

CS258 Information Theory Homework 9

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Exercise 1 (Channel with two independent looks at Y) Let Y_1 and Y_2 be conditionally independent and conditionally identically distributed given X

1. Show that $I(X; Y_1, Y_2) = 2I(X; Y_1) - I(Y_1; Y_2)$
2. Conclude that the capacity of the channel $X \mapsto Y_1, Y_2$ is less than twice the capacity of the channel $X \mapsto Y_1$.

Proof.

1.

$$\begin{aligned} I(X; Y_1, Y_2) &= H(Y_1, Y_2) - H(Y_1, Y_2|X) \\ &= H(Y_1, Y_2) - H(Y_1|X) - H(Y_2|X) \\ &= H(Y_1) + H(Y_2) - I(Y_1; Y_2) - H(Y_1|X) - H(Y_2|X) \\ &= I(X; Y_1) + I(X; Y_2) - I(Y_1, Y_2) \\ &= 2I(X; Y_1) - I(Y_1, Y_2) \end{aligned} \tag{1}$$

2.

$$\begin{aligned} C_1 &= \max_{p(x)} I(X; Y_1, Y_2) \\ &= \max_{p(x)} (2I(X; Y_1) - I(Y_1, Y_2)) \\ &\leq \max_{p(x)} 2I(X; Y_1) \\ &= 2C_2 \end{aligned} \tag{2}$$

□

Exercise 2 (Two-look Gaussian channel) Given $X \mapsto Y_1, Y_2$. Consider the ordinary Gaussian channel with two correlated looks at X , that is, $Y = (Y_1, Y_2)$, where

$$\begin{aligned} Y_1 &= X + Z_1 \\ Y_2 &= X + Z_2 \end{aligned}$$

with a power constraint P on X , and $(Z_1, Z_2) \sim \mathcal{N}_2(0, K)$, where

$$K = \begin{bmatrix} N & N\rho \\ N\rho & N \end{bmatrix}$$

Find the capacity C for

1. $\rho = 1$
2. $\rho = 0$
3. $\rho = -1$

Solution. From Theorem 8.6.5 [Cover] we know that the Gaussian distribution maximizes the entropy over all distributions with the same variance. Hence it is clear that normally distributed $X \sim \mathcal{N}(0, P)$ will maximize the mutual information. In this case $(Y_1, Y_2) \sim \left(0, \begin{bmatrix} P+N & P+\rho N \\ P+\rho N & P+N \end{bmatrix}\right)$

$$\begin{aligned} \max I(X; Y_1, Y_2) &= h(Y_1, Y_2) - h(Y_1, Y_2|X) \\ &= h(Y_1, Y_2) - h(Z_1, Z_2) \\ &= \frac{1}{2} \log (2\pi e)^2 \left| \begin{bmatrix} P+N & P+\rho N \\ P+\rho N & P+N \end{bmatrix} \right| - \frac{1}{2} \log (2\pi e)^2 \left| \begin{bmatrix} N & N\rho \\ N\rho & N \end{bmatrix} \right| \\ &= \frac{1}{2} \log \left(1 + \frac{2P}{(1+\rho)N} \right) \end{aligned} \quad (3)$$

1. $\rho = 1, C = \frac{1}{2} \log \left(1 + \frac{P}{N} \right)$
2. $\rho = 0, C = \frac{1}{2} \log \left(1 + \frac{2P}{N} \right)$
3. $\rho = -1, C = +\infty$.

□

Exercise 3 (Output power constraint) Consider an additive white Gaussian noise channel with an expected output power constraint P . Thus, $Y = X + Z, Z \sim \mathcal{N}(0, \sigma^2)$, Z is independent of X , and $EY^2 \leq P$. Find the channel capacity.

Solution.

$$\begin{aligned} I(X; Y) &= h(Y) - h(Y|X) \\ &= h(Y) - h(Z) \\ &= h(Y) - \frac{1}{2} \log (2\pi e \sigma^2) \\ &\leq \frac{1}{2} \log (2\pi e P) - \frac{1}{2} \log (2\pi e \sigma^2) \\ &= \frac{1}{2} \log \frac{P}{\sigma^2} \end{aligned} \quad (4)$$

The equality holds when Y is normally distributed. In this case $X \sim \mathcal{N}(0, P - \sigma^2)$.

□

Exercise 4 (Exponential noise channels) $Y_i = X_i + Z_i$, where Z_i is i.i.d. exponentially distributed noise with mean μ . Assume that we have a mean constraint on the signal (i.e., $EX_i \leq \lambda$). Show that the capacity of such a channel is $C = \log \left(1 + \frac{\lambda}{\mu} \right)$

Proof.

$$\begin{aligned} I(X; Y) &= h(Y) - h(Y|X) \\ &= h(Y) - h(Z) \\ &= h(Y) - \sum_i h(Z_i) \\ &\leq \sum_i (h(Y_i) - h(Z_i)) \end{aligned} \quad (5)$$

The equality holds when Y_i s are independent, which can be obtained if X_i s are independent. Hence we can only consider the channel for the input and output to be single-valued. Still, $I(X; Y) = h(Y) - h(Z)$ holds.

Note for exponentially distributed Z ,

$$\begin{aligned}
h(Z) &= - \int_0^{+\infty} g(z) \ln \frac{1}{\mu} e^{-\frac{z}{\mu}} dz \\
&= - \int_0^{+\infty} g(z) \ln \frac{1}{\mu} dz - \int_0^{+\infty} g(z) \frac{z}{\mu} dz \\
&= 1 + \ln \mu
\end{aligned} \tag{6}$$

Note that $EY = EX + EZ \leq \lambda + \mu$. For mean-value bounded Y , by Theorem 12.1.1 and Example 12.2.5 [Cover], the maximizing differential entropy is $h^*(Y) = 1 + \ln(\lambda + \mu)$, with distribution $p^*(y) = \frac{1}{\lambda + \mu} e^{-\frac{y}{\lambda + \mu}}$. Therefore

$$I(X; Y) \leq \sum_{i=1}^n ((1 + \ln(\lambda + \mu)) - (1 + \ln(\mu))) = n \ln \frac{\lambda + \mu}{\mu} \tag{7}$$

The equality holds when X_i are independent with mean value λ and $Y_i \sim \exp\left(\frac{1}{\lambda + \mu}\right)$. We need to find such distribution for X_i . Since X_i and Z_i are independent and $Y_i = X_i + Z_i$, it follows that the characteristic functions hold the following relation.

$$\phi_Y(t) = \phi_X(t) \cdot \phi_Z(t) \tag{8}$$

Therefore

$$\begin{aligned}
\phi_X(t) &= \frac{\phi_Y(t)}{\phi_Z(t)} \\
&= \frac{(1 - i(\lambda + \mu)t)^{-1}}{(1 - i\mu t)^{-1}} \\
&= \frac{1}{\lambda + \mu} \frac{[\mu - i\mu(\lambda + \mu)t] + \lambda}{1 - i(\lambda + \mu)t} \\
&= \frac{\mu}{\lambda + \mu} + \frac{\lambda}{\lambda + \mu} (1 - i(\lambda + \mu)t)^{-1}
\end{aligned} \tag{9}$$

The characteristic function above is a linear combination of two kinds of distribution. We can set every X_i to be 0 with the probability of $\frac{\mu}{\lambda + \mu}$, and to be exponentially distributed with mean value $\lambda + \mu$ by the probability of $\frac{\lambda}{\lambda + \mu}$. Then the channel capacity $n \ln \frac{\lambda + \mu}{\mu}$ can be obtained. □