instead of predicting the label y of a new $x \in \mathcal{X}$, we provide a **score** $S(x)$ with
 - **large** values if we think that x is 1;
 - **small** values if we think that x is -1.

Perfect and random scores



Perfect and random scores



Definition

- **Perfect score:** S is **perfect** if there exists s^\star such that

$$P(Y = 1 | S(X) \geq s^\star) = 1 \quad \text{and} \quad P(Y = -1 | S(X) < s^\star) = 1.$$

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$$P(Y = 1 | S(X) \geq s^*) = 1 \quad \text{and} \quad P(Y = -1 | S(X) < s^*) = 1.$$

- **Random score:** S is **random** if $S(X)$ and Y are independent.

Link between a score and a classification rule

- For a given **score** S and a **threshold** s , we obtain a **classification rule**:

$$f_s(x) = \begin{cases} 1 & \text{if } S(x) \geq s \\ -1 & \text{otherwise.} \end{cases}$$

- We have

	$f_s(X) = -1$	$f_s(X) = 1$
$Y = -1$	OK	E_1
$Y = 1$	E_2	OK

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- For any threshold s , we can define **2 errors**:

$$\alpha(s) = \mathbf{P}(f_s(X) = 1 | Y = -1) = \mathbf{P}(S(X) \geq s | Y = -1)$$

and

$$\beta(s) = \mathbf{P}(f_s(X) = -1 | Y = 1) = \mathbf{P}(S(X) < s | Y = 1).$$

We can also define

- **Specificity:** $sp(s) = \mathbf{P}(S(X) < s | Y = -1) = 1 - \alpha(s)$;
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Performance of a score

Visualize errors $\alpha(s)$ and $\beta(s)$ on a same graph for all thresholds s .

ROC curve

- **Idea:** define a 2-dimensionnel graph to represent errors $\alpha(s)$ and $\beta(s)$ for all values of s .

ROC curve

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Definition

The **ROC curve** of a score S is the **parametrized curve** defined by

$$\begin{cases} x(s) = \alpha(s) = 1 - sp(s) = \mathbf{P}(S(X) > s | Y = -1) \\ y(s) = 1 - \beta(s) = se(s) = \mathbf{P}(S(X) \geq s | Y = 1) \end{cases}$$

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Remark

- For any score S : $x(-\infty) = y(-\infty) = 1$ and $x(+\infty) = y(+\infty) = 0$.

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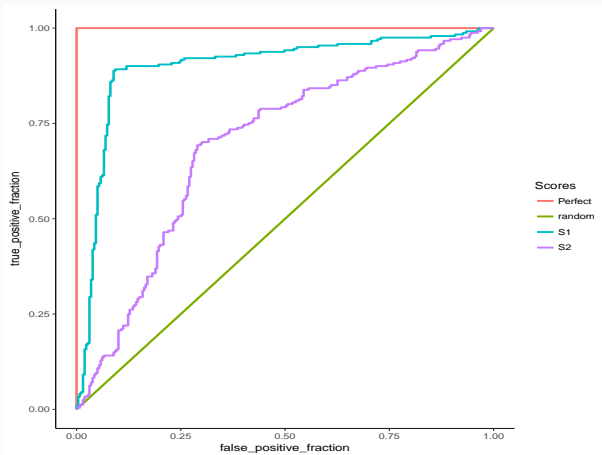
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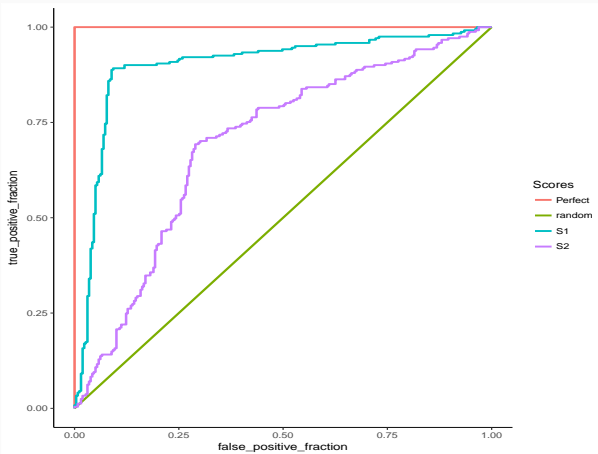
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- For a **perfect score**: $x(s^*) = 0$ and $y(s^*) = 1$.
- For a **random score**: $x(s) = y(s) \ \forall s$.





Interpretation

We measure performance of a score by its ability to approach the line $y = 1$ as fast as possible.

Definition

- **Area Under ROC** for a score S , denoted $AUC(S)$ is often used to measure performance of a S .
- **Perfect** score: $AUC(S) = 1$. **Random** score: $AUC(S) = 1/2$.

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Proposition

- *Let (X_1, Y_1) et (X_2, Y_2) be 2 i.i.d. observations. Then*

$$AUC(S) = \mathbf{P}(S(X_1) \geq S(X_2) | (Y_1, Y_2) = (1, -1)).$$

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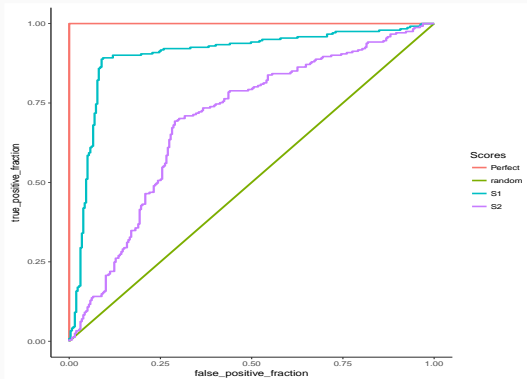
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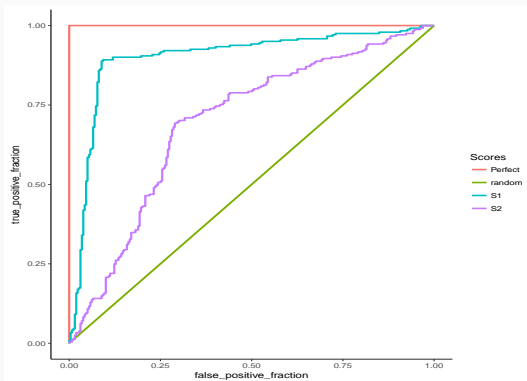
Conclusion

$AUC(S)$ measures the probability that S **correctly orders** two observations with **different labels**.

Example



Example



```
> df1 %>% group_by(Scores) %>% summarize(auc(D,M))  
1 Perfect    1.000000  
2 random    0.500000  
3 S1        0.8999824  
4 S2        0.6957177
```

Optimal score

- $AUC(S)$ can be seen as a **cost function** for a score S ;
- **Question**: is there an **optimal score** S^* for this cost function?

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Consequence

We have to find a "good" estimate $S_n(x) = S_n(x, \mathcal{D}_n)$ of

$$S^*(x) = P(Y = 1|X = x).$$

Summary

	Cost $\ell(y, f(x))$	Risk $\mathbf{E}[\ell(Y, f(X))]$	Winner f^\star
Regression	$(y - f(x))^2$	$\mathbf{E}[Y - f(X)]^2$	$\mathbf{E}[Y X = x]$
Binary class.	$\mathbf{1}_{y \neq f(x)}$	$\mathbf{P}(Y \neq f(X))$	Bayes rule
Scoring		$AUC(S)$	$\mathbf{P}(Y = 1 X = x)$

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Reminder

- n observations $(X_1, Y_1), \dots, (X_n, Y_n)$ i.i.d in $\mathcal{X} \times \mathcal{Y}$.

Goal

Given a cost function $\ell : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}^+$, we search a **machine** $f_n(x) = f_n(x, \mathcal{D}_n)$ closed to the **optimal machine** f^\star defined by

$$f^\star \in \operatorname{argmin}_f \mathcal{R}(f)$$

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Question

Given a machine f_n , **what can we say about its risk** $\mathcal{R}(f_n)$?

Empirical risk

- Since the distribution of (X, Y) is **unknown**, we **can't compute**
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One solution

Cross validation or **bootstrap** approaches.

Validation hold out

- The simplest approach.
- It consists in splitting the data \mathcal{D}_n into:
 1. a learning or training set $\mathcal{D}_{n,train}$ used to learn a machine f_n ;
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Inputs. \mathcal{D}_n : data, $\{\mathcal{T}, \mathcal{V}\}$: a partition of $\{1, \dots, n\}$.

1. Learn the machine with $\mathcal{D}_{n,train} = \{(X_i, Y_i) : i \in \mathcal{T}\} \implies f_{n,train}$;
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Comments

n_{train} and n_{test} should be **large enough** to

1. fit $f_{n,train}$;
2. estimate its risk $\mathcal{R}(f_{n,train})$.

K fold cross-validation

- Idea: repeat validation hold out algorithm on each element of a data partition.

Algorithm - CV

Inputs. \mathcal{D}_n : data, K an integer ;

1. Define a random partition $\{\mathcal{I}_1, \dots, \mathcal{I}_K\}$ of $\{1, \dots, n\}$;
2. For $k = 1, \dots, K$
 - 2.1 $\mathcal{I}_{train} = \{1, \dots, n\} \setminus \mathcal{I}_k$ and $\mathcal{I}_{test} = \mathcal{I}_k$;
 - 2.2 Learn the machine with $\mathcal{D}_{n,app} = \{(X_i, Y_i) : i \in \mathcal{I}_{app}\} \implies f_{n,k}$;
 - 2.3 Let $f_n(X_i) = f_{n,k}(X_i)$ for $i \in \mathcal{I}_{test}$;
3. Output

$$\widehat{\mathcal{R}}_n(f_n) = \frac{1}{n} \sum_{i=1}^n \ell(Y_i, f_n(X_i)).$$

Comments

- More useful than validation hold out when n is small.
- More accurate but more time consuming.
- K has to be chosen by the user (we often set $K = 10$).

Leave one out

- When $K = n$, we obtain leave one out cross validation.
- Risk is estimated by

$$\widehat{\mathcal{R}}_n(f_n) = \frac{1}{n} \sum_{i=1}^n \ell(Y_i, f_n^i(X_i))$$

where f_n^i stands for the machine defined on \mathcal{D}_n after deleted the i th observation.

- Exercises 1-3, IML1.

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Examples

- number of input variables in linear and logistic models.
- penalty parameters for lasso and ridge regressions.
- depth for tree algorithms.
- number of nearest neighbors.
- bandwidth of kernel regression estimators.
- number of iterations for boosting algorithms.
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- The choice of these parameters reveals crucial for the performance of the machine.

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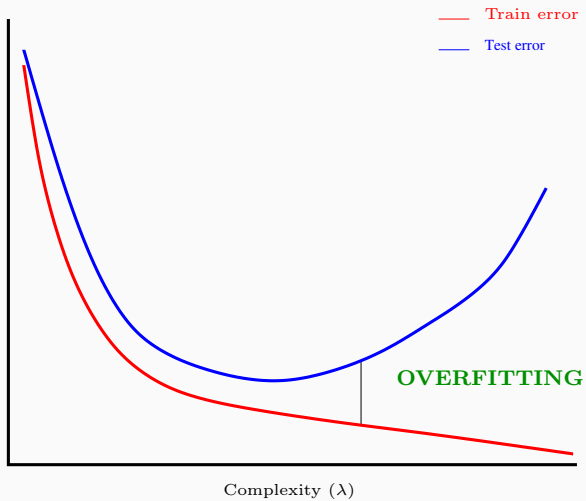
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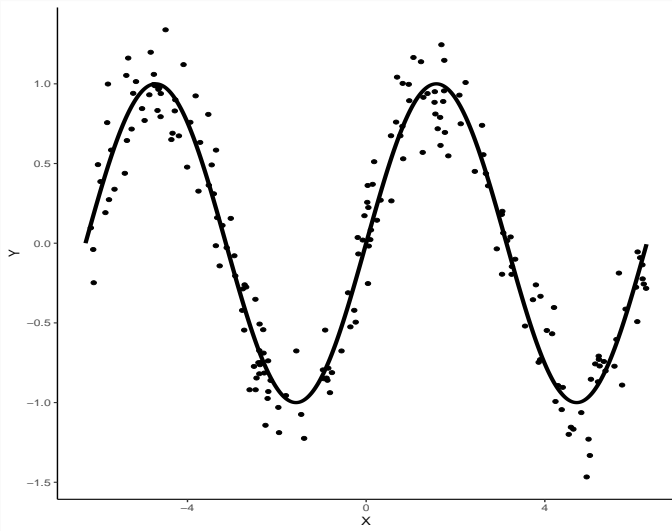
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Overfitting

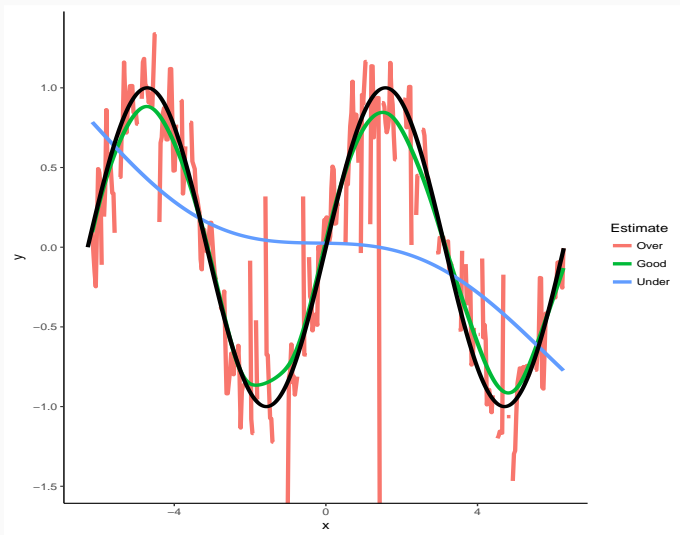
Good fitting on the training data (i.e. $f(X_i) = Y_i$) but **poor predictive** performances on **new individuals**.



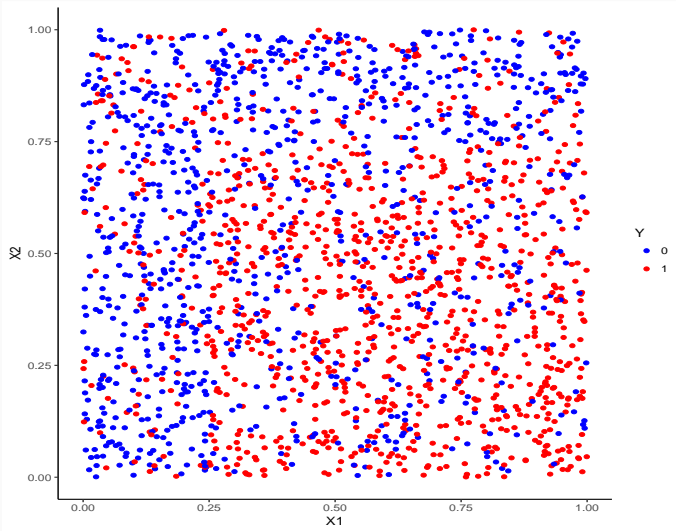
Overfitting for regression



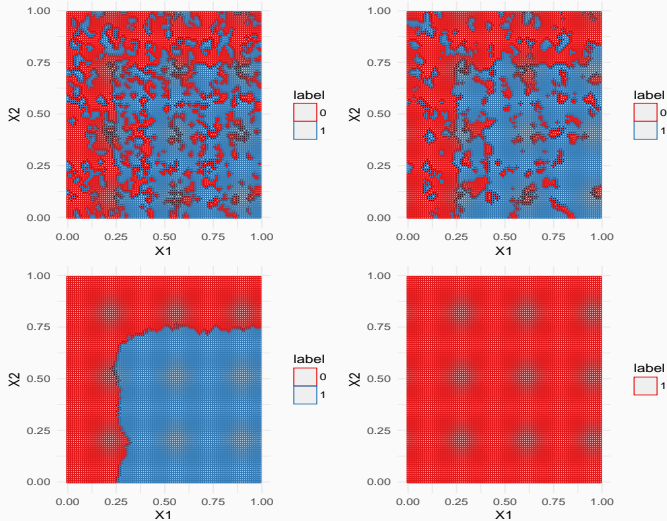
Overfitting for regression



Overfitting for supervised classification



Overfitting for supervised classification



- Run application [overfitting.app](#).

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Part II

Parametric versus nonparametric approaches

Outline

1. Some parametric methods

Linear and logistic models

Linear discriminant analysis

Just one explanatory variable

LDA: general case

2. Some nonparametric methods

Kernel and nearest neighbors methods

The curse of dimensionality

3. Empirical risk minimization

Setting

Caret package

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Mathematical framework

- n i.i.d observations $(X_1, Y_1), \dots, (X_n, Y_n)$ in $\mathcal{X} \times \mathcal{Y}$.
- $\ell : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}^+$ cost function.

Problem

Find a good estimate $f_n(\cdot) = f_n(\cdot, \mathcal{D}_n)$ of

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- $\ell : \mathcal{Y} \times \mathcal{Y} \rightarrow \mathbb{R}^+$ cost function.

Problem

Find a good estimate $f_n(\cdot) = f_n(\cdot, \mathcal{D}_n)$ of

$$f^\star \in \operatorname{argmin}_f \mathcal{R}(f) = \mathbf{E}[\ell(Y, f(X))].$$

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- Modelize remains to fix a class of functions \mathcal{F} and to assume that $f^\star \in \mathcal{F}$.

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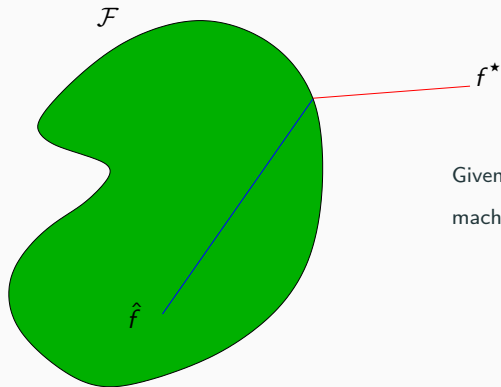
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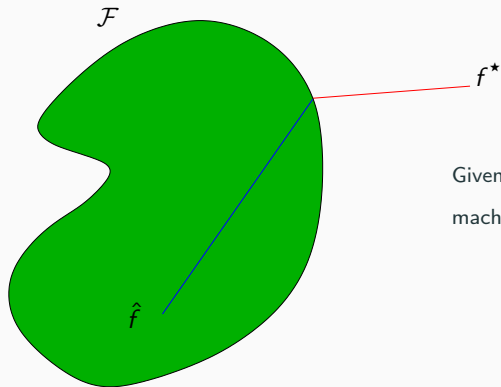
$$f^* \in \operatorname{argmin}_f \mathcal{R}(f) = \mathbf{E}[\ell(Y, f(X))].$$

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- Modelize remains to fix a class of functions \mathcal{F} and to assume that $f^* \in \mathcal{F}$.
- Modelize = make an assumption.

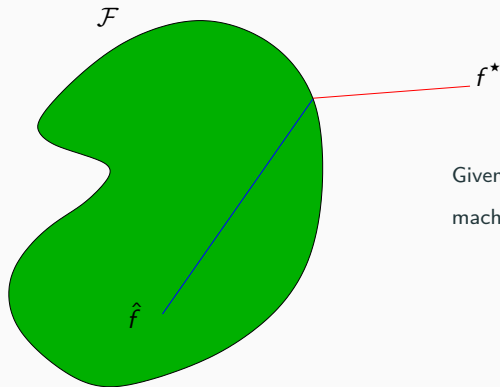


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Remarks

- These two terms vary in **opposite directions**.
- Statistician's job: **trade-off** between these two terms.

Parametric and non parametric

Definition

- If $\mathcal{F} = \{f_\theta : \theta \in \Theta\}$ with Θ of finite dimension, then the model is **parametric**.
- If \mathcal{F} is an infinite dimensional space, then the model is **non-parametric**.

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- **Loss of accuracy** in NP models. In this part, we will study this loss.

Outline

1. Some parametric methods

Linear and logistic models

Linear discriminant analysis

Just one explanatory variable

LDA: general case

2. Some nonparametric methods

Kernel and nearest neighbors methods

The curse of dimensionality

3. Empirical risk minimization

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The linear model

- In regression with $\mathcal{X} = \mathbb{R}^d$, the linear model is the parametric reference model.
- This model makes the assumption that the regression function is linear:

$$m^{\star}(x) = \mathbf{E}[Y|X = x] = \beta_1 x_1 + \cdots + \beta_d x_d.$$

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$$m^*(x) = \mathbf{E}[Y|X = x] = \beta_1 x_1 + \cdots + \beta_d x_d.$$

- Or equivalently

$$Y = \beta_1 X_1 + \cdots + \beta_d X_d + \varepsilon$$

where $\mathbf{E}[\varepsilon|X = x] = 0$ and $\mathbf{V}[\varepsilon|X = x] = \sigma^2$.

Remark

Estimate $m^* \iff$ **estimate $\beta \in \mathbb{R}^d$** (finite dimension \implies **parametric** model).

Some properties

- Least squares estimates minimize

$$\sum_{i=1}^n \varepsilon_i^2 = \sum_{i=1}^n (Y_i - (\beta_1 X_{i1} + \cdots + \beta_d X_{id}))^2.$$

The solution is given by

$$\hat{\beta}_n = (\mathbb{X}^t \mathbb{X})^{-1} \mathbb{X}^t \mathbb{Y}.$$

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- Regression function m^\star is thus estimated by

$$\hat{m}_n(x) = \hat{\beta}_1 x_1 + \cdots + \hat{\beta}_d x_d.$$

Assumption

Under some **technical assumptions**, we prove that

- $\mathbf{E}[\hat{\beta}] = \beta$ and $\mathbf{V}[\hat{\beta}] = (\mathbb{X}^t \mathbb{X})^{-1} \sigma^2$.

