

0 Fundamentals

Normal: $\exp(-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu))$
Beta: $\text{Beta}(\theta; \alpha, \beta) \propto \theta^{\alpha-1} (1-\theta)^{\beta-1}$
Laplace: $\frac{1}{2l} \exp(-\frac{|\mathbf{x}-\mu|}{l})$
Properties of Expectation:
 $\mathbb{E}[\mathbf{g}(\mathbf{X})] = \int_{\mathbf{X}(\Omega)} \mathbf{g}(\mathbf{x}) \cdot p(\mathbf{x}) d\mathbf{x}$ (if \mathbf{g} nice and \mathbf{X} cont.) (**LOTUS**)
 $\mathbb{E}_{\mathbf{Y}}[\mathbb{E}_{\mathbf{X}}[\mathbf{X}|\mathbf{Y}]] = \mathbb{E}[\mathbf{X}]$ (**Tower rule**)
Covariance:
 $\text{Cov}[\mathbf{X}, \mathbf{Y}] \doteq \mathbb{E}[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{Y} - \mathbb{E}[\mathbf{Y}])^T]$
Correlation: $\text{Cor}[\mathbf{X}, \mathbf{Y}](i, j) \doteq \frac{\text{Cov}[X_i, Y_j]}{\sqrt{\text{Var}[X_i] \text{Var}[Y_j]}}$
Variance: $\text{Var}[\mathbf{X}] \doteq \text{Cov}[\mathbf{X}, \mathbf{X}]$
Properties of variance:
 $\text{Var}[\mathbf{A}\mathbf{X} + \mathbf{b}] = \mathbf{A} \text{Var}[\mathbf{X}] \mathbf{A}^T$
 $\text{Var}[\mathbf{X} + \mathbf{Y}] = \text{Var}[\mathbf{X}] + \text{Var}[\mathbf{Y}] + 2\text{Cov}[\mathbf{X}, \mathbf{Y}]$
 $\text{Var}[\mathbf{X}] = \mathbb{E}_{\mathbf{Y}}[\text{Var}_{\mathbf{X}}[\mathbf{X}|\mathbf{Y}]] + \text{Var}_{\mathbf{Y}}[\mathbb{E}_{\mathbf{X}}[\mathbf{X}|\mathbf{Y}]]$ (**LOTV**)
Jensen: Given g convex: $g(\mathbb{E}[X]) \leq \mathbb{E}[g(X)]$
Change of variables formula $\mathbf{Y} = g(\mathbf{X}) \Rightarrow p_{\mathbf{Y}}(\mathbf{y}) = p_{\mathbf{X}}(g^{-1}(\mathbf{y})) \cdot |\det(\mathbf{D}g^{-1}(\mathbf{y}))|$
Bayes' rule: $p(\mathbf{x}|\mathbf{y}) = \frac{p(\mathbf{y}|\mathbf{x}) \cdot p(\mathbf{x})}{p(\mathbf{y})}$

Posterior $p(\mathbf{x}|\mathbf{y})$,
Prior $p(\mathbf{x})$, (Conditional) likelihood $p(\mathbf{y}|\mathbf{x})$, Joint likelihood $p(\mathbf{x}, \mathbf{y})$, Marginal likelihood $p(\mathbf{y})$.
Marginal and conditional of Gaussians:
Given $A, B \subseteq \{1, \dots, n\}$: $\mathbf{X}_A \sim \mathcal{N}(\mu_A, \Sigma_A)$ and $\mathbf{X}_A | \mathbf{x}_B \sim \mathcal{N}(\mu_A + \Sigma_{AB} \Sigma_B^{-1}(\mathbf{x}_B - \mu_B), \Sigma_A - \Sigma_{AB} \Sigma_B^{-1} \Sigma_{BA})$

conjugate iff prior and posterior from same family of distributions.
MLE: $\hat{\theta}_{\text{MLE}} \doteq \arg\max_{\theta \in \Theta} p(y_{1:n} | \mathbf{x}_{1:n}, \theta)$

MAP estimate: $\hat{\theta}_{\text{MAP}} \doteq \arg\max_{\theta \in \Theta} p(\theta | \mathbf{x}_{1:n}, y_{1:n})$
RM conditions Given a function $M(\theta)$ and random variables $N(\theta)$ with $\mathbb{E}[N(\theta)] = M(\theta)$ $\theta_{n+1} \leftarrow \theta_n - a_n(N(\theta_n) - \alpha)$ converges to $M(\theta_*) = \alpha$ if $a_t > 0, \sum_{t=0}^{\infty} a_t = \infty, \sum_{t=0}^{\infty} a_t^2 < \infty$. + some niceness conditions
Woodbury: $(\mathbf{A} + \mathbf{UCV})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1} \mathbf{U}(\mathbf{C}^{-1} + \mathbf{V} \mathbf{A}^{-1} \mathbf{U})^{-1} \mathbf{V} \mathbf{A}^{-1}$

1 Bayesian Linear Regression
Setting: $\mathbf{y} = \mathbf{X}\mathbf{w} + \epsilon, \epsilon \sim \mathcal{N}(\mathbf{0}, \sigma_n^2 \mathbf{I})$
Prior: $\mathbf{w} \sim \mathcal{N}(\mathbf{0}, \sigma_p^2 \mathbf{I})$
Posterior: $\mathbf{w} | \mathcal{D} \sim \mathcal{N}(\mu, \Sigma)$, with $\Sigma \doteq (\sigma_n^{-2} \mathbf{X}^T \mathbf{X} + \sigma_p^{-2} \mathbf{I})^{-1}$ and $\mu \doteq \sigma_n^{-2} \Sigma \mathbf{X}^T \mathbf{y}$

MAP: $\hat{\mathbf{w}}_{\text{MAP}} = \arg\min_{\mathbf{w}} \|\mathbf{y} - \mathbf{X}\mathbf{w}\|_2^2 + \frac{\sigma_p^2}{\sigma_n^2} \|\mathbf{w}\|_2^2$, identical to ridge regression with $\lambda \doteq \frac{\sigma_p^2}{\sigma_n^2}$.
A **Laplace prior** on the weights is equivalent to **lasso regression** with decay $\lambda \doteq \frac{\sigma_n^2}{n}$.
Inference: $\mathbf{y}^* | \mathbf{x}^*, \mathcal{D} \sim \mathcal{N}(\mu^T \mathbf{x}^*, \sigma_n^2 \mathbf{I} + \sigma_n^2 \Sigma \mathbf{x}^* \mathbf{x}^{*T})$.
 $\text{Var}[\mathbf{y}^* | \mathbf{x}^*] = \mathbb{E}_{\theta}[\text{Var}_{\mathbf{y}^*}[\mathbf{y}^* | \mathbf{x}^*, \theta]] + \text{Var}_{\theta}[\mathbb{E}_{\mathbf{y}^*}[\mathbf{y}^* | \mathbf{x}^*, \theta]]$.