We deduce that (**exercise 2, IML0**)

$$\mathbf{E}[\|\hat{\beta} - \beta\|^2] = O\left(\frac{1}{n}\right) \quad \text{and} \quad \mathbf{E}[(\hat{m}_n(x) - m^\star(x))^2] = O\left(\frac{1}{n}\right).$$

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Remark

- Least squares estimates achieve the **parametric rate** ($1/n$).
- Moreover, if **errors terms** $\varepsilon_i, i = 1 \dots, n$ are **Gaussian**, we can compute **the distribution of the least squares estimates** (confidence intervals, test statistics...).
- See [[Grob, 2003](#), [Cornillon and Matzner-Løber, 2011](#)] for more information.

Example

- Linear model to explain ozone concentration.

```
> model_lin <- lm(V4~V5+V6+V7+V8+V9+V10+V11+V12+V13,data=Ozone)
> summary(model_lin)

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)  59.9517553  38.3286940   1.564 0.119421
V5           -0.0139111   0.0072511  -1.918 0.056527 .
V6             0.0276862   0.1741433   0.159 0.873847
V7             0.0808740   0.0237694   3.402 0.000812 ***
V8             0.1503404   0.0692994   2.169 0.031272 *
V9             0.5253439   0.1247136   4.212 3.87e-05 ***
V10          -0.0010052   0.0003944  -2.549 0.011586 *
V11             0.0049796   0.0147772   0.337 0.736501
V12          -0.1543882   0.1192917  -1.294 0.197140
V13          -0.0033951   0.0048963  -0.693 0.488883
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Logistic model

- **Logistic model** is the "brother" of the linear model in the context of **binary classification** ($\mathcal{Y} = \{-1, 1\}$).
- This model makes the **assumption** that (the logit transformation of) the probability $p(x) = \mathbf{P}(Y = 1|X = x)$ is **linear**:

$$\text{logit } p(x) = \log \frac{p(x)}{1 - p(x)} = \beta_1 x_1 + \cdots + \beta_d x_d = x^t \beta.$$

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- $\beta = (\beta_1, \dots, \beta_d) \in \mathbb{R}^d \implies$ **parametric model**.
- Unknown parameters β_1, \dots, β_d are estimated by **maximizing the (log)-likelihood**:

$$\mathcal{L}_n(\beta) = \sum_{i=1}^n \left\{ y_i x_i^t \beta - \log(1 + \exp(x_i^t \beta)) \right\}.$$

Some properties

Theorem [Fahrmeir and Kaufmann, 1985]

Under technical assumptions we have

1. the ML estimate $\{\hat{\beta}_n\}_n$ is **consistent**: $\hat{\beta}_n \xrightarrow{\mathbf{P}} \beta$;
2. the ML estimate $\{\hat{\beta}_n\}_n$ is **asymptotically gaussian**:

$$\sqrt{n}(\hat{\beta}_n - \beta) \xrightarrow{\mathcal{L}} \mathcal{N}(0, \mathcal{I}^{-1}(\beta)).$$

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$$\mathbf{E}[\|\hat{\beta} - \beta\|^2] = \mathcal{O}\left(\frac{1}{n}\right).$$

Important remark

Again, the ML estimate **achieves the parametric rate** $(1/n)$.

Example

- In R, we can fit a logistic model with the `glm` function.

```
> model_log <- glm(type~.,data=spam,family=binomial)
> summary(model_log)$coefficients[1:5,]
```

	Estimate	Std. Error	z value	Pr(> z)
(Intercept)	-1.5686144	0.1420362	-11.043767	2.349719e-28
make	-0.3895185	0.2314521	-1.682933	9.238799e-02
address	-0.1457768	0.0692792	-2.104194	3.536157e-02
all	0.1141402	0.1103011	1.034806	3.007594e-01
num3d	2.2515195	1.5070099	1.494031	1.351675e-01

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- Logistic regression **directly models** the parameter of the distribution of $Y|X = x$.
- **Linear discriminant analysis** do the opposite. It consists in
 - **modeling** the distributions of $X|Y = j$ for $j = 1, \dots, K$ by gaussian distributions $f_j(x)$.

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- **Linear discriminant analysis** do the opposite. It consists in
 - **modelizing** the distributions of $X|Y = j$ for $j = 1, \dots, K$ by gaussian distributions $f_j(x)$.
 - calculating the posterior distribution $Y|X = x$ with **Bayes formula** :

$$\mathbf{P}(Y = j|X = x) = \frac{\pi_j f_j(x)}{\sum_{\ell=1}^K \pi_{\ell} f_{\ell}(x)}$$

where $\pi_j = \mathbf{P}(Y = j), j = 1, \dots, K$.

Example: Fisher's iris problem

- Explain iris species by lengths and widths of petals and sepals.

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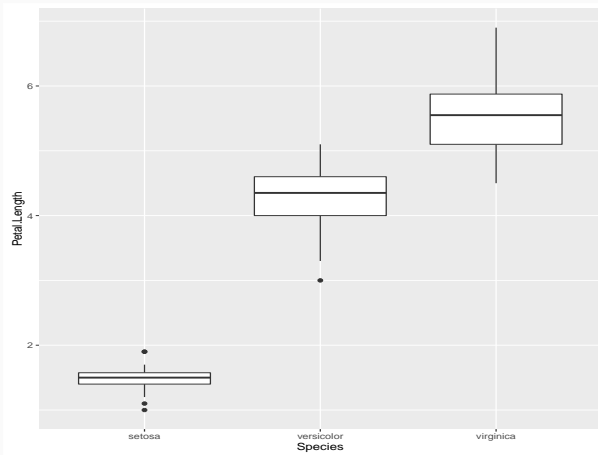
- Explain **iris species** by **lengths and widths of petals and sepals**.
- 5 variables :
 - the target variable **species** (categorical).
 - lengths and widths of petals and sepals.

```
> summary(iris)
  Sepal.Length   Sepal.Width   Petal.Length   Petal.Width
Min.   :4.300   Min.   :2.000   Min.   :1.000   Min.   :0.100
1st Qu.:5.100   1st Qu.:2.800   1st Qu.:1.600   1st Qu.:0.300
Median :5.800   Median :3.000   Median :4.350   Median :1.300
Mean   :5.843   Mean   :3.057   Mean   :3.758   Mean   :1.199
3rd Qu.:6.400   3rd Qu.:3.300   3rd Qu.:5.100   3rd Qu.:1.800
Max.   :7.900   Max.   :4.400   Max.   :6.900   Max.   :2.500

  Species
setosa   :50
versicolor:50
virginica :50
```

- We first want to explain **Species** by
- We can draw the following **boxplot**.

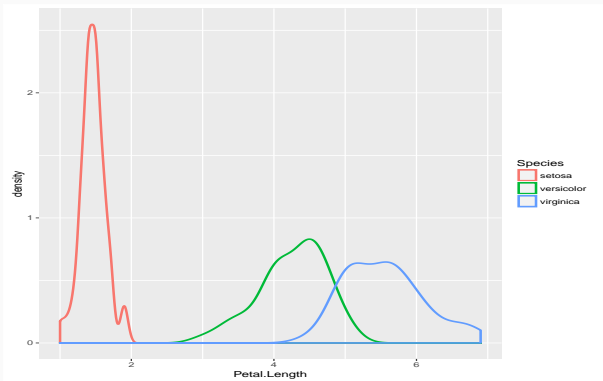
```
> ggplot(iris)+aes(x=Species,y=Petal.Length)+geom_boxplot()+theme_bw()
```



Visualize densities

- `geom_density` allows to visualize conditional distributions of $X|Y = j$, $j = 1, 2, 3$.

```
> ggplot(iris)+aes(x=Petal.Length,color=Species)+geom_density(size=1)
```



A model

- The three densities on the graph look like **Gaussian** densities.

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- Let $X = \text{Petal.Length}$ and $Y = \text{Species}$. We assume that distributions of X given $Y = k$ are Gaussians $\mathcal{N}(\mu_k, \sigma^2)$, $k = 1, 2, 3$.

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- Let $X = \text{Petal.Length}$ and $Y = \text{Species}$. We assume that distributions of X given $Y = k$ are Gaussians $\mathcal{N}(\mu_k, \sigma^2)$, $k = 1, 2, 3$.
- Densities of $X|Y = k$ are thus given by

$$f_{X|Y=k}(x) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{(x - \mu_k)^2}{2\sigma^2}\right).$$

Estimation

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Estimators

These quantities are naturally **estimated** by

$$\hat{\mu}_k = \frac{1}{n_k} \sum_{i: Y_i=k} X_i, \quad \hat{\sigma}^2 = \frac{1}{n-2} \sum_{k=1}^K \sum_{i: Y_i=k} (X_i - \hat{\mu}_k)^2$$

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$$\hat{\pi}_k = \frac{n_k}{n} \quad \text{where} \quad n_k = \sum_{i=1}^n \mathbf{1}_{\{Y_i=k\}}.$$

Example with R

```
> library(MASS)
> model <- lda(Species~Petal.Length,data=iris)
> model
```

Call:

```
lda(Species ~ Petal.Length, data = iris)
```

Prior probabilities of groups:

	setosa	versicolor	virginica
	0.3333333	0.3333333	0.3333333

Group means:

	Petal.Length
setosa	1.462
versicolor	4.260
virginica	5.552

Coefficients of linear discriminants:

	LD1
Petal.Length	2.323774

Making predictions

- **predict** function allows to **predict** species of **new iris**:

```
> don_pred
  Sepal.Length Sepal.Width Petal.Length Petal.Width
          5.0         3.6         1.4         0.2
          5.5         2.4         3.7         1.0
          7.1         3.0         5.9         2.1
          6.7         3.3         5.7         2.5
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```

- We just have to enter

```
> predict(model,newdata=don_pred)
$class
[1] setosa      versicolor virginica  virginica
Levels: setosa versicolor virginica
$posterior
      setosa  versicolor  virginica
1.000000e+00 2.589892e-10 6.170197e-21
3.123152e-06 9.997752e-01 2.217125e-04
1.113402e-23 9.723296e-04 9.990277e-01
9.198362e-22 3.913109e-03 9.960869e-01
```

- **Goal:** explain iris specie by the 4 explanatory variables Sepal.Length, Sepal.Width, Petal.Length, Petal.Width. We denote by X_1, X_2, X_3, X_4 these 4 variables and $X = (X_1, X_2, X_3, X_4)$.

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- The approach is **similar** to the previous case (1 variable)
 1. We model distributions of $X|Y = k$ by **Gaussian multivariate** distributions.
 2. We use **Bayes formula** to obtain posterior probabilities $P(Y = k|X = x)$.

LDA: general case

- Distributions of $X|Y = k$ are **are assumed to be Gaussians** $\mathcal{N}(\mu_k, \Sigma)$ where $\mu_k \in \mathbb{R}^p$ and Σ is a $p \times p$ definite positive matrix. Densities of $X|Y = k$ are thus given by:

$$f_{X|Y=k}(x) = \frac{1}{(2\pi\det(\Sigma))^{p/2}} \exp\left(-\frac{1}{2}(x - \mu_k)^t \Sigma^{-1}(x - \mu_k)\right).$$

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- Posterior probabilities $\mathbf{P}(Y = k|X = x)$ are obtained thanks to the **Bayes formula**

$$\mathbf{P}(Y = k|X = x) = \frac{\pi_k f_{X|Y=k}(x)}{f(x)}$$

where $f(x)$, the density of X , is computed from $f_{X|Y=k}(x)$ and from prior probabilities $\pi_k = \mathbf{P}(Y = k)$.

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$$\hat{\pi}_k = \frac{n_k}{n} \quad \text{with} \quad n_k = \sum_{i=1}^n \mathbf{1}_{\{Y_i=k\}}.$$

Example with R

```
> full_model<- lda(Species~.,data=iris)
```

```
> full_model
```

Call:

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lda(Species ~ ., data = iris)
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Prior probabilities of groups:

	setosa	versicolor	virginica
	0.3333333	0.3333333	0.3333333

Group means:

	Sepal.Length	Sepal.Width	Petal.Length	Petal.Width
setosa	5.006	3.428	1.462	0.246
versicolor	5.936	2.770	4.260	1.326
virginica	6.588	2.974	5.552	2.026

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```
> don_pred
```

Sepal.Length	Sepal.Width	Petal.Length	Petal.Width
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```
> predict(model_complet,newdata=don_pred)
$class
[1] setosa      versicolor virginica  virginica
Levels: setosa versicolor virginica

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82 9.648075e-16 9.999997e-01 3.266704e-07
103 1.231264e-42 2.592826e-05 9.999741e-01
145 4.048249e-46 2.524984e-07 9.999997e-01
```

Classification rule

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$$P(Y = k|X = x).$$

Classification rule

- **Reminder:** LDA allows to estimate posterior probabilities:

$$\mathbf{P}(Y = k|X = x).$$

- **Classification rule:** we choose the group which maximizes these probabilities

$$\hat{g}(x) = k \quad \text{if and only if} \quad \mathbf{P}(Y = k|X = x) \geq \mathbf{P}(Y = j|X = x), \quad j \neq k.$$

- **Boundary between 2 groups:** set of points x such that $\mathbf{P}(Y = k|X = x) = \mathbf{P}(Y = j|X = x)$.

- Or

$$\begin{aligned}\log \frac{\mathbf{P}(Y = k|X = x)}{\mathbf{P}(Y = \ell|X = x)} &= \log \frac{f_k(x)}{f_\ell(x)} + \log \frac{\pi_k}{\pi_\ell} \\ &= \log \frac{\pi_k}{\pi_\ell} - \frac{1}{2}(\mu_k + \mu_\ell)^t \Sigma^{-1}(\mu_k - \mu_\ell) \\ &\quad + x^t \Sigma^{-1}(\mu_k - \mu_\ell)\end{aligned}\tag{1}$$

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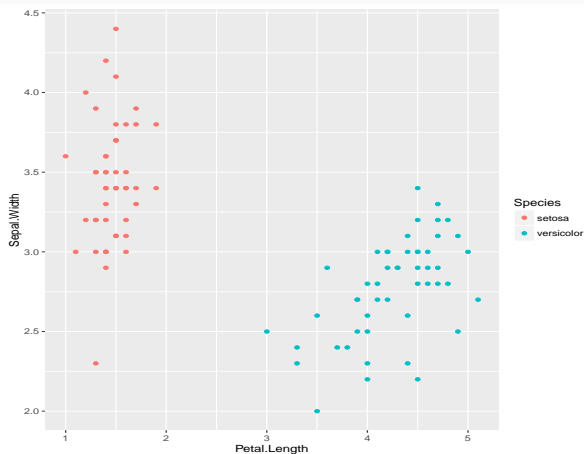
Conclusion

Bondary between 2 groups is linear!

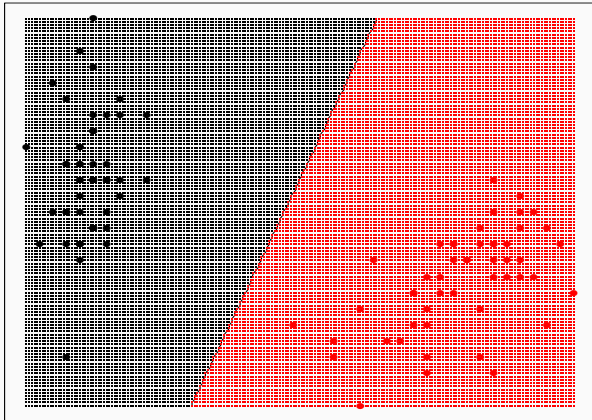
Example

- **Boundary** between "Setosa" and "Versicolor" for 2 variables.

```
> iris1 <- iris[iris$Species%in%c("setosa","versicolor"),c(3,2,5)]  
> ggplot(iris1)+aes(x=Petal.Length,y=Sepal.Width,color=Species)+geom_point()
```



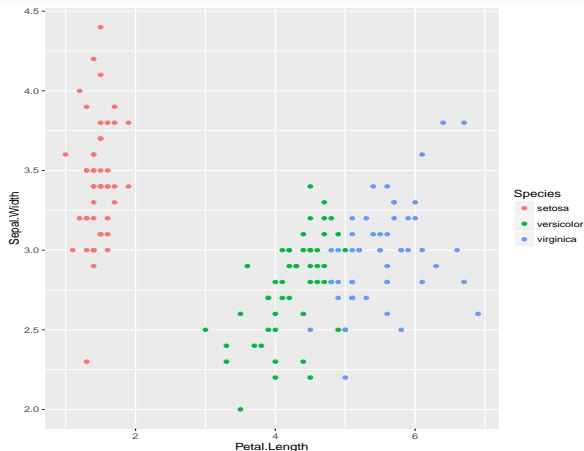
Boundary two classes



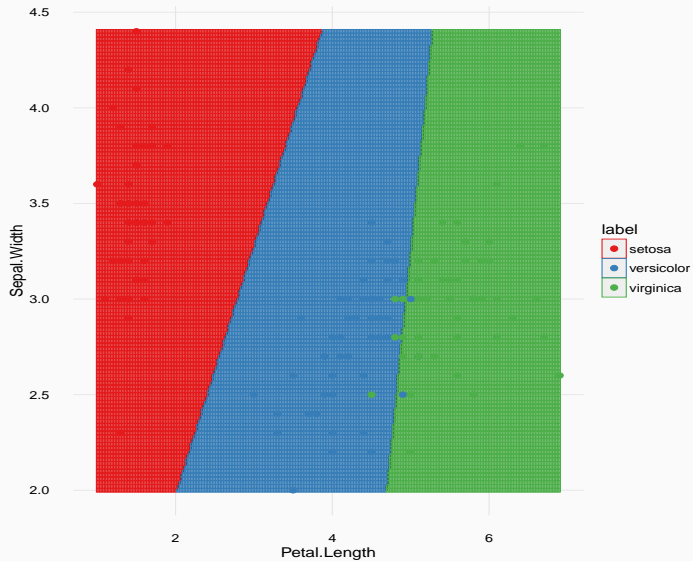
Example - 3 labels

- We do the same for the 3 species (3 classes).

```
> ggplot(iris)+aes(x=Petal.Length,y=Sepal.Width,color=Species)+geom_point()
```



Boundaries



Linear discriminant functions

Definition

Linear discriminant functions are defined by

$$\delta_k(x) = x^t \Sigma^{-1} \mu_k - \frac{1}{2} \mu_k^t \Sigma^{-1} \mu_k + \log \pi_k, \quad k = 1, \dots, K.$$

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Propriété

Thanks to (1), we deduce

$$\operatorname{argmax}_k \mathbf{P}(Y = k | X = x) = \operatorname{argmax}_k \delta_k(x).$$

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Propriété

Thanks to (1), we deduce

$$\operatorname{argmax}_k \mathbf{P}(Y = k | X = x) = \operatorname{argmax}_k \delta_k(x).$$

Conclusion

Maximising posterior probabilities is similar to maximising linear discriminant functions.

Outline

1. Some parametric methods

Linear and logistic models

Linear discriminant analysis

Just one explanatory variable

LDA: general case

2. Some nonparametric methods

Kernel and nearest neighbors methods

The curse of dimensionality

3. Empirical risk minimization

Setting

Caret package

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Local averaging

Idea

- Parametric models require **strong assumptions** on the function to estimate.
- Nonparametric approaches try to be less **restrictive**.

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- Nonparametric approaches try to be less **restrictive**.
- These methods consists of studying the data on a **neighborhood** of the points where we want to estimate the target function.

- For both regression and supervised classification, nonparametric approaches rely on **local averaging**:

$$\hat{f}_n(x) = \sum_{i=1}^n W_{ni}(x) Y_i$$

where the weights W_{ni} depend on the algorithm.

- W_{ni} **large** if X_i is **closed** to x .

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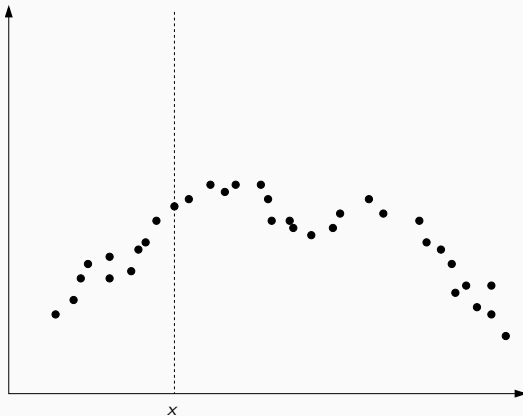
Kernel estimate

- $(X_1, Y_1), \dots, (X_n, Y_n)$ i.i.d. with the same law as (X, Y) .
- **Goal:** estimate $m^*(x) = \mathbf{E}[Y|X = x]$.



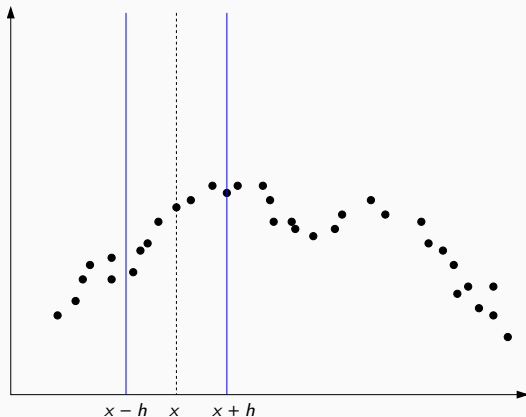
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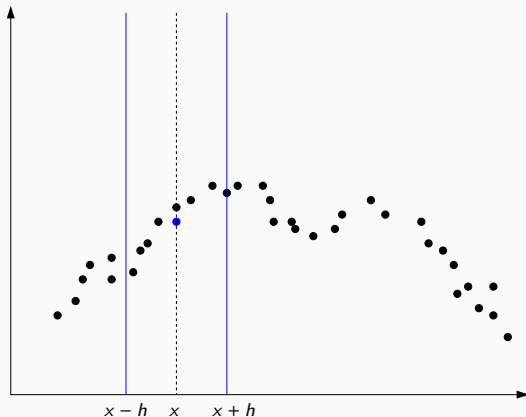
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- The estimator

$$\widehat{m}_n(x) = \text{Average}(Y_i : X_i \in [x - h, x + h]) = \frac{\sum_{i=1}^n \mathbf{1}_{x-h \leq X_i \leq x+h} Y_i}{\sum_{i=1}^n \mathbf{1}_{x-h \leq X_i \leq x+h}}.$$

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Definition

Let $h > 0$ and $K : \mathcal{X} \rightarrow \mathbb{R}^+$. The kernel estimate with bandwidth h and kernel K is defined by

$$\widehat{m}_n(x) = \frac{\sum_{i=1}^n K\left(\frac{X_i - x}{h}\right) Y_i}{\sum_{i=1}^n K\left(\frac{X_i - x}{h}\right)}.$$

Choice of the bandwidth

- Usual kernels when $\mathcal{X} = \mathbb{R}^d$:

1. Uniform: $K(x) = \mathbf{1}_{\|x\| \leq 1}$;
2. Gaussian: $K(x) = \exp(-\|x\|^2)$;
3. Epanechnikov: $K(x) = \frac{3}{4}(1 - \|x\|^2)\mathbf{1}_{\|x\| \leq 1}$.

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- The choice of the bandwidth h reveals crucial for the performance of the estimate:
 1. h large: steady estimator, low variance, large bias;
 2. h small: unsteady estimator ("overfitting"), large variance, small bias.

Conclusion

h governs the complexity of the estimate.

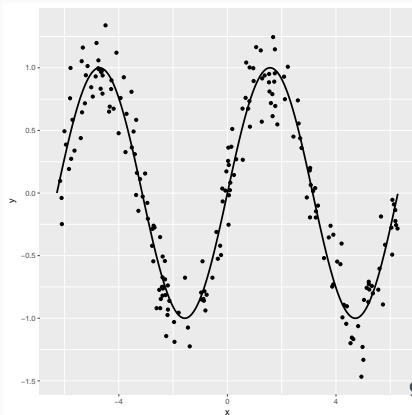
Example

- We generate data $(X_i, Y_i), i = 1, \dots, n = 200$ according to the model

$$Y_i = \sin(X_i) + \varepsilon_i, \quad i = 1, \dots, n$$

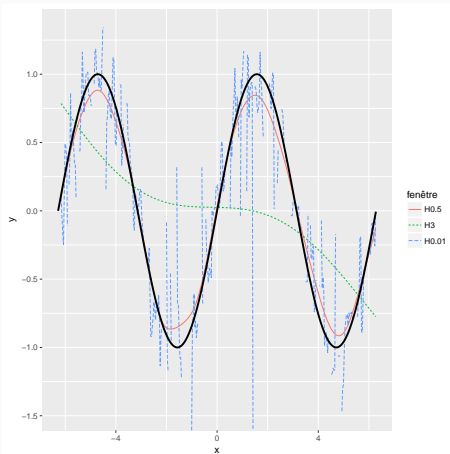
where X_i has a uniform distribution on $[-2\pi, 2\pi]$, ε_i has a Gaussian distribution $\mathcal{N}(0, 0.2^2)$.

```
> n <- 200; set.seed(1234)
> X <- runif(n, -2*pi, 2*pi)
> set.seed(5678)
> eps <- rnorm(n, 0, 0.2)
> Y <- sin(X) + eps
> df <- data.frame(X=X, Y=Y)
> x <- seq(-2*pi, 2*pi, by=0.01)
> df1 <- data.frame(x=x, y=sin(x))
> ggplot(df1) + aes(x=x, y=y) +
  geom_line(size=1) +
  geom_point(data=df, aes(x=X, y=Y))
```



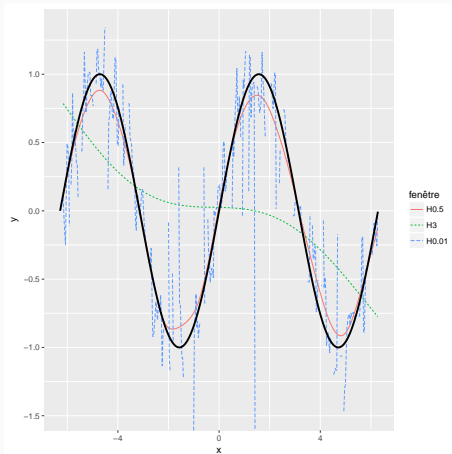
- **locpoly** function from **kernSmooth** package allows to fit kernel estimates.

```
> h1 <- 0.5; h2 <- 3; h3 <- 0.01
> fx1 <- locpoly(X,Y,bandwidth=h1)
> fx2 <- locpoly(X,Y,bandwidth=h2)
> fx3 <- locpoly(X,Y,bandwidth=h3)
> df1 <- data.frame(x=x,y=sin(x))
> df2 <- data.frame(x=fx1$x,
  "H0.5"=fx1$y, "H3"=fx2$y,
  "H0.01"=fx3$y)
> df22 <- melt(df2,id.vars=1)
> names(df22)[2:3] <- c("fenêtre",
  "y")
> ggplot(df22)+aes(x=x,y=y)+
  geom_line(aes(color=fenêtre,
    lty=fenêtre))+geom_line
  (data=df1,aes(x=x,y=y),size=1)
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- **Exercise 4-IML1.**

Nearest neighbors algorithm

Definition

Let $k \leq n$ an integer. The **k -nearest neighbors** estimate is defined by

$$\hat{m}_n(x) = \frac{1}{k} \sum_{i \in \text{knn}(x)} Y_i$$

where for $x \in \mathcal{X}$

$$\text{knn}(x) = \{i : X_i \text{ is among the knn of } x \text{ among } \{X_1, \dots, X_n\}\}.$$

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Remark

Once again, k reveals **crucial** for the performance of the estimate:

1. **k large**: steady estimate, low variance, high bias;
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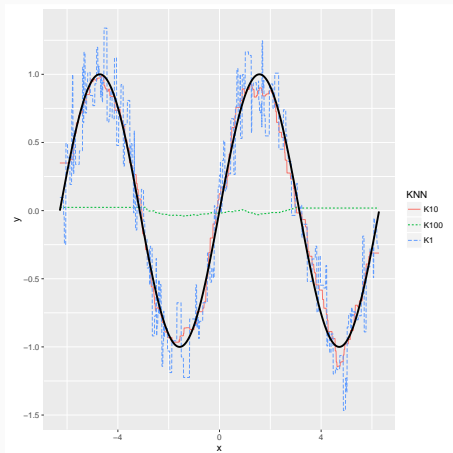
1. **k large**: steady estimate, low variance, high bias;
2. **k small**: "overfitting", large variance, small bias.

$\implies k$ governs the **complexity** of the model.

Example

- `knn.reg` function from **FNN** package allows to fit k -nearest neighbors estimate.

```
> k1 <- 10; k2 <- 100; k3 <- 1
> fx1 <- knn.reg(X,as.matrix(x),y=Y,k=k1)
> fx2 <- knn.reg(X,as.matrix(x),y=Y,k=k2)
> fx3 <- knn.reg(X,as.matrix(x),y=Y,k=k3)
> df1 <- data.frame(x=x,y=sin(x))
> df2 <- data.frame(x=x,"K100"=fx1$pred,
+                   "K100"=fx2$pred,"K1"=fx3$pred)
> df22 <- melt(df2,id.vars=1)
> names(df22)[2:3] <- c("KNN","y")
> ggplot(df22)+aes(x=x,y=y)+
+   geom_line(aes(color=KNN,lty=KNN))+
+   geom_line(data=df1,aes(x=x,y=y),size=1)
```



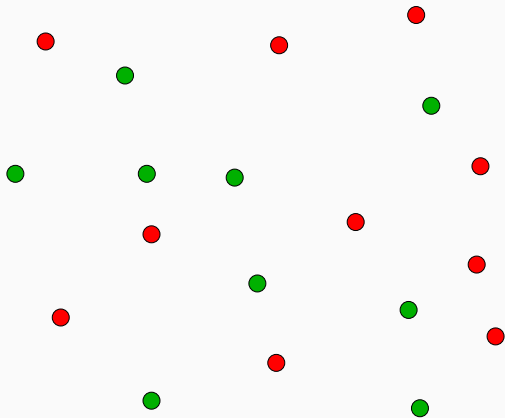
Supervised classification

- Kernel and nearest neighbors estimates have been presented in regression ($\mathcal{Y} = \mathbb{R}$).
- Approaches are similar in supervised classification:
 1. neighborhoods are defined in the same way;
 2. (only) change:

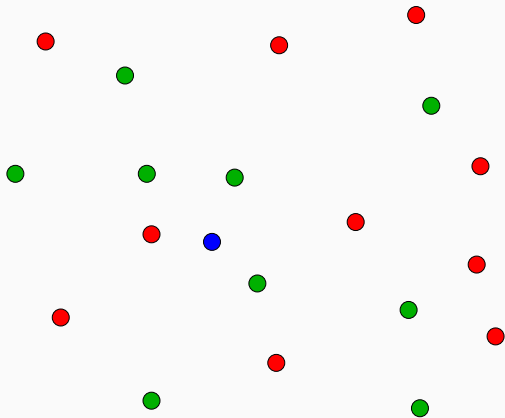
Supervised classification

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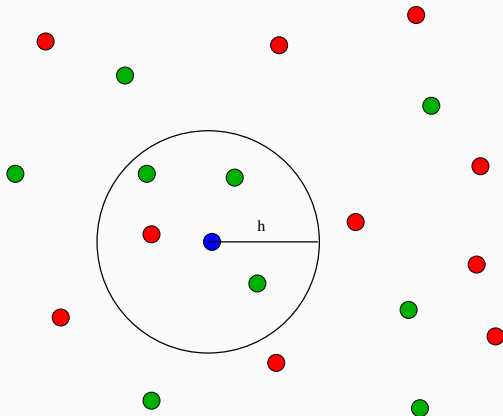
Kernel for supervised classification



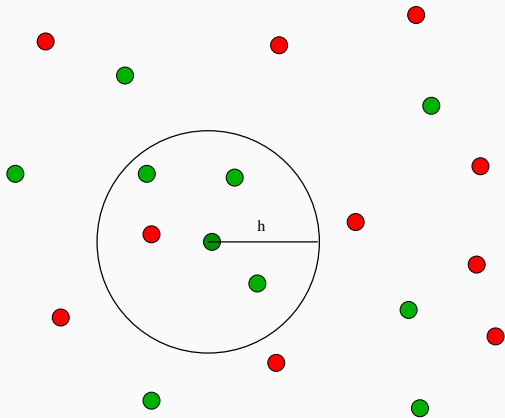
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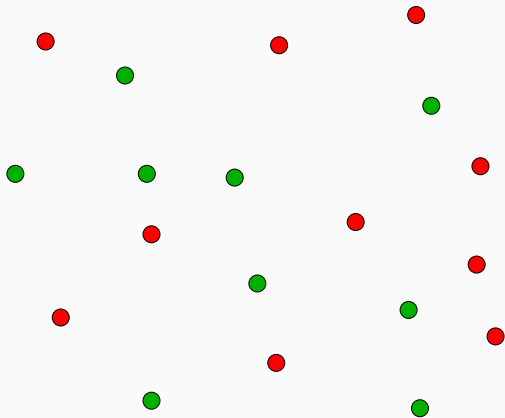
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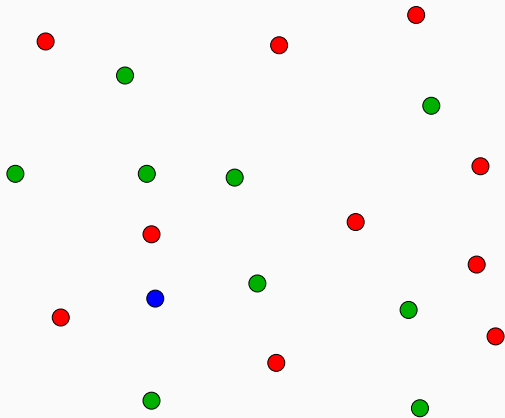
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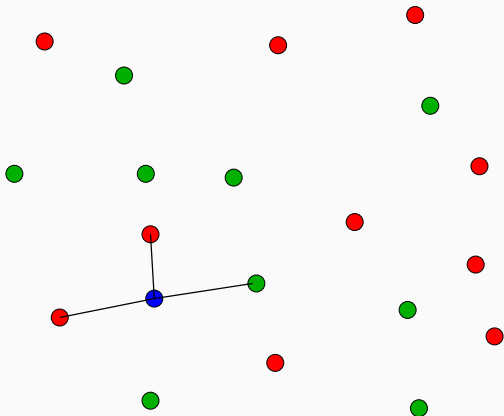
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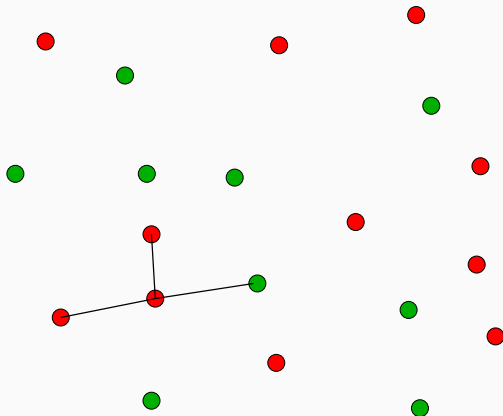
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The k -nn rule

- Let $k \leq n$, the k -nn rule apply a **majority vote** to assess the group of new individuals:

$$\hat{g}_n(x) = MV(Y_i : i \in knn(x)) = \operatorname{argmax}_{k \in \mathcal{Y}} \sum_{i \in knn(x)} \mathbf{1}_{Y_i=k}$$

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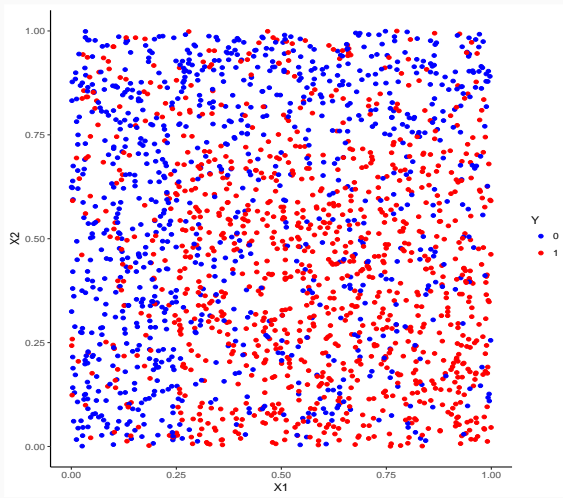
Remark

As for regression, the choice of k reveals **crucial** for the performance of the estimate:

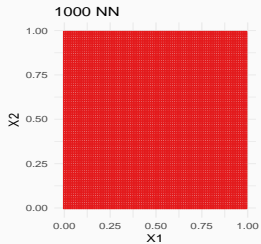
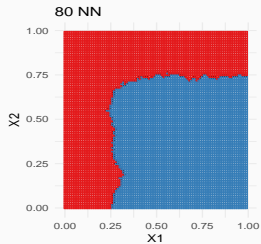
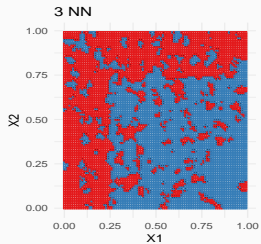
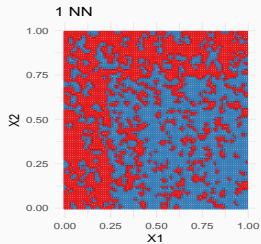
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Example

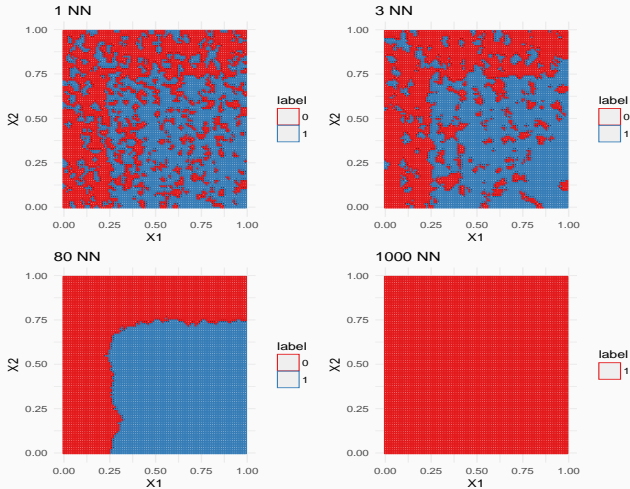
- **Goal:** explain a binary variable Y by 2 continuous variables X_1 and X_2 . We have $n = 2\,000$ observations.



k -nn rules



k -nn rules



Conclusion

We clearly visualize how the choice of k is important.

Consistency [Györfi et al., 2002]

- For both regression and supervised classification, kernel rules and nearest neighbors rules are universally consistent (under weak assumptions).

Theorem [Stone, 1977]

If $k \rightarrow \infty$ and $k/n \rightarrow 0$, then the k -nn rule is universally consistent.

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If $k \rightarrow \infty$ and $k/n \rightarrow 0$, then the k -nn rule is universally consistent.

Theorem [Devroye and Krzyżak, 1989]

If $h \rightarrow 0$ and $nh^d \rightarrow +\infty$, then the kernel rule is universally consistent.

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Nonparametric methods (always) suffer from **the curse of dimensionality**: as the dimension d increases, we have less and less observations in the neighborhoods of $x \implies$

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Theorem

We consider the regression problem (explain Y by X_1, \dots, X_d) and denote by m_n the k -nn estimate. Under technical assumptions, the quadratic risk of m_n satisfies (see exercise 3-IML0)

$$\mathcal{R}(m_n) = O\left(n^{-\frac{2}{d+2}}\right).$$

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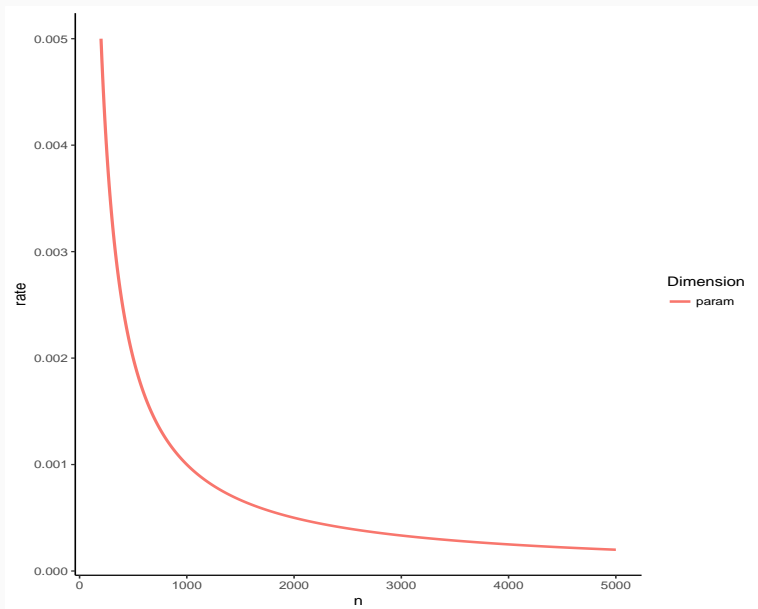
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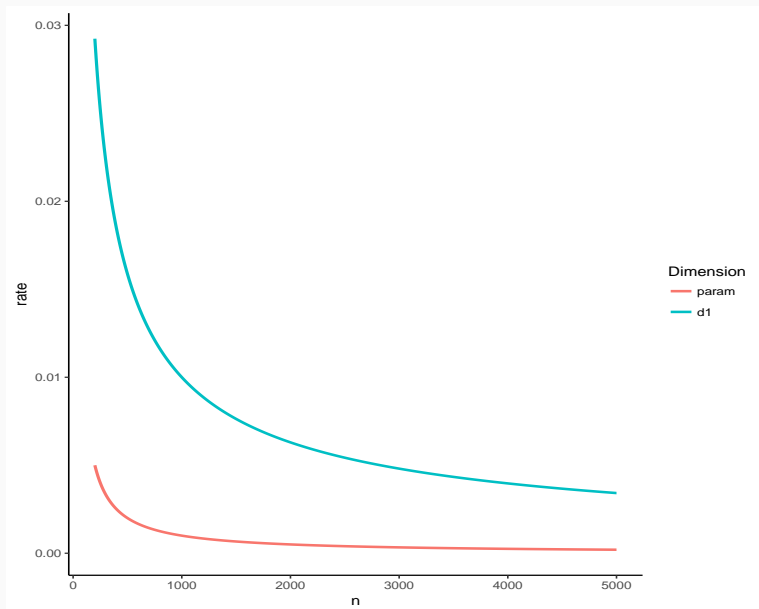
Consequence

- $d = 1$: rate $n^{-2/3}$, $d = 5$: rate $n^{-2/7}$.
- In practice, nonparametric estimates are **not efficient** in **high dimensional spaces**.

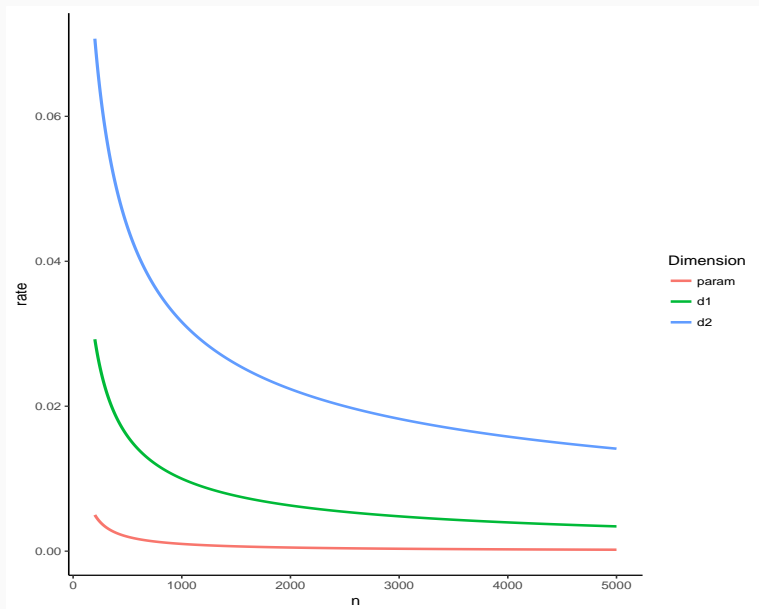
Curse of dimensionality (Illustration)



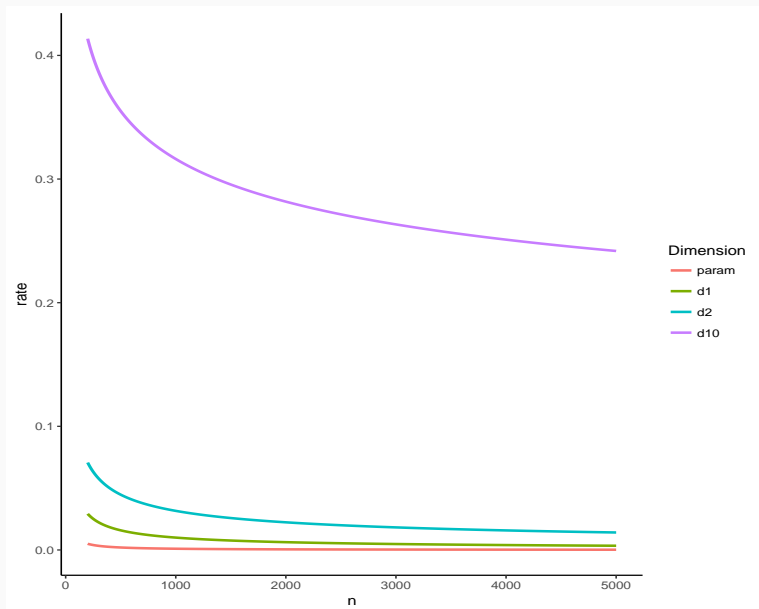
Curse of dimensionality (Illustration)



Curse of dimensionality (Illustration)



Curse of dimensionality (Illustration)



Outline

1. Some parametric methods
 - Linear and logistic models
 - Linear discriminant analysis
 - Just one explanatory variable
 - LDA: general case
2. Some nonparametric methods
 - Kernel and nearest neighbors methods
 - The curse of dimensionality
3. Empirical risk minimization
 - Setting
 - Caret package
4. Bibliography

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Choosing parameters

- Most of the machines depends on parameters.

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Rules	Parameters
<i>k</i> -nn	<i>k</i> : number of neighbors
kernel	<i>h</i> : bandwidth
trees	depth
boosting	number of iterations
...	...

Choosing parameters

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- Selection of these parameters reveals **crucial for the performances of the estimates**.

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Rules	Parameters
k -nn	k : number of neighbors
kernel	h : bandwidth
trees	depth
boosting	number of iterations
...	...

- Selection of these parameters reveals **crucial for the performances of the estimates**.
- Goal:**
 - define procedures which allow to **automatically** select these parameters;
 - establish theoretical guarantees for these procedures (GB lecture).

Framework

- \mathcal{F} a collection of machines.
- **Risk** for a machine f : $\mathcal{R}(f) = \mathbf{E}[\ell(Y, f(X))]$.
- **Goal**: select \hat{f} in \mathcal{F} such that

$$\mathcal{R}(\hat{f}) \approx \inf_{f \in \mathcal{F}} \mathcal{R}(f).$$

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$$\mathcal{R}(\hat{f}) \approx \inf_{f \in \mathcal{F}} \mathcal{R}(f).$$

ERM

- **Estimate the risk** of the machines in \mathcal{F} (validation hold out, cross validation...) $\implies \hat{R}_n(f)$.
- Choose the machine \hat{f} which **minimizes the estimated risk** $\hat{R}_n(f)$.

Selecting k (k -nn rule)

- Data splitting:
 - A learning or train set $\mathcal{D}_m = \{(X_1, Y_1), \dots, (X_m, Y_m)\}$;
 - A test set $\mathcal{D}_\ell = \{(X_{m+1}, Y_{m+1}), \dots, (X_n, Y_n)\}$ with $m + \ell = n$.
- Candidates: $\mathcal{G}_m = \{g_k, 1 \leq k \leq m\} \rightarrow k$ -nn rules using \mathcal{D}_m .
- Risk: $L(g) = \mathbf{P}(g(X) \neq Y)$.

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ERM Strategy

Choose \hat{g}_n which minimizes

$$\frac{1}{\ell} \sum_{i=m+1}^n \mathbf{1}_{g_k(X_i) \neq Y_i}.$$

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- This package allows to select machines and to estimate their performances.
- More than 230 algorithms are available on caret:
<http://topepo.github.io/caret/index.html>

- Classification and regression training.
- This package allows to select machines and to estimate their performances.
- More than 230 algorithms are available on caret:
<http://topepo.github.io/caret/index.html>
- We just have to specify:
 - the method (logistic, k-nn, trees, randomForest...)
 - a grid for the values of parameters (number of NN...)
 - the risk or the cost function (error probability, AUC, quadratic risk...)
 - how to estimate the risk (validation hold out, cross validation, bootstrap...).

Validation hold out i