$\text{aleatoric uncertainty}$ $\text{epistemic uncertainty}$
 $\mathbf{f} | \mathbf{X} \sim \mathcal{N}(\Phi \mathbb{E}[\mathbf{w}], \Phi \text{Var}[\mathbf{w}] \Phi^T) = \mathcal{N}(\mathbf{0}, \mathbf{K})$, with $\mathbf{K} \doteq \Phi \Phi^T$

Kernel-Linear: $k(\mathbf{x}, \mathbf{x}') \doteq \sigma_p^2 \cdot \phi(\mathbf{x})^T \phi(\mathbf{x}') = \text{Cov}[\mathbf{f}(\mathbf{x}), \mathbf{f}(\mathbf{x}')]$.
Linear: $k(\mathbf{x}, \mathbf{x}') = \mathbf{l}^T \mathbf{x}'$
RBF/Gaussian: $k(\mathbf{x}, \mathbf{x}') = \exp(-\frac{\|\mathbf{x} - \mathbf{x}'\|_2^2}{2\sigma_p^2})$
Polynomial: $k(\mathbf{x}, \mathbf{x}') = (1 + \mathbf{x}^T \mathbf{x}')^d$
Laplacian: $k(\mathbf{x}, \mathbf{x}') = \exp(-\alpha \|\mathbf{x} - \mathbf{x}'\|)$
Matern: $\frac{2^{1-\nu}}{\Gamma(\nu)} \left(\sqrt{2\nu} \frac{\|\mathbf{x} - \mathbf{x}'\|_2}{\sigma_p} \right)^{\nu} K_{\nu} \left(\sqrt{2\nu} \frac{\|\mathbf{x} - \mathbf{x}'\|_2}{\sigma_p} \right)$
Stationary: $k(\mathbf{x}, \mathbf{x}') = k(\mathbf{x} - \mathbf{x}')$
Isotropic: $k(\mathbf{x}, \mathbf{x}') = \tilde{k}(\|\mathbf{x} - \mathbf{x}'\|_2)$.
Properties of Kernels:
K_{AA} is symmetric and **p.s.d.**
Composition: addition, multiplication, and composition with a function f with positive coefficients in Taylor expansion. **Bochner's Theorem:** A continuous kernel on \mathbb{R}^d is p.s.d iff its Fourier transform $p(\omega)$ is non-negative.
Cost: $\mathcal{O}(d^3 + nd^2)$, can be performed online with cost $\mathcal{O}(d^2)$ per iteration.

2 Filtering
Kalman filter: $X_0 \sim \mathcal{N}(\mu, \Sigma)$
 $X_{t+1} = F X_t + \epsilon_t, \epsilon_t \sim \mathcal{N}(\mathbf{0}, \Sigma_x)$
 $Y_t = H X_t + \eta_t, \eta_t \sim \mathcal{N}(\mathbf{0}, \Sigma_y)$
Conditioning: compute $p(x_{t+1} | y_{1:t})$ from observing y_t
Prediction: compute $p(x_{t+1} | y_{1:t})$
 $X_{t+1} | y_{1:t+1} \sim \mathcal{N}(\mu_{t+1}, \Sigma_{t+1})$, with $\mu_{t+1} = F \mu_t + \mathbf{K}_{t+1}(y_{t+1} - H F \mu_t)$ and $\Sigma_{t+1} = (\mathbf{I} - \mathbf{K}_{t+1} H)(F \Sigma_t F^T + \Sigma_x)$
Kalman gain: $\mathbf{K}_{t+1} \doteq (F \Sigma_t F^T + \Sigma_x) H^T (H(F \Sigma_t F^T + \Sigma_x) H^T + \Sigma_y)^{-1}$

3 Gaussian Processes
Gaussian process = infinite set of random variables s.t. any finite number of them are jointly Gaussian.
 $f \sim \mathcal{GP}(\mu, k)$ with $\mu: \mathcal{X} \rightarrow \mathbb{R}, k: \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$
 $\forall A \subseteq \{\mathbf{x}_1, \dots, \mathbf{x}_m\} \subseteq \mathcal{X}: \mathbf{f}_A \sim \mathcal{N}(\mu_A, \mathbf{K}_{AA})$
Posterior: $f | \mathcal{D} \sim \mathcal{GP}(\mu', k')$,
 $\mu'(\mathbf{x}) \doteq \mu(\mathbf{x}) + \mathbf{k}_{\mathbf{x}, A}^T (\mathbf{K}_{AA} + \sigma_n^2 \mathbf{I})^{-1} (\mathbf{y}_A - \mu_A)$,
 $k'(\mathbf{x}, \mathbf{x}') \doteq k(\mathbf{x}, \mathbf{x}') - \mathbf{k}_{\mathbf{x}, A}^T (\mathbf{K}_{AA} + \sigma_n^2 \mathbf{I})^{-1} \mathbf{k}_{\mathbf{x}', A}$
Cost: $\mathcal{O}(n^3)$
Maximize Marginal Likelihood:
 $\hat{\theta}_{\text{MLE}} \doteq \arg\max_{\theta} p(y_{1:n} | \mathbf{x}_{1:n}, \theta) = \arg\max_{\theta} \int p(y_{1:n} | \mathbf{x}_{1:n}, f, \theta) p(f | \theta) df$.
Write $\mathbf{K}_{\mathbf{y}, \theta} \doteq \mathbf{K}_{f, \theta} + \sigma_n^2 \mathbf{I}$, and obtain:
 $\hat{\theta}_{\text{MLE}} = \arg\min_{\theta} \frac{1}{2} \mathbf{y}^T \mathbf{K}_{\mathbf{y}, \theta}^{-1} \mathbf{y} + \frac{1}{2} \log \det(\mathbf{K}_{\mathbf{y}, \theta})$.

$\frac{\partial}{\partial \theta_j} \log p(\mathbf{y} | \mathbf{X}, \theta) = \frac{1}{2} \text{tr} \left(\left(\alpha \mathbf{T} - \mathbf{K}_{\mathbf{y}, \theta}^{-1} \right) \frac{\partial \mathbf{K}_{\mathbf{y}, \theta}}{\partial \theta_j} \right), \alpha = \mathbf{K}_{\mathbf{y}, \theta}^{-1} \mathbf{y}$
Local methods: Only condition on \mathbf{x}' where $|k(\mathbf{x}, \mathbf{x}')| \geq \tau$.
Kernel Approximation: Construct a low dimensional feature map $\phi: \mathbb{R}^d \rightarrow \mathbb{R}^m$ that approximates the kernel: $k(\mathbf{x}, \mathbf{x}') \approx \phi(\mathbf{x})^T \phi(\mathbf{x}')$, then apply BLR.
Random Fourier features: given k stationary, $p(\omega) = \int_{\mathbb{R}^d} k(\xi) e^{-i \xi^T \omega} d\xi$
 $k(\mathbf{x} - \mathbf{x}') = \int_{\mathbb{R}^d} p(\omega) e^{i \omega^T (\mathbf{x} - \mathbf{x}')} d\omega = \mathbb{E}_{\omega \sim p, b \sim \mathcal{U}([0, 2\pi])} [2 \cos(\omega^T \mathbf{x} + b) \cos(\omega^T \mathbf{y} + b)]$.
 $\phi(\mathbf{x}) = \frac{1}{\sqrt{m}} [z_{\omega_1, b_1}(\mathbf{x}), \dots, z_{\omega_m, b_m}(\mathbf{x})]^T$ with $z_{\omega, b}(\mathbf{x}) = \sqrt{2} \cos(\omega^T \mathbf{x} + b)$