```
> K_cand <- seq(1,500,by=20)
> library(caret)
> ctrl1 <- trainControl(method="LGOCV",number=1,index=list(1:1500))
> KK <- data.frame(k=K_cand)
> e1 <- train(Y~.,data=donnees,method="knn",trControl=ctrl1,tuneGrid=KK)
> e1
```

k-Nearest Neighbors

2000 samples
2 predictor
2 classes: '0', '1'

No pre-processing

Resampling: Repeated Train/Test Splits Estimated (1 reps, 75%)

Summary of sample sizes: 1500

Resampling results across tuning parameters:

k	Accuracy	Kappa
---	----------	-------

Validation hold out ii

1	0.620	0.2382571
21	0.718	0.4342076
41	0.722	0.4418388
61	0.718	0.4344073
81	0.720	0.4383195
101	0.714	0.4263847
121	0.716	0.4304965
141	0.718	0.4348063
161	0.718	0.4348063
181	0.718	0.4348063
201	0.720	0.4387158
221	0.718	0.4350056
241	0.718	0.4350056
261	0.722	0.4428232
281	0.714	0.4267894
301	0.714	0.4269915
321	0.710	0.4183621
341	0.696	0.3893130

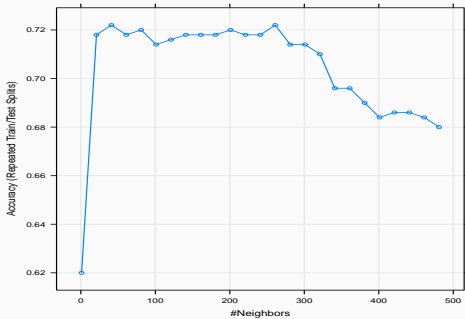
Validation hold out iii

361	0.696	0.3893130
381	0.690	0.3767090
401	0.684	0.3645329
421	0.686	0.3686666
441	0.686	0.3679956
461	0.684	0.3638574
481	0.680	0.3558050

Accuracy was used to select the optimal model using the largest value.
The final value used for the model was $k = 261$.

```
> plot(e1)
```

Validation hold out iv



Cross validation i

```
> library(doMC)
> registerDoMC(cores = 3)
> ctrl2 <- trainControl(method="cv",number=10)
> e2 <- train(Y~.,data=dapp,method="knn",trControl=ctrl2,tuneGrid=KK)
> e2
```

k-Nearest Neighbors

1500 samples

2 predictor

2 classes: '0', '1'

No pre-processing

Resampling: Cross-Validated (10 fold)

Summary of sample sizes: 1350, 1350, 1350, 1350, 1350, 1350, ...

Resampling results across tuning parameters:

k	Accuracy	Kappa
---	----------	-------

Cross validation ii

1	0.6280000	0.2519051
21	0.7333333	0.4623213
41	0.7273333	0.4503384
61	0.7360000	0.4682891
81	0.7353333	0.4673827
101	0.7313333	0.4596395
121	0.7306667	0.4584747
141	0.7366667	0.4703653
161	0.7340000	0.4654675
181	0.7306667	0.4585136
201	0.7313333	0.4597224
221	0.7333333	0.4638243
241	0.7333333	0.4637789
261	0.7306667	0.4581189
281	0.7320000	0.4604955
301	0.7246667	0.4452185
321	0.7166667	0.4283226
341	0.7120000	0.4183438

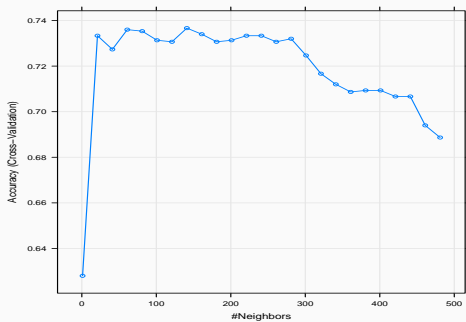
Cross validation iii

361	0.7086667	0.4109784
381	0.7093333	0.4121146
401	0.7093333	0.4117108
421	0.7066667	0.4057889
441	0.7066667	0.4047529
461	0.6940000	0.3782209
481	0.6886667	0.3662798

Accuracy was used to select the optimal model using the largest value.
The final value used for the model was $k = 141$.

```
> plot(e2)
```

Cross validation iv



Repeated cross-validation i

```
> ctrl3 <- trainControl(method="repeatedcv",repeats=5,number=10)
> e3 <- train(Y~.,data=dapp,method="knn",trControl=ctrl3,tuneGrid=KK)
> e3
```

k-Nearest Neighbors

```
1500 samples
  2 predictor
  2 classes: '0', '1'
```

No pre-processing

Resampling: Cross-Validated (10 fold, repeated 5 times)

Summary of sample sizes: 1350, 1350, 1350, 1350, 1350, 1350, ...

Resampling results across tuning parameters:

k	Accuracy	Kappa
1	0.6222667	0.2416680
21	0.7352000	0.4661220

Repeated cross-validation ii

41	0.7312000	0.4580125
61	0.7310667	0.4580882
81	0.7321333	0.4606022
101	0.7329333	0.4626718
121	0.7326667	0.4623496
141	0.7328000	0.4628236
161	0.7345333	0.4663240
181	0.7344000	0.4660110
201	0.7322667	0.4616271
221	0.7324000	0.4619926
241	0.7326667	0.4624912
261	0.7310667	0.4591799
281	0.7282667	0.4530797
301	0.7248000	0.4454653
321	0.7170667	0.4292033
341	0.7118667	0.4181330
361	0.7112000	0.4163210
381	0.7109333	0.4154893

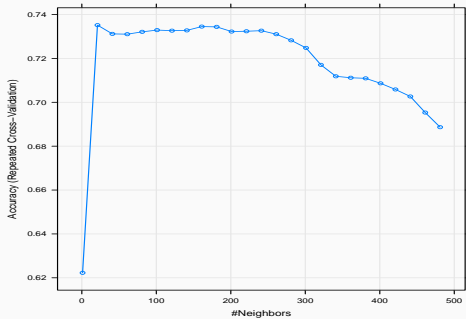
Repeated cross-validation iii

401	0.7086667	0.4104291
421	0.7058667	0.4043432
441	0.7026667	0.3972028
461	0.6953333	0.3813444
481	0.6886667	0.3664347

Accuracy was used to select the optimal model using the largest value.
The final value used for the model was $k = 21$.

```
> plot(e3)
```

Repeated cross-validation iv



Minimizing AUC i

```
> donnees1 <- donnees
> names(donnees1)[3] <- c("Class")
> levels(donnees1$Class) <- c("G0","G1")
> ctrl11 <- trainControl(method="LGOCV",number=1,index=list(1:1500),
                        classProbs=TRUE,summary=twoClassSummary)
> e4 <- train(Class~.,data=donnees1,method="knn",trControl=ctrl11,
              metric="ROC",tuneGrid=KK)
> e4
```

k-Nearest Neighbors

2000 samples

2 predictor

2 classes: 'G0', 'G1'

No pre-processing

Resampling: Repeated Train/Test Splits Estimated (1 reps, 75%)

Summary of sample sizes: 1500

Minimizing AUC ii

Resampling results across tuning parameters:

k	ROC	Sens	Spec
1	0.6190866	0.5983264	0.6398467
21	0.7171484	0.6903766	0.7432950
41	0.7229757	0.6861925	0.7547893
61	0.7200500	0.6945607	0.7394636
81	0.7255567	0.6945607	0.7432950
101	0.7319450	0.6903766	0.7356322
121	0.7382452	0.6945607	0.7356322
141	0.7353757	0.7029289	0.7318008
161	0.7308549	0.7029289	0.7318008
181	0.7351272	0.7029289	0.7318008
201	0.7340050	0.7029289	0.7356322
221	0.7324099	0.7071130	0.7279693
241	0.7349028	0.7071130	0.7279693
261	0.7365780	0.7071130	0.7356322
281	0.7349749	0.6987448	0.7279693

Minimizing AUC iii

301	0.7356963	0.7029289	0.7241379
321	0.7341493	0.6861925	0.7318008
341	0.7343898	0.6527197	0.7356322
361	0.7306385	0.6527197	0.7356322
381	0.7301816	0.6359833	0.7394636
401	0.7270957	0.6276151	0.7356322
421	0.7255487	0.6317992	0.7356322
441	0.7258933	0.6192469	0.7471264
461	0.7220619	0.6150628	0.7471264
481	0.7236330	0.6108787	0.7432950

ROC was used to select the optimal model using the largest value.

The final value used for the model was $k = 121$.

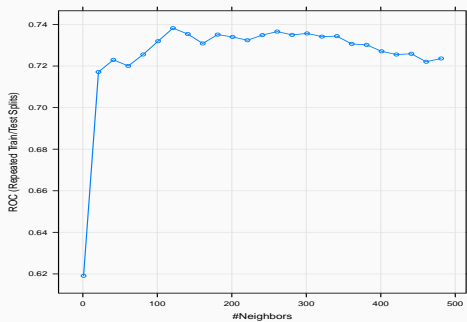
```
> getTrainPerf(e4)
```

	TrainROC	TrainSens	TrainSpec	method
--	----------	-----------	-----------	--------

1	0.7382452	0.6945607	0.7356322	knn
---	-----------	-----------	-----------	-----

```
> plot(e4)
```

Minimizing AUC iv



Summary

- Parametric: strong assumption but fast rates ($1/n$).
- Non parametric: less restrictive but slow rates plus curse of dimensionality ($1/n^{2/(d+2)}$).

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Summary

- Parametric: strong assumption but fast rates ($1/n$).
- Non parametric: less restrictive but slow rates plus curse of dimensionality ($1/n^{2/(d+2)}$).
- ERM strategy: select (automatically) parameters which minimizes the estimated risk.
- Exercise 5, IML1.

Outline

1. Some parametric methods

- Linear and logistic models

- Linear discriminant analysis

 - Just one explanatory variable

 - LDA: general case

2. Some nonparametric methods

- Kernel and nearest neighbors methods

- The curse of dimensionality




3. Empirical risk minimization




- Setting

- Caret package

4. Bibliography

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Part III

Linear model: variable selection and et regularization

1. Subset selection
2. Penalized regression
 - Ridge regression
 - Lasso regression
 - Supervised classification
3. Bibliography

Framework

- $(X_1, Y_1), \dots, (X_n, Y_n)$ i.i.d. observations with the same distribution as (X, Y) which takes values in $\mathcal{X} \times \mathcal{Y}$;
- In this **part**, we assume $\mathcal{X} = \mathbb{R}^d$ and $\mathcal{Y} = \mathbb{R}$ or $\{-1, 1\}$.

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Linear and logistic models

1. If $\mathcal{Y} = \mathbb{R}$,

$$m(x) = \mathbf{E}[Y|X = x] = \beta_0 + \beta_1 x_1 + \dots + \beta_d x_d = x^t \beta.$$

2. If $\mathcal{Y} = \{-1, 1\}$,

$$\text{logit } p(x) = \beta_0 + \beta_1 x_1 + \dots + \beta_d x_d = x^t \beta$$

where $p(x) = \mathbf{P}(Y = 1|X = x)$.

Some limits

- 2 drawbacks in some situations:

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 1. prediction accuracy: LS and ML estimates can have large variance (especially when d is large) and thus poor prediction accuracy.

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 1. prediction accuracy: LS and ML estimates can have large variance (especially when d is large) and thus poor prediction accuracy.
 2. interpretation: when d is large, we don't know what are the most important variables.

Goals

- Since we have more and more data, these drawbacks are occurring more and more often.
- We need to develop new automatic procedures to select important variables.

An example

- We generate observations $(x_i, y_i), i = 1, \dots, 500$ according to

$$Y = 1X_1 + 0X_2 + \dots + 0X_{q+1} + \varepsilon$$

where $X_2, X_{q+1}, \dots, \varepsilon$ are i.i.d. with law $\mathcal{N}(0, 1)$.

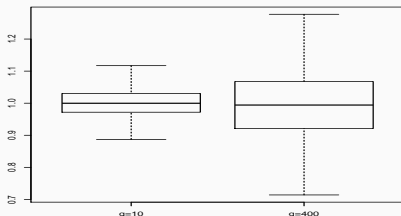
An example

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where $X_2, X_{q+1}, \dots, \varepsilon$ are i.i.d. with law $\mathcal{N}(0, 1)$.

- We compute the **LS estimator of β_1** for 1000 replications. We draw boxplot of these estimators for $q = 10$ and $q = 400$.



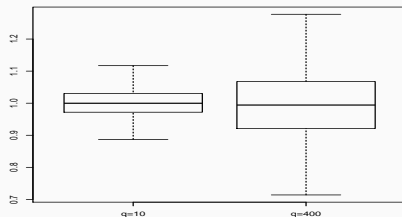
An example

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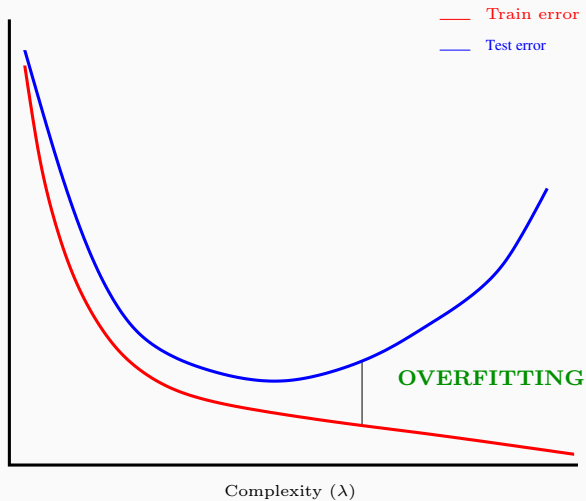
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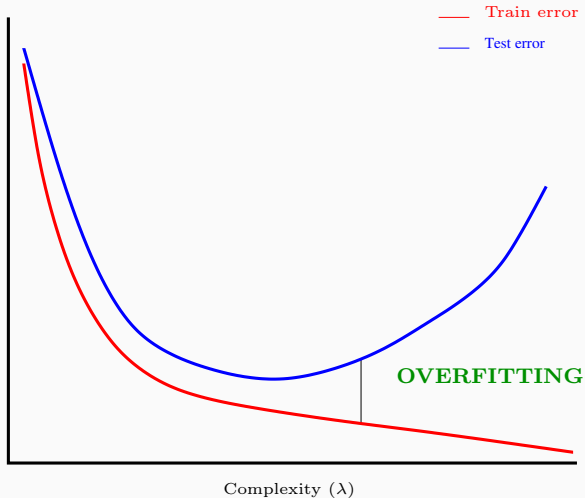
Conclusion

Large **variance** (thus **loss of accuracy**) when the number of **unnecessary variables increases**.

Size of the model



Size of the model



Conclusion

The **size** of the model governs the bias/variance trade-off.

1. Subset selection
2. Penalized regression
 - Ridge regression
 - Lasso regression
 - Supervised classification
3. Bibliography

Best subset selection

- $(X_1, Y_1), \dots, (X_n, Y_n)$ i.i.d. with the same law as (X, Y) which takes values in $\mathbb{R}^d \times \mathbb{R}$;
- d input variables \implies

Best subset selection

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The idea

1. Fit the 2^d models;
2. Choose the one which optimizes a given criterion.

Best subset selection

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The idea

1. Fit the 2^d models;
2. Choose the one which optimizes a given criterion.

Algorithm : best subset selection

1. for $k = 0, \dots, d$:
 - 1.1 Fit the $\binom{d}{k}$ linear models with k variables;
 - 1.2 Choose the model with the higher R^2 . Denote \mathcal{M}_k this model.
2. Select, among $\mathcal{M}_0, \mathcal{M}_1, \dots, \mathcal{M}_d$, the best model according to a given criterion.

Some criteria

- **AIC**: Akaike Information Criterion

$$-2\mathcal{L}_n(\hat{\beta}) + 2d.$$

- **BIC**: Bayesian Information Criterion

$$-2\mathcal{L}_n(\hat{\beta}) + \log(n)d.$$

- **Adjusted R^2** :

$$R_a^2 = 1 - \frac{n-1}{n-d+1}(1-R^2) \quad \text{where} \quad R^2 = \frac{SSR}{SST} = \frac{\|\hat{\mathbf{Y}} - \bar{\mathbf{Y}}\mathbf{1}\|^2}{\|\mathbf{Y} - \bar{\mathbf{Y}}\mathbf{1}\|^2}.$$

- **Mallows's C_p** :

$$C_p = \frac{1}{n} \left(\sum_{i=1}^n (Y_i - \hat{Y}_i)^2 + 2d\hat{\sigma}^2 \right).$$

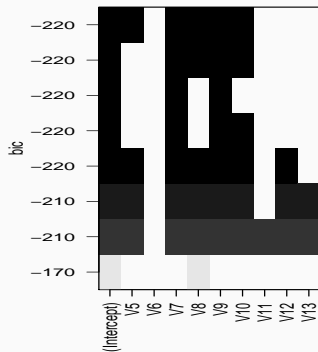
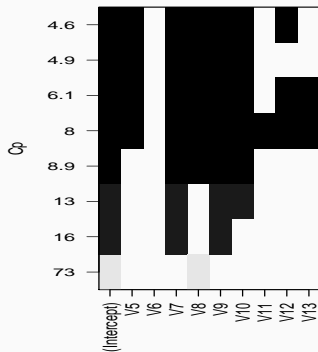
- `regsubsets` from `leaps` package allows to make best subset selection.