Inducing points and approximate the kernel matrix with a low rank approximation. **SoR:** assume 0 covariance
FITC: assume diagonal covariance
Runtime: $\mathcal{O}(nk^2)$
4 Variational Inference
Approximate $p(\theta | \mathbf{x}_{1:n}, y_{1:n})$ with $q_{\lambda}(\theta) \in \mathcal{Q}$
Laplace Approximation:
 $q(\theta) \doteq \mathcal{N}(\hat{\theta}; \hat{\theta}, \Lambda^{-1}) \propto \exp(\hat{\psi}(\theta))$, with $\hat{\theta}$ the mode and $\Lambda \doteq -\mathbf{H}_{\psi}(\hat{\theta}) = -\mathbf{H}_{\theta} \log p(\theta | \mathcal{D})|_{\theta=\hat{\theta}}$.
Inference: $p(\mathbf{y}^* | \mathbf{x}^*, \mathcal{D}) \approx \int p(\mathbf{y}^* | \mathbf{x}^*, \theta) q_{\lambda}(\theta) d\theta$.
Suprise: $S[\mathbf{u}] \doteq -\log u$ (convex).
Entropy: $H[p] \doteq \mathbb{E}_{\mathbf{x} \sim p}[S[p(\mathbf{x})]]$.
Cross-entropy: $H[p||q] \doteq \mathbb{E}_{\mathbf{x} \sim p}[S[q(\mathbf{x})]]$.
KL divergence: $\text{KL}(p||q) \doteq H[p||q] - H[p] = \mathbb{E}_{\mathbf{x} \sim p} \left[\log \frac{p(\mathbf{x})}{q(\mathbf{x})} \right]$.
Gaussian: $H[\mathcal{N}(\mu, \Sigma)] = \frac{1}{2} \log((2\pi e)^d \det(\Sigma))$.
 $\text{KL}(p||q) \geq 0$ (Gibbs); $\text{KL}(p||q) = 0$ iff $p = q$ almost everywhere. $\mathcal{N}(\mu, \Sigma)$ has the **highest entropy** among all distributions mean μ and variance Σ .
 $\text{KL}(\text{Bern}(p) || \text{Bern}(q)) = p \log \frac{p}{q} + (1-p) \log \frac{(1-p)}{(1-q)}$
Gaussian KL: $p \sim \mathcal{N}(\mu_p, \Sigma_p), q \sim \mathcal{N}(\mu_q, \Sigma_q)$:
 $\text{KL}(p||q) = \frac{1}{2} (\text{tr}(\Sigma_q^{-1} \Sigma_p) + (\mu_p - \mu_q)^T \Sigma_q^{-1} (\mu_p - \mu_q) - d + \log \frac{\det(\Sigma_q)}{\det(\Sigma_p)})$.
Forward KL: $q_1^* \doteq \arg\min_{q \in \mathcal{Q}} \text{KL}(p||q)$; minimize forward KL by **moment matching**
Reverse KL: $q_2^* \doteq \arg\min_{q \in \mathcal{Q}} \text{KL}(q||p)$. Reverse KL tends to greedily select the mode and underestimate the variance.
Evidence lower bound (ELBO)
 $L(q, p; \mathcal{D}) = p(\mathcal{D}) - \text{KL}(q||p(\cdot | \mathcal{D})) = \mathbb{E}_{\theta \sim q} [p(\mathcal{D} | \theta)] - \text{KL}(q||p)$

Reparametrization trick: For $\epsilon \sim \phi$ independent of λ s.t. $\theta = \mathbf{g}(\epsilon; \lambda)$, then: $\mathbb{E}_{\theta \sim q_{\lambda}}[\mathbf{f}(\theta)] = \mathbb{E}_{\epsilon \sim \phi}[\mathbf{f}(\mathbf{g}(\epsilon; \lambda))]$. For ELBO: $\nabla_{\lambda} \mathbb{E}_{\theta \sim q_{\lambda}}[\mathbf{f}(\theta)] = \mathbb{E}_{\epsilon \sim \phi}[\nabla_{\lambda} \mathbf{f}(\mathbf{g}(\epsilon; \lambda))]$.
Gaussian: $q_{\lambda}(\theta) \doteq \mathcal{N}(\theta; \mu, \Sigma)$; $\epsilon \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$, set: $\theta = \mathbf{g}(\epsilon; \lambda) \doteq \Sigma^{1/2} \epsilon + \mu$, then: $\phi(\epsilon) = q_{\lambda}(\theta) \cdot |\det(\Sigma^{1/2})|$ and $\epsilon = \mathbf{g}^{-1}(\theta; \lambda) = \Sigma^{-1/2}(\theta - \mu)$
5 Markov Chains
A Markov Chain over $S \doteq \{0, \dots, n-1\}$, is a sequence $(X_t)_{t \in \mathbb{N}_0} \in S$, such that the **Markov property:** $X_{t+1} \perp X_{0:t-1} | X_t$ is satisfied.
Time-homogeneous if there exists $p(\mathbf{x}' | \mathbf{x}) \doteq \mathbb{P}(X_{t+1} = \mathbf{x}' | X_t = \mathbf{x})$, with transition matrix $P_{ij} = p(\mathbf{x}_j | \mathbf{x}_i)$. Each row sums up to 1. The state of a MC at t is a probability distribution q_t ; $q_{t+1} = q_t \mathbf{P}$. A distribution π is **stationary** iff $\pi = \pi \mathbf{P}$.
Irreducible: $\forall \mathbf{x}, \mathbf{x}' \in S \exists t \in \mathbb{N}: p^{(t)}(\mathbf{x}' | \mathbf{x}) > 0$.
Aperiodic: $\forall \mathbf{x} \in S \exists t_0 \in \mathbb{N} \forall t \geq t_0: p^{(t)}(\mathbf{x} | \mathbf{x}) > 0$.
Ergodic: $\exists t \in \mathbb{N} \forall \mathbf{x}, \mathbf{x}' \in S: p^{(t)}(\mathbf{x}' | \mathbf{x}) > 0$.
Irreducible MC \rightarrow ergodic MC use: $\mathbf{P}^* \doteq \frac{1}{2} \mathbf{P} + \frac{1}{2} \mathbf{I}$
An ergodic MC has a unique stat. dist. π (with full support) and $\lim_{t \rightarrow \infty} q_t = \pi$, independently of q_0 .
Detailed balance equation: $\forall \mathbf{x}, \mathbf{x}' \in S: \pi(\mathbf{x}) p(\mathbf{x}' | \mathbf{x}) = \pi(\mathbf{x}') p(\mathbf{x} | \mathbf{x}')$ \Rightarrow MC is reversible w.r.t. $\pi \Rightarrow \pi$ is stationary.
Ergodic theorem For an ergodic MC and a stat. dist. π as well as $f: S \rightarrow \mathbb{R}$: $\frac{1}{n} \sum_{i=1}^n f(x_i) \xrightarrow{\text{a.s.}} \sum_{\mathbf{x} \in S} \pi(\mathbf{x}) f(\mathbf{x}) = \mathbb{E}_{\mathbf{x} \sim \pi}[f(\mathbf{x})]$, for $n \rightarrow \infty$ where $x_i \sim X_i | x_{i-1}$.