```
> library(leaps)
> reg.fit <- regsubsets(V4~V5+V6+V7+V8+V9+V10+V11+V12+V13,data=Ozone)
> summary(reg.fit)

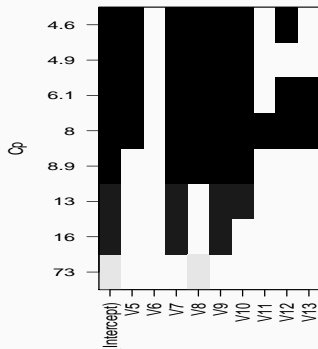
1 subsets of each size up to 8
Selection Algorithm: exhaustive
```

		V5	V6	V7	V8	V9	V10	V11	V12	V13
1	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
2	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
3	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
4	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
5	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
6	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
7	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "
8	(1)	" "	" "	" "	" "	" "	" "	" "	" "	" "

```
> plot(reg.fit,scale="Cp")
> plot(reg.fit,scale="bic")
```



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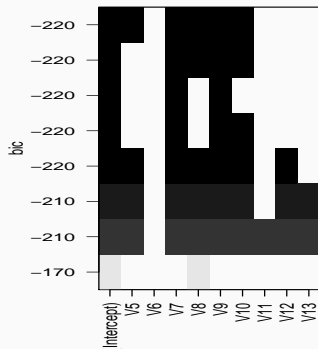


- Mallows's C_p selects:

$$Y = \beta_0 + \beta_1 V_5 + \beta_2 V_7 + \beta_3 V_8 + \beta_4 V_9 + \beta_5 V_{10} + \beta_6 V_{12} + \varepsilon.$$

- BIC selects:

$$Y = \beta_0 + \beta_1 V_5 + \beta_2 V_7 + \beta_3 V_8 + \beta_4 V_9 + \beta_5 V_{10} + \varepsilon.$$



Stepwise selection

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- When d is large, we can seek a good path through all possible subsets.
- Stepwise selection procedures define recursive models by adding or deleting one variable at each step.

Forward stepwise selection

1. Let \mathcal{M}_0 the null model (only the intercept);
2. for $k = 0, \dots, d - 1$:
 - 2.1 Define the $d - k$ models by adding one variable in \mathcal{M}_k ;
 - 2.2 Choose, among those $d - k$ models, the one which maximizes the R^2 .
Denote \mathcal{M}_{k+1} this model.
3. Select, among $\mathcal{M}_0, \dots, \mathcal{M}_d$, the best model according to a given criterion.

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Backward stepwise selection

1. Let \mathcal{M}_d the **full model** (d variables);
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 - 2.1 Define the k models by deleting one variable in \mathcal{M}_k ;
 - 2.2 Choose, among those k models, the one which maximizes R^2 . Denote \mathcal{M}_{k-1} this model.
3. Select, among $\mathcal{M}_0, \dots, \mathcal{M}_d$, the best model according to a given **criterion**.

- We just have to add the argument `method="forward"` or `method="backward"` in `regsubsets` to make subset selection.

```
> reg.fit.for <- regsubsets(V4~V5+V6+V7+V8+V9+V10+V11+V12+V13,data=Ozone,
  method="forward")
> reg.fit.back <- regsubsets(V4~V5+V6+V7+V8+V9+V10+V11+V12+V13,data=Ozone,
  method="backward")
```

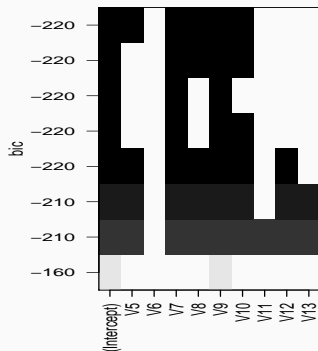
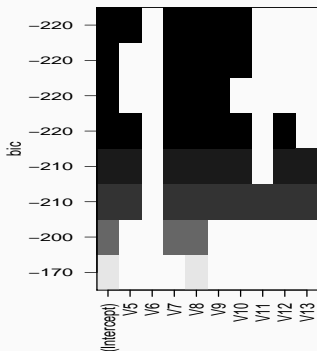
```
> summary(reg.fit.for)
```

		V5	V6	V7	V8	V9	V10	V11	V12	V13
1	(1)	"	"	"	"	"	"	"	"	"
2	(1)	"	"	"	"	"	"	"	"	"
3	(1)	"	"	"	"	"	"	"	"	"
4	(1)	"	"	"	"	"	"	"	"	"
5	(1)	"	"	"	"	"	"	"	"	"
6	(1)	"	"	"	"	"	"	"	"	"
7	(1)	"	"	"	"	"	"	"	"	"
8	(1)	"	"	"	"	"	"	"	"	"

```
> summary(reg.fit.back)
```

		V5	V6	V7	V8	V9	V10	V11	V12	V13
1	(1)	"	"	"	"	"	"	"	"	"
2	(1)	"	"	"	"	"	"	"	"	"
3	(1)	"	"	"	"	"	"	"	"	"
4	(1)	"	"	"	"	"	"	"	"	"
5	(1)	"	"	"	"	"	"	"	"	"
6	(1)	"	"	"	"	"	"	"	"	"
7	(1)	"	"	"	"	"	"	"	"	"
8	(1)	"	"	"	"	"	"	"	"	"

```
> plot(reg.fit.for,scale="bic")
> plot(reg.fit.back,scale="bic")
```



Remark

For this example, forward and backward selection provide the same model (it's not always the case).

- Best subset and stepwise selection have been proposed for regression ($\mathcal{Y} = \mathbb{R}$).

Binary classification

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- These approaches are exactly the same for binary classification ($\mathcal{Y} = \{-1, 1\}$).
- With R, we can use:
 - `bestglm` function from the `bestglm` package for best subset selection.
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- Exercise 1-2, IML2.

1. Subset selection
2. Penalized regression
 - Ridge regression
 - Lasso regression
 - Supervised classification
3. Bibliography

- For large values of d , least square estimates in the linear model

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- **How?** By imposing a constraint on the size of the coefficients:

$$\hat{\beta}^{pen} = \underset{\beta}{\operatorname{argmin}} \sum_{i=1}^n \left(y_i - \sum_{j=1}^d x_{ij} \beta_j \right)^2$$

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subject to $\|\beta\|_? \leq t$.

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- Which **norm** for the constraint?
- How should we **select** t ?
 - t small \implies **strong** constraint ($\hat{\beta}_j \approx 0$) ;
 - t large \implies **small** constraint ($\hat{\beta}_j \approx \hat{\beta}_{j,LS}$).

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2. or equivalently by imposing a penalty on the size of the coefficients

$$\hat{\beta}^R = \underset{\beta}{\operatorname{argmin}} \left\{ \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^d x_{ij} \beta_j \right)^2 + \lambda \sum_{j=1}^d \beta_j^2 \right\}. \quad (3)$$

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- (2) are (3) the same in the sense that there is a one-to-one correspondence between t and λ .
- Ridge estimate depends on t (or λ) : $\hat{\beta}^R = \hat{\beta}^R(t) = \hat{\beta}^R(\lambda)$.
- Input variables are generally standardized to make the variables at the same scale (it is automatic in classical softwares).

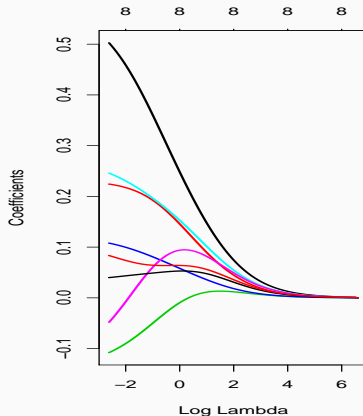
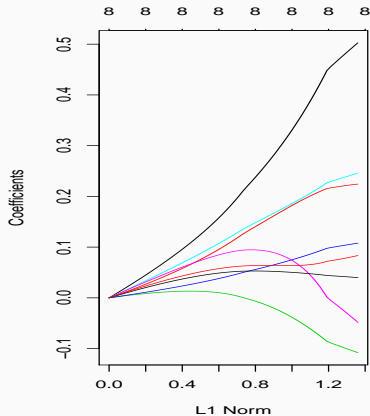
An example

- The problem: explain the level of prostate specific antigen by a number (8) of clinical measures.
- $n = 100$ data available at <https://web.stanford.edu/~hastie/ElemStatLearn/>

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- The problem: explain the level of prostate specific antigen by a number (8) of clinical measures.
- $n = 100$ data available at <https://web.stanford.edu/~hastie/ElemStatLearn/>
- Package `glmnet` allows to make ridge regression on R.

```
> reg.ridge <- glmnet(prostate.data2[,2:9],prostate.data2[,10],alpha=0)
> plot(reg.ridge,label=TRUE)
> plot(reg.ridge,xvar="lambda",label=TRUE,lwd=2)
```



Some properties of ridge estimates

Proposition

1. *Solution of (3) is given by*

$$\hat{\beta}^R = \hat{\beta}^R(\lambda) = (\mathbb{X}^t \mathbb{X} + \lambda \mathbb{I})^{-1} \mathbb{X}^t \mathbb{Y}.$$

2. *It follows that*

$$\text{bias}(\hat{\beta}^R) = -\lambda(\mathbb{X}^t \mathbb{X} + \lambda \mathbb{I})^{-1} \beta$$

and

$$\mathbf{V}(\hat{\beta}^R) = \sigma^2 (\mathbb{X}^t \mathbb{X} + \lambda \mathbb{I})^{-1} \mathbb{X}^t \mathbb{X} (\mathbb{X}^t \mathbb{X} + \lambda \mathbb{I})^{-1}.$$

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Remarks

- For $\lambda = 0$, we obtain LS estimates.
- $\lambda \nearrow \implies \text{bias} \nearrow$ and $\text{variance} \searrow$ and conversely as $\lambda \searrow$.

Choice of λ

- This choice of λ reveals **crucial** for the performance: if $\lambda \approx 0$ then $\hat{\beta}^R \approx \hat{\beta}^{MCO}$, if λ "large" then $\hat{\beta}^R \approx 0$.

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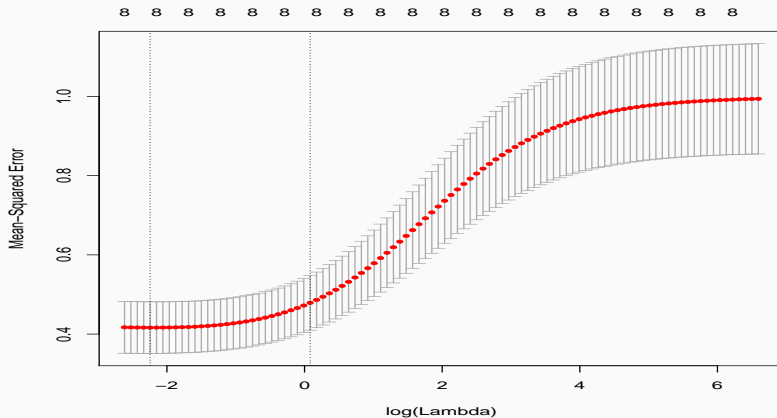
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- The procedure to select λ is usual:
 1. **Estimation of a criterion** for a grid of λ ;
 2. We choose the value of λ which **minimizes** the estimated criterion.
- **Example:** `cv.glmnet` selects the value of λ which minimizes the **quadratic risk**:

$$\mathbb{E}[(Y - X^t \hat{\beta}^R(\lambda))^2]$$

estimated by **cross validation**.

```
> reg.cvridge <- cv.glmnet(prostate.data2[,2:9],prostate.data2[,10],alpha=0)
> bestlam <- reg.cvridge$lambda.min
> bestlam
[1] 0.1060069
> plot(reg.cvridge)
```



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Definition [Tibshirani, 1996]

1. Lasso estimates $\hat{\beta}^L$ minimize

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Comparison Ridge-Lasso

- If \mathbb{X} is an orthonormal input matrix, we have an explicit solution for ridge and lasso.

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Proposition

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$$\hat{\beta}_j^R = \frac{\hat{\beta}_j}{1 + \lambda} \quad \text{and} \quad \hat{\beta}_j^L = \begin{cases} \text{sign}(\hat{\beta}_j)(|\hat{\beta}_j| - \lambda) & \text{if } |\hat{\beta}_j| \geq \lambda \\ 0 & \text{otherwise.} \end{cases}$$

where $\hat{\beta}_j$ is the LS of β_j .

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where $\hat{\beta}_j$ is the LS of β_j .

Comments

- Ridge does a proportional shrinkage;

Comparison Ridge-Lasso

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Proposition

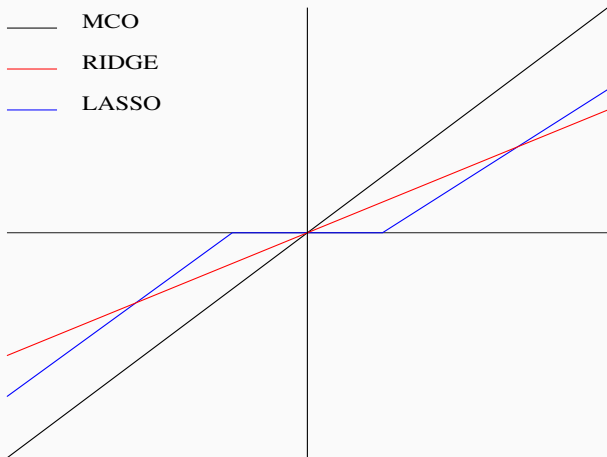
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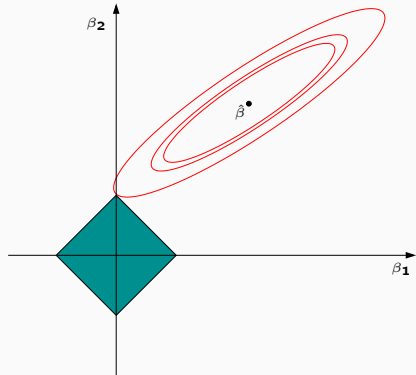
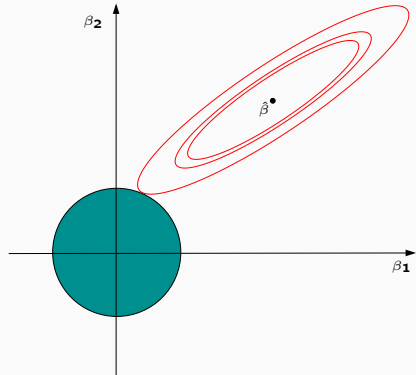
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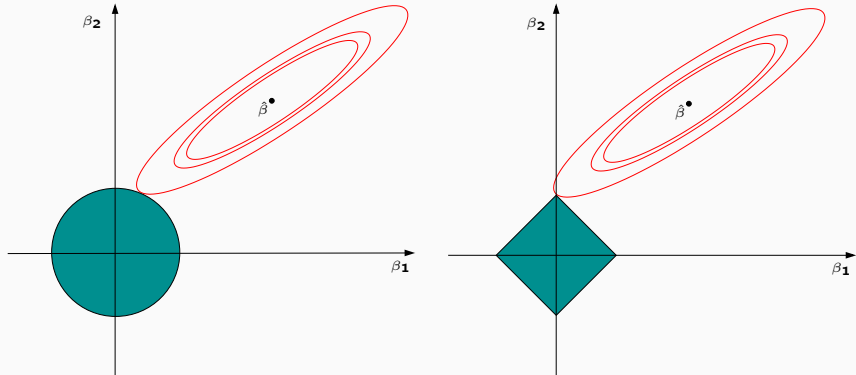
- **Ridge** does a **proportional shrinkage**;
- **Lasso** **translates** each coefficient by a factor λ , **truncating at 0** (when it is small).



Conclusion

Lasso put **small coefficients to 0** \Rightarrow variables with small coefficients are **excluded** from the model.





Relationship between ridge and lasso

Both methods find the first point where the **elliptical contours** hit the **constraint region**:

1. L_2 for **ridge** and L_1 norm for lasso.
2. The diamonds (L_1) has corner \implies the constraint region is often **hit at a corner**.

Some remarks

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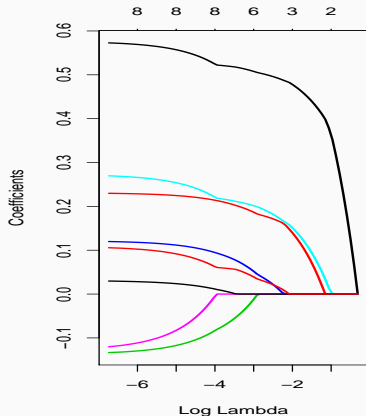
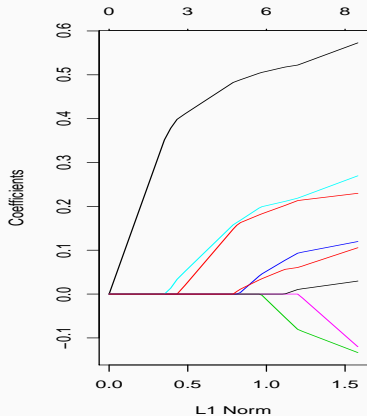
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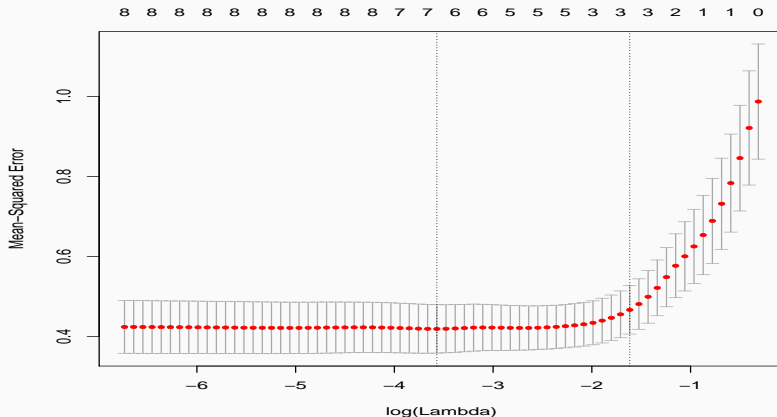
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 - input variables X_1, \dots, X_d are generally **standardized** before the analysis.
 - $\lambda \nearrow \implies$ bias \nearrow and variance \searrow and reciprocally as $\lambda \searrow$.
 - Choice of λ reveals **crucial** (minimization of an estimated criterion).
- **BUT**, unlike ridge: $\lambda \nearrow \implies$ **some estimated parameters equal 0 for lasso** ([Bühlmann and van de Geer, 2011]).

```
> reg.lasso <- glmnet(prostate.data2[,2:9],prostate.data2[,10],alpha=1)
> plot(reg.lasso,label=TRUE)
> plot(reg.lasso,xvar="lambda",label=TRUE,lwd=2)
```



Choice of λ

```
> reg.cvlasso <- cv.glmnet(prostate.data2[,2:9],prostate.data2[,10],alpha=1)
> bestlam <- reg.cvlasso$lambda.min
> bestlam
[1] 0.02815637
> plot(reg.cvlasso)
```



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- It is not difficult to adjust these methods to the logistic model $\mathcal{Y} = \{-1, 1\}$.
- Penalty terms are the same.
- Only change: least square criterion is replaced by likelihood.

Lasso and Ridge for logistic regression

Definition

Let $\tilde{y}_i = (y_i + 1)/2$ ($\tilde{y}_i = 0$ or 1).

- **Ridge estimates** for logistic regression are defined by

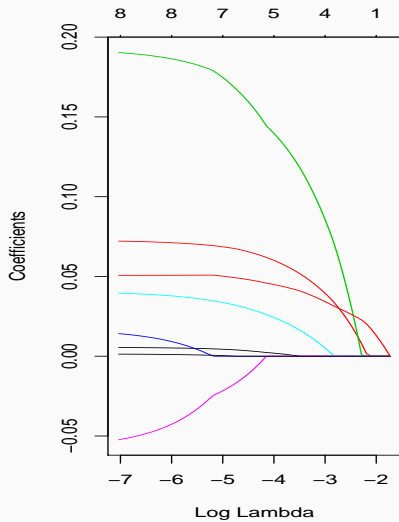
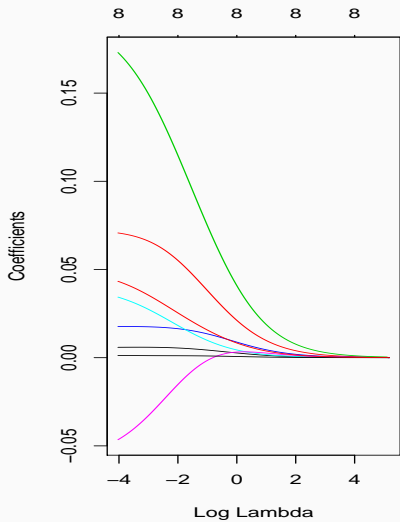
$$\hat{\beta}^R = \underset{\beta}{\operatorname{argmin}} \left\{ - \sum_{i=1}^n (\tilde{y}_i x_i^t \beta - \log(1 + \exp(x_i^t \beta))) + \lambda \sum_{j=1}^d \beta_j^2 \right\}.$$

- **Lasso estimates** for logistic regression are defined by

$$\hat{\beta}^L = \underset{\beta}{\operatorname{argmin}} \left\{ - \sum_{i=1}^n (\tilde{y}_i x_i^t \beta - \log(1 + \exp(x_i^t \beta))) + \lambda \sum_{j=1}^d |\beta_j| \right\}.$$

- To make ridge or lasso for logistic regression, we just have to add **family=binomial** in **glmnet** function.
- **It is the only change** (coefficient paths, choice of λ are the same...).

```
> colnames(donnees)
[1] "sbp"          "tobacco"      "ldl"          "adiposity"    "typea"        "obesity"
[7] "alcohol"      "age"          "chd"
> log.ridge <- glmnet(donnees[,1:8],donnees[,9],family="binomial",alpha=0)
> log.lasso <- glmnet(donnees[,1:8],donnees[,9],family="binomial",alpha=1)
> plot(log.ridge,xvar="lambda")
> plot(log.lasso,xvar="lambda")
```



Elastic net

- [Zou and Hastie, 2005] have proposed to combine ridge and lasso with the following penalty term (called elastic net penalty)

$$\lambda \sum_{j=1}^d ((1 - \alpha)\beta_j^2 + \alpha|\beta_j|)$$

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 - This parameter corresponds (obviously) to the alpha parameter in glmnet function.

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 - This parameter corresponds (obviously) to the alpha parameter in glmnet function.
- Advantage: more flexible since elastic net includes ridge and lasso.

- [Zou and Hastie, 2005] have proposed to combine ridge and lasso with the following penalty term (called elastic net penalty)

$$\lambda \sum_{j=1}^d ((1 - \alpha)\beta_j^2 + \alpha|\beta_j|)$$

where $\alpha \in [0, 1]$.

- α measures the trade-off ridge/lasso :
 - $\alpha = 1 \implies$ Lasso;
 - $\alpha = 0 \implies$ Ridge.
 - This parameter corresponds (obviously) to the alpha parameter in glmnet function.
- **Advantage**: more flexible since elastic net includes ridge and lasso.
- **Drawback**: we have to select both α and λ (you can use caret to do that).

Summary

- LASSO and ridge regressions allow to make efficient linear models when the classical linear model is defective:

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Summary

- LASSO and ridge regressions allow to make **efficient** linear models when the **classical linear model is defective**:
 - **high correlations** between inputs;
 - **high dimension** (large number of inputs).
- When the linear model is efficient, we don't need to use these methods.
- **Exercise 3-4, IML2.**

1. Subset selection
2. Penalized regression
 - Ridge regression
 - Lasso regression
 - Supervised classification
3. Bibliography



Bühlmann, P. and van de Geer, S. (2011).

Statistics for high-dimensional data.

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Hastie, T., Tibshirani, R., and Friedman, J. (2009).

The Elements of Statistical Learning: Data Mining, Inference, and Prediction.

Springer, second edition.



Tibshirani, R. (1996).

Regression shrinkage and selection via the lasso.

Journal of the Royal Statistical Society, Series B, 58:267–288.

 Zou, H. and Hastie, T. (2005).

Regularization and variable selection via the elastic net.

Journal of the Royal Statistical Society, Series B, 67:301–320.

Part IV

Trees

Outline

1. Binary trees
2. Choice of the split
 - Regression
 - Supervised classification
3. Pruning a tree
4. Appendix: pruning algorithm
5. Bibliography

- Tree algorithms are statistical learning algorithms for both regression and supervised classification.
- Popular method, not (too) difficult to understand, visualization tool.

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Presentation

- Tree algorithms are statistical learning algorithms for both regression and supervised classification.
- Popular method, not (too) difficult to understand, visualization tool.
- Tree algorithms are not generally the most performant algorithms... but a lot of efficient algorithms are defined from trees (random forest, gradient tree boosting...).
- There are different ways to build trees.
- We focus on the CART algorithm [Breiman et al., 1984] which is the most widely used algorithm to define trees.

Outline

1. Binary trees
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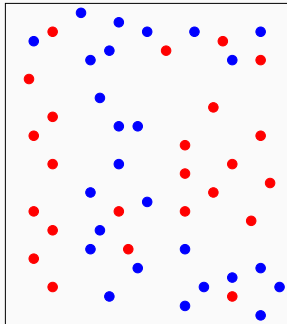
- The problem: explain output Y by p inputs X_1, \dots, X_p .

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- Y might be categorical (binary or not) or continuous and X_1, \dots, X_p categorical or continuous.
- For simplicity (to make figures), we first assume that Y is binary (-1 ou 1) and that $p = 2$ (2 inputs X_1 and X_2 continuous).

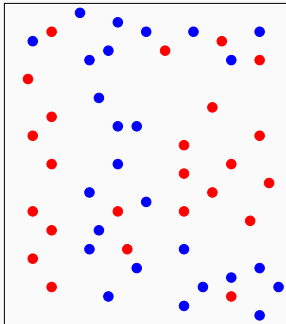
Data visualization

- n observations $(X_1, Y_1), \dots, (X_n, Y_n)$ where $X_i \in \mathbb{R}^2$ and $Y_i \in \{-1, 1\}$.



Data visualization

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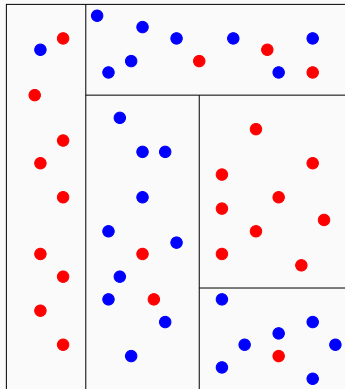
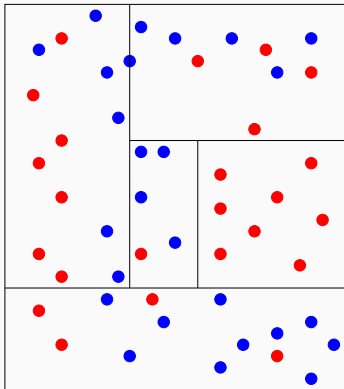


Tree partitions

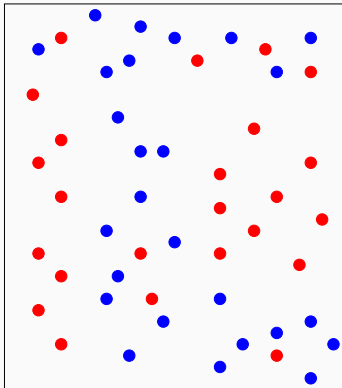
Find a **partition** of the feature space into a set of rectangles which **divides** points according to their color.

Binary partitions

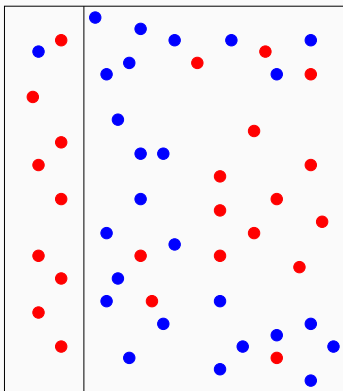
- CART algorithm restricts attention to **recursive** binary partitions.
- 2 examples:



- At each step, the method splits the data into two regions according to a **split variable** and a **split point**.

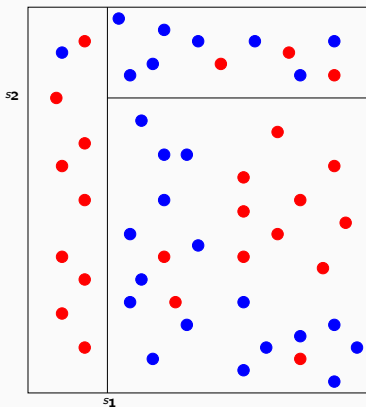


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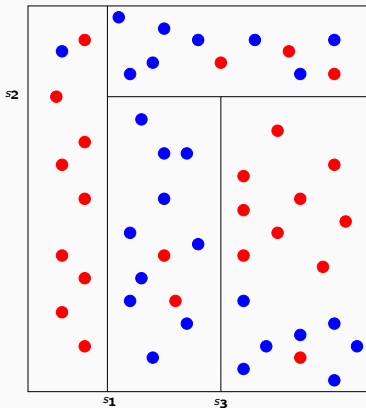


51

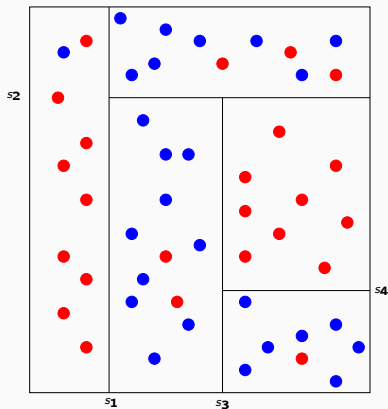
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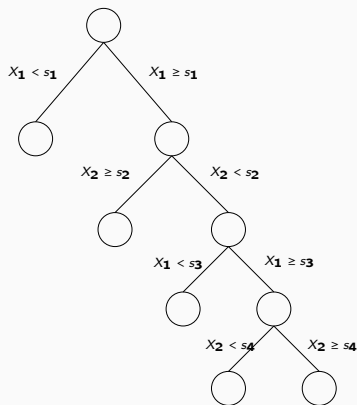
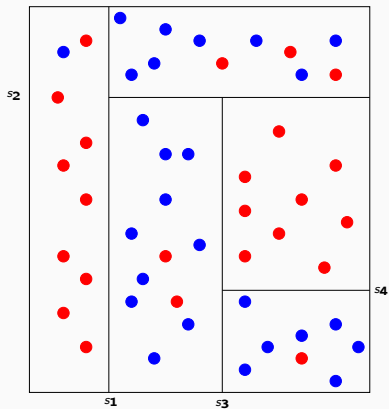
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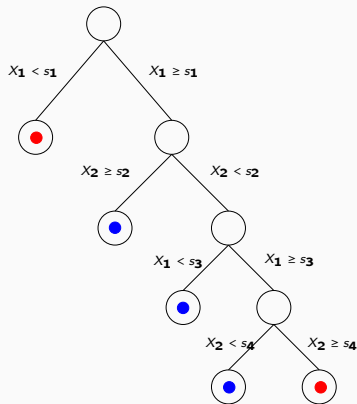
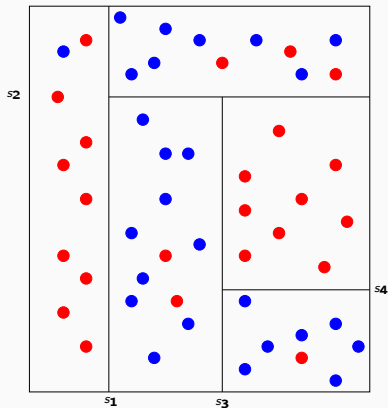
- At each step, the method splits the data into two regions according to a **split variable** and a **split point**.



A tree partition



A tree partition



Classification rule

At the end, we do a **majority vote** in each cell of the partition (in each rectangle).

Definitions

- Each elements of the partition are called **terminal nodes**.
- \mathbb{R}^p (the first node) is the **root node**.
- Each split (each question) defines two child nodes, the **left and right child nodes**.

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- \mathbb{R}^p (the first node) is the **root node**.
- Each split (each question) defines two child nodes, the **left and right child nodes**.

Question

- Tree process is **recursive**: we just have to know how to split a node.
- How to define a good split (or find a good question)?

Outline

1. Binary trees
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Question

How to choose a split?

Question

How to choose a split?

- At each step, we have to find (j, s) which split a node \mathcal{N} into two children nodes

$$\mathcal{N}_1(j, s) = \{X \in \mathcal{N} | X_j \leq s\} \quad \text{and} \quad \mathcal{N}_2(j, s) = \{X \in \mathcal{N} | X_j > s\}.$$

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- (j, s) is selected by minimizing a criterion which measures the impurity of the two children nodes.

- Impurity of a node should be
 1. **small** when the node is homogeneous: values of Y are **closed** to each other in the node.
 2. **large** when the node is heterogeneous: values of Y are **different** from each other in the node.

Impurity

- Impurity of a node should be
 1. **small** when the node is homogeneous: values of Y are **closed** to each other in the node.
 2. **large** when the node is heterogeneous: values of Y are **different** from each other in the node.

The idea

For a given impurity measure \mathcal{I} , we choose the split (j, s) which minimizes

$$\mathbf{P}(\mathcal{N}_1)\mathcal{I}(\mathcal{N}_1(j, s)) + \mathbf{P}(\mathcal{N}_2)\mathcal{I}(\mathcal{N}_2(j, s))$$

where $\mathbf{P}(\mathcal{N}_k)$ stands for the proportion of observations in \mathcal{N}_k , $k = 1, 2$

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- In regression (Y continuous), we usually use the **variance** to measure the impurity in the node

$$\mathcal{I}(\mathcal{N}) = \frac{1}{|\mathcal{N}|} \sum_{i: X_i \in \mathcal{N}} (Y_i - \bar{Y}_{\mathcal{N}})^2,$$

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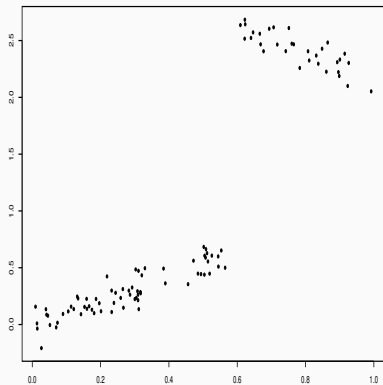
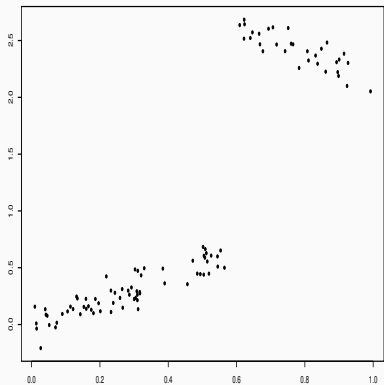
Split for regression

At each step, we choose (j, s) which minimizes

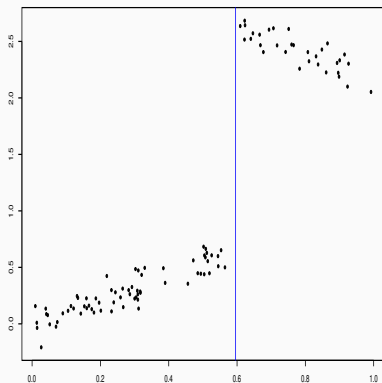
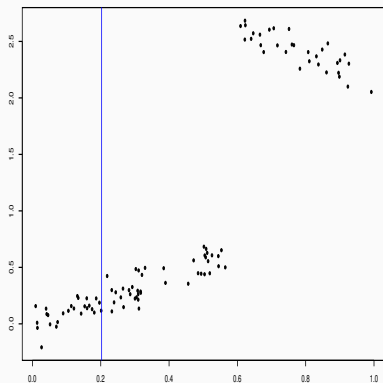
$$\sum_{X_i \in \mathcal{N}_1(j,s)} (Y_i - \bar{Y}_1)^2 + \sum_{X_i \in \mathcal{N}_2(j,s)} (Y_i - \bar{Y}_2)^2$$

where $\bar{Y}_k = \frac{1}{|\mathcal{N}_k(j,s)|} \sum_{X_i \in \mathcal{N}_k(j,s)} Y_i$, $k = 1, 2$.

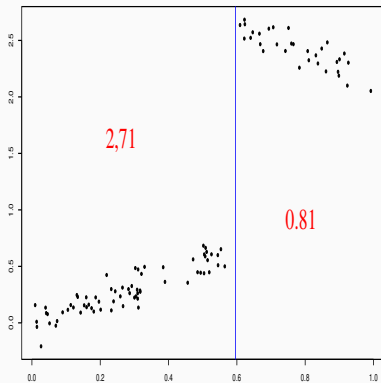
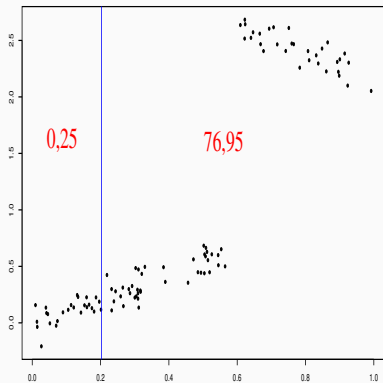
Example



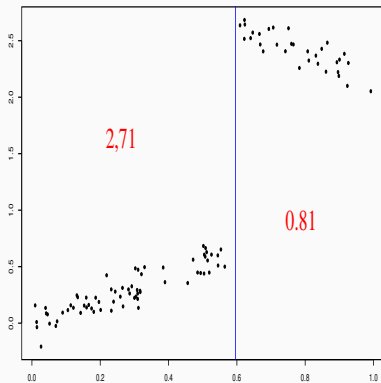
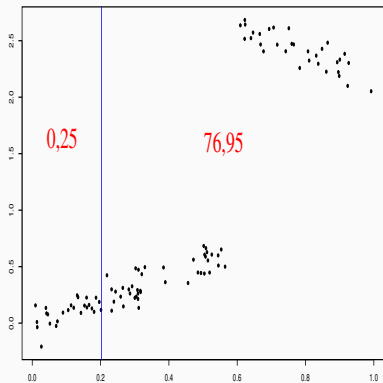
Example



Example



Example



Conclusion

We choose the **right** split.

Outline

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Definition

Impurity of \mathcal{N} is defined by

$$\mathcal{I}(\mathcal{N}) = \sum_{j=1}^K f(p_j(\mathcal{N}))$$

where

- $p_j(\mathcal{N})$ stands for the proportion of class j in \mathcal{N} .
- f is a concave function $[0, 1] \rightarrow \mathbb{R}^+$ such that $f(0) = f(1) = 0$.

Examples of functions f

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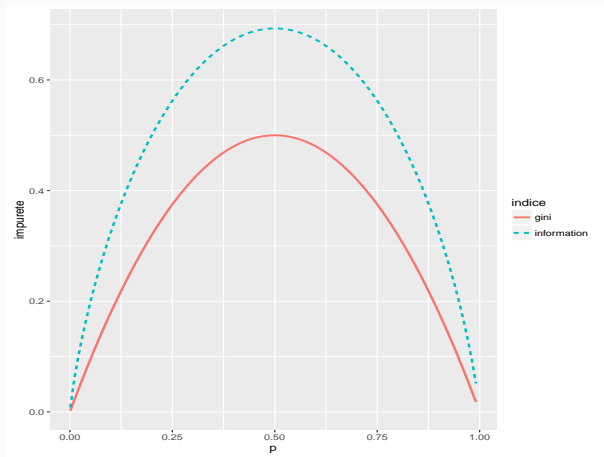
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Binary case

We have

1. $\mathcal{I}(\mathcal{N}) = 2p(1 - p)$ for Gini
2. $\mathcal{I}(\mathcal{N}) = -p \log p - (1 - p) \log(1 - p)$ for Information

where p stands for the proportion of 1 (or -1) in \mathcal{N} .



Split for supervised classification

- Recall that for a given node \mathcal{N} and (j, s) , the two child nodes are defined by

$$\mathcal{N}_1(j, s) = \{X \in \mathcal{N} | X_j \leq s\} \quad \text{and} \quad \mathcal{N}_2(j, s) = \{X \in \mathcal{N} | X_j > s\}.$$

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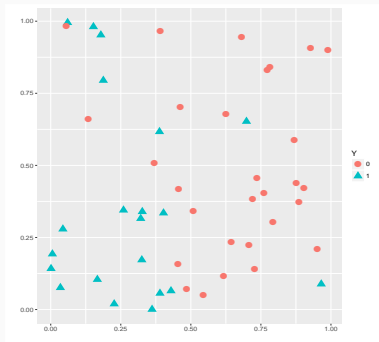
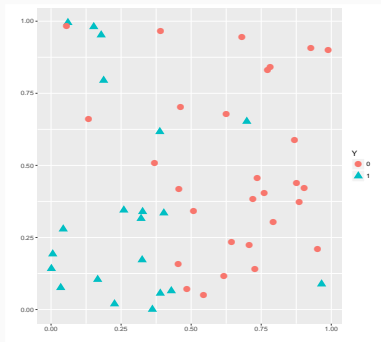
Choice of (j, s)

For a given impurity measure \mathcal{I} , we choose (j, s) which **minimizes**:

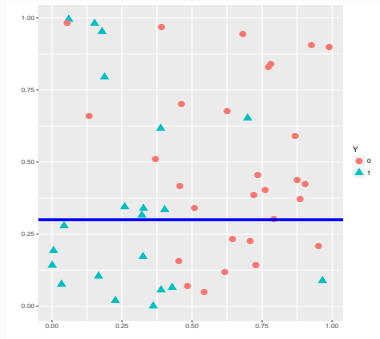
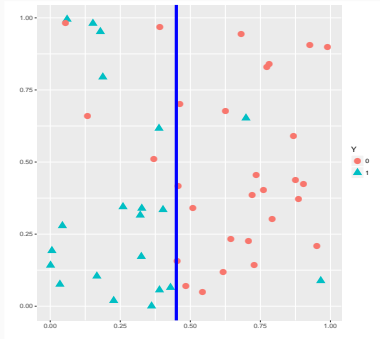
$$\mathbf{P}(\mathcal{N}_1)\mathcal{I}(\mathcal{N}_1(j, s)) + \mathbf{P}(\mathcal{N}_2)\mathcal{I}(\mathcal{N}_2(j, s)).$$

Example

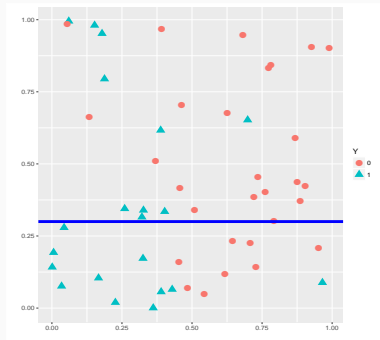
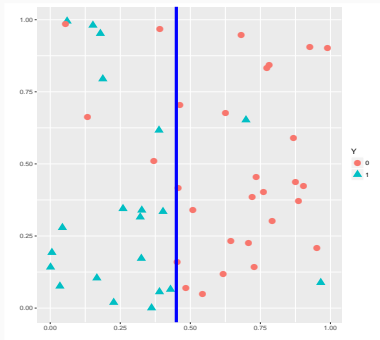
$$\mathcal{I}(\mathcal{N}) = 0.4872$$



Example

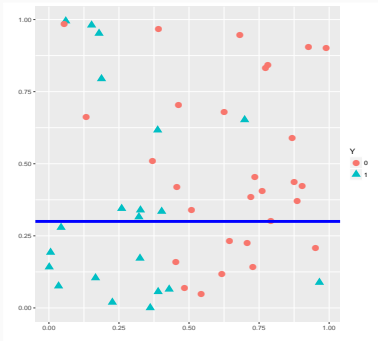
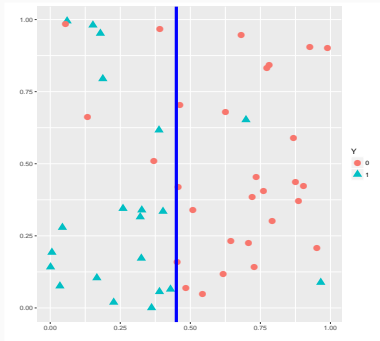


Example



	$\mathcal{I}(\mathcal{N}_1)$	$\mathcal{I}(\mathcal{N}_2)$	Crit.
Left	0.287	0.137	0.2061
Right	0.488	0.437	0.4562

Example



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Left	0.287	0.137	0.2061
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Conclusion

We select the **left** split. (Exercise 1,2,3-IML3.)

Outline

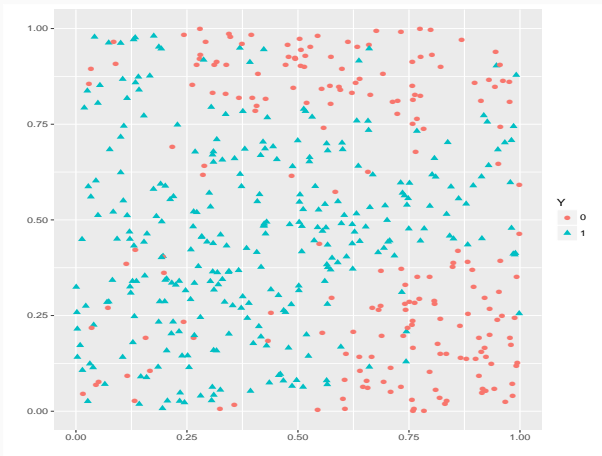
1. Binary trees
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- How to select an efficient tree?

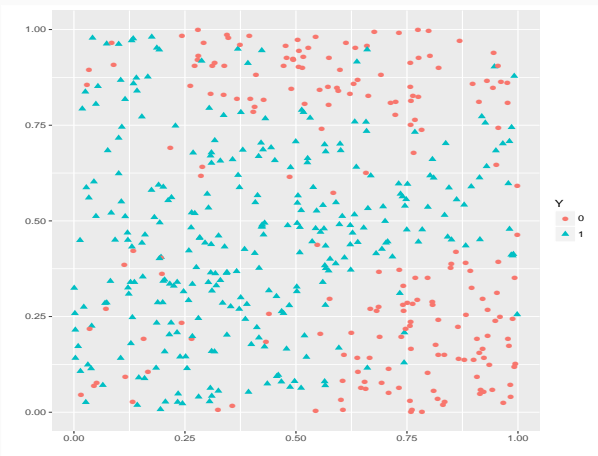
- How to select an **efficient tree**?
- Do we choose the **maximum or deeper** tree? (split the nodes until one observation by node).

- How to select an **efficient tree**?
- Do we choose the **maximum or deeper** tree? (split the nodes until one observation by node).
- Grow a **large tree** and then **prune** this tree (select a subtree of this large tree)?

An example for binary classification



An example for binary classification

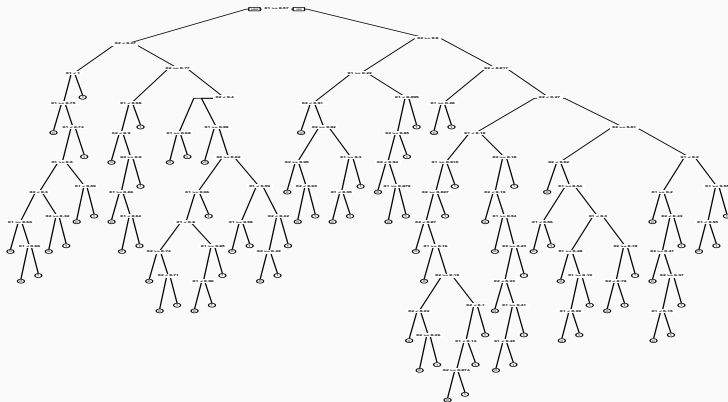


Optimal tree?

Intuitively, we are tempted to choose 5 or 6 terminal nodes.

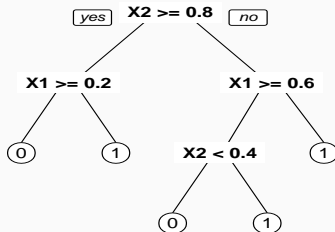
"Deeper" tree

```
> library(rpart)
> library(rpart.plot)
> tree1 <- rpart(Y~.,data=my_data,cp=0.0001,minsplit=2)
> prp(tree1)
```



A smaller tree

```
> tree2 <- rpart(Y~.,data=my_data)  
> prp(tree2)
```



Comparison

- We estimate the **misclassification error** of these two trees on a **test set**.

```
> prev1 <- predict(tree1,newdata=dtest,type="class")
> prev2 <- predict(tree2,newdata=dtest,type="class")
> round(mean(prev1!=dtest$Y),3)
[1] 0.157
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Conclusion

- Performance is **not always improved** by the **size** of the tree.

Comparison

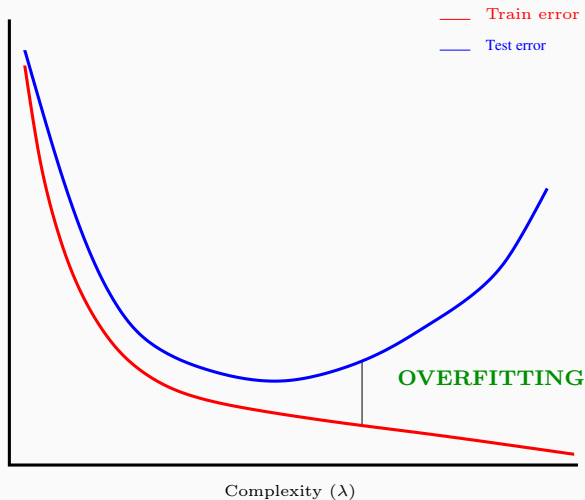
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Conclusion

- Performance is **not always improved** by the **size** of the tree.
- Tree size is a tuning parameter which governs the **model's complexity**. We have to **select** this parameter.

Overfitting



Remark

Complexity is governed by the **depth** (or **size**) of the tree.

Bias and variance

Depth controls the tradeoff bias/variance :

1. Small tree \implies steady (robust) tree \implies small variance... but... large bias.
2. Large tree \implies unsteady tree \implies small bias... but... large variance (overfitting).

Bias and variance

Depth controls the tradeoff bias/variance :

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Pruning [Breiman et al., 1984]

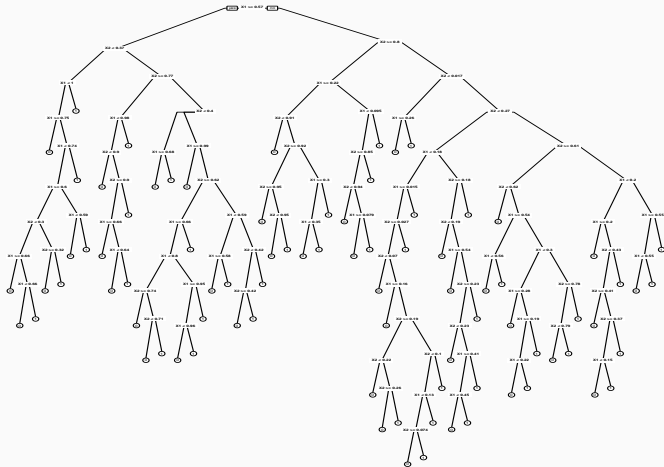
Instead of stopping the splitting process, we

1. grow a large tree (very deep tree) \mathcal{T}_{max} ;
2. then select a sequence of nested subtrees (see Appendix 4.4):

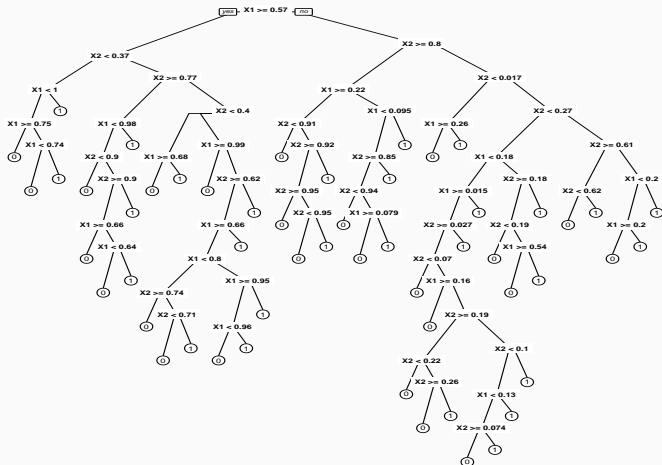
$$\mathcal{T}_{max} = \mathcal{T}_0 \supset \mathcal{T}_1 \supset \cdots \supset \mathcal{T}_K.$$

3. finally select one subtree in this sequence.

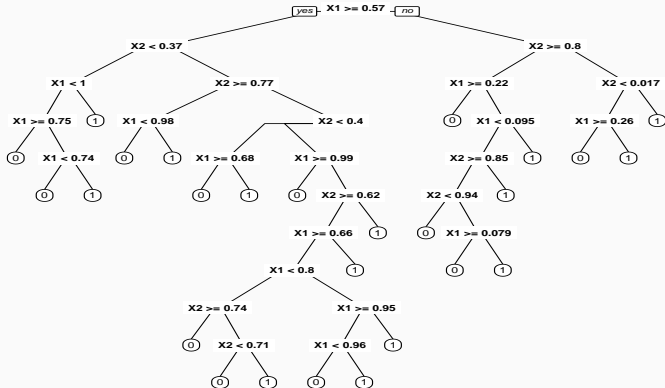
Nested trees



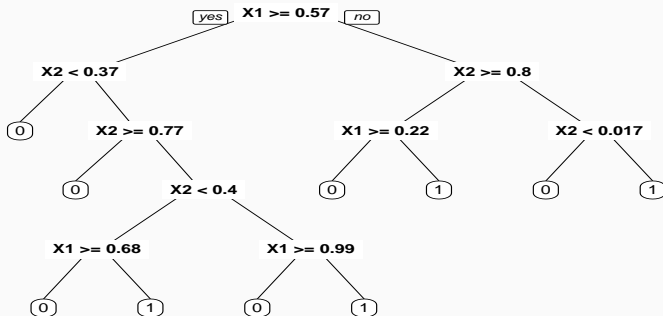
Nested trees



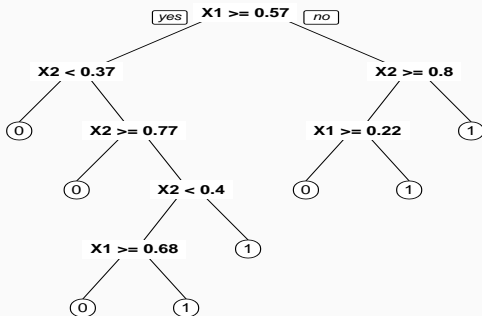
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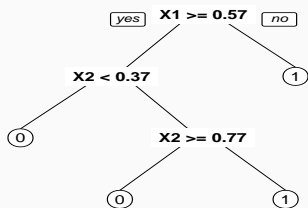
Nested trees



Nested trees



Nested trees



①

Example

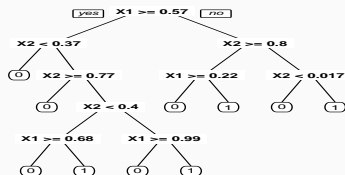
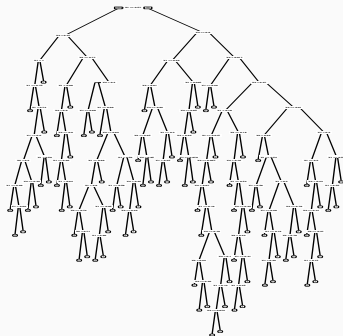
```
> printcp(tree)
Classification tree:
rpart(formula = Y ~ ., data = my_data, cp = 1e-04, minsplit = 2)
Variables actually used in tree construction:
[1] X1 X2
Root node error: 204/500 = 0.408
n= 500
```

	CP	nsplit	rel error	xerror	xstd
1	0.2941176	0	1.000000	1.00000	0.053870
2	0.1225490	1	0.705882	0.71569	0.049838
3	0.0931373	3	0.460784	0.49020	0.043844
4	0.0637255	4	0.367647	0.43627	0.041928
5	0.0122549	5	0.303922	0.34314	0.038034
6	0.0098039	7	0.279412	0.34314	0.038034
7	0.0049020	9	0.259804	0.36275	0.038923
8	0.0040107	25	0.181373	0.34804	0.038260
9	0.0036765	41	0.112745	0.39216	0.040184
10	0.0032680	49	0.083333	0.40196	0.040586
11	0.0024510	52	0.073529	0.41176	0.040980
12	0.0001000	82	0.000000	0.43137	0.041742

```

> arbre1 <- prune(tree,cp=0.005)
> prp(tree)
> prp(tree1)

```



Remark

We have to **select** one tree in the sequence

$$T_{max} = T_0 \supset T_1 \supset \dots \supset T_M.$$

The final tree

Risk estimation

We choose the **final tree** by minimizing a risk $\mathcal{R}(T_m) = \mathbf{E}[\ell(Y, T_m(X))]$ (as usual). For instance,

1. **quadratic risk** $\mathbf{E}[(Y - T_m(X))^2]$ in **regression** ;
2. **misclassification error** $\mathbf{P}(Y \neq T_m(X))$ in **supervised classification**.

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Select the optimal tree

The approach consists in

1. **estimating the risk** for each subtree.
2. selecting the subtree which **minimizes the estimated risk**.

- Estimations of $\mathcal{R}(m)$ are in the column **xerror** of the function **printcp**:

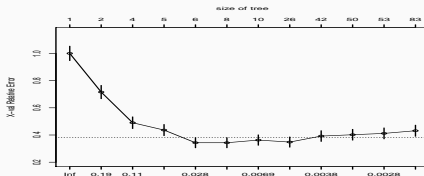
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- We can look at the estimated error for each subtree with **plotcp**

```
> plotcp(tree3)
```

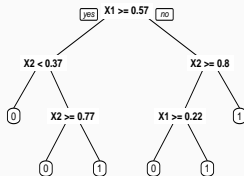


Conclusion

We choose the tree with **5 splits**.

Visualisation of the final tree

```
> alpha_opt <- arbre$cptable[which.min(tree$cptable[, "xerror"]), "CP"]  
> tree_final <- prune(tree, cp=alpha_opt)  
> prp(tree_final)
```



Classification rule and score for a tree

- Final tree \mathcal{T} consists of a partition of \mathbb{R}^p into $|\mathcal{T}|$ terminal nodes $\mathcal{N}_1, \dots, \mathcal{N}_{|\mathcal{T}|}$.

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- **Classification rule:**

$$\hat{g}(x) = \begin{cases} 1 & \text{if } \sum_{i: X_i \in \mathcal{N}(x)} \mathbf{1}_{Y_i=1} \geq \sum_{i: X_i \in \mathcal{N}(x)} \mathbf{1}_{Y_i=0} \\ 0 & \text{otherwise,} \end{cases}$$

where $\mathcal{N}(x)$ stands for the terminal node which contains x .

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- Score:

$$\hat{S}(x) = \hat{\mathbf{P}}(Y = 1 | X = x) = \frac{1}{n} \sum_{i: X_i \in \mathcal{N}(x)} \mathbf{1}_{Y_i=1}.$$

Predict function

- `predict` function (or `predict.rpart`) allows to estimate the label or the score of a new observation:

```
> x_new <- data.frame(X1=0.5,X2=0.85)
> predict(arbre_final,newdata=x_new)
      0      1
1 0.9 0.1
> predict(arbre_final,newdata=x_new,type="class")
1
0
Levels: 0 1
```

Conclusion

- "Simple" method for both regression and supervised classification.
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- This drawback will become an advantage for bootstrap aggregating \implies random forest.
- Exercise 4-IML3.

Outline

1. Binary trees
2. Choice of the split
 - Regression
 - Supervised classification
3. Pruning a tree
4. Appendix: pruning algorithm
5. Bibliography

Construction of the sequence

- Let T be a tree with $|T|$ terminal nodes $\mathcal{N}_1, \dots, \mathcal{N}_{|T|}$.
- Define $R(\mathcal{N})$ the risk (error) in node \mathcal{N} :

- **Regression:**

$$R(\mathcal{N}) = \frac{1}{|\mathcal{N}|} \sum_{i: X_i \in \mathcal{N}} (Y_i - \bar{Y}_{\mathcal{N}})^2.$$

- **Classification:**

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Definition

For $\alpha > 0$,

$$C_{\alpha}(T) = \sum_{m=1}^{|T|} N_m R(\mathcal{N}_m) + \alpha |T|$$

is the **cost complexity criterion** of T .

The idea

- $C_\alpha(T)$ measures both the **fitting** and the **complexity** of the tree.
- The **idea** is to find the subtree T_α which minimizes $C_\alpha(T)$ for a safe choice of α .

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Remark

- $\alpha = 0 \implies T_\alpha = T_0 = T_{max}$.
- $\alpha = +\infty \implies T_\alpha = T_{+\infty} = \text{tree without split}$.
- α is called the **complexity parameter**.

Theorem [Breiman et al., 1984]

There exists a finite sequence $\alpha_0 = 0 < \alpha_1 < \dots < \alpha_M$ with $M < |T_{max}|$ and a sequence of nested trees

$$T_{max} = T_0 \supset T_1 \supset \dots \supset T_M$$

such that $\forall \alpha \in [\alpha_m, \alpha_{m+1}[$

$$T_m = \operatorname{argmin}_T C_\alpha(T).$$

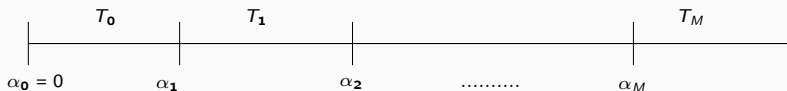
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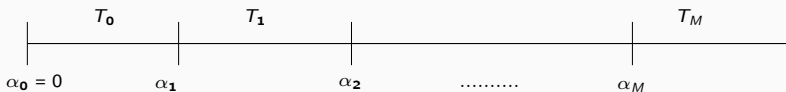
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such that $\forall \alpha \in [\alpha_m, \alpha_{m+1}[$

$$T_m = \underset{T}{\operatorname{argmin}} C_\alpha(T).$$




Important consequence

- We now are faced with a **finite** sequence of nested trees.
- We have to choose one tree in this sequence (or **one value of α**).