Acceptance distribution (Metropolis-Hastings): $\text{Bern}(\alpha(\mathbf{x}' | \mathbf{x}))$ where $\alpha(\mathbf{x}' | \mathbf{x}) \doteq \min \left\{ 1, \frac{q(\mathbf{x}) r(\mathbf{x} | \mathbf{x}')}{q(\mathbf{x}') r(\mathbf{x}' | \mathbf{x})} \right\}$ to decide whether to follow the proposal yields a Markov chain with stationary distribution $p(\mathbf{x}) = \frac{1}{Z} q(\mathbf{x})$. **Gibbs distribution:** $p(\mathbf{x}) = \frac{1}{Z} \exp(-f(\mathbf{x}))$, f is the **energy function**. f convex $\Rightarrow p$ **log-concave**.
 $\alpha(\mathbf{x}' | \mathbf{x}) = \min \left\{ 1, \frac{r(\mathbf{x} | \mathbf{x}')}{r(\mathbf{x}' | \mathbf{x})} \exp(f(\mathbf{x}) - f(\mathbf{x}')) \right\}$.
 $S[p(\mathbf{x})] = f(\mathbf{x}) + \log Z$
MALA/LMC: Shift the proposal distribution perpendicularly to the gradient of the energy function: $r(\mathbf{x}' | \mathbf{x}) = \mathcal{N}(\mathbf{x}'; \mathbf{x} - \eta_t \nabla f(\mathbf{x}), 2\eta_t \mathbf{I})$.
ULA: Unadjusted Langevin Algorithm (MALA) with $\alpha(\mathbf{x}' | \mathbf{x}) = 1$.
SGLD: approximate gradient of ULA with unbiased estimator
HMC: lift samples up to a higher dimension and use Hamiltonian dynamics to sample from the target distribution.

Diffusion: Simulate Gaussian noising process as MC and learn backward process.
 $q(x_t | x_{t-1}) = \mathcal{N}(x_t; \sqrt{1 - \beta_t} x_{t-1}, \beta_t \mathbf{I})$
 $q(x_t | x_0) = \mathcal{N}(x_t; \sqrt{\bar{\alpha}_t} x_0, 1 - \bar{\alpha}_t), \bar{\alpha}_t = \prod_{j=1}^t (1 - \beta_j)$
 $q(x_{t-1} | x_t, x_0) = \mathcal{N}(x_{t-1}; \mu_t'(x_t, x_0), \beta_t' \mathbf{I})$
 $\mu_t'(x_t, x_0) = \frac{(1 - \bar{\alpha}_{t-1}) \sqrt{\alpha_t} x_t}{1 - \bar{\alpha}_t} + \frac{1 - \alpha_t \sqrt{\bar{\alpha}_{t-1}}}{1 - \bar{\alpha}_t} x_0$
 $= \frac{1}{\alpha_t} \left(x_t + \frac{1 - \alpha_t}{\sqrt{1 - \bar{\alpha}_t}} \epsilon_{\lambda}(x_t, t) \right); \beta_t' = \sigma_t^2$
 $L_t = \frac{(1 - \alpha_t)^2}{\sigma_t^2 (1 - \bar{\alpha}_t) \alpha_t} \|\epsilon - \epsilon_{\lambda}(x_t, t)\|^2$

6 Bayesian Deep Learning
Bayesian neural networks: Gaussian prior on weights $\theta \sim \mathcal{N}(\mathbf{0}, \sigma_p^2 \mathbf{I})$, and Gaussian likelihood to describe how well the data is described by the model:
 $y | \mathbf{x}, \theta \sim \mathcal{N}(f(\mathbf{x}; \theta), \sigma_n^2)$. The MAP estimate is: $\hat{\theta}_{\text{MAP}} = \arg\min_{\theta} \frac{1}{2\sigma_p^2} \|\theta\|_2^2 + \frac{1}{2\sigma_n^2} \sum_{i=1}^n (y_i - f(\mathbf{x}_i; \theta))^2$. Update rule: $\theta \leftarrow \theta(1 - \frac{\eta_t}{\sigma_p}) + \eta_t \sum_{i=1}^n \nabla \log p(y_i | \mathbf{x}_i, \theta)$

Heteroscedastic Noise: Use a neural network with 2 outputs f_1, f_2 , and define: $y | \mathbf{x}, \theta \sim \mathcal{N}(\mu(\mathbf{x}; \theta), \sigma^2(\mathbf{x}; \theta))$ where $\mu(\mathbf{x}; \theta) \doteq f_1(\mathbf{x}; \theta)$ and $\sigma^2(\mathbf{x}; \theta) \doteq \exp(f_2(\mathbf{x}; \theta))$.
Approximate inference:
Variational inference: $p(\mathbf{y}^* | \mathbf{x}^*, \mathcal{D}) \approx \mathbb{E}_{\theta \sim q_{\lambda}} [p(\mathbf{y}^* | \mathbf{x}^*, \theta)] \approx \frac{1}{m} \sum_{i=1}^m p(\mathbf{y}^* | \mathbf{x}^*, \theta^{(i)})$.
MCMC/SWA: store T snapshots $\theta^{(1)}, \dots, \theta^{(T)}$ and sample from Gaussian approximation: $\theta \sim \mathcal{N}(\mu, \Sigma)$ with $\mu = \frac{1}{T} \sum_{i=1}^T \theta^{(i)}$ and $\Sigma = \frac{1}{T-1} \sum_{i=1}^T (\theta^{(i)} - \mu)(\theta^{(i)} - \mu)^T$.
Probabilistic ensembles: run m models on m independently sampled datasets and average the predictions.

Dropout/Dropconnect: we also need to perform dropout/dropconnect during inference.

Algorithm 7.3: Stein variational gradient descent, SVGD
1 initialize particles $\{\theta^{(i)}\}_{i=1}^m$
2 repeat
3 for each particle $i \in [m]$ do
4 $\theta^{(i)} \leftarrow \theta^{(i)} + \eta \hat{\phi}_{\theta^{(i)}, p}(\theta^{(i)})$ where $\hat{\phi}_{\theta, p}(\theta) \doteq \frac{1}{m} \sum_{j=1}^m [k(\theta, \theta^{(j)}) \nabla_{\theta} \log p(\theta) + \nabla_{\theta^{(j)}} k(\theta, \theta^{(j)})]$
5 until converged

Expected calibration error: For m bins $\ell_{\text{ECE}} \doteq \sum_{m=1}^M \frac{|B_m|}{n} |\text{freq}(B_m) - \text{conf}(B_m)|$
Maximum Calibration Error: $\ell_{\text{MCE}} \doteq \max_m |\text{freq}(B_m) - \text{conf}(B_m)|$

Histogram binning: calculate $q_m = \text{freq}(B_m)$ on validation set and return q_m when confidence is in B_m during inference.
Platt scaling: replace logits z_i with $\sigma(az_i + b)$ and find the optimal a, b .
Temperature scaling: Platt scaling with $b=0$ and $a = \frac{1}{T}$.