Outline

1. Binary trees
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-  Breiman, L., Friedman, J., Olshen, R., and Stone, C. (1984).
Classification and regression trees.
Wadsworth & Brooks.

Part V

Bagging and random forests

1. Bagging
2. Random forests
 - The algorithm
 - OOB error
 - Variable importance
3. Bibliography

1. Bagging

2. Random forests

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3. Bibliography

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The idea

- Instead of fitting one "sophisticated" machine, fit a lot of **simple** machines and **aggregate** them.
- **Example:**

$$\widehat{m}(x) = \frac{1}{B} \sum_{k=1}^B \widehat{m}_k(x)$$

where $\widehat{m}_1(x), \dots, \widehat{m}_B(x)$ are simple machines.

Questions

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- How to define the **simple** machines?
- Do we choose **efficient** simple machines? **Not efficient** (large bias, large variance) machines?
- How many machines?

- **One constraint:** we want to fit simple machines in a similar way (only trees for instance).

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- **Problem:** if you run the same algorithm on the same dataset $(X_1, Y_1), \dots, (X_n, Y_n)$, all simple machines will be the same and

$$\widehat{m}(x) = \frac{1}{B} \sum_{k=1}^B \widehat{m}_k(x) = \widehat{m}_1(x)$$

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\implies aggregation is **useless**.

- **Solution:** run the same algorithm on **different** datasets.

- We have at hand **one** dataset $\mathcal{D}_n = (X_1, Y_1), \dots, (X_n, Y_n)$.

Bootstrap sample

- We have at hand **one** dataset $\mathcal{D}_n = (X_1, Y_1), \dots, (X_n, Y_n)$.
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Bootstrap

- Define new datasets by **randomly** draw dataset with **replacement** from the training data.

Bootstrap: example

- The sample:

1	2	3	4	5	6	7	8	9	10
---	---	---	---	---	---	---	---	---	----

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1	2	3	4	5	6	7	8	9	10
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- Bootstrap samples:

3	4	6	10	3	9	10	7	7	1	m_1
2	8	6	2	10	10	2	9	5	6	m_2
2	9	4	4	7	7	2	3	6	7	m_3
6	1	3	3	9	3	8	10	10	1	m_4
3	7	10	3	2	8	6	9	10	2	m_5
	\vdots								\vdots	
7	10	3	4	9	10	10	8	6	1	m_B

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3	7	10	3	2	8	6	9	10	2	m_5
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- We finally aggregate:

$$\widehat{m}_B(x) = \frac{1}{B} \sum_{k=1}^B m_k(x).$$

Bagging algorithm

- Estimates m_k are not fitted on the original dataset $\mathcal{D}_n = (X_1, Y_1), \dots, (X_n, Y_n)$ but on **bootstrap samples**.

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Inputs:

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- B a positive integer.

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Bagging

Inputs:

- a "simple machine" (a tree, 1NN rule...)
- B a positive integer.

For $k = 1, \dots, B$:

1. Draw a bootstrap sample from \mathcal{D}_n .
2. **Fit the simple machine** on this bootstrap sample: $m_k(x)$.

Output: the aggregate estimate $\widehat{m}_B(x) = \frac{1}{B} \sum_{k=1}^B m_k(x)$.

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- As B increases, \widehat{m}_B stabilizes.

Important conclusion

- B is not an important parameter, we have to choose it as large as possible (often 500).
- Bagging is random but it is less random when B is large.

Some properties

Bias and variance

For regression, we have $\mathbf{E}[\widehat{m}_B(x)] = \mathbf{E}[m_k(x)]$, $\forall k = 1, \dots, B$ and

$$\mathbf{V}[\widehat{m}_B(x)] \approx |\rho(x)| \mathbf{V}[m_k(x)]$$

where $\rho(x) = \text{corr}(m_k(x), m_{k'}(x))$ for $k \neq k'$.

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- **Consequence**: we need simple machines **sensitive to small disturbances of the data**.
- **Trees** are known to satisfy this property (**drawback becomes an advantage...**).

1. Bagging

2. Random forests

The algorithm

OOB error

Variable importance

3. Bibliography

1. Bagging

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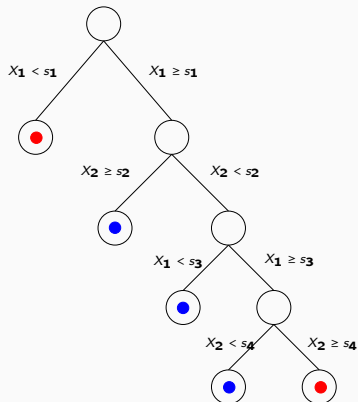
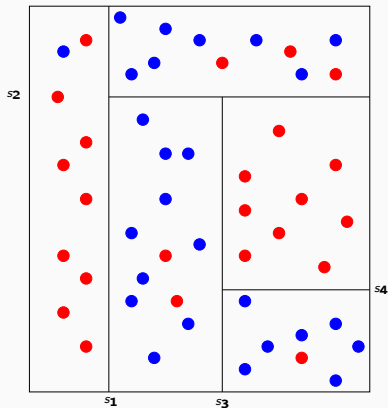
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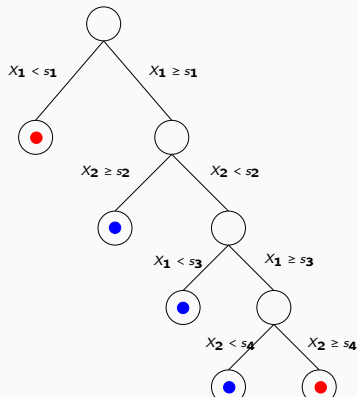
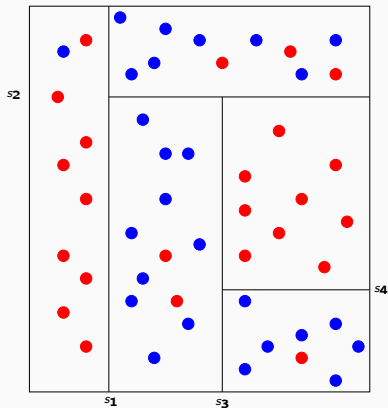
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3. Bibliography

Tree (reminder)



Tree (reminder)



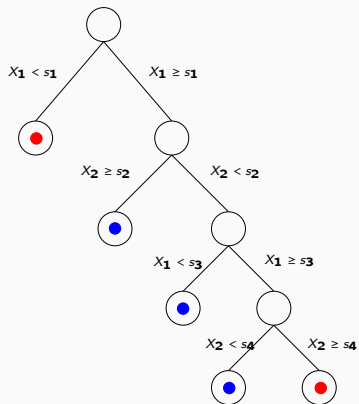
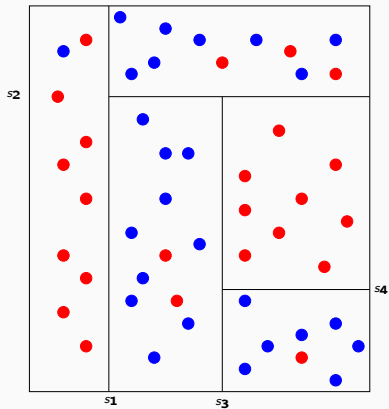
Important parameter: depth

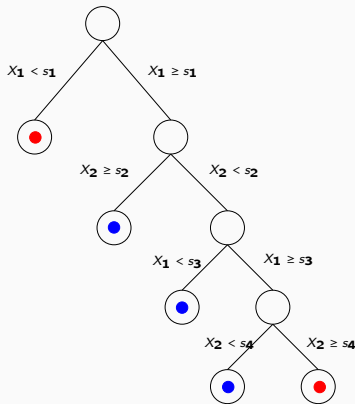
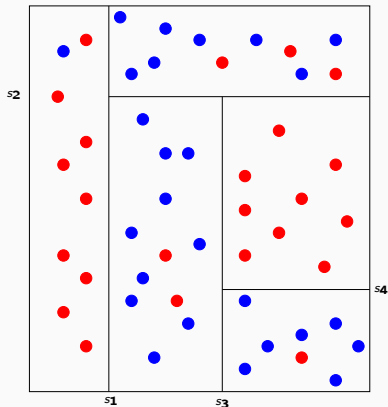
- small: bias ↗, variance ↘
- large: bias ↘, variance ↗

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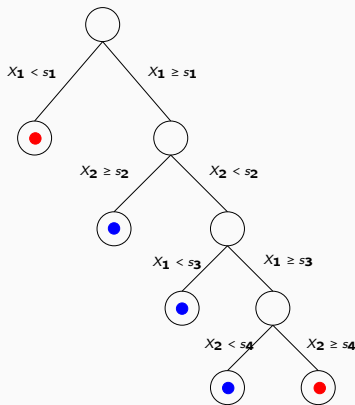
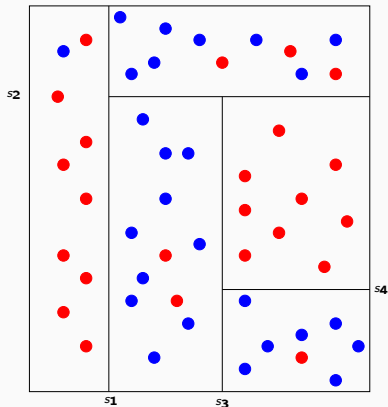
- A random forest = a collection of trees.
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- References
 - <http://www.stat.berkeley.edu/~breiman/RandomForests/>
 - Robin Genuer's phd thesis [Genuer, 2010].
- Trees are fitted as for the CART process (no pruning) with only one small variation.





Trees for the forest

- At each step, the best split is selected among $mtry \leq d$ inputs randomly chosen among the d inputs.



Trees for the forest

- At each step, the best split is selected among $mtry \leq d$ inputs randomly chosen among the d inputs.
- Goal:** try to **reduce correlations** between the trees, to make the trees more different from each other.

Random forest algorithm

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- B size of the forest;
- $mtry \in \{1, \dots, d\}$ number of candidate inputs for each split.

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For $k = 1, \dots, B$:

1. Draw a **bootstrap** sample from \mathcal{D}_n ;
2. Fit a tree according to the CART process, each split is chosen among $mtry$ variables randomly chosen among the d input variables. Denote by $T_k(x)$ the tree.

Output: the random forest $\widehat{T}_B(x) = \frac{1}{B} \sum_{k=1}^B T_k(x)$.

- The algorithm is for both regression and binary classification:
 1. for regression, the RF estimates $m^*(x) = \mathbf{E}[Y|X = x]$;
 2. for binary classification, the RF estimates $S^*(x) = \mathbf{P}(Y = 1|X = x)$.

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- Simple algorithm. On R, you can use randomForest function from the randomForest package or the ranger function from the ranger package.
- Estimate known to be efficient for complex data and robust (wrt to the choice of its parameter).

Choice of the parameter

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Remind

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- Trees for forest are deep trees with a small number of observations in each terminal node.
- By default randomForest fit trees with (only) 5 observations in terminal nodes for regression and 1 for supervised classification.

- This parameter (slightly) governs the bias/variance trade-off of the forest.

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Conclusion

- We can look at the performances of the forest for many values of $mtry$.
- By default $mtry = d/3$ for regression and \sqrt{d} for supervised classification.

Application on the spam dataset

```
> library(randomForest)
> forest1 <- randomForest(type~.,data=spam)
> forest1
```

Call:

```
randomForest(formula = type ~ ., data = spam)
```

```
      Type of random forest: classification
```

```
      Number of trees: 500
```

```
No. of variables tried at each split: 7
```

```
      OOB estimate of  error rate: 5.26%
```

Confusion matrix:

	0	1	class.error
0	1352	42	0.03012912
1	79	827	0.08719647

Outline

1. Bagging

2. Random forests

The algorithm

OOB error

Variable importance

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Random forest performance

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 - Misclassification error $\mathbf{P}(Y \neq \widehat{T}_B(X))$ for supervised classification.

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- These criteria can be estimated by **validation hold out** or **cross validation**.
- **Bootstrap step** in bagging algorithms proposes another way to estimate these criteria: **OOB (Out Of Bag)**.

Out Of Bag error

- For each (X_i, Y_i) , construct its random forest predictor by averaging only those trees corresponding to bootstrap samples in which (X_i, Y_i) does not appear:

$$\hat{Y}_i = \frac{1}{|\mathcal{I}_B|} \sum_{k \in \mathcal{I}_B} T_k(X_i)$$

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Out Of Bag estimates

- OOB quadratic risk: $\frac{1}{n} \sum_{i=1}^n (\hat{Y}_i - Y_i)^2$.
- OOB misclassification error: $\frac{1}{n} \sum_{i=1}^n \mathbf{1}_{\hat{Y}_i \neq Y_i}$.

Example

3	4	6	10	3	9	10	7	7	1	m_1
2	8	6	2	10	10	2	9	5	6	m_2
2	9	4	4	7	7	2	3	6	7	m_3
6	1	3	3	9	3	8	10	10	1	m_4
3	7	10	3	2	8	6	9	10	2	m_5
7	10	3	4	9	10	10	8	6	1	m_6

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7	10	3	4	9	10	10	8	6	1	m_6

- (X_1, Y_1) does not appear in bootstrap samples 2, 3 and 5, thus

$$\hat{Y}_1 = \frac{1}{3}(m_2(X_1) + m_3(X_1) + m_5(X_1)).$$

- We do the same for all the observations $\implies \hat{Y}_2, \dots, \hat{Y}_n$.

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- We do the same for all the observations $\implies \hat{Y}_2, \dots, \hat{Y}_n$.
- We obtain the OOB quadratic risk:

$$\frac{1}{n} \sum_{i=1}^n (\hat{Y}_i - Y_i)^2.$$

Example

- Spam dataset with $mtry = 1$:

```
> forest2 <- randomForest(Y~.,data=spam,mtry=1)
> forest2
```

Call:

```
randomForest(formula = Y ~ ., data = dapp, mtry = 1)
```

```
      Type of random forest: classification
```

```
      Number of trees: 500
```

```
      No. of variables tried at each split: 1
```

```
      OOB estimate of  error rate: 8.04%
```

Confusion matrix:

```
      0    1 class.error
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```

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Conclusion

OOB misclassification error: 8.04% for $mtry = 1$ and 5.26% for $mtry = 7$.

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- Single trees are highly interpretable.
- Linear combinations of trees (random forests) loose this important features.

- **Single trees** are highly interpretable.
- Linear combinations of trees (random forests) **lose** this important features.
- There exists a **score** which measures **importance** of each inputs.
- As for OOB error, this score is based on the fact for some observations **does not appear in bootstrap samples**.

- Let OOB_k denotes the OOB sample of the k -th tree.

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- Let E_{OOB_k} the quadratic error of the k -th tree measured on OOB_k :

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- **Permute** (randomly) the values of input j in $OOB_k \implies OOB_k^j$ and compute the quadratic error on this dataset:

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Definition

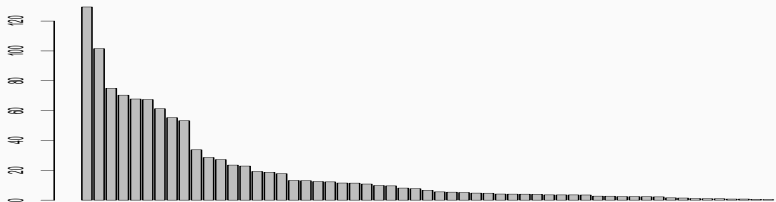
The **variable importance score** for the j variable is defined by

$$Imp(X_j) = \frac{1}{B} \sum_{k=1}^B (E_{OOB_k}^j - E_{OOB_k}).$$

Example

- It is easy to obtain variable importance score with randomForest

```
> imp <- importance(forest1)
> imp1 <- sort(imp,decreasing=TRUE)
> ord <- order(imp,decreasing=TRUE)
> ord
 [1] 52 53 55  7 56 16 21 25 57  5 24 19 26 23 46 27 11  8 50 12 37  3 18  6 45
[26] 17 10  2 28 42 49 35  1 36 39 13 54  9 30 33 22 51 29 14 43 44 31 20 48 15
[51] 40  4 41 34 32 38 47
> barplot(imp1,beside=TRUE)
```



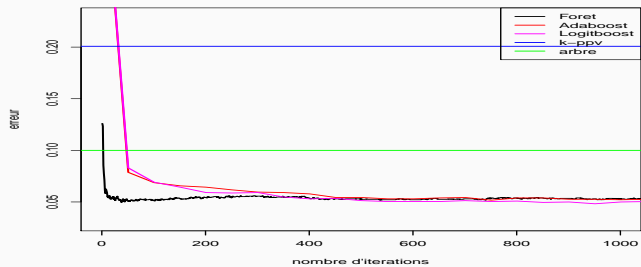
Comparison - spam dataset

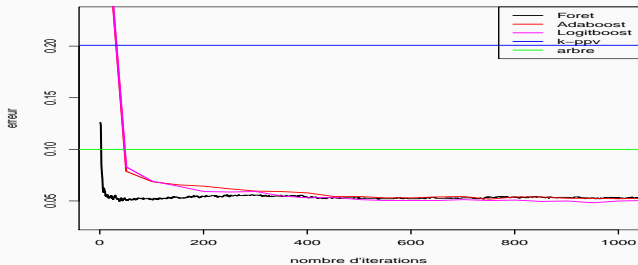
- We make a comparison between some statistical learning algorithms on the spam dataset.

Comparison - spam dataset

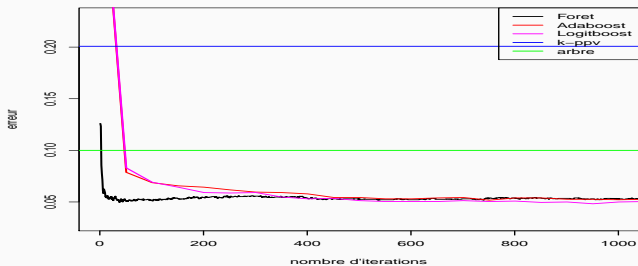
- We make a comparison between some statistical learning algorithms on the spam dataset.
- To do that, we split the data into a
 - a **training set** of size 2300 to **fit and calibrate** the models;
 - a **test set** of size 2301 to **estimate misclassification error** of each model

$$L_n(\hat{g}) = \frac{1}{n_{\text{test}}} \sum_{i \in \mathcal{D}_{\text{test}}} \mathbf{1}_{\hat{g}(X_i) \neq Y_i}.$$






Method	M. error
Random Forest	0.050
Adaboost	0.052
Logitboost	0.048
k -NN	0.200
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- Exercise 5-IML3

1. Bagging
2. Random forests
 - The algorithm
 - OOB error
 - Variable importance
3. Bibliography

 Breiman, L. (1996).
Bagging predictors.
Machine Learning, 26(2):123–140.

 Genuer, R. (2010).
Forêts aléatoires : aspects théoriques, sélection de variables et applications.
PhD thesis, Université Paris XI.

Test - Instructions

- **Document allowed:** 1 sheet A4 format (single sided). No calculators, no laptops, no tablets, no mobile phone...
- Questions using the sign ♣ may have **one or several correct answers**. Other questions have a single correct answer.
- Only the last sheet (answer sheet page 9) is to be returned. You can keep all the other pages.
- Squares corresponding to good answers have to be **colored with a black pen**. Cross or circle marks are not sufficient! It is not possible to correct (once a square has been colored).

Scoring process

- No answer to one question \implies 0 point for the question.
- Questions with a single correct answer: positive score for a good answer, negative score for a bad answer.
- Questions with several correct answers (sign ♣): positive score for each good answer, negative or nul score for each bad answer.

Mistake in exercise 1

- Many question in the exercices, they are not in the same order.
- **be careful**: Exercise 1 should start with: We consider the following tibbles:

```
df1  A tibble ...
```

```
df2  A tibble ...
```

- But in some subjects, these tibbles could be presented:
 - Between Question 1 and Question 2
 - Between Question 2 and Question 3
 - After Question 3

Solution

You have to find the tibbles df1 and df2 before answering to Question 1, Question 2 and Question 3.

Project (machine learning part)

- Find a dataset for a **supervised learning problem** (explain one variable by other variables). This dataset should contain at least 800 individuals and 30 variables (continuous or categorical).
- **Descriptive part:** present data (individuals and variables) and use efficient R tools (dplyr, ggplot...) for data manipulation and visualization.
⇒ **not** a list of graph or summaries! You have to comment each graph and statistical summaries.

Machine learning part

- Identify the **practical problem**;
- **Translate** the practical problem into a **mathematical problem** (Y , X , loss function, risk).
- Propose many **machine learning algorithms** (k -nn, linear/logistic, ridge, lasso, tree, random forest...)
- Define a way to **compare these algorithms** (validation hold out, cross validation...).
- **Be careful**: you have also to **select parameters** for each algorithms... You can look at exercise 6 of the third tutorial.
- **Conclusion**: choice of the **best method** and analysis of its **performances**.

- **Deadline:** December, 15th (11:59 pm).
- Each group should provide a **notebook** (.rmd file) and put on blackboard (you will receive instructions):
 - the **dataset** (.txt, .csv)
 - the **rmd** file and the **html** output file (with figures, R commands, R output...)
- **Be careful** (again): I will **test your codes** by running all the chunks of the notebook (the notebook should be complete!), in case of **problem** with some chunks, you will be **penalized**.

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- More than an Introduction to machine learning.
- Propose a solid mathematical framework to make machine learning.

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THANK YOU