7 Active Learning
 $H[\mathbf{X} | \mathbf{Y}] = \mathbb{E}_{\mathbf{y} \sim p} [H[\mathbf{X} | \mathbf{Y} = \mathbf{y}]] = \mathbb{E}_{\mathbf{x}, \mathbf{y} \sim p} [-\log p(\mathbf{x} | \mathbf{y})]$
 $H[\mathbf{X}, \mathbf{Y}] \doteq \mathbb{E}_{\mathbf{x}, \mathbf{y} \sim p} [-\log p(\mathbf{x}, \mathbf{y})]$
 $H[\mathbf{X}, \mathbf{Y}] = H[\mathbf{Y}] + H[\mathbf{X} | \mathbf{Y}] = H[\mathbf{X}] + H[\mathbf{Y} | \mathbf{X}]$
 $H[\mathbf{X} | \mathbf{Y}] = H[\mathbf{Y} | \mathbf{X}] + H[\mathbf{X}] - H[\mathbf{Y}]$ (Bayes Rule)
 $H[\mathbf{X} | \mathbf{Y}] \leq H[\mathbf{X}]$ (Information never hurts)
 $\mathbf{I}(\mathbf{X}; \mathbf{Y}) \doteq H[\mathbf{X}] + H[\mathbf{Y}] - H[\mathbf{X}, \mathbf{Y}] = \text{KL}(p(\mathbf{x}, \mathbf{y}) || p(\mathbf{x}) p(\mathbf{y}))$
 $\mathbf{I}(\mathbf{X}; \mathbf{Y} | \mathbf{Z}) = H[\mathbf{X} | \mathbf{Z}] - H[\mathbf{X} | \mathbf{Y}, \mathbf{Z}] = \text{KL}(p(\mathbf{x}, \mathbf{y} | \mathbf{z}) || p(\mathbf{x} | \mathbf{z}) p(\mathbf{y} | \mathbf{z}))$
 $\mathbf{I}(\mathbf{X}; \mathbf{Y}; \mathbf{Z}) = \mathbf{I}(\mathbf{X}; \mathbf{Z}) + \mathbf{I}(\mathbf{X}; \mathbf{Y} | \mathbf{Z})$

MI of dependent Gaussians: given $X \sim \mathcal{N}(\mu, \Sigma)$ and $Y = X + \epsilon$ with $\epsilon \sim \mathcal{N}(\mathbf{0}, \sigma^2 \mathbf{I})$: $\mathbf{I}(X; Y) = \frac{1}{2} \log |\sigma^{-2} \Sigma + \mathbf{I}|$
Given a (discrete) function $F: \mathcal{P}(\mathcal{X}) \rightarrow \mathbb{R}$:
Marginal gain: $\Delta_F(\mathbf{x} | A) \doteq F(A \cup \{\mathbf{x}\}) - F(A)$
Submodular: $\forall \mathbf{x} \in \mathcal{X} \forall A \subseteq B \subseteq \mathcal{X}: \Delta_F(\mathbf{x} | A) \geq \Delta_F(\mathbf{x} | B)$.
Monotone: $\forall A \subseteq B: F(A) \leq F(B)$.
I is monotone submodular.
Uncertainty sampling: $\mathbf{x}_{t+1} \doteq \arg\max_{\mathbf{x} \in \mathcal{X}} \Delta_F(\mathbf{x} | S_t) = \arg\max_{\mathbf{x} \in \mathcal{X}} \mathbf{I}(\mathbf{f}_{\mathbf{x}}; \mathbf{y}_{\mathbf{x}} | \mathbf{y}_{S_t})$

$= \arg\max_{\mathbf{x} \in \mathcal{X}} \mathbf{x} \cdot \log \left(1 + \frac{\sigma_{\mathbf{x}}^2(\mathbf{x})}{\sigma_n(\mathbf{x})^2} \right)$ **Greedy** maximization of **I** is a $(1 - 1/e)$ -approximation of the optimum. Does not work with heteroscedastic noise: fails to distinguish between sources of uncertainty.

Bayesian active learning by disagreement (BALD): $\mathbf{x}_{t+1} \doteq \arg\max_{\mathbf{x} \in \mathcal{X}} \mathbf{I}(\theta; \mathbf{y}_{\mathbf{x}} | \mathbf{x}_{1:t}, y_{1:t}) = \arg\max_{\mathbf{x} \in \mathcal{X}} \mathcal{H}[\mathbf{y}_{\mathbf{x}} | \mathbf{x}_{1:t}, y_{1:t}] - \mathbb{E}_{\theta | \mathbf{x}_{1:t}, y_{1:t}} [\mathcal{H}[\mathbf{y}_{\mathbf{x}} | \theta]]$

Transductive learning: $\mathbf{x}_{t+1} \doteq \arg\max_{\mathbf{x} \in \mathcal{X}} \mathbf{I}(f_{\mathbf{x}}^*; \mathbf{y}_{\mathbf{x}} | \mathbf{x}_{1:t}, y_{1:t})$
8 Bayesian Optimization
MAB: We are given a set of k actions, and want to maximize reward.

The **Regret** for a time horizon T associated with choices $\{\mathbf{x}_t\}_{t=1}^T$ is defined as: $R_T \doteq \sum_{t=1}^T \left(\max_{\mathbf{x}} f^*(\mathbf{x}) - f^*(\mathbf{x}_t) \right)$.
instantaneous regret

Goal: **sublinear regret:** $\lim_{T \rightarrow \infty} \frac{R_T}{T} = 0$.
Acquisition function used to greedily pick the next point to sample based on the current model
Well-calibrated confidence intervals: with probability $\geq 1 - \delta$: $\forall \mathbf{x} \in \mathcal{X}: f^*(\mathbf{x}) \in \mathcal{C}_t(\mathbf{x}) = [\mu_t(\mathbf{x}) - \beta_t \sigma_t(\mathbf{x}), \mu_t(\mathbf{x}) + \beta_t \sigma_t(\mathbf{x})]$.
UCB: $\mathbf{x}_{t+1} \doteq \arg\max_{\mathbf{x} \in \mathcal{X}} \mu_t(\mathbf{x}) + \beta_t + 1 \sigma_t(\mathbf{x})$
Choosing $\beta_t(\delta) \in \mathcal{O}(\sqrt{\log(|\mathcal{X}|t/\delta)})$

approximation we get: $R_T = \mathcal{O}(\sqrt{T\gamma_T})$,
 where $\gamma_T \doteq \max_{S \subseteq \mathcal{X}} \mathbb{I}(\mathbf{f}_S; \mathbf{y}_S) = \frac{1}{|S|=T}$
 $\max_{S \subseteq \mathcal{X}} \frac{1}{2} \log \det \left(\mathbf{I} + \sigma_n^{-2} \mathbf{K}_{SS} \right)$, is
 the maximum information gain after T rounds.
Information gain of some kernels:
 Linear: $\gamma_T = \mathcal{O}(\log T)$
 Gaussian: $\gamma_T = \mathcal{O}((\log T)^{d+1})$

Improvement: $I_t(x) \doteq \max\{f(x) - \hat{f}_t, 0\}$
PI: $x_{t+1} = \operatorname{argmax}_{x \in \mathcal{X}} \mathbb{P}(I_t(x) > 0)$
EI: $x_{t+1} = \operatorname{argmax}_{x \in \mathcal{X}} \mathbb{E}[I_t(x)]$
Thompson Sampling: sample $\tilde{f}_{t+1} \sim p(\cdot | x_{1:t}, y_{1:t})$ and select $\mathbf{x}_{t+1} \doteq \operatorname{argmax}_{\mathbf{x} \in \mathcal{X}} \tilde{f}_{t+1}(\mathbf{x})$.
Information ratio: $\Psi_t(x) \doteq \frac{\Delta(x)^2}{I_t(x)}$,
 with $\Delta(x) \doteq \max_{x'} f^*(x') - f^*(x)$
IDS: $x_{t+1} = \operatorname{argmin}_{x \in \mathcal{X}} \left\{ \Psi_t(x) = \frac{\hat{\Delta}_t(x)^2}{I_t(x)} \right\}$
 with $\hat{\Delta}_t(x) \doteq \max_{x'} \hat{f}_t(x') - \hat{f}_t(x)$

9 Markov Decision Processes
 A (finite) Markov decision process is specified by a (finite) set of states $X \doteq \{1, \dots, n\}$; a (finite) set of actions $A \doteq \{1, \dots, m\}$; transition probabilities $p(x' | x, a) \doteq \mathbb{P}(X_{t+1} = x' | X_t = x, A_t = a)$; a reward function $r : X \times A \rightarrow \mathbb{R}$ which maps the current state x and an action a to some reward.
 r induces a sequence of rewards: $R_t \doteq r(X_t, A_t)$.
 A policy is a function that maps each state $x \in X$ to a probability distribution over the actions. That is, for any $t > 0$: $\pi(a | x) \doteq \mathbb{P}(A_t = a | X_t = x)$.

A policy induces a MC $(X_t^\pi)_{t \in \mathbb{N}_0}$
Discounted payoff: $G_t \doteq \sum_{m=0}^{\infty} \gamma^m R_{t+m}$,
 $\gamma \in [0, 1)$ is the discount factor.

State value function: $v_t^\pi(x) \doteq \mathbb{E}_\pi[G_t | X_t = x]$
State-action value function (Q-function): $q_t^\pi(x, a) \doteq \mathbb{E}_\pi[G_t | X_t = x, A_t = a]$
Bellman Expectation Equation:
 $v^\pi(x) = r(x, \pi(x)) + \gamma \mathbb{E}_{x' | x, \pi(x)}[v^\pi(x')]$
 For stochastic policies: $v^\pi(x) = \mathbb{E}_{a \sim \pi(x)}[q^\pi(x, a)]$

Can be used to find v^π given policy π ,
 by solving linear system of equations in $\mathcal{O}(n^3)$.
Fixed point iteration: $\mathbf{B}^\pi \mathbf{v} \doteq \mathbf{r}^\pi + \gamma \mathbf{P}^\pi \mathbf{v}$.
 \mathbf{B}^π is contraction with contraction factor $\gamma < 1 \implies$ unique optimal value function v^* .

Greedy policy: $\pi(x) \doteq \operatorname{argmax}_{a \in A} q_t^\pi(x, a)$
Bellman's Theorem: A policy π^* is optimal iff it is greedy w.r.t. its own value function.

Bellman optimality equations:
 $v^*(x) = \max_{a \in A} q^*(x, a)$
 $q^*(x, a) = r(x, a) + \gamma \mathbb{E}_{x' | x, a}[\max_{a' \in A} q^*(x', a')]$
Algorithm 10.17: Policy iteration
 initialize π (arbitrarily)
repeat
 compute v^π
 compute π_{opt}
 $\pi \leftarrow \pi_{\text{opt}}$
until converged
 For finite MDPs, policy iteration converges to an optimal policy (monotonic improvement).

Algorithm 10.20: Value iteration
 initialize $v(x) \leftarrow \max_{a \in A} r(x, a)$ for each $x \in X$
for $t = 1$ **to** ∞ **do**
 $v \leftarrow (B^* v)(x) = \max_{a \in A} q(x, a)$ for each $x \in X$
 choose π_v
 Value iteration to an ϵ -optimal policy in polynomial time, as v^* and q^* are a fixed-points of the Bellman update B^* .
 v_t corresponds to the optimal value function assuming only t steps are ever taken.
A Partially observable Markov decision process (POMDP) is a Markov process with hidden states, a set of supplementary observations Y , and observation probabilities $o(y | x) \doteq \mathbb{P}(Y_t = y | X_t = x)$. Given a POMDP, the corresponding Belief-state Markov decision process is a Markov decision process specified by the belief space $\mathcal{B} \doteq \Delta^X$; the set of actions A ; transition probabilities $\tau(b' | b, a) \doteq \mathbb{P}(B_{t+1} = b' | B_t = b, A_t = a)$; and rewards $\rho(b, a) \doteq \mathbb{E}_{x \sim b}[r(x, a)] = \sum_{x \in X} b(x) r(x, a)$.
 $b_{t+1}(x) = \mathbb{P}(X_{t+1} = x | y_{1:t+1}, a_{1:t})$ is deterministic.
 $\mathbb{P}(y_{t+1} | b_t, a_t) = \mathbb{E}_{x \sim b_t}[\mathbb{E}_{x' | x, a_t}[\mathbb{P}(y_{t+1} | X_{t+1} = x')]] = \sum_{x \in X} b_t(x) \sum_{x' \in X} p(x' | x, a_t) \cdot o(y_{t+1} | x')$.

10 Tabular Reinforcement Learning
The reinforcement learning problem: probabilistic planning in unknown environments. A trajectory τ is a sequence: $\tau \doteq (r_0, \tau_1, \tau_2, \dots)$, with $\tau_i \doteq (x_i, a_i, r_i, x_{i+1})$.
On-policy (On): Agent chooses policy.
Off-policy (Off): No choice of policy. More sample efficient, less stable.
Model-based (MB): Learn underlying MDP. More sample efficient, allows for planning and transfers well to new tasks.
Model-free (MF): Learn value function directly. Simpler, doesn't suffer from model bias, tends to perform better.
Value estimation (VE): Learn value function given policy.
Control (C): Determine optimal policy
 A sequence $(\pi_t)_{t \in \mathbb{N}_0}$ of policies is greedy in the limit of infinite exploration (GLIE) if: All pairs (x, a) are visited infinitely often.
 $\lim_{t \rightarrow \infty} \pi_t(a | x) = \mathbf{1}\{a = \operatorname{argmax}_{a' \in A} Q_t^*(x, a')\}$, where Q_t^* is the optimal action-value function for the estimated MDP at time t .

Model-based MLE: $\hat{p}(x' | x, a) = \frac{N(x' | x, a)}{N(a | x)}$
 $\hat{r}(x, a) = \frac{1}{N(a | x)} \sum_{i=0, x_t=x, a_t=a}^{\infty} r_t$
 ϵ -greedy: With probability ϵ , choose a random action, otherwise choose the action with the highest value. $(\epsilon_t)_{t \in \mathbb{N}_0}$ satisfies RM \implies GLIE \implies convergence. **softmax exploration:** $\pi(a | x) \propto \exp(Q(x, a) / \lambda)$ with temperature $\lambda > 0$.

Algorithm 11.6: R_{\max} algorithm
 add the fairy-tale state x^* to the Markov decision process
 set $\hat{r}(x, a) = R_{\max}$ for all $x \in X$ and $a \in A$
 set $\hat{p}(x^* | x, a) = 1$ for all $x \in X$ and $a \in A$
 compute the optimal policy $\hat{\pi}$ for \hat{r} and \hat{p}
for $t = 0$ **to** ∞ **do**
 execute policy $\hat{\pi}$ (for some number of steps)
 for each visited state-action pair (x, a) , update $\hat{r}(x, a)$
 estimate transition probabilities $\hat{p}(x' | x, a)$
 after observing "enough" transitions and rewards, recompute the optimal policy $\hat{\pi}$ according to the current model \hat{r} and \hat{p} .

With probability at least $1 - \delta$, R_{\max} reaches an ϵ -optimal policy in a number of steps that

is polynomial in $|X|, |A|, T, 1/\epsilon, 1/\delta$, and R_{\max} .
TD learning: On/MF/VE
 $V(x) \leftarrow V(x) + \alpha_t(r + \gamma V(x') - V(x))$
SARSA: On/MF/VE
 $Q(x, a) \leftarrow Q(x, a) + \alpha_t(r + \gamma Q(x', a') - Q(x, a))$
 Off-policy version (expected SARSA): $Q(x, a) \leftarrow Q(x, a) + \alpha_t(r + \gamma \mathbb{E}_{a' \sim \pi(x')} [Q(x', a')] - Q(x, a))$
Q learning: Off/MF/C $Q(x, a) \leftarrow (1 - \alpha_t)Q(x, a) + \alpha_t \left(r + \gamma \max_{a' \in A} Q(x', a') \right)$
 $(\alpha_t)_{t \in \mathbb{N}_0}$ satisfies RM + GLIE \implies convergence for TD, SARSA and Q learning.
 All 3 methods can be initialized arbitrarily.

Algorithm 11.14: Optimistic Q-learning
 1 initialize $Q^*(x, a) = V_{\max} \prod_{i=1}^{\text{init}} (1 - a_i)^{-1}$
 2 **for** $t = 0$ **to** ∞ **do**
 3 pick action $a_t = \operatorname{argmax}_{a \in A} Q^*(x, a)$ and observe the transition (x, a_t, r, x')
 4 $Q^*(x, a_t) \leftarrow (1 - a_t)Q^*(x, a_t) + a_t(r + \gamma \max_{a' \in A} Q^*(x', a'))$ // (11.27)

With probability at least $1 - \delta$, Q learning converges to an ϵ -optimal policy in a number of steps that is polynomial in $|X|, |A|, T, 1/\epsilon, 1/\delta$, and R_{\max} .

11 Model-free Reinforcement Learning
Parametric value function approximation: learn approximation $V(\mathbf{x}; \theta)$ or $Q(\mathbf{x}, \mathbf{a}; \theta)$ parametrized by θ . Can view TD-learning as SGD on the squared loss $\ell(\theta; x, r, x') \doteq \frac{1}{2} (r + \gamma V^{\text{old}}(x') - \theta(x))^2$.
Q-learning with function approximation: scaling to large state spaces (Off/MF/C)
Bellman error: $\delta_B \doteq r + \gamma \max_{a' \in A} Q^*(x', a'; \theta^{\text{old}}) - Q^*(x, a; \theta)$.
 Update: $\theta \leftarrow \theta + \alpha_t \delta_B \nabla_\theta Q^*(x, a)$ with $\theta^{\text{old}} = \theta$ being treated as constant w.r.t. θ .

DDQN: stabilizing targets
 Train 2 separate networks: target network and online network. $\ell_{\text{DDQN}} \doteq \frac{1}{2} (r + \gamma \max_{a' \in A} Q^*(x', a'; \theta^{\text{old}}) - Q^*(x, a; \theta))^2$.
 Update target network with hard updates or Polyak averaging: $\theta^{\text{old}} \leftarrow \alpha \theta + (1 - \alpha) \theta^{\text{old}}$
DDQN: avoiding maximization bias
 Choose maximum action from online network and evaluate it with target network.

Policy optimization/Policy gradient methods: directly optimize policy π_φ instead of value function.
Trajectory distribution:
 $\Pi_\varphi(\tau) \doteq p(x_0) \prod_{t=0}^{T-1} p(x_{t+1} | x_t, a_t) \pi_\varphi(a_t | x_t)$
Policy value function: $j(\varphi) = j(\pi_\varphi) = \mathbb{E}_{\pi_\varphi}[G_0] = \mathbb{E}_{\pi_\varphi}[\sum_{t=0}^{\infty} \gamma^t R_t]$
Bounded variant: $j_T(\pi) \doteq \mathbb{E}_\pi[G_{0:T}]$

Score function trick:
 $\nabla_\varphi \mathbb{E}_{\tau \sim \Pi_\varphi}[G_0] = \mathbb{E}_{\tau \sim \Pi_\varphi}[G_0 \nabla_\varphi \log \Pi_\varphi(\tau)] = \mathbb{E}_{\tau \sim \Pi_\varphi}[G_0 \sum_{t=0}^{T-1} \nabla_\varphi \log \pi_\varphi(a_t | x_t)]$.
 Has high variance unlike the reparametrization trick.
Baseline:
 $\mathbb{E}_{\tau \sim \Pi_\varphi}[G_0 \nabla_\varphi \log \Pi_\varphi(\tau)] = \mathbb{E}_{\tau \sim \Pi_\varphi}[\sum_{t=0}^{T-1} (G_0 - b(\tau_0:t-1)) \nabla_\varphi \log \pi_\varphi(a_t | x_t)]$
REINFORCE (On/MF/C): Select baseline $b_t = g_{0:t-1}$: $\nabla_\varphi j_T(\varphi) = \mathbb{E}_{\tau \sim \Pi_\varphi}[\sum_{t=0}^{T-1} \gamma^t g_{t:T} \nabla_\varphi \log \pi_\varphi(a_t | x_t)]$

Policy gradient theorem:
 $\nabla j(\varphi) = \sum_{t=0}^{\infty} \mathbb{E}_{x_t, a_t} [\gamma^t q^\pi \varphi(x_t, a_t) \nabla_\varphi \log \pi_\varphi(a_t | x_t)]$
 $\propto \mathbb{E}_{x \sim \rho_\varphi^\infty} \mathbb{E}_{a \sim \pi_\varphi(a|x)} [q^\pi \varphi(x, a) \nabla_\varphi \log \pi_\varphi(a | x)]$
Actor-Critic methods: scaling to large action spaces
 Parameterized policy $\pi(a | \mathbf{x}; \varphi) \doteq \pi_\varphi$ (actor)
 Value function approximation $q^\pi \varphi(\mathbf{x}, \mathbf{a}) \approx Q^\pi \varphi(\mathbf{x}, \mathbf{a}; \theta)$ (critic).
On-policy AC: learn critic through SARSA and actor through policy gradient methods

Algorithm 12.10: Online actor-critic
 1 initialize parameters φ and θ
 2 **repeat**
 3 use π_φ to obtain transition (x, a, r, x')
 4 $\delta = r + \gamma Q(x', \pi_\varphi(x'); \theta) - Q(x, a; \theta)$ // actor update
 5 $\varphi \leftarrow \varphi + \eta \gamma Q(x, a; \theta) \nabla_\varphi \log \pi_\varphi(a | x)$ // critic update
 6 $\theta \leftarrow \theta + \eta \delta \nabla_\theta Q(x, a; \theta)$
 7 **until** converged

Advantage: $a^\pi(\mathbf{x}, \mathbf{a}) \doteq q^\pi(\mathbf{x}, \mathbf{a}) - v^\pi(\mathbf{x})$
 π is optimal $\iff \forall \mathbf{x} \in \mathcal{X}, \mathbf{a} \in A: a^\pi(\mathbf{x}, \mathbf{a}) \leq 0$
Advantage actor-critic (A2C):
 replace Q with advantage function A (predicting sign is easier than predicting absolute quantity).
 Advantage isn't directly parametrized: we parametrize V^π and approximate Q with $\sum_{t=k}^T \gamma^{t-k} r_t + \gamma^{T-k} V^\pi(x_{T+1})$.
 When compared to REINFORCE, actor-critic methods have lower variance and higher bias.
TRPO: improving sample efficiency in on-policy AC (On/MF/C)
 $\varphi_{k+1} \leftarrow \operatorname{argmax}_\varphi J(\varphi)$ subject to $\text{KL}(\pi_{\varphi_k}(\cdot | x) || \pi_\varphi(\cdot | x)) \leq \delta$
 $J(\varphi) \doteq \mathbb{E}_{x \sim \rho_{\varphi_k}^\infty, a \sim \pi_{\varphi_k}(\cdot | x)} [w_k(\varphi; x, a) A^{\pi_{\varphi_k}}(x, a)]$
 $w_k(\varphi; x, a) \doteq \frac{\pi_\varphi(a | x)}{\pi_{\varphi_k}(a | x)}$ are the importance sampling weights.
PPO: unconstrained objective $\operatorname{argmax}_\varphi J(\varphi) - \frac{1}{\varphi}$
 $\lambda \mathbb{E}_{x \sim \rho_{\varphi_k}^\infty} \text{KL}(\pi_{\varphi_k}(\cdot | x) || \pi_\varphi(\cdot | x))$

GRPO: improving compute efficiency
 PPO with heuristic approximation of advantage $\hat{A}_{t,i} = \frac{g_{t:T} - \text{mean}(g_{t:T})}{\text{std}(g_{t:T})}$
Off-policy AC: Parametrize maximum over actions with π_φ .
 Objective (deterministic policy gradients/hill-climbing): $\varphi^* = \operatorname{argmax}_\varphi J_\mu(\varphi) = \operatorname{argmax}_\varphi \mathbb{E}_{x \sim \mu} [Q^*(x, \pi_\varphi(x); \theta)]$
 Learn critic through bootstrapped Q-learning
 The exploration distribution $\mu(x)$ is typically selected as uniform sampling from replay buffer.
 Exploration can be achieved through Gaussian dithering as in DDPG.
TD3 = DDPG with 2 critic networks for evaluating policy and calculating maximum.

For randomized policies like SVG, we get: $J_\mu(\varphi) = \mathbb{E}_{x \sim \mu} \mathbb{E}_{a \sim \pi_\varphi(\cdot | x)} [Q^*(x, a; \theta)]$
Reparametrize to get the gradient: $\nabla_\varphi J_\mu(\varphi) = \mathbb{E}_{x \sim \mu} \mathbb{E}_{\epsilon \sim \varphi} [D_a Q^*(x, a) |_{a=g(\epsilon; \varphi)} D_\varphi g(\epsilon; \varphi)]$
MERL: encourage exploration through new objective: $j_\lambda(\varphi) = j(\varphi) + \lambda \mathbb{H}[\Pi_\varphi]$

$= \sum_{t=0}^{\infty} \mathbb{E}_{x_t, a_t} \mathbb{E}_{\pi_\varphi} [r(x_t, a_t) + \lambda \mathbb{H}[\pi_\varphi(\cdot | x_t)]]$
 Assuming HMM defined by $p(O_t | x_t, a_t) \propto \exp(\frac{1}{\lambda} r(x_t, a_t))$, we get:
 $\Pi_\star(\tau) = p(\tau | \mathcal{O}_{1:T}) \propto \left[p(x_1) \prod_{t=1}^{T-1} p(x_{t+1} | x_t, a_t) \right] \exp\left(\frac{1}{\lambda} \sum_{t=1}^T r(x_t, a_t)\right)$
 The objective $\text{KL}(\Pi_\star || \Pi_\varphi)$ is equivalent to the MERL objective.
Soft-Actor-Critic and **MAP optimization** are off-policy actor-critic methods with entropy regularization
Soft-value function: $q^*(x, a) = \frac{1}{\lambda} r(x, a) + \mathbb{E}_{x' \sim x, a} [\log \int_A \exp(q^*(x', a')) da']$
 Changes the value function, not just the objective
Finetuning LLMs:
Bradley-Terry model: $p(y_A \succ y_B | x, r) = \sigma(r(y_A | x) - r(y_B | x))$
RLHF: - 1. calculate reward with MLE ($\theta = \operatorname{argmax}_\theta p(D | r_\theta)$) 2. calculate policy with PPO
Optimal policy: $\Pi_\star \propto \Pi_{\text{init}} \exp(\frac{1}{\lambda} r(y | x))$
DPO: optimize over $r_\varphi(y | x) = \log \frac{\Pi_\varphi}{\Pi_{\text{init}}} + \text{const.}$

12 Model-based Reinforcement Learning
Strict generalization of model-free RL.
Algorithm 13.1: Model-based reinforcement learning (outline)
 start with an initial policy π and no (or some) initial data \mathcal{D}
 for several episodes **do**
 roll out policy π to collect data
 learn a model of the dynamics f and rewards r from data
 plan a new policy π based on the estimated models

Assuming the dynamics are known and deterministic ($x_{t+1} = f(x_t, a_t)$):
MPC: plan over finite time horizon H
 To solve the problem of sparse rewards, add $\gamma^H V(x_H)$ to the reward
Target shooting:
 generate random sequences of actions and choose the best one. (primitive tree search method)
Trajectory sampling = MPC with stochastic dynamics
closed-loop control: planning done online
open-loop control: policy precomputed offline \implies apply model-free techniques to the new objective
Learning dynamics: $x_{t+1} \sim f(x_t, a_t; \psi)$ (approximate) greedy optimization:
PILCO for GPs, **PETS** for neural networks
Thompson sampling: sample f from posterior and maximize
Optimistic exploration: optimize over set $\mathcal{M}(\mathcal{D})$ of "plausible" models
H-UCRL:
 $\pi_{t+1} = \operatorname{argmax}_\pi \max_{\eta(\cdot) \in [-1, 1]^d} J_H(\pi; \hat{f}_t)$ with $\hat{f}_t(x, a) = \mu_t(x, a) + \beta_t \eta(x, a) \sigma_t(x, a)$
Constrained optimization: $\max_\pi J_\mu(\pi; f)$ subject to $J_\mu^c(\pi; f) = \mathbb{E}_{x \sim \mu, x_{1:\infty} \sim \pi, f} [\sum_{t=0}^{\infty} \gamma^t c(x_t)] \leq \delta$
 Assuming a family of plausible models $\mathcal{M}(\mathcal{D})$, we can be optimistic w.r.t. rewards and pessimistic w.r.t. constraints.
 $\max_\pi \max_{f \in \mathcal{M}(\mathcal{D})} J_\mu(\pi; f)$
 subject to $\max_{f \in \mathcal{M}(\mathcal{D})} J_\mu^c(\pi; f) \leq \delta$
 By Leo Schmidt-Traub
 - based off of Nils Jensen's notes